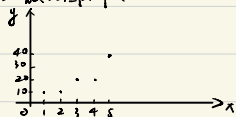


2.14

解: (1) 散点图如下:



(2) 由图像:

大致有线性关系

(3) 由题意知:  $L_{xx} = \sum (x_i - \bar{x})^2 = 10$ ,  $L_{yy} = \sum (y_i - \bar{y})^2 = 600$

$$L_{xy} = \sum (x_i - \bar{x})(y_i - \bar{y}) = 70$$

令回归方程为  $\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i$ , 其中

$$\hat{\beta}_1 = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sum (x_i - \bar{x})^2} = 7$$

$$\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x} = -1$$

$$\Rightarrow \hat{y}_i = -1 + 7x_i$$

$$(4) \sum e_i^2 = \sum (y_i - \hat{y}_i)^2 = \sum (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2$$

$$= (10-6)^2 + (13-13)^2 + (20-20)^2 + (27-27)^2 + (34-34)^2$$

$$= 110$$

$$\hat{\sigma}^2 = \frac{1}{n-2} \sum e_i^2$$

$$= \frac{110}{3}$$

$$\Rightarrow \hat{\sigma} = 6.055$$

$$(5) t_{0.05}(3) = 3.182$$

$$\textcircled{1} \hat{\beta}_1 \sim N(\beta_1, \frac{\hat{\sigma}^2}{L_{xx}})$$

$$t_1 = \frac{\hat{\beta}_1 - \beta_1}{\hat{\sigma} / \sqrt{L_{xx}}} = \frac{(\hat{\beta}_1 - \beta_1) \sqrt{L_{xx}}}{\hat{\sigma}} \sim t(3)$$

$$\Rightarrow P(|t| < t_{0.05}(3)) = 1 - \alpha$$

$$\Rightarrow P(\hat{\beta}_1 - t_{0.05}(3) \frac{\hat{\sigma}}{\sqrt{L_{xx}}} < \beta_1 < \hat{\beta}_1 + t_{0.05}(3) \frac{\hat{\sigma}}{\sqrt{L_{xx}}}) = 1 - \alpha$$

$\Rightarrow \beta_1$  的置信度 95% 的置信区间为:

$$(\hat{\beta}_1 - t_{0.05}(3) \frac{\hat{\sigma}}{\sqrt{L_{xx}}}, \hat{\beta}_1 + t_{0.05}(3) \frac{\hat{\sigma}}{\sqrt{L_{xx}}})$$

$$\text{即: } (0.91, 13.09)$$

$$\textcircled{2} \hat{\beta}_0 \sim N(\beta_0, (\frac{1}{n} + \frac{(\bar{x})^2}{L_{xx}}) \hat{\sigma}^2)$$

$$\text{即: } \hat{\beta}_0 \sim N(\beta_0, 1.1 \hat{\sigma}^2)$$

同理可求:

$\beta_0$  的置信度 95% 的置信区间为:

$$(\hat{\beta}_0 - t_{0.05}(3) \sqrt{1.1} \hat{\sigma}, \hat{\beta}_0 + t_{0.05}(3) \sqrt{1.1} \hat{\sigma})$$

$$\text{即: } (-21.21, 17.21)$$

$$(6) r^2 = \frac{SSR}{SST} = \frac{ESS}{TSS} = \frac{L_{xy}^2}{L_{xx} L_{yy}} = \frac{70^2}{10 \times 600} = 0.817$$

$$(7) SSR = ESS = \sum (\hat{y}_i - \bar{y})^2 = (6-20)^2 + (13-20)^2 + 0^2 + (27-20)^2 + (34-20)^2 = 490$$

$$SSE = RSS = \sum (y_i - \hat{y}_i)^2 = 110$$

$$SST = TSS = \sum (y_i - \bar{y})^2 = L_{yy} = 600$$

$$SSR/1 = 490$$

$$SSE/(n-2) = 36.67$$

$$F = \frac{SSR/1}{SSE/(n-2)} = 13.36$$

$$P(F > F_{\alpha}(1, 3)) = \alpha \xrightarrow{\text{查表}} 0.025 < \alpha < 0.05$$

故方差分析如表:

	df	SS	MS	F	p-value
回归	1	490	490	13.36	0.025~0.05
残差	3	110	36.67		
总计	4	600			

$$(8) H_0: \beta_1 = 0, H_1: \beta_1 \neq 0$$

$$\Rightarrow \hat{\beta}_1 \sim N(0, \frac{\hat{\sigma}^2}{L_{xx}})$$

$$\text{令 } t = \frac{\hat{\beta}_1}{\hat{\sigma} / \sqrt{L_{xx}}} = \frac{\hat{\beta}_1 \sqrt{L_{xx}}}{\hat{\sigma}} \sim t(3)$$

$$P(|t| < t_{0.05}(3))$$

$$= P(-t_{0.05}(3) \cdot \frac{\hat{\sigma}}{\sqrt{L_{xx}}} < \hat{\beta}_1 < t_{0.05}(3) \cdot \frac{\hat{\sigma}}{\sqrt{L_{xx}}})$$

$$= P(-6.09 < \hat{\beta}_1 < 6.09)$$

$$= 1 - \alpha = 0.95$$

由  $\hat{\beta}_1 = 7$

故拒绝  $H_0$ , 不能说明  $\beta_1$  在 95% 的置信水平上为 0

$\Rightarrow \beta_1$  显著不为 0

$$(9) r = \frac{L_{xy}}{\sqrt{L_{xx} L_{yy}}} = 0.904$$

$$t = \frac{\sqrt{n-2} r}{\sqrt{1-r^2}} = 3.656$$

$$t_{0.05}(3) = 3.182$$

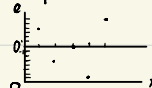
$$\Rightarrow |t| > t_{0.05}(3)$$

$\Rightarrow$  可以认为  $y$  与  $x$  的简单回归系数显著不为 0

(10) 由题:  $e_1 = 10 - 6 = 4$ ,  $e_2 = 13 - 13 = 0$ ,  $e_3 = 20 - 20 = 0$ ,

$$e_4 = 27 - 27 = 0, e_5 = 34 - 34 = 0$$

残差图如下:



残差在  $e=0$  附近随机变化且无明显同线性异常

但由于样本量太小, 无法得出准确判断

$$(11) \hat{y}_0 = -1 + 7 \times 4.2 = 28.4$$

$\Rightarrow$  销售收入将达到 28.4 万元

$$\text{由 } \hat{y}_0 \sim N(\beta_0 + \beta_1 x_0, (\frac{1}{n} + \frac{(x_0 - \bar{x})^2}{L_{xx}}) \hat{\sigma}^2) \text{ 可知:}$$

$$\hat{y}_0 - \hat{\beta}_0 \sim N(0, (1 + \frac{1}{n} + \frac{(x_0 - \bar{x})^2}{L_{xx}}) \hat{\sigma}^2)$$

$$\Rightarrow \text{令 } t = \frac{\hat{y}_0 - \hat{\beta}_0}{\hat{\sigma} \sqrt{1 + \frac{1}{n} + \frac{(x_0 - \bar{x})^2}{L_{xx}}}} \sim t(3)$$

$$\text{故 } P(|t| < t_{0.05}(3)) = 0.95$$

$\Rightarrow \hat{y}_0$  的置信概率 95% 的置信区间为

$$(\hat{y}_0 - t_{0.05}(3) \hat{\sigma} \sqrt{1 + \frac{1}{n} + \frac{(x_0 - \bar{x})^2}{L_{xx}}}, \hat{y}_0 + t_{0.05}(3) \hat{\sigma} \sqrt{1 + \frac{1}{n} + \frac{(x_0 - \bar{x})^2}{L_{xx}}})$$

$$\text{即: } (24.71, 32.01)$$

Ch3

3.3

$$\text{证明: } \hat{\sigma}^2 = \frac{1}{n-p-1} SSE = \frac{1}{n-p-1} e'e = \frac{1}{n-p-1} \sum e_i^2$$

$$\hat{Y} = X\hat{\beta} = X(X'X)^{-1}X'Y = HY$$

$$e = Y - \hat{Y} = (I - H)Y$$

$$D(e) = Cov(e, e) = Cov((I - H)Y, (I - H)Y) = \sigma^2(I - H)$$

$$\Rightarrow E(\sum e_i^2) = \sum D(e_i) = \sum (1 - h_{ii}) \sigma^2$$

$$\text{此时, } \text{tr}(Y) = \text{tr}(X(X'X)^{-1}X')$$

$$= \text{tr}((X'X)^{-1}X'X)$$

$$(\text{由 } X'X = (n-p+1)I) = \text{tr}(I_{p+1})$$

$$= p+1$$

$$\Rightarrow E(\sum e_i^2) = \sum (1-h_{ii})\sigma^2$$

$$= n\sigma^2 - (p+1)\sigma^2$$

$$\Rightarrow E(\hat{\sigma}^2) = \frac{1}{n-p-1} \cdot (n-(p+1))\sigma^2$$

$$= \sigma^2$$

无偏估计。