Probability Rules

Set Properties

 $S \cap (T \cup U) = (S \cap T) \cup (S \cup U)$ $S \cup (T \cap U) = (S \cup T) \cap (S \cup U)$

De Morgan's Law
$$(S \cap T)^{C} = S^{C} \cup T^{C}$$

$$S^{C} \cap T^{C} = (S \cup T)^{c}$$

Permutation - ordered arrangement of object (N!)

Combination - unordered arrangement of object $\binom{N}{n}$

Multinomial Probability

$$\frac{n!}{n_1! \, n_2! \dots n_r!} p_1^{\ n_1} p_2^{\ n_2} \dots p_r^{\ n_r}$$

Multiplication Rule

$$P_{X,Y,Z} = P_X(x)P_{X|Y}(x|y)P_{Z|X,Y}(z|x,y)$$

Marginal PDF

$$f_X(x) = \int f(x, y) dy$$

Conditional PDF

$$f_{X|Y} = \frac{f(x,y)}{f(y)}$$

Total Probability Theorem

Discrete
$$P_X(x) = \sum_{y} P_Y(y) P_{(X|Y)}(x|y)$$
Continuous

$$f_X(x) = \int_{-\infty}^{\infty} f_Y(y) f_{X|Y}(x|y) \, dy$$

Total Expectation Theorem

$$E[X|Y = y] = \sum_{x} x P_{(X|Y)}(x|y)$$
$$E[X] = \sum_{y} p_{Y}(y) E[X|Y = y]$$

$$E[X] = \sum_{x} p_{Y}(y)E[X|Y = y]$$
Continuous
$$E[X|Y = y] = \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx$$

$$E[X] = \int_{-\infty}^{\infty} f_{Y}(y)E[X|Y = y] dy$$

Law of Iterated Expectation

$$E\big[E[Y|X]\big] = E[Y]$$

Conditional Variance/ Law of Total Variance

$$Var(X) = E[Var(X|Y)] + Var(E[X|Y])$$

Mean and Variance of the Sum of a **Random Number of Random**

$$E[Y] = E[N]E[X]$$

$$Var(Y) = E[N]Var(X)$$

$$+ E[X]^{2}Var(N)$$

Transformation of Random Variable

$$f_y(y) = \frac{1}{|a|} f_x(\frac{y-b}{a})$$

$$f_{y}(y) = \frac{dx(y)}{dy} f_{X}(x(y))$$

Cookbook: For monotonic functions Find the appropriate limits for Y Find CDF $F_r(x(y))$ Differentiate with respect to y

Convolution

If X and Y are independent and Z =X + Y, then

$$f_z(z) = \int_{-\infty}^{\infty} f_x(z - y) f_y(y) dy$$

Linear Regression

Correlation Coefficient

$$\rho(X,Y) = \frac{Cov(X,Y)}{\sigma_X \sigma_Y}$$

Linear Least Mean squares

$$\widehat{\theta_L} = E[\theta] + \frac{Cov(X,Y)}{Var(X)}(X - E[X])$$

$$a = \frac{Cov(X,Y)}{Var(X)} = \rho \frac{\sigma_{\theta}}{\sigma_{x}} = \bar{Y} - \hat{b}\bar{X}$$

$$b = E[\theta] - \frac{Cov(X,Y)}{Var(X)} (E[X])$$

$$= \frac{\overline{XY} - \overline{X}\overline{Y}}{\overline{X^2} - \overline{X}^2}$$
MSE of the Linear Estimator
$$E\left[\left(\widehat{\Theta_L} - \theta\right)^2\right] = (1 - p^2)Var(\theta)$$

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Multivariate Linear Regression Model $Y_i = \beta_i^{TX} + \epsilon_i = X_1^T \beta^\star + \epsilon_1$

$$Y_i = \beta_i^{TX} + \epsilon_i = X_1^T \beta^* + \epsilon_1$$

 β^* -intercpet, Y_i response/dependent variable, X_i —independent/explanatory variable

Least Square Estimator

$$\hat{\beta} = argmin \sum_{i=1}^{n} (Y_i - X_i^T \beta)^2$$
$$= (X^T X)^{-1} X^T Y$$

X- Design Matrix with rank p

$$\begin{aligned} &\epsilon \sim Nu_n(0,\sigma^2I_n) \\ &\hat{\beta} \sim N_n(\beta^*,\sigma^2(X^TX)^{-1} \\ &quadratic\ risk = E\left[\left|\left|\hat{\beta} - \beta\right|\right|_2^2\right] \\ &= \sigma^2 trace((X^TX)^{-1}) \\ &prediction\ error = \sigma^2(n-p) \end{aligned}$$

$$\widehat{\sigma^2} = \frac{\left| \left| Y - X \, \widehat{\beta} \, \right| \right|^2}{n - p} = \frac{\sum_{i=1}^n \widehat{\epsilon^2}}{n - p}$$

CLT, Convergence, Inequalities, Delta Method

Central Limit Theorem

$$Z = \frac{\sqrt{(n) |X - \mu|}}{\sigma}$$

Covariance Matrix
$$\Sigma = \begin{array}{cc} Cov(X,X) & Cov(X,Y) \\ Cov(Y,X) & Cov(Y,Y) \end{array}$$

$$Cov(AX + B) = ACov(X)A^{T}$$

Multivariate CLT

$$\begin{array}{l} \sqrt{n}(\overline{X_n} - \mu) \xrightarrow{distribution} \mathrm{N}(0, \Sigma) \\ \sqrt{n}\Sigma^{-\frac{1}{2}}(\overline{X_n} - \mu) \xrightarrow{distribution} \mathrm{N}(0, \mathrm{I_d}) \end{array}$$

Markov Inequality
$$P(X \ge a) \le \frac{E[X]}{a}$$

Chebyshev Inequality

$$P(|X - \mu| \ge c) \le \frac{\sigma^2}{c^2}$$
Chernoff Bound

$$P(|\mu_n - \mu| \ge a) \le e^{-nh(a)}$$

Hoeffding's Inequality
$$P(|\overline{X_n} - \mu| \ge \epsilon) \le 2e^{\frac{2n\epsilon^2}{(b-a)^2}}$$

Jensen's Inequality

$$E[f(x)] \geq f(E[x])$$

Convergence

Almost Surely

$$P\left(w:T_{n(w)\xrightarrow[n\to\infty]{T(w)}}\right)=1$$

In Probability
$$P(|T_n-T| \geq \ \epsilon) \underset{n \to \infty}{\longrightarrow} 0$$

In Distribution
$$E[f(T_n)] \underset{n \to \infty}{\longrightarrow} E[f(t)]$$

Pointwise Convergence

$$\lim_{n\to\infty}g_n(x)=g(x)$$

Uniform Convergence

$$\sup|g_n(x) - g(x)| = 0$$

Delta Method

If estimator Z_n is asymptotically normal and the function g is differentiable everywhere, then

$$\sqrt{n}(g(Z_n)) \xrightarrow{distribution} N(0, g'(\theta)^2 \sigma^2)$$

Multivariate Delta Method
$$\sqrt{n} \big(g(T_n) - g(\theta) \big) \xrightarrow{distribution} N(0, \nabla g(\theta)^T \Sigma \nabla g(\theta)$$

Estimator properties

Asymptotic Normality
$$\sqrt{\{n\}} (\widehat{\theta_{-}} n - \theta) \xrightarrow{distribution} N(0, \sigma^2)$$

$$bias(\widehat{\theta_n}) = E[\widehat{\theta_n}] - \theta$$

Efficiency

The smaller the variance the more efficient is the estimator

Consistency

As $n \to \infty$ the estimator goes towards a single point

Quadratic Risk

Quadratic Risk
$$R\left(\widehat{\theta_n}\right) = E\left[\left|\widehat{\theta_n} - \theta\right|^2\right] \\ = Var\left(\widehat{\theta_n}\right) \\ + bias\left(\widehat{\theta_n}\right)^2$$
 Total Variation

Total Variation

$$TV(P_{\theta},P_{\theta'}) = \frac{1}{2}\sum |P_{\theta}(X) - P_{\theta'}(X)|$$
 same for continuous but with the integral

Kullback-Leibler Divergence(KL)

$$KL(P,Q) = \sum P(X) \ln(\frac{p(X)}{q(X)})$$

Hypothesis Tests

Fisher Information

$$I(\theta) = -E[H(l(\theta))] = Var(l'(\theta))$$
$$= \frac{1}{Var(\theta)}$$

Wald's Test

$$\sqrt{n}(I(\theta_0))^{\frac{1}{2}}(\widehat{\theta_n^{MLE}} - \theta_0) \stackrel{d}{\rightarrow} N_d(0, I_d)$$

$$T_n = \frac{n(\hat{\theta}^{MLE} - \theta_0)^2}{Var(\hat{\theta})} \sim \chi_d^2$$

Multiple Parameters

$$n(\widehat{\theta_n^{MLE}} - \theta_0)^T I(\theta_0)(\widehat{\theta_n^{MLE}} - \theta_0) \xrightarrow{d} \chi_d^2$$

Likelihood Ratio Test

$$T_n = 2(ln(\widehat{ heta_n^{MLE}}) - ln(\widehat{ heta_n^c}))$$
 $T_n o \chi_{d-r}^2$, by Wilk's Theorem

Chi Squared Test

$$n\sum_{j=1}^k \frac{\left(\widehat{p}_j - p_j^0\right)^2}{p_j^0} \stackrel{d}{\to} \chi_{k-1}^2$$

Where k is the number of categories

Chi Squared Test for Distribution

$$n\sum_{i=0}^{k} \frac{\left(\frac{N_{j}}{n} - f_{\widehat{\theta}}(j)\right)^{2}}{f_{\widehat{\theta}}(j)} \xrightarrow{d} \chi_{k-d-1}^{2}$$

Where d is the number of parameters of the distribution

One Sided Two Sample Test

$$\frac{\{\overline{X_n} - \overline{Y_n} - (\mu_1 - \mu_2)\}}{\sqrt{\frac{\widehat{\sigma_X^2}}{n} + \frac{\widehat{\sigma_Y^2}}{m}}}$$

Welch-Satterthwaite Formula

Calculates the degrees of freedom for multiple sample test

$$N = \frac{\left(\frac{\widehat{\sigma_d^2}}{n} + \frac{\widehat{\sigma_c^2}}{m}\right)^2}{\frac{\widehat{\sigma_d^4}}{n^2(n-1)} + \frac{\widehat{\sigma_c^4}}{m^2(m-1)}} \ge \min(n, m)$$

Kolmogorov-Smirnov Test

$$\begin{split} T_n &= sup\sqrt{n}|F_n(t) - F^0(t)| \\ &= \sqrt{n} \max \left\{ \max \left(\left| \frac{i-1}{n} - F^0(X_i) \right| \right) \right\} \end{split}$$

$$d^{2}(F_{n}, F) = \int [F_{n}(t) - F(t)]^{2} dF(t)$$

Anderson-Darling
$$d^2(F_n,F) = \int \frac{[F_n(t) - F(t)]^2}{F(t) (1 - F(t))} dF(t)$$

Kolmogorov-Lilliefors test

$$\sup |F_n(t) - \Phi_{\widehat{\mu},\widehat{\sigma^2}}(t)|$$

Linear Regression Significance Test

$$\frac{\widehat{\beta_{j}}-\beta_{j}}{\sqrt{\widehat{\sigma^{2}}\gamma_{j}}} \sim t_{n-p},$$

where γ_j is the jth diagonal coefficient

$$T_n^j = \frac{\widehat{\beta}_j}{\sqrt{\widehat{\sigma}^2 \gamma_j}}$$

Bonferroni's Test

$$R_{s,\alpha} = \bigcup R_{j,\frac{\alpha}{k}}$$

So in order to test that a group of explanatory variables affect the response variable we have to make the level α much more stringent by dividing it by the number of explanatory variables

Required Sample Size

$$N = \frac{\left(\Phi^{-1}(1-\beta) + \Phi^{-1}\left(1-\frac{\alpha}{2}\right)\right)^2}{\frac{\tau^2}{\sigma^2}\gamma(1-\gamma)}$$

Where γ is the proportion of people assigned to treatment and τ is a chosen parameter

Confidence Intervals

Methods

Conservative bound

We use the biggest variance

Quadratic Equation

$$(\theta - \overline{R_n})^2 \le \frac{q_{\alpha/2}^2 \, \sigma^2}{n}$$

Just plug in estimated mean and asymptotic variance into Gaussian CI

CI Solve Method

$$\left| \hat{\theta} - \theta \right| \le \frac{q_{\alpha/2}\sigma}{\sqrt{n}}$$

Gaussian CI

$$\widehat{\theta_n} \mp \frac{Z_{value}\sigma}{\sqrt{n}}$$

$$\widehat{\theta_n} \mp \frac{Z_{value}\lambda}{\sqrt{n}}$$

$$\Pi_{\psi} = \inf \left(1 - \mathsf{B}_{\psi}(\theta) \right)$$

Inference

	H_0 True	H_1 True
Reject	Type 1	
	Error (α)	
	(False	
	Positive)	
Fail to		Type 2
Reject		Error (β)
		(False
		Negative)

Bayesian Statistics

Bayes Formula for Pdfs

$$\begin{split} & \prod (\theta | X_1, \dots, X_n) \\ &= \frac{L(X_1, \dots, X_n | \theta) \prod(\theta)}{\int L(X_1, \dots, X_n | \theta) \prod(\theta) \, d\theta} \end{split}$$

Conjugate Prior happens when prior and posterior distributions are the same but with different parameters.

Improper prior on θ is a measurable nonnegative function $\Pi()$ defined on θ that is not integrable

Jeffrey's Prior

$$\Pi_{j}(\theta) \propto \sqrt{\det(I(\theta))}$$

$$iff \eta = \phi(\theta)$$

$$I(\eta) = I(\phi^{-1}(\eta)) * \frac{d\theta}{d\eta} \text{ and } \Pi_{j}(\eta)$$

$$\propto \sqrt{\det(\hat{I}(\eta))}$$

Conditional Quantile/ Median

$$Median = \int_{-\infty}^{m(x)} h(y|x)dy = \frac{1}{2}$$

Bayes Confidence Region

We need to take the inverse of the posterior likelihood function with a chosen q_{α}

Distributions, Likelihoods

Exponential Family

$$f_{\theta}(y) = exp\left(\sum_{i=1}^{k} \eta_{i}(\theta) T_{i}(y) - B(\theta)\right) h(y)$$

Canonical Exponential Family

$$f_{\theta}(y) = \exp\left(\frac{y\theta - b(\theta)}{\phi} + C(y, \phi)\right)$$
$$E[x] = b'(\theta)$$
$$Var(x) = \phi b''(\theta)$$

 ϕ is called a dispersion parameter

Link Function

If
$$g(m(x)) = X^T \beta + \epsilon$$

Then g is the link function

Canonical Link

$$g(m) = (b')^{-1}(m)$$

d-dimensional Gaussian distribution

$$\frac{\exp\left(-\frac{1}{2}(X-\mu)^T\Sigma^{-1}(X-\mu)\right)}{\sqrt{(2\Pi)^d\det(\Sigma)}}$$

First Order Statistic

$$f_{y^1} = n(1 - F_x(y))^{n-1} f_x(y)$$

Nth Order Statistic

$$f_{y^n} = n(F_x(y))^{n-1} f_x(y)$$

LOGIT

$$\log\left(-\frac{m(x)}{1-m(x)}\right) = X^T \beta$$

PROBIT

$$\Phi^{-1}\big(m(x)\big) = X^T\beta$$

Cauchy Distribution

$$\frac{1}{\Pi(1+(X-\mu)^2)}$$

Laplace Distribution

$$\frac{1}{2}e^{-|X-\mu|}$$

Pascal Distribution

Pascal distribution
$$P_{Y_k}(t) = {t-1 \choose k-1} p^k (1-p)^{t-k}$$

$$E[Y_k] = \frac{k}{p}$$

$$Var(Y_k) = \frac{k(1-p)}{p^2}$$

Huber's Loss

$$\begin{cases} \frac{a^2}{2}, & \text{if } |a| \le \delta \\ \delta\left(|a| - \frac{\delta}{2}\right), & \text{otherwise} \end{cases}$$

Kurtosis

$$E[\left(\frac{X-E[X]}{\sigma_d(X)}\right)^4-3]$$

It measures how spread out the tails are, if kurtosis is positive we know that the tails are heavier than Gaussian and vice versa

Likelihoods

Uniform([0,b])

$$\frac{1}{h^n}$$

Categorical Likelihood

 $L_n(X_1,\dots,X_n;P_1,\dots,P_k) = P_1^{N_1}\dots P_k^{N_k}$ It's a generalised Bernoulli model

Gauccia

$$\frac{1}{\sigma^n (2\Pi)^{\frac{n}{2}}} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (X_i - \mu)^2\right)$$

Remouli

$$p^{\wedge}(\sum_{i=1}^{n}X_{i})(1-p)^{\wedge}(n-\sum_{i=1}^{n}X_{i})$$

Poisso

$$\frac{e^{-n\lambda}\lambda^{\wedge}(\sum_{i=1}^{n}X_{i})}{X_{1!}\dots X_{n!}}$$

Exponential

$$\lambda^n e^{-\lambda \sum_{i=1}^n X_i}$$

Theorems

Slutsky's Theorem

$$T_n + U_n \xrightarrow{distribution} T + U$$

Same for other operations

Continuous Mapping Theorem

If there is convergence in probability for T, then

$$f(T_n) \xrightarrow{distribution} f(T)$$

Cochran's Theorem

For X_-1,\ldots,X_-n random variables with $X_n \sim \mathrm{N}(\mu,\sigma^2)$, if S_n is the sample variance, then $\frac{nS_n}{\sigma^2} \sim \chi_{n-1}^2$ and variance is $S_n = \frac{\sum_{i=1}^n ((X_i - \bar{X})^2)}{n-1}$, multivariate variance is $S_n = \frac{\sum_{i=1}^n X_i^2}{n} - \frac{\sum_{i=1}^n X_i^2}{n}$

Glivenko-Cantelli Theorem

$$\sup |F_n(t) - F(t)| \stackrel{a.s.}{\longrightarrow} 0,$$

where $F_n(t)$ is the empirical CDF

Cox Proportional Hazard Model

$$f(t) = 1 - \exp\left(-\int_0^t h(u) e^{\beta^T X} du\right)$$

Data Analysis

Belmont Principles

- Respect for Persons
- 2. Beneficence
- 3. Justice

Histogram

Trade off in the binsize of histogram is between variance and bias (detail and noise)