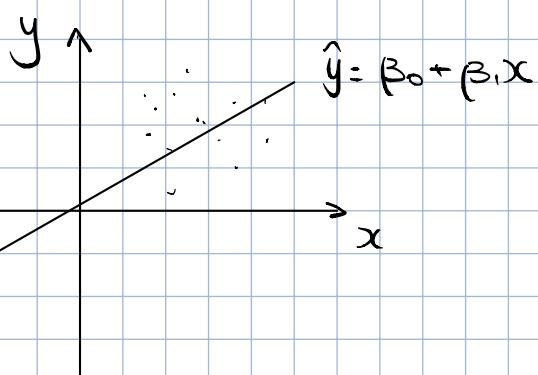
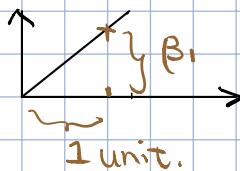


## Intuition on Stats.

Start from a simple linear regression, which has 1 predictor.



$\beta_1$ : slope.



$\Delta x \Rightarrow \Delta y : \beta_1$

model:  $\hat{y} = \beta_0 + \beta_1 x$ .

model assumption: we expect / model predicts

$\hat{y}$ : on average

slope:  $\Delta x \rightarrow \Delta y : \beta_1$

"1"  $\Delta x$ : - For numeric var, 1 unit.

- For categorical var with  $K$  levels.

"1" can represent  $1 \underbrace{2}_{\text{any gap}} 3 \dots K$

any gap between two neighbors

"distance" between neighbors may vary!

$|K-1|$  different representation of "1"  $\Delta x$

need  $|K-1|$  different  $\beta$ 's to estimate  $\Delta y$  resulted from different "1" in  $\Delta x$ .

method 1. 1 2 3 4 ... k  $\Rightarrow$  baseline change.

Method 2. 1 2 3 4 ... k  $\Rightarrow$  baseline fix:  


$$1 \Delta x \rightarrow \Delta y = \underbrace{\beta_0}_{\text{baseline}} + \underbrace{\beta_{1,1} I(\text{level 2}) + \beta_{1,2} I(\text{level 3}) + \dots + \beta_{1,k-1} I(\text{level k})}_{\text{increment}}$$

level 1  
level 2  
level 3  
⋮  
level k

Interpret  $\beta_{1,1}, \dots, \beta_{1,k-1}$  as "increment":

- "slope":  $\Delta x \rightarrow \Delta y$

Compare to level 1 (baseline)  $\rightarrow$  level 2 has  $\Delta y: \beta_{1,1}$

**model:**  $\hat{y} = \beta_0 + \beta_{1,1} I(\text{level 2}) + \beta_{1,2} I(\text{level 3}) + \dots + \beta_{1,k-1} I(\text{level k})$

**model assumption:** we expect / model predicts

$\hat{y}$ : on average

**slope:**  $\Delta x \rightarrow \Delta y: \beta_1$

Compared to level 1, level 2 will have  $\Delta y: \beta_1$

For model with 1 predictor, (categorical)

$$\hat{y} = \beta_0 + \beta_{1,1} I(\text{level 2}) + \beta_{1,2} I(\text{level 3}) + \dots + \beta_{1,k-1} I(\text{level k})$$

	$I(\text{level 2})$	$I(\text{level 3})$	$I(\text{level 4})$	...	$I(\text{level k})$
level 1	0	0	0	...	0
2	1	0	0	...	0
3	0	1	0	...	0
⋮	⋮	⋮	⋮	⋮	⋮
k	0	0	0	...	1

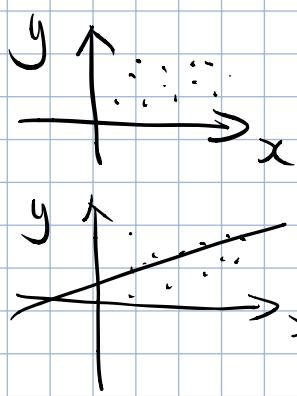
Already multivariate linear regression !

Special case ,

One observation can only have 1 level ,

# SLR.

$y \cdot x$



EDA

→ linear trend.



Linear regression

$$\text{Assumption: } y = \beta_0 + \beta_1 x + \varepsilon$$

$$\text{Model fitting: } \hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x$$

Interpretation of  $\hat{\beta}_0$ ,  $\hat{\beta}_1$

$$\text{Prediction: } \hat{y}^* = \hat{\beta}_0 + \hat{\beta}_1 x^*$$

extrapolation.

Evaluation:  $R^2$

RMSE.

Inference:

$$\begin{cases} \text{CI} & \text{--- true.} \\ \text{HT} & \beta_0 = 0 \quad \text{vs} \quad \beta_0 \neq 0 \end{cases}$$

method

{ simulation

{ CI: bootstrap

{ HT: permutation

math. ( $\varepsilon \stackrel{\text{iid}}{\sim} N(0, \sigma_\varepsilon^2)$ )

$$\begin{cases} \text{CI} & \hat{\beta} \pm t^* \times \text{se}(\hat{\beta}) \xrightarrow{\text{degree of freedom}} \\ \text{HT: } t^* = \frac{\hat{\beta}}{\text{se}(\hat{\beta})} & p\text{-val} = \Pr(|t| > t^*) \end{cases}$$

Interpretation.

## Predictive Interval.

Diagnosis:

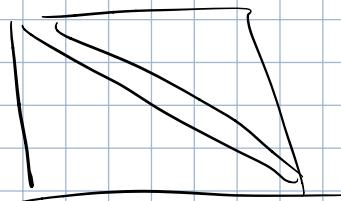
Condition (4) ← residual plot

Influential pts:

{  
leverage  
standardized residuals  
Cook's Distance

MLR.

EDA : ggplot



Model assumption:  $y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_p x_p + \varepsilon$ .

Model fitting:  $\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \dots + \hat{\beta}_p x_p$

Interpretation: - Hold constant - on average

- expect  $\Delta x \rightarrow \Delta y$

$$\hat{y}^* = f(x^*)$$

Type of predictors:

- mean-center
- categorical : dummy variable
- Interaction :  $\Delta \beta$
- Transformation :  $\log(\cdot)$ ;  $x^2 + x$

Model comparison :

ANOVA

Adj R<sup>2</sup>

AIC

BIC

Feature engineering -

split data

recipe.

workflow

fit model

prediction.

CV

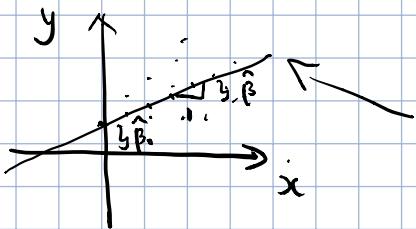
conditions : residual plot

+  
multicollinearity.

# Simple Linear Regression.

$$y \sim x$$

EDA.



$$\text{Model Assumption: } Y = \beta_0 + \beta_1 x + \varepsilon$$

$$\text{Fit: } \hat{Y} = \hat{\beta}_0 + \hat{\beta}_1 x \rightarrow$$

`tidy()`: point. se. t-stat. p-val.

$$\text{meaningful?} \rightarrow \hat{\beta}_0 = x=0 \Rightarrow \hat{Y} = \hat{\beta}_0$$

$$\begin{matrix} \uparrow \\ \text{center } x. \end{matrix} \quad \begin{matrix} \uparrow \\ \hat{\beta}_1 : \Delta x \rightarrow \Delta \hat{Y} : \hat{\beta}_1 \end{matrix}$$

Interpret.  
\* we expect / predict.

\* on average

\* 1 unit.  $\rightarrow \hat{\beta}_1$ .

$$\text{prediction: } \hat{Y}^* = \hat{\beta}_0 + \hat{\beta}_1 x^*$$

$x^*$   $\rightarrow$  extrapolation.

$$\text{evaluation: } R^2 \quad \uparrow$$

$$\text{RMSE.} \quad \downarrow$$

Inference.  $\left\{ \begin{array}{l} \text{CI: } \underline{x} \underline{x} \underline{x} \underline{x} \underline{x} \\ \text{HT} \end{array} \right.$

Interpret. CI: -  $(95\%)$

- model predict / expect.

- on average.

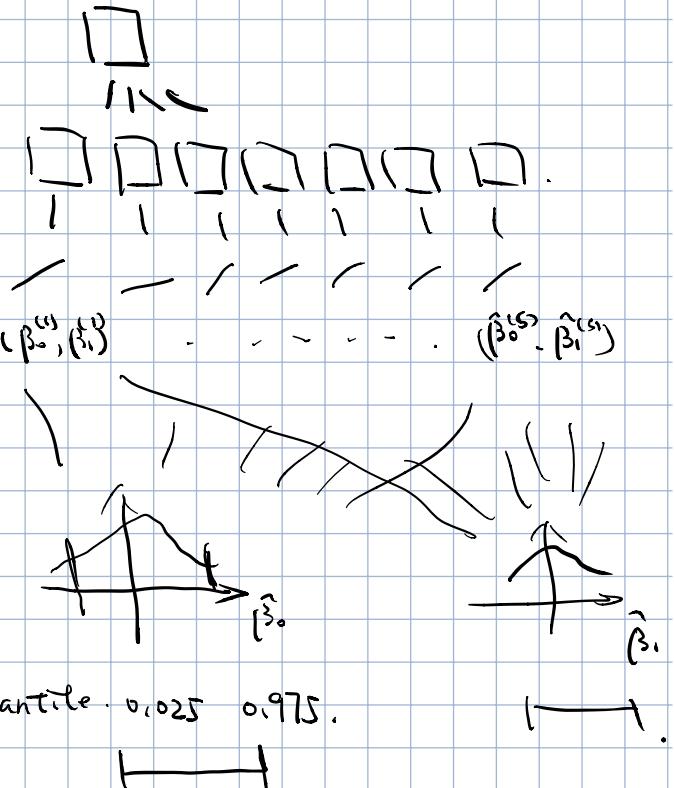
-  $\Delta x \uparrow \cdot [ \underline{x}, \overline{x} ]$ .

- CI for  $\beta$

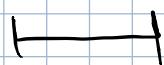
- CI for  $\bar{Y}$  individual.  $\rightarrow$  average mean.

- prediction interval.  $> \text{CI. } \bar{Y}$ .

{ simulation. Bootstrap.: sample w/ replacement  
math.



Quantile: 0.025 0.975.



$$\text{math: } Y = \beta_0 + \beta_1 X + \varepsilon$$

$$\varepsilon \sim N(0, \sigma_\varepsilon^2)$$

$$\hat{\beta} \pm t^* \times \text{se}(\hat{\beta})$$

$\uparrow$   
 $t_{df}$

$$df = n - p - 1 = n - 2.$$

↑  
# predictor

HT:

$$H_0: \beta_1 = 0 \quad \text{vs: } H_1: \beta_1 \neq 0$$

Assume  $H_0$  is true.

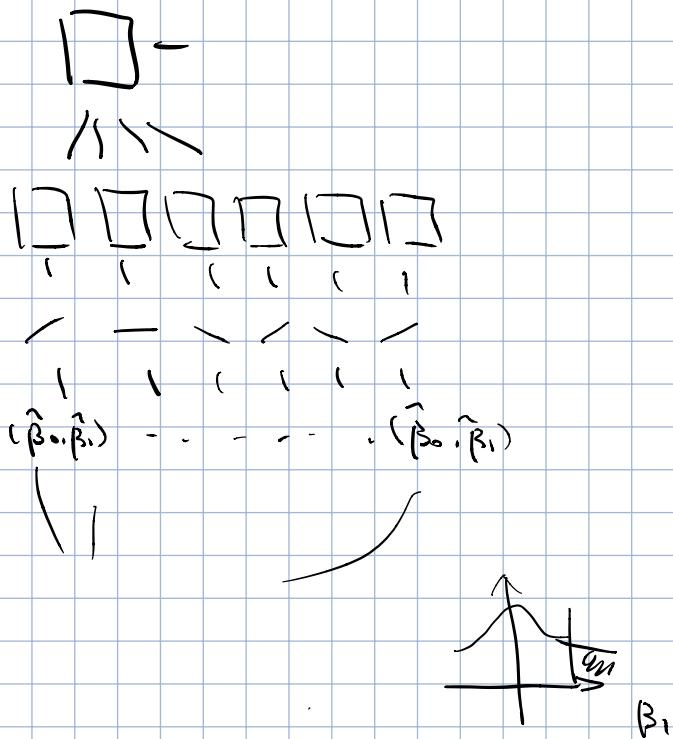
check obs. is reasonable.

$$\text{Math: t-stat. } \frac{\hat{\beta}}{\text{se}(\hat{\beta})} \sim t_{n-2}$$



$\Pr(|t^*| > t)$  p-val. small.  
reject  $H_0$ .

simulation : permutation. (sample w/o replacement)



calculate p-val.

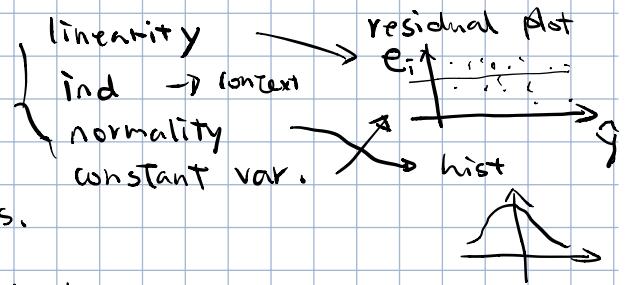
Diagnosis:

- check conditions.

- influential points.

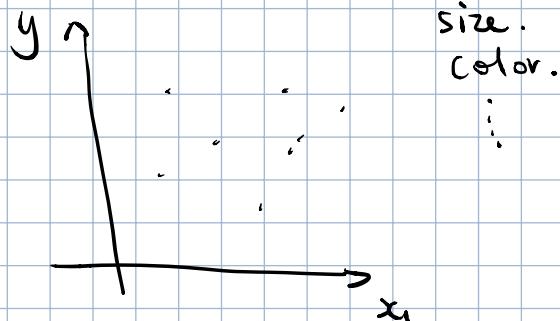
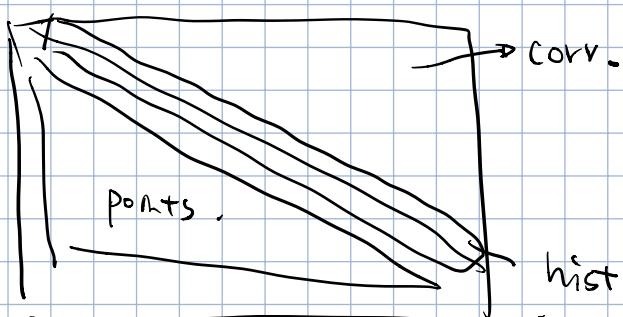
} outlier. ← standardize res. > 3  
} high leverage. ←

→ Cook dist. : 0.5, 1.



MLR.

EDA:: ggpairs(·)



$$\text{Model : } Y = \beta_0 + \beta_1 x_1 + \dots + \beta_p x_p + \epsilon.$$

$$\text{Fit : } \hat{Y} = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \dots + \hat{\beta}_p x_p + \epsilon.$$

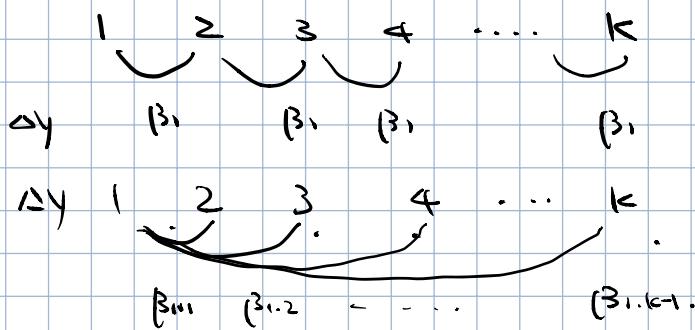
hold other constant.

Interpretation.  $\hat{\beta}_1$  (4) - on avg - we expect. - 1 unit  $\rightarrow \Delta Y : \hat{\beta}_1$ .  
 - hold other constant.

Δx : numeric var  $\Delta x \rightarrow \Delta y : \hat{\beta}_1$

$\Delta x$ : 1 unit increase.

$\Delta x$  : categorical var.  $\leftarrow$  level.  $\Delta$



$k-1$  dummy.  $I(x_i=1)$   $I(x_i=2)$   $\dots$   $I(x_i=k)$ .

1	0	0	...	0
2	1	0	...	0
:				
k	0	0	...	1

$$Y = \beta_0 + \beta_1 x_1 + \dots$$

$\downarrow$   $x_1$  is k-level cate...

$$Y = \underbrace{\beta_0}_{=} + \underbrace{\beta_{1,1} I(x_1=1)}_{=} + \dots + \underbrace{\beta_{1,k-1} I(x_1=k)}_{=}$$

$$\begin{cases} \beta_0 + \beta_{0,1} \\ \vdots + \dots \\ \beta_0 + \beta_{0,k-1} \\ \downarrow \\ \text{baseline.} \end{cases}$$

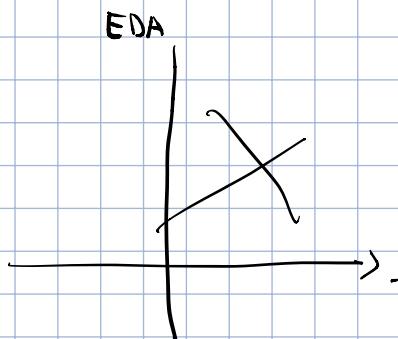
$\beta_0$  increment. - Δx  $\Rightarrow \Delta y$  = compare to level 1.

we expect the  $y$  of obs w/ level 2. will increase by  $\beta_{1,1}$ , on avg. hold other constant.

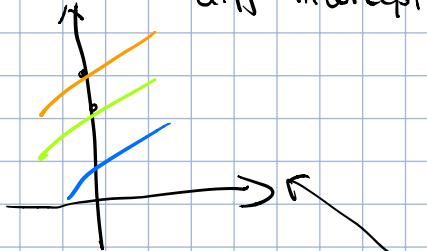
$\hat{\beta}_1$  : obs. with level ...

Interpretation : 1. numeric vs. categorical.

2. interaction.



w/o int. same slope,  
diff intercept



$$Y = (\beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_1 x_2)$$

$x_1$ : numeric       $x_2$  = categorical.

$$Y = (\beta_0 + (\beta_2 x_2) + (\beta_1 x_1))$$

↑  
intercept.

$$Y = (\underbrace{\beta_0 + (\beta_2 x_2)}_{\text{inter}}) + (\underbrace{(\beta_1 + (\beta_3 x_2)) x_1}_{\text{slope}})$$

$$\begin{matrix} \beta_0 \\ \beta_0 + \beta_{2,1} \\ \vdots \\ \beta_0 + \beta_{2,k-1} \end{matrix}$$

$$\begin{matrix} \beta_1 \\ \vdots \\ \beta_1 + \beta_{3,1} \\ \vdots \\ \beta_1 + \beta_{3,k-1} \end{matrix}$$

baseline + increment.

↓

$\Delta Y$

$\Delta \beta$ :

the potential influence  
of

- $\Delta x_2 \rightarrow \Delta \beta$  :  $\hat{\beta}_{3,1}$
- expect.
- on avg
- hold other const.

Compare to level 1. we expect. obs w/

(level 2. on. outcome. increase by  $\hat{\beta}_{3,1}$ )

on avg. hold other constant.

- mean center.

$\Delta x \rightarrow \Delta y$

- log. transformation.

$$y = \beta_0 + \beta_1 \log x_1$$

$$\log(x_1 + \Delta x)$$

$\Delta x \rightarrow \Delta y$ .

$$y = \beta_0 + \beta_1 \log(1+x) = (\beta_0 + \beta_1 \log 1 + \beta_1 \log x)$$

x increase by 1%.

$$y = \beta_1 \log 1.1$$

$$x + 0.1x$$

$$\log(x_1 + c) = \log x_1 + \log c$$

$$y = \beta_0 + \beta_1 \log x + \epsilon$$

? linear regression.

Yes,

linearity wrt.  $\beta$ .

Type. of predictors.

Model comparison.

} anova.  $H_0: \beta_1 = \dots = \beta_k = 0$ .  
 p-val. small, reject  $H_0$ .  
 Adj.  $R^2$ : ↑  
 AIC, BIC = b.

Feature engineer: extract feature from predict.

represent predictor.

Spend data: split  $\begin{cases} \text{training} \\ \text{test} \end{cases}$  + prediction.



create recipe:

- `recipe(...)`
- `step-...`

} recipe.

} model.

recipe + model.

} workflow.

- - - ,

" "

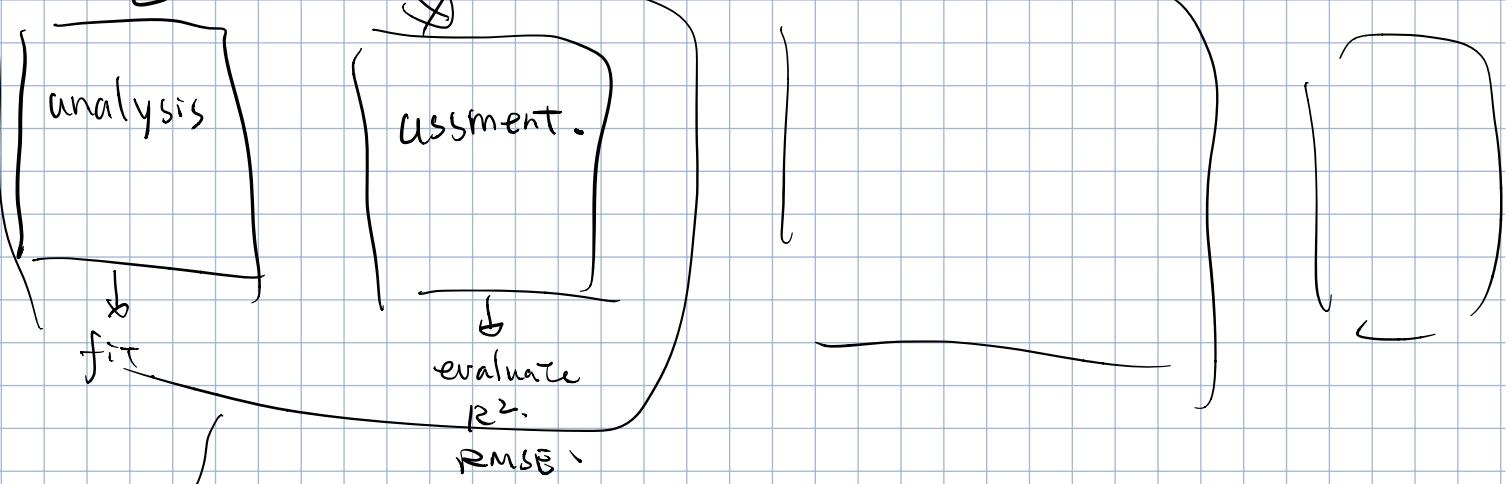
} fit. model.

} prediction.

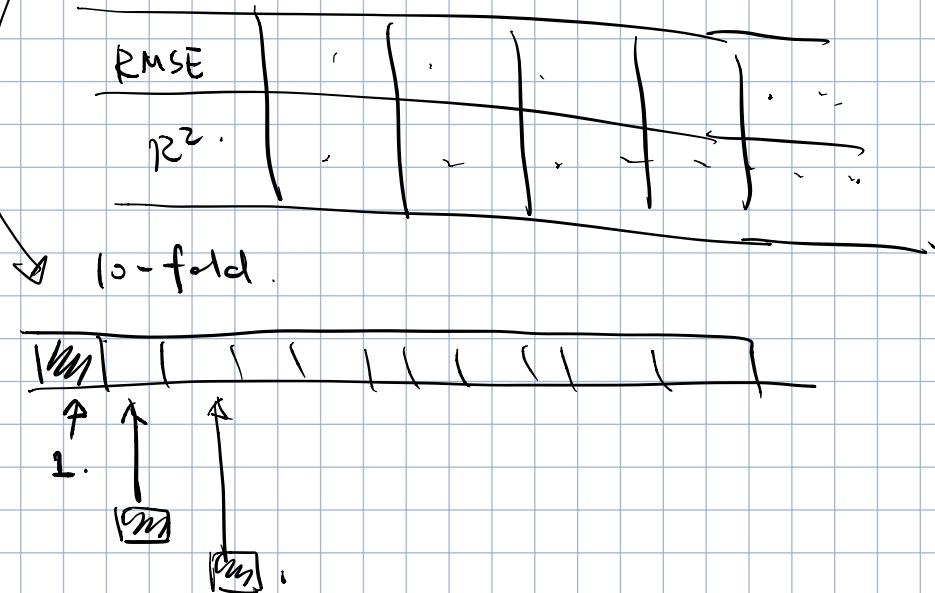
$\begin{cases} ? \\ CV \end{cases} \rightarrow \text{test.}$

$\begin{cases} \text{training} \\ + \text{test.} \end{cases}$





CV.



cv. }

Inference. | CI.

H.T. =  $H_0: \beta_1 = 0$  hold other constant

Diagnosis : conditions : (4). residual plot. - SLR.

. multicollinearity.  $\rightarrow$  non-identifiable.

$\downarrow$   
 $se(\cdot) \uparrow$ .

detect:  $\uparrow$   
Inference imprecise.  
corr.  $> 0.9$ .

{ abnormal. coeff.  
VIF.  $> 10$ .

Select one to remove.



4.

Model Comparison