

- 6.1 1. a. I use the \bar{x} to do this: $\hat{\mu} = \bar{x} = \frac{\sum x_i}{27} = 8.14$ ✓
 b. We have $\hat{\mu}' = \frac{219.8 - 8.1 - 8.2 - 8.1 - 7.8 - 7.9 - 7.8}{27 - 6} = 8.16$, using sample mean ✓
 c. We use the sample variance: $\hat{\sigma}^2 = \frac{\sum x_i^2}{27} - 8.14^2 = 2.66 \Rightarrow \sigma = \sqrt{2.66} = 1.63$ ✓
 d. Since there are 4 out of 27 that exceeds 10 mpa, so we have $\hat{p} = \frac{4}{27}$. ✓
 e. We use the sample mean and the sample variance: $C = \frac{\sigma}{\mu} = \frac{1.63}{8.14} = 0.2$ ✓

8. a. We used the sample to estimate the population: $\hat{p} = \frac{80 - 12}{80} = 0.85$

b. $P(\text{system works}) = p^2 = 0.7225$ ✓

9. a. Since for Poisson distribution, $E(X) = \mu$, however,

$$E(\bar{x}) = E\left(\frac{1}{n} \sum_{i=1}^n x_i\right) = \frac{1}{n} \sum_{i=1}^n E(x_i) = \frac{1}{n} \cdot n \cdot E(X) = \mu$$

Therefore, \bar{x} is an unbiased estimator of μ , and $E(\bar{x}) = \mu = 2.11$ ✓

b. $V(\bar{x}) = V\left(\frac{1}{n} \sum_{i=1}^n x_i\right) = \frac{1}{n^2} \sum_{i=1}^n V(x_i) = \frac{1}{n^2} \cdot n \cdot V(X) = \frac{\sigma^2}{n} = \frac{\mu}{n} = 0.0141 \Rightarrow \hat{\sigma} = 0.119$ ✓

13. Since $E(\hat{\theta}) = E(3\bar{x}) = 3E(\bar{x}) = 3E\left(\frac{1}{n} \sum_{i=1}^n x_i\right) = 3 \cdot \frac{1}{n} \sum_{i=1}^n E(x_i) = \frac{3}{n} \cdot n \cdot \mu = 3\mu$

and $\mu = E(X) = \int_{-1}^1 (0.5x + 0.5\theta x^2) dx = \left(\frac{1}{4}x^2 + \frac{1}{6}\theta x^3\right) \Big|_{-1}^1 = \frac{1}{4} + \frac{1}{6}\theta - \frac{1}{4} + \frac{1}{6}\theta = \frac{1}{3}\theta$

So $E(\hat{\theta}) = 3\mu = \theta$, therefore, $\hat{\theta} = 3\bar{x}$ is an unbiased estimator of θ . ✓

6.2 20. a. Since $f(p) = \binom{n}{x} p^x (1-p)^{n-x}$, so $\ln f(p) = \ln \binom{n}{x} + x \ln p + (n-x) \ln(1-p)$
 $\frac{d \ln f(p)}{dp} = \frac{x}{p} - \frac{n-x}{1-p}$, let $\frac{x}{p} = \frac{n-x}{1-p} \Rightarrow p = \frac{x}{n}$, take $x=3, n=20$ we know $\hat{p} = 0.15$.

b. Since $E(\hat{p}) = E\left(\frac{x}{n}\right) = \frac{1}{n} \sum_{i=1}^n \frac{x_i}{n} = \frac{1}{n} \cdot n \cdot \frac{1}{n} \cdot x = \frac{x}{n} = p$, so \hat{p} is unbiased. ✓

c. That is $h(\hat{p}) = (1-p)^5 = 0.85^5 = 0.444$ ✓

21. a. $E(X^1) = \frac{1}{n} \sum_{i=1}^n x_i$, $E(X^2) = \frac{1}{n} \sum_{i=1}^n x_i^2$, and $V(X) = E(X^2) - [E(X)]^2$, so we have
 $\frac{\Gamma^2(1+\frac{1}{\alpha})}{\Gamma(1+\frac{2}{\alpha})} = \frac{1}{n} \cdot \frac{(\sum x_i)^2}{\sum x_i^2}$, $\beta = \frac{\sum x_i}{n \Gamma(1+\frac{1}{\alpha})}$, in line with the problem description.

b. $\frac{\Gamma^2(1+\frac{1}{\alpha})}{\Gamma(1+\frac{2}{\alpha})} = 0.95 \Rightarrow \hat{\alpha} = 5$, then $\hat{\beta} = \frac{23}{\Gamma(1.2)}$. ✓

Since $f(x; \lambda, \theta) = \lambda e^{-\lambda(x-\theta)}$, thus the joint pdf $f(x_1, x_2, \dots, x_n; \lambda, \theta) = \lambda^n e^{-\lambda \sum (x_i - \theta)}$, then $\ln f(x_1, \dots, x_n; \lambda, \theta) = n \ln \lambda + (-\lambda) \sum (x_i - \theta)$
We have $\frac{\partial \ln f(x_1, \dots, x_n; \lambda, \theta)}{\partial \lambda} = \frac{n}{\lambda} - \sum (x_i - \theta) \Rightarrow \hat{\lambda} = \frac{n}{\sum (x_i - \hat{\theta})}$

and $\frac{\partial \ln f(x_1, \dots, x_n; \lambda, \theta)}{\partial \theta} = -n\lambda \Rightarrow \hat{\theta} = -\frac{n^2}{\sum (x_i - \hat{\theta})}$ ✓

b. Then $\hat{\theta} = 7.007, \hat{\lambda} = -0.7007$ ✗

32. a. $F(y) = P(Y \leq y) = P(\max(x_i) \leq y) = P(x_1 \leq y) P(x_2 \leq y) \dots P(x_n \leq y)$

$= \int_0^y \frac{1}{\theta} dy \int_0^y \frac{1}{\theta} dy \dots$
 $M = (X) = \left(\frac{y}{\theta}\right)^n$

$f_Y(Y) = F'(Y) = \frac{n y^{n-1}}{\theta^n}, 0 \leq y \leq \theta$, otherwise it's 0. ✓

b. $E(Y) = \int_0^\theta \frac{n y^n}{\theta^n} dy = \frac{n}{n+1} \cdot \frac{y^{n+1}}{\theta^n} \Big|_0^\theta = \frac{n\theta}{n+1}$, so $E(\hat{\theta}) \neq E(Y)$,

but $E\left(\frac{n+1}{n}\right) = E(Y)$, so $\frac{n+1}{n} \max(x_i)$ is unbiased. ✓