



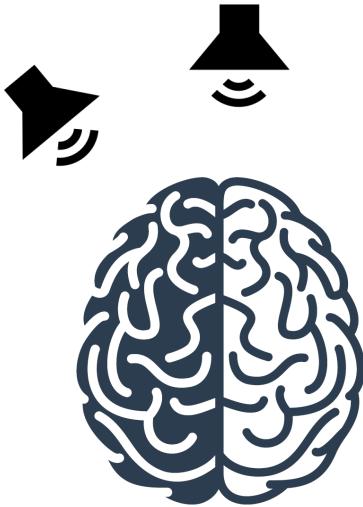
Detection of Code Juggling with Spike Counts Data Analysis

Yunran Chen and Surya T. Tokdar

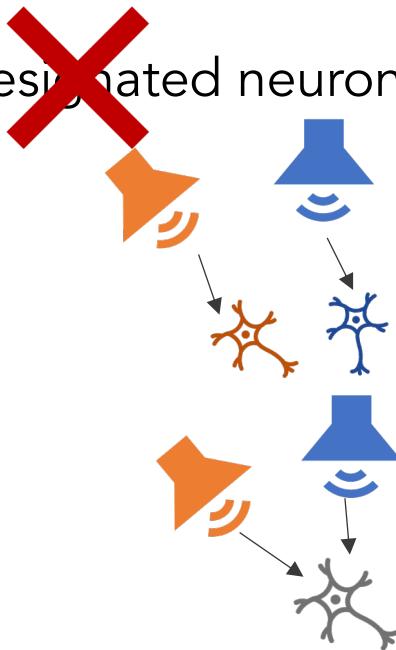
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Motivation

How does brain preserve information about multiple simultaneous items ?



- Designated neurons for each signal



- Small **receptive field** (perceptual sensitivity)



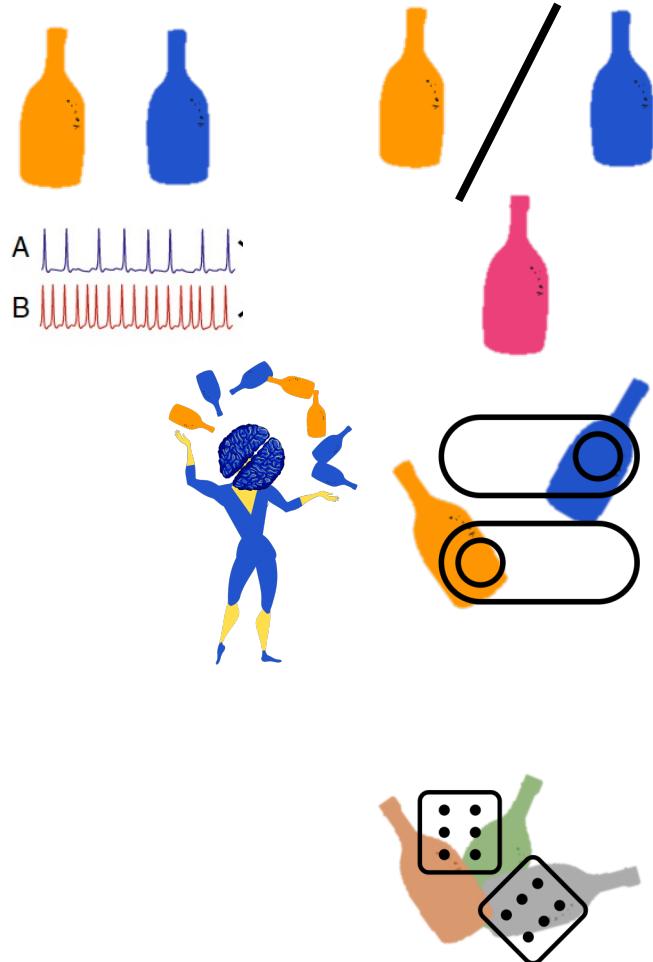
- Receptive fields too large

- Visual (Alonso and Chen, 2009; Keliris et al., 2019)
- Auditory (Groh et al., 2003; Werner-Reiss and Groh, 2008; Bulkin and Groh, 2011)

A single neuron will be exposed to multiple simultaneous stimuli !

- **Dynamics** in presentation in a neural level
- Can a single neuron preserve info from both and

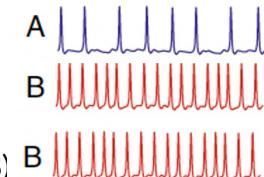
Motivation: potential dynamics



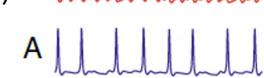
- Always encode A (or B)
- Treat it as a new stimulus
- Switch between A and B
- **Code Juggling**
- **2nd order stochasticity: constituent stimuli**
 - Greater efficiency on information preserving
 - Overdispersed + bounded by A & B
- Stochasticity not related A nor B
 - Overdispersed
 - Not related to A nor B



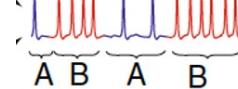
- 1st order stochasticity
 - Poisson distribution
- (Ventura et al., 2002; Kass et al., 2005)



- Across trials
- (Caruso et al., 2018)



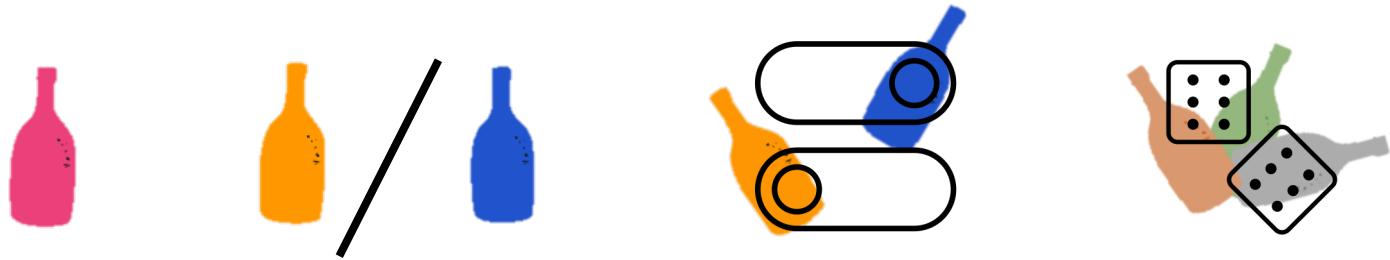
- Within a trial
- (Glynn et al., 2021)



Higher trial-to-trial
variability
(Semedo et al., 2019)

Research question:

How to **model** these biologically meaningful encoding schemes?



Poisson Mixtures

How to **select** the most likely scheme and quantify the uncertainty?

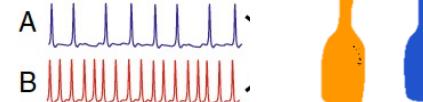


Hypothesis Testing
(Bayes Factor)

Statistical Analysis Framework

Assumptions:

- Two-stimuli cases
- Single stimulus spike counts follow Poisson distribution



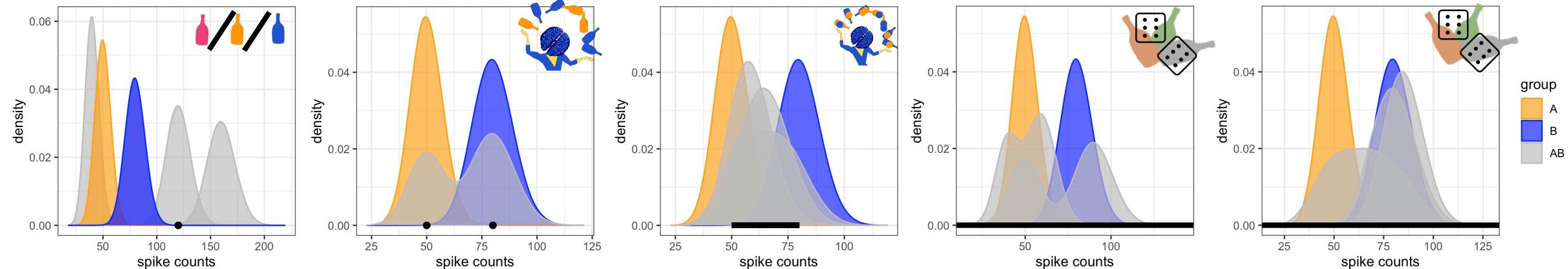
$$P^A = \text{Poi}(\mu^A), \quad P^B = \text{Poi}(\mu^B), \quad \mu^A, \mu^B > 0.$$

Data: $(Y_j^e : j \in \{1, \dots, n_e\}, e \in \{A, B, AB\})$

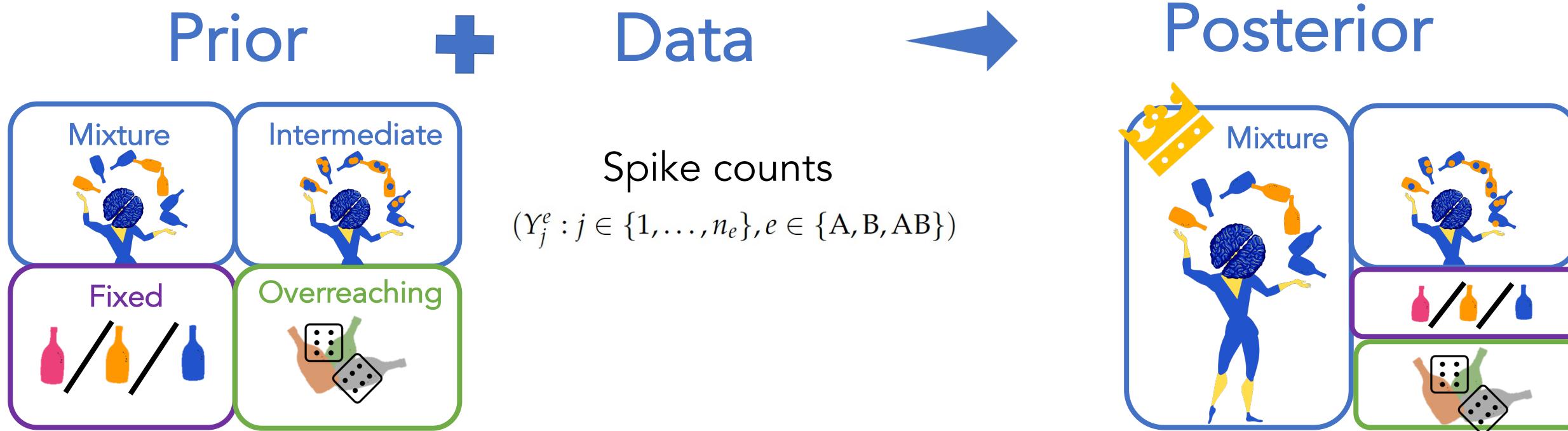
Hypotheses:

- (whole-trial) FIXED: $P^{AB} = \text{Poi}(\mu), \quad \mu \in (0, \infty),$
- (sub-trial) MIXTURE: $P^{AB} = \rho \text{Poi}(\mu^A) + (1 - \rho) \text{Poi}(\mu^B), \quad \rho \in (0, 1),$
- INTERMEDIATE: $P^{AB} = \int \text{Poi}(r\mu^A + (1 - r)\mu^B) f(r) dr, \quad f$ is a density on $(0, 1),$
- OVERREACHING: $P^{AB} = \int \text{Poi}(m) g(m) dm, \quad g$ is a density on $(0, \infty),$

Probability distributions (support + continuity):



Hypothesis Testing in a Bayesian framework



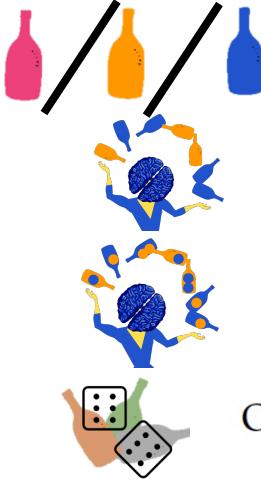
Challenges: nonparametric density estimation and fair competition

Predictive recursion marginal likelihood method (PRML)

(Newton et al., 1998; Newton, 2002; Tokdar et al., 2009; Martin and Tokdar, 2011)

Statistical Analysis Framework

Hypotheses:



- Straightforward
- FIXED: $P^{AB} = \text{Poi}(\mu)$, $\mu \in (0, \infty)$,
unknown
- MIXTURE: $P^{AB} = \rho \text{Poi}(\mu^A) + (1 - \rho) \text{Poi}(\mu^B)$, $\rho \in (0, 1)$,
 μ^A μ^B
- INTERMEDIATE: $P^{AB} = \int \text{Poi}(r\mu^A + (1 - r)\mu^B) f(r) dr$, f is a density on $(0, 1)$,
 r μ^A μ^B
- OVERREACHING: $P^{AB} = \int \text{Poi}(m) g(m) dm$, g is a density on $(0, \infty)$,
 m $g(m)$

Statistical Analysis Framework

Hypotheses:



FIXED: $P^{AB} = \text{Poi}(\mu)$,

$\mu \in (0, \infty)$,

Marginal likelihood: $p(Y^A, Y^B, Y^{AB}) = p(Y^A)p(Y^B)p(Y^{AB}|Y^A, Y^B)$

$$p(Y^{AB}|Y^A, Y^B) = p(Y^{AB}) = \int p(Y^{AB}|\mu)\pi(\mu)d\mu$$

Set a prior and can obtain closed form

Statistical Analysis Framework



OVERREACHING: $P^{AB} = \int_{-\infty}^{\infty} \text{Poi}(m)g(m)dm,$ g is a density on $(0, \infty),$

Marginal likelihood: $p(Y^{AB}|Y^A, Y^B) = p(Y^{AB})$

Predictive Recursion (Smith and Markov, 1978; Newton et al., 1998; Newton, 2002)

Data Y_1, \dots, Y_n from mixture distribution

$$m_f(y) = \int \kappa(y | u) f(u) d\nu(u)$$

For $i \in \{1, \dots, n\}$

$$f_i(u) = (1 - w_i)f_{i-1}(u) + w_i \frac{\kappa(Y_i | u)f_{i-1}(u)}{m_{i-1}(Y_i)}, \quad u \in \mathcal{U},$$

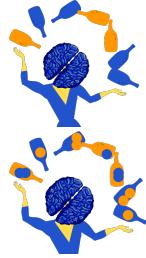
$$m_{i-1}(y) = \int \kappa(y | u)f_{i-1}(u)d\nu(u), \quad y \in \mathcal{Y}$$

Theoretical guarantee (Tokdar et al., 2009) :

$f_n \quad m_n$ converge to the truth asymptotically (provided $\sum_{i=1}^{\infty} w_i = \infty \quad \sum_{i=1}^{\infty} w_i^2 < \infty$)

Bayesian inferential paradigm interpretation (Tokdar et al., 2009)

Statistical Analysis Framework



MIXTURE: $P^{AB} = \rho Poi(\mu^A) + (1 - \rho) Poi(\mu^B), \quad \rho \in (0, 1),$

INTERMEDIATE: $P^{AB} = \int Poi(r\mu^A + (1 - r)\mu^B) f(r) dr, \quad f \text{ is a density on } (0, 1),$

Marginal likelihood: $p(Y^{AB}|Y^A, Y^B) = \int p(Y^{AB}|\theta) p(\theta|Y^A, Y^B) d\theta \quad \theta = (\mu^A, \mu^B)$

First-stage prior: $\pi_0(\mu^A, \mu^B) = 1 / \sqrt{\mu^A \mu^B},$

Second-stage prior: $\pi(\mu^A, \mu^B) = \text{Gam}(\mu^A | 0.5 + \sum_{j=1}^{n_A} Y_j^A, n_A) \times \text{Gam}(\mu^B | 0.5 + \sum_{j=1}^{n_B} Y_j^B, n_B)$

$$m_f(y) = \int \kappa_\theta(y \mid u) f(u) d\nu(u)$$

PRML score (Martin and Tokdar, 2011)

Predictive Recursion Marginal Likelihood

Data Y_1, \dots, Y_n from mixture distribution

$$i \in \{1, \dots, n\}$$

$$m_f(y) = \int \kappa_\theta(y \mid u) f(u) d\nu(u)$$

$$f_i(u) = (1 - w_i) f_{i-1}(u) + w_i \frac{\kappa_\theta(Y_i \mid u) f_{i-1}(u)}{m_{i-1}(Y_i)}, \quad u \in \mathcal{U},$$

$$m_{i-1,\theta}(y) = \int \kappa_\theta(y \mid u) f_{i-1}(u) d\nu(u), \quad y \in \mathcal{Y}$$

$$\kappa = \kappa_\theta$$

$$L_n(\theta) := \prod_{i=1}^n m_{i-1,\theta}(Y_i)$$

Predictive Recursion

(Newton et al., 1998; Newton, 2002)

PRML score

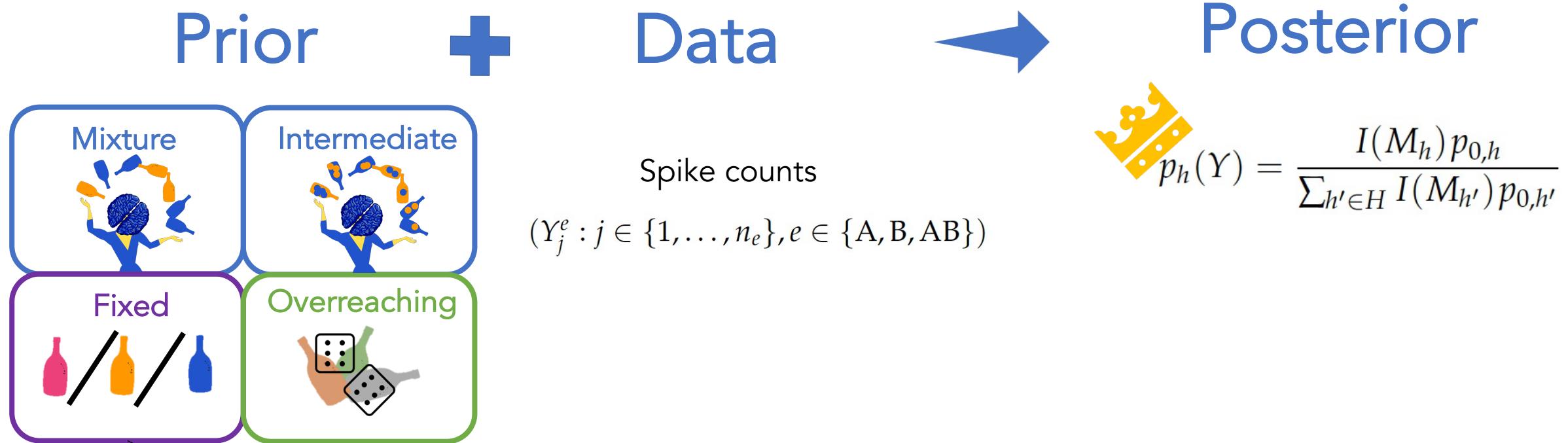
(Martin and Tokdar, 2011)

- PRML score estimate true marginal density: $p(Y_1, \dots, Y_n \mid \theta) = \prod_{i=1}^n p(Y_i \mid Y_1, \dots, Y_{i-1}, \theta)$
- Asymptotic consistency: $n^{-1} \log\{L_n(\theta)/L_n(\theta^\star)\} \rightarrow -\inf_{f \in \mathbb{F}} d_{\text{KL}}(m^\star, m_{f,\theta})$.
- Interpretable: expectation filtration approximation to a fully Bayesian estimation $w_i = (1 + a)^{-1}$, $a > 0$.

Laplace Approximation: $I(M_h) := \int_{\Theta_h} L_{n,h}(\theta_h) \pi_h(\theta_h) d\theta_h, \quad h \in H,$

$$\hat{I}(M_h) = L_{n,h}(\hat{\theta}_h) (2\pi)^{-d_h/2} |\Sigma_h|^{1/2},$$

Fair competition: choice of prior

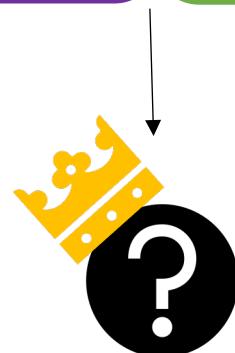
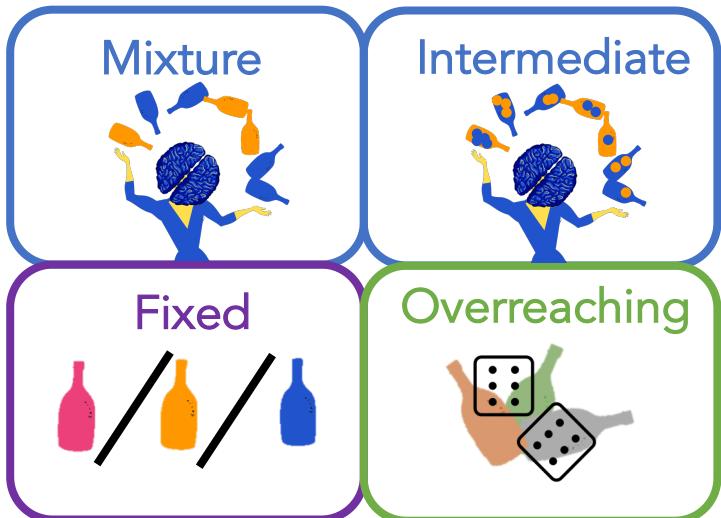


- Sensitive to the choice of prior
 - Bounded uniform (same as overreaching)
 - Jeffreys' prior (improper) - - > tune the constant
 - Jeffreys' prior (intrinsic Bayes factor) (Berger & Pericchi, 1996)



Two potential frameworks

Full Bayesian



Jeffreys' prior: tune constant to match the confidence level

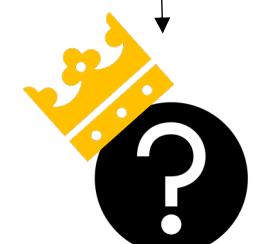
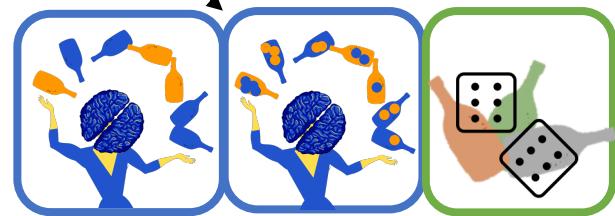
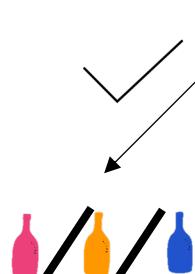
$$p(Y^{AB}|Y^A, Y^B) = p(Y^{AB}) = \int p(Y^{AB}|\mu)\pi(\mu)d\mu$$

$$\pi(\mu) \propto \mu^{-1/2}$$

Two-stage procedure

Poisson variance test
(Brown and Zhao, 2002)

→ Is it Fixed?



Performance Assessment

100 experiment sets: $p^A = \text{Poi}(50)$ $p^B = \text{Poi}(80)$ $n_A = n_B = n_{AB}$

FIXED: $P^{AB} = \text{Poi}(80m)$

Multiplier

MIXTURE: $P^{AB} = m\text{Poi}(50) + (1 - m)\text{Poi}(80)$

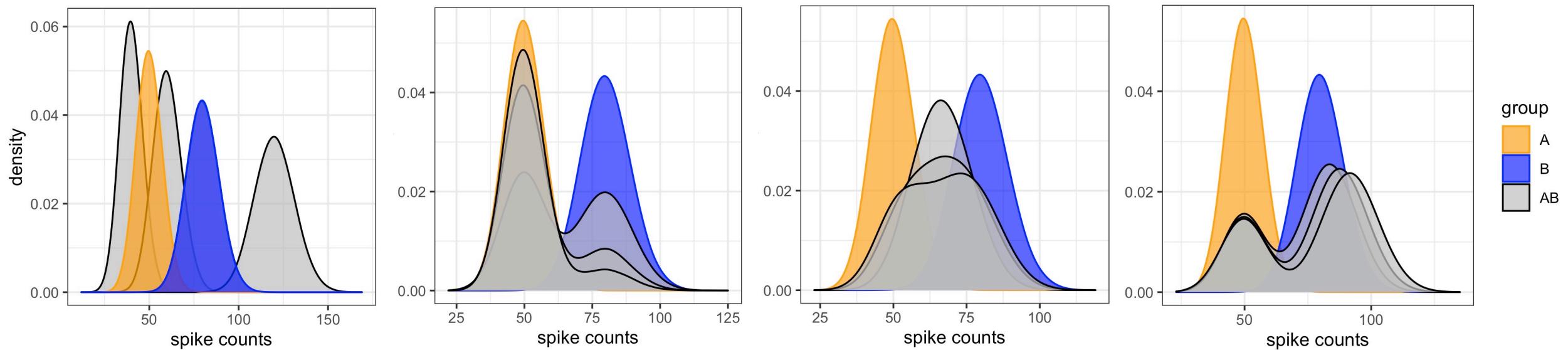
Skewness

INTERMEDIATE: $P^{AB} = \int \text{Poi}(r50 + (1 - r)80)\text{Beta}(r; a, b)dr$

Fano factor

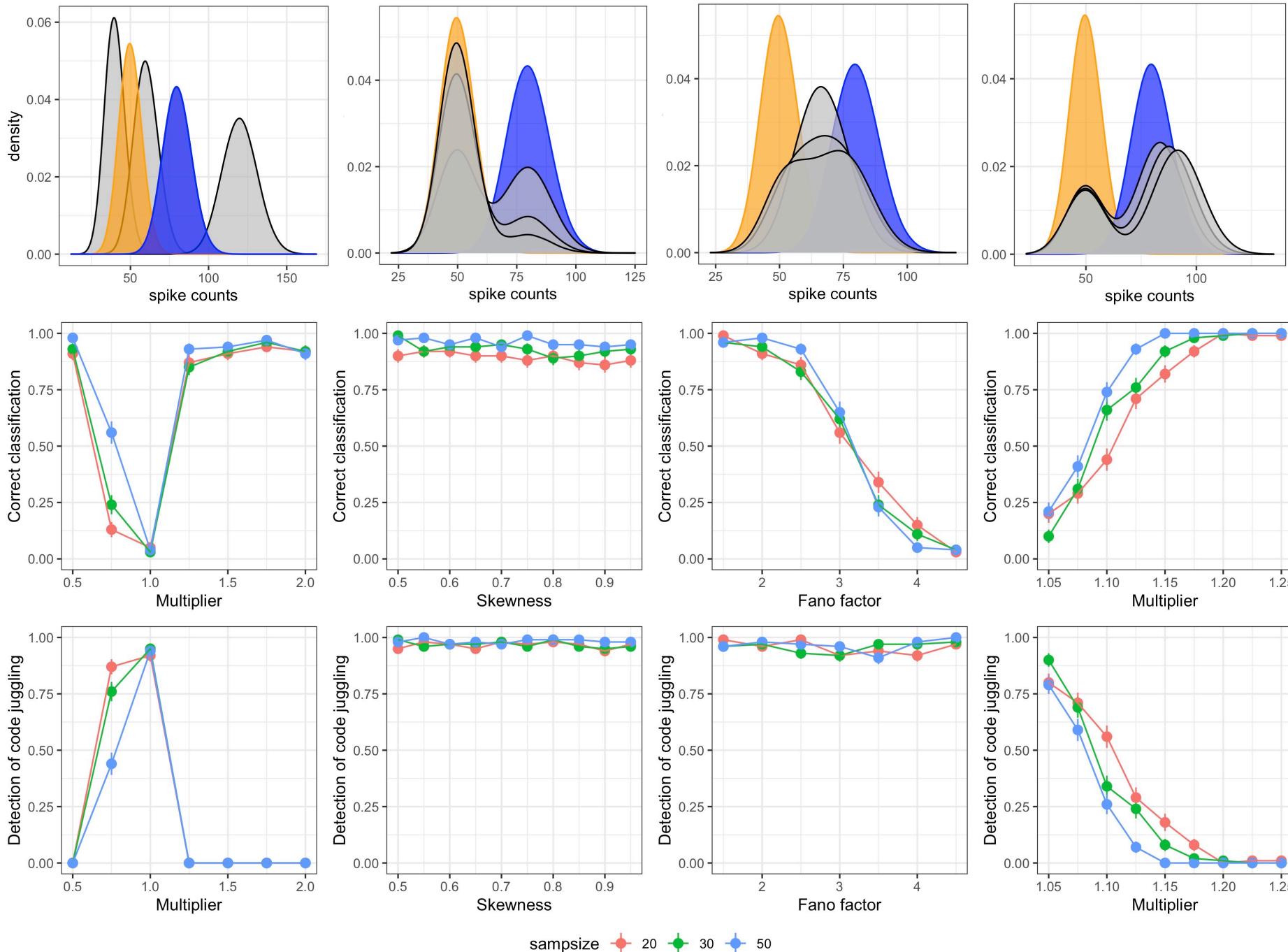
OVERREACHING: $P^{AB} = 1/3\text{Poi}(50) + 2/3\text{Poi}(80m)$

Multiplier



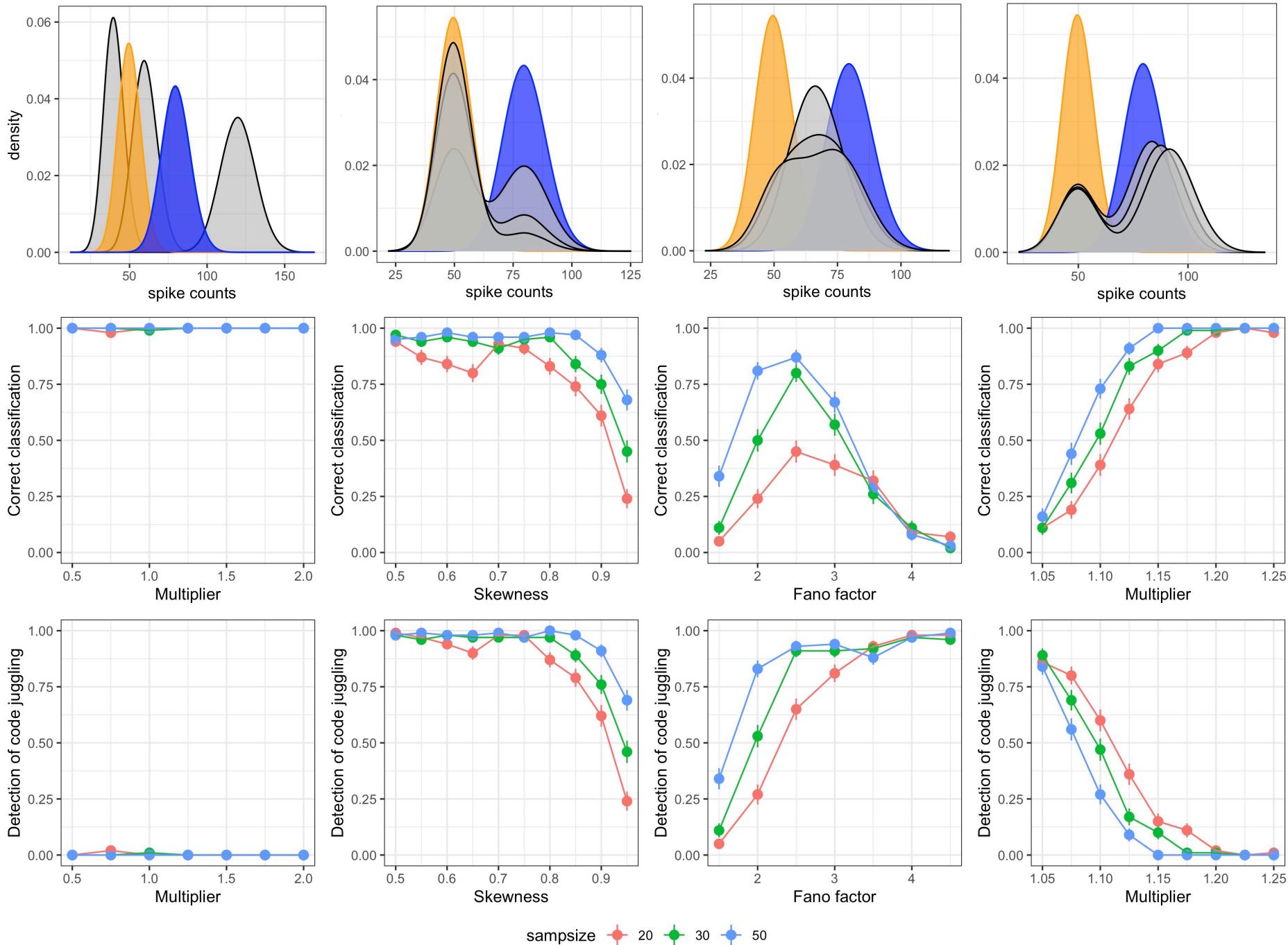
Performance Assessment

Bounded Uniform



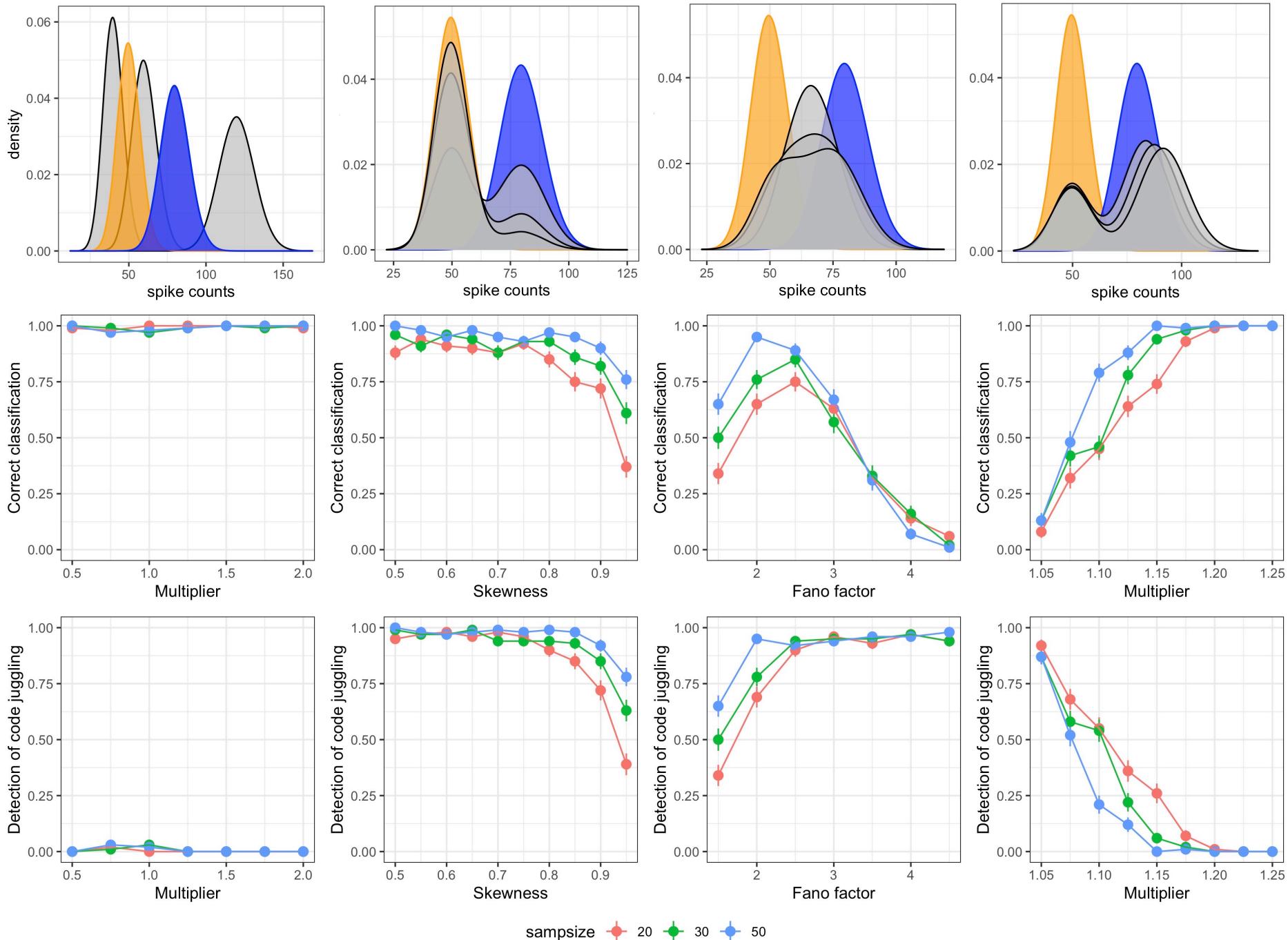
Performance Assessment

Jeffreys' prior
(intrinsic Bayes factor)



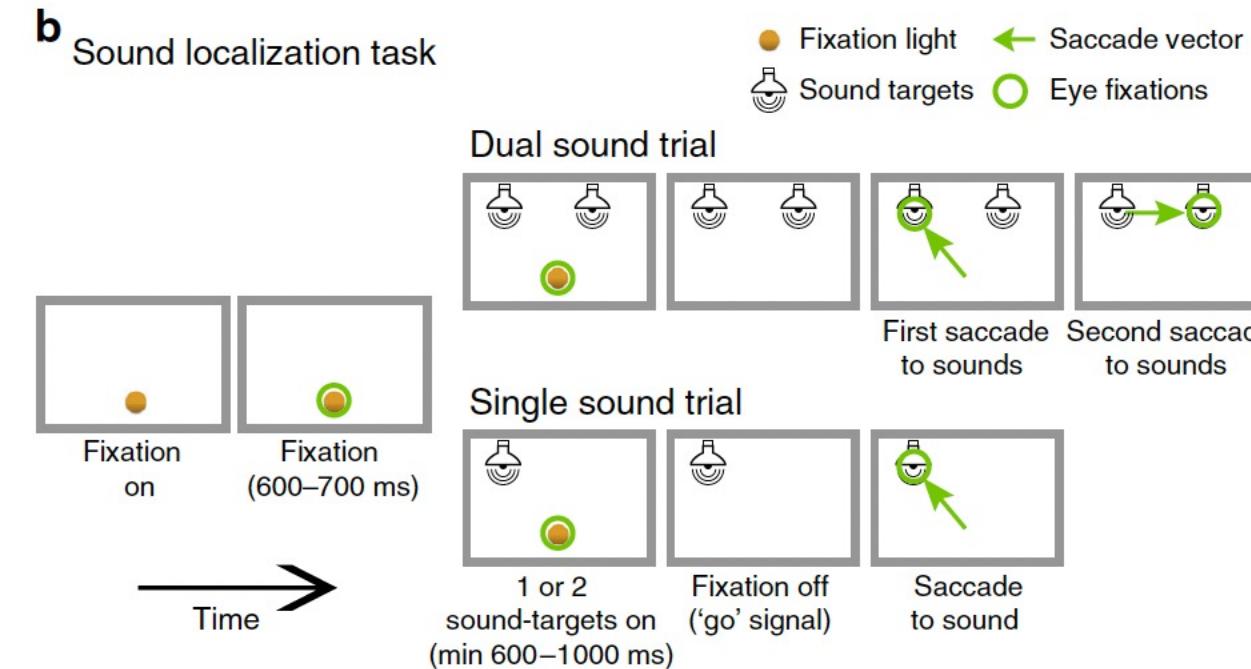
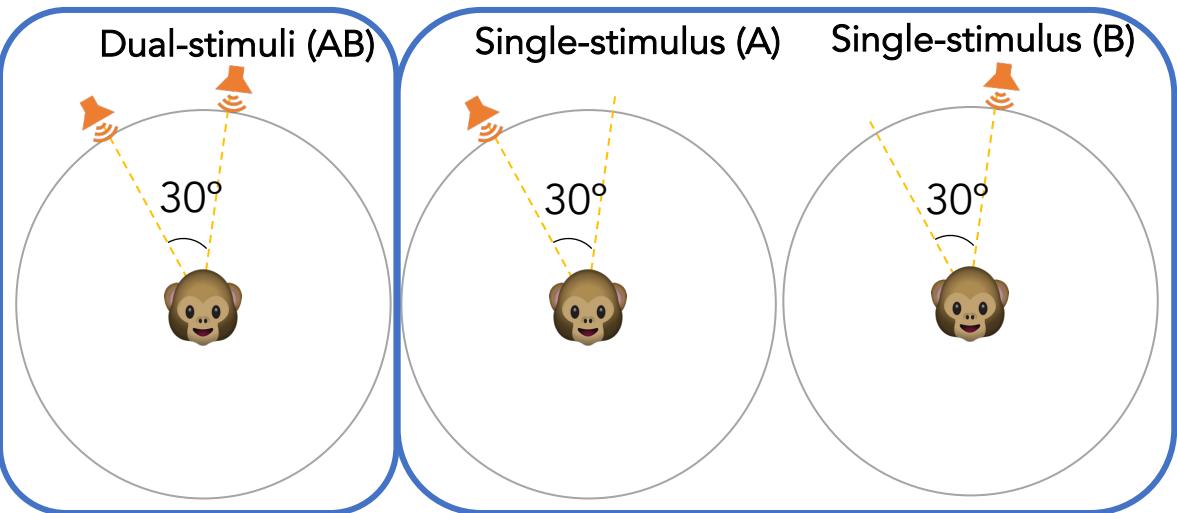
Performance Assessment

Jeffreys' prior
(w/ constant 1)



sampsizes: 20 (red dot), 30 (green diamond), 50 (blue circle)

Application in IC data

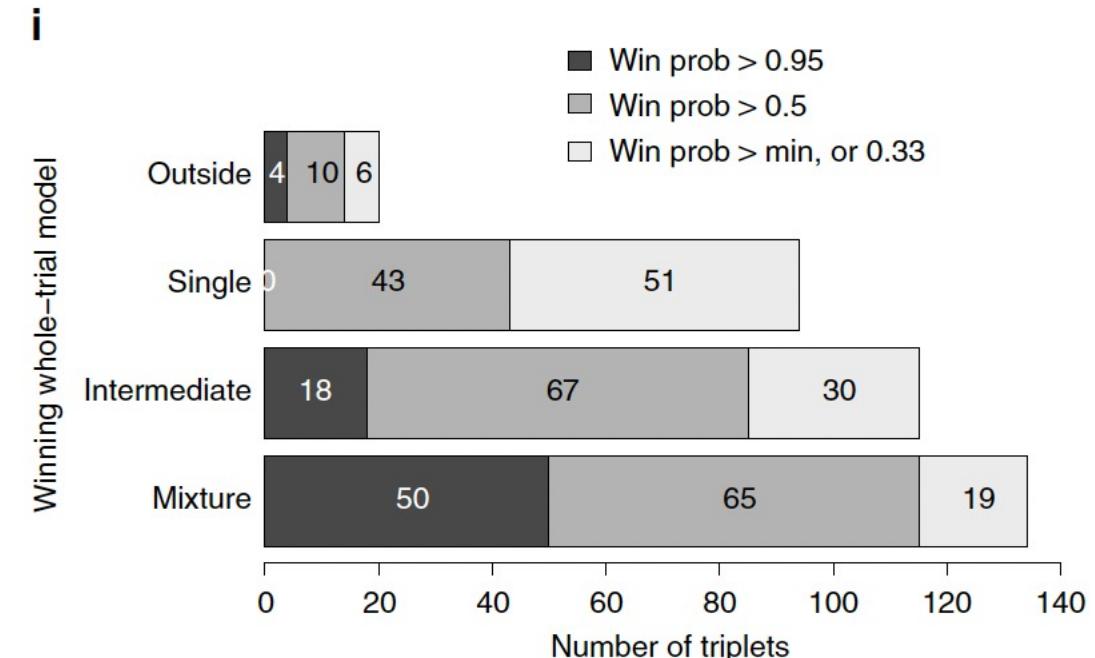


Experiment design (Caruso et al., 2018)

- Localization task: eye movements to sound (saccades)
- Single cell recording in Inferior colliculus (IC): accurate sound localization

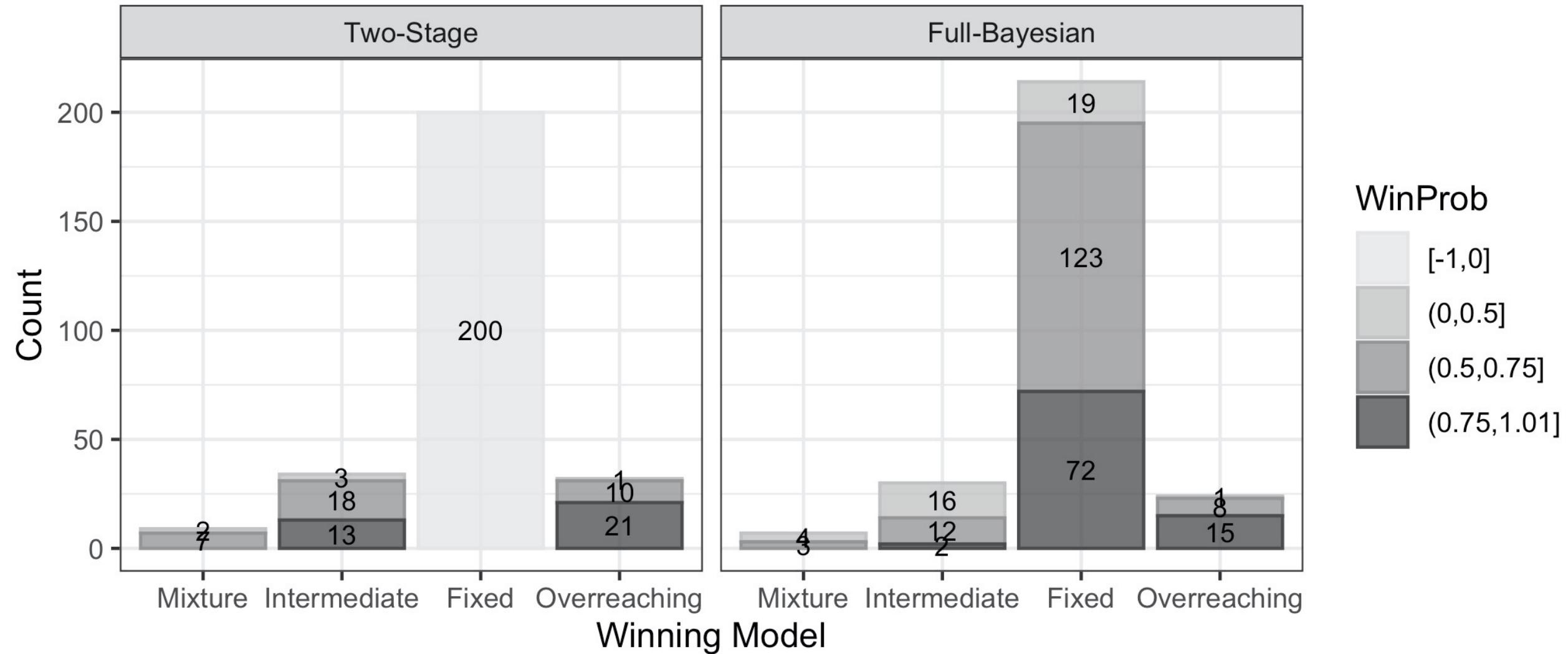
Results from Caruso et al. (2018)

- Chi-square goodness of fit test: 363 triplets

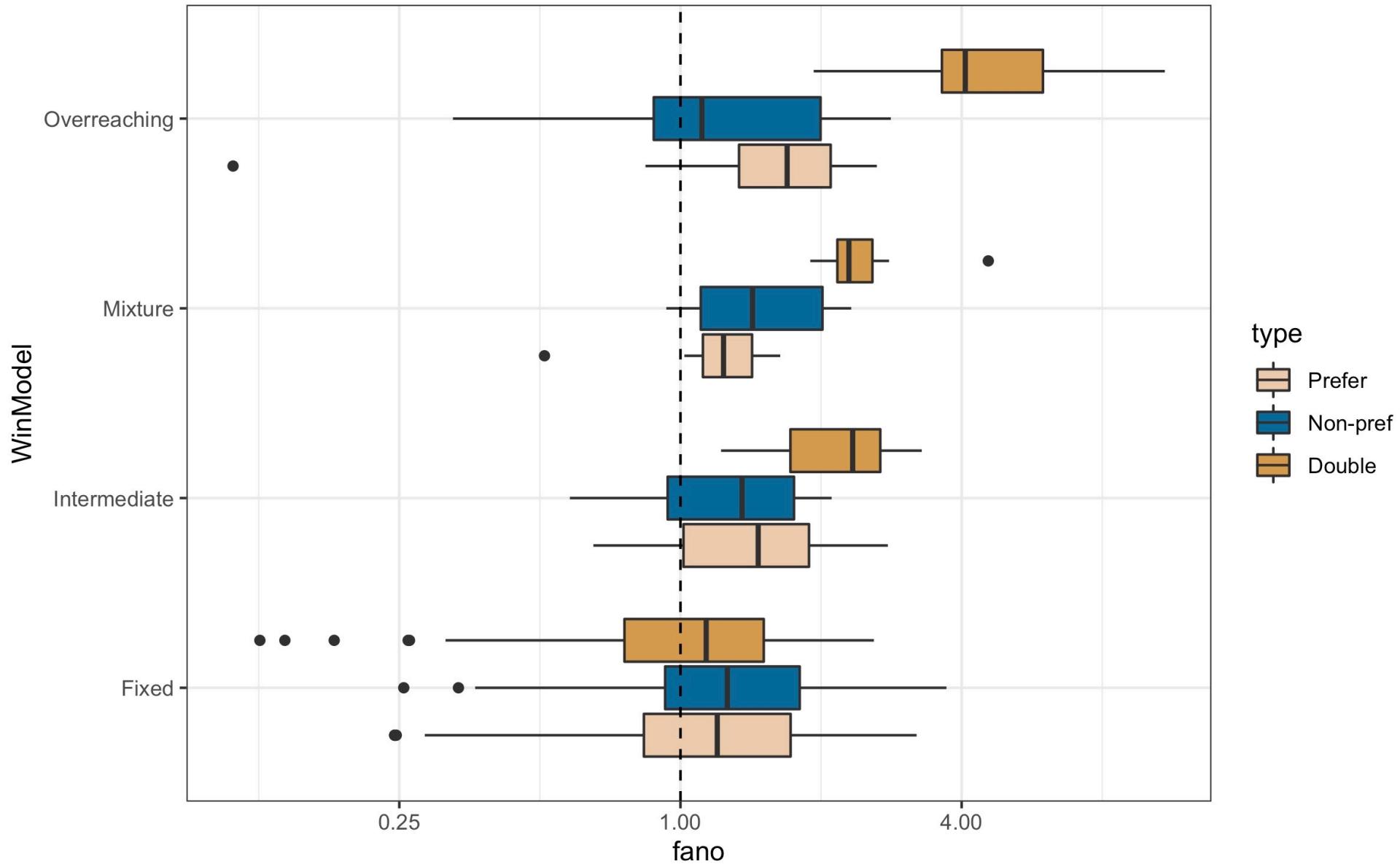


Results

Preprocessing: Poisson variance test (Brown and Zhao, 2002) ensure Poisson-like distribution for single stimulus

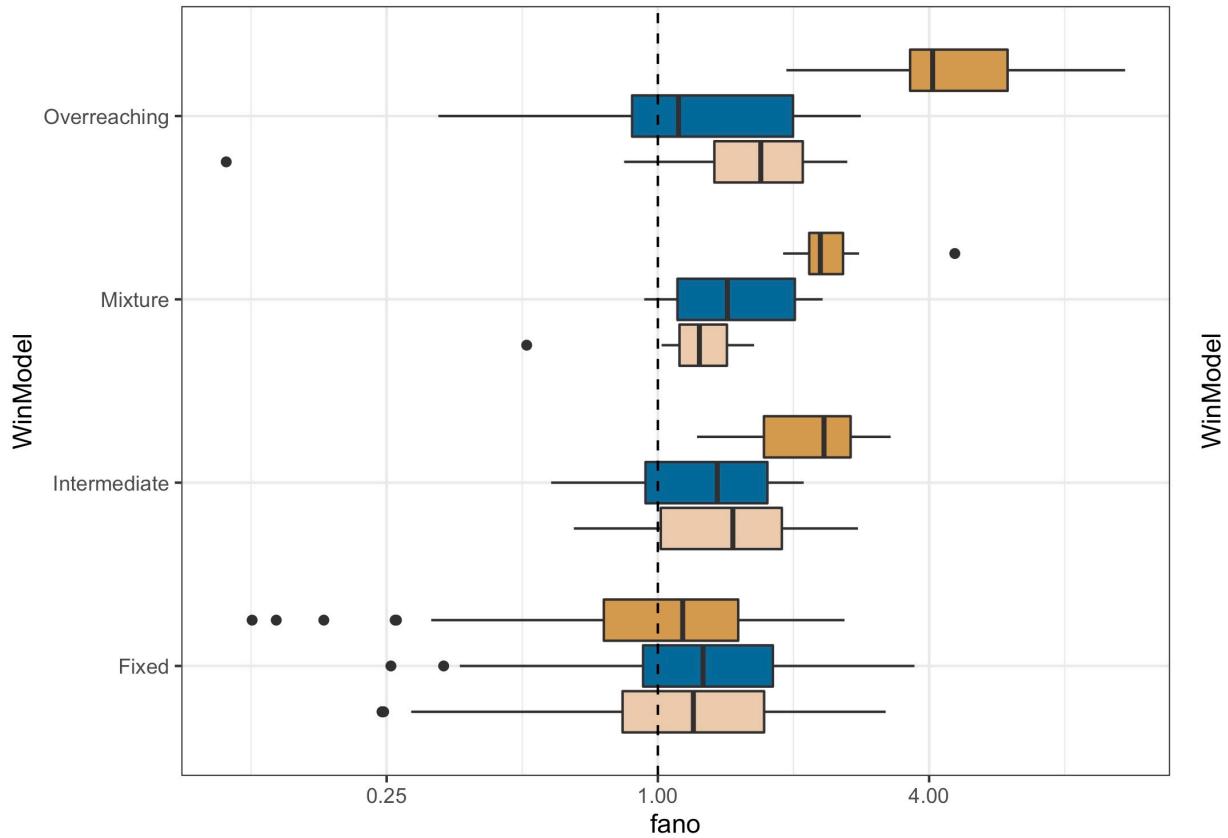


Results

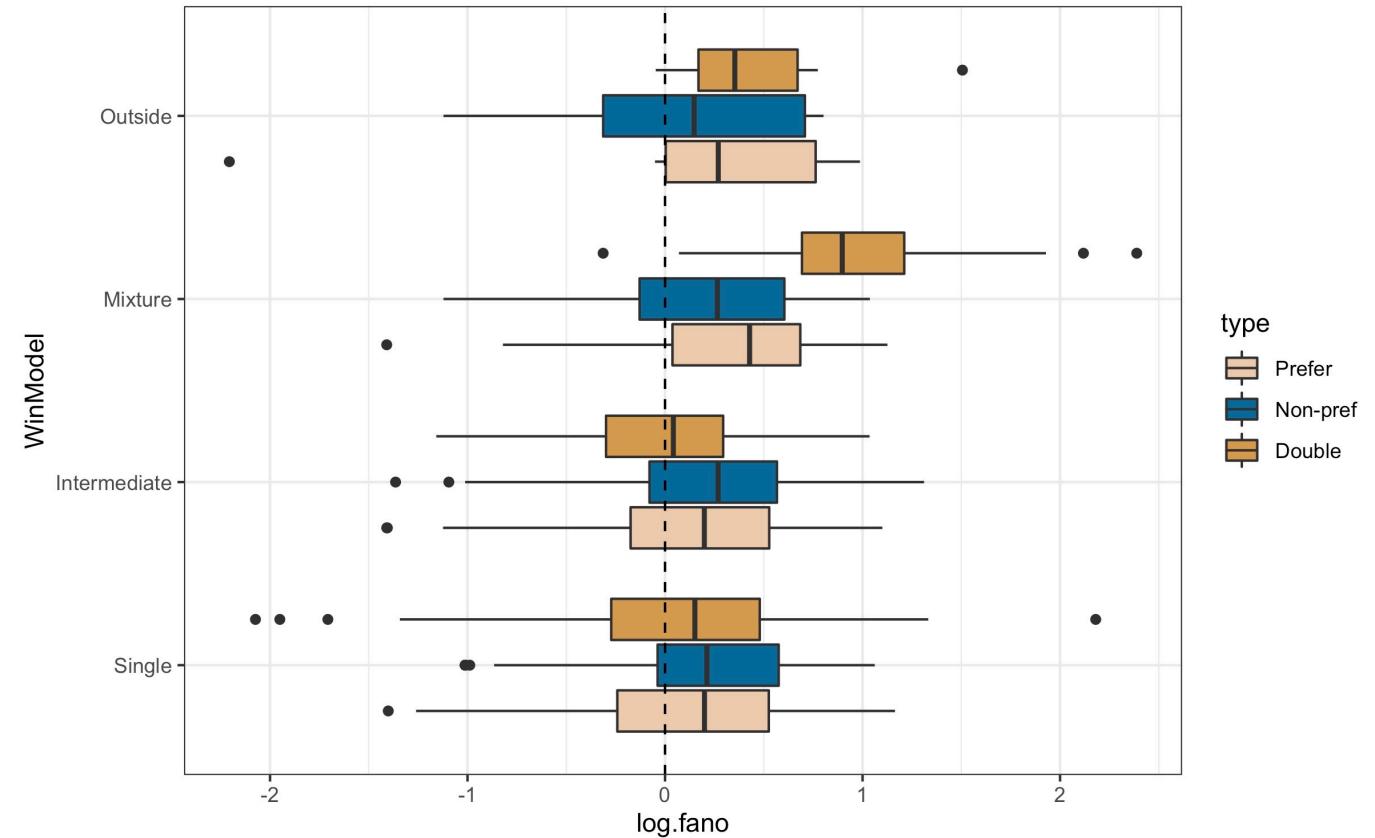


Results

New Method



Old Poisson





Thank you

Reference

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PRML Algorithm

Input: i.i.d observations Y_1, \dots, Y_n

Output: marginal likelihood $L_n(\theta) = \prod_{i=1}^n m_{i-1,\theta}(Y_i)$, gradient

$\nabla \log L_n(\theta) = \sum_{i=1}^n \nabla \log m_{i-1,\theta}(Y_i)$ and mixing density $f_{n,\theta}$

Initialize: $f_{0,\theta} \in \mathbb{F}$ of f (usually uniform); compute $\nabla f_{0,\theta}(u)$;

weights $w_1, \dots, w_n \in (0, 1)$ (usually $w_i = (i+1)^{-\gamma}$)

for $i = 1, \dots, n$ **do**

$$m_{i-1,\theta}(Y_i) = \int p(Y_i|\theta, u') f_{i-1,\theta}(u') d\mu(u') \quad (13)$$

$$f_{i,\theta}(u) = (1 - w_i) f_{i-1,\theta}(u) + w_i \frac{p(Y_i|\theta, u) f_{i-1,\theta}(u)}{m_{i-1,\theta}(Y_i)} \quad (14)$$

For $\nabla f_{i,\theta}(u)$:

$$G(\theta, u) = p(Y_i|\theta, u) \nabla f_{i-1,\theta}(u) + \nabla p(Y_i|\theta, u) f_{i-1,\theta}(u) \quad (15)$$

$$\nabla \log m_{i-1,\theta}(Y_i) = \frac{\int G(\theta, u) d\mu(u)}{m_{i-1,\theta}(Y_i)} \quad (16)$$

$$\nabla f_{i,\theta}(u) = (1 - w_i) \nabla f_{i-1,\theta}(u) + w_i \left\{ \frac{G(\theta, u) - p(Y_i|\theta, u) f_{i-1,\theta}(u) \nabla \log m_{i-1,\theta}(Y_i)}{m_{i-1,\theta}(Y_i)} \right\} \quad (17)$$

end