

Nonparametric estimation of the continuous treatment effect with measurement error

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Background and focus of the paper

- Identifying and estimating the causal effect of a treatment or policy from observational studies is of great interest to economics, social science, and public health researchers.
- Early studies focused on whether an individual receives the treatment or not
- More recently, as opposed to such binary treatments, researchers have been investigating the causal effect of a continuously valued treatment, where the effect depends not only on the introduction of the treatment but also on the intensity
- However, all these methods require the treatment data to be measured without errors. This paper propose a broad class of novel and robust nonparametric estimators for the average dose–response function (ADRF) as in Ai et al. (2021)

Basic Setup

- This paper consider a continuously valued treatment in which the observed treatment variable is denoted by T with the probability density function $f_T(t)$ and support $\mathcal{T} \subset \mathbb{R}$.
- $Y^*(t)$ denote the potential outcome if one was treated at level t for $t \in \mathcal{T}$. In practice, each individual can only receive one treatment level T and we only observe the corresponding outcome $Y := Y^*(T)$
- $X \in \mathbb{R}^r$ be a vector of given covariates related to both T and $Y^*(t)$ for $t \in \mathcal{T}$
- Instead of observing T we observe S so that

$$S = T + U \tag{1}$$

assuming error U independent of T, X and $\{Y^*(t)\}_{t \in \mathcal{T}}$ and c.f ϕ_U is known

Basic set up

- Now we have some basic assumptions relevant in most of treatment effect literature.

(i) (*Unconfoundedness*) $T \perp \{Y^*(t)\}_{t \in \mathcal{T}} | X$

(ii) (*No Intereference*) for $i = 1(1)N$ outcome of individual i is not affected by treatment assigned to any other individual i.e $Y_i^*(T_i, \mathbf{T}_{(-i)}) = Y_i^*(T_i, \mathbf{T}'_{(-i)})$ for any $\mathbf{T}_{(-i)}, \mathbf{T}'_{(-i)}$, where $Y_i^*(T_i, \mathbf{T}_{(-i)})$ is the potential outcome of individual i given the treatment T_i and others with $\mathbf{T}_{(-i)} := (T_1, \dots, T_{i-1}, T_{i+1}, \dots, T_N)$, respectively.

(iii) (*Consistency*) $Y = Y^*(t)$ a.s. if $T = t$

(iv) (*Positivity*) $f_{T|X}(t|X) > 0$ a.s. for all $t \in \mathcal{T}$.

Unconditional ADRF

- As mentioned earlier the goal of paper is to estimate unconditional ADRF i.e $\mu(t) = \mathbb{E}\{Y^*(t)\}$ for fixed $t \in \mathbb{T}$ based on i.i.d $\{S_i, X_i, Y_i\}_{i=1}^N$
- Under following assumptions one can write $\mu(t)$ as

$$\begin{aligned}\mu(t) &= \mathbb{E}[\mathbb{E}\{Y^*(t)|X\}] \\ &= \mathbb{E}\{Y^*(t)|X, T = t\} \\ &= \mathbb{E}\{Y|X, T = t\} \\ &= \int_{\mathcal{X}} \int_{\mathcal{Y}} \frac{f_T(t)}{f_{T|X}(t|x)} y f_{Y|X,T}(y|x, t) f_{X|T}(x|t) dy dx \\ &= \mathbb{E}\{\pi_0(t, X)Y|T = t\}\end{aligned}\tag{2}$$

where,

$$\pi_0(t, X) = \frac{f_T(t)}{f_{T|X}(t|x)}\tag{3}$$

Unconditional ADRF

- **GOOD NEWS** If T is fully observable and $\pi_0(t, x)$ is known a well known consistent estimator for $\mu(t)$ is Nadaraya-Watson estimator i.e

$$\mu_{NW} = \frac{\sum_{i=1}^N \pi_0(t, X_i) Y_i L(\{h^{-1}(t - T_i)\})}{\sum_{i=1}^N L(\{h^{-1}(t - T_i)\})} \quad t \in \mathcal{T} \quad (4)$$

where L is univariate Kernel function $\int_{-\infty}^{\infty} L(x) dx = 1$, and h is bandwidth

- **BAD NEWS** Unfortunately we do not observe T but observe S and also $\pi_0(t, x)$ is unknown in practice.

De-Convolution Kernel Approach

- Clearly from Fourier's inversion theorem we have

$$f_T(t) = \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-iwt} \phi_T(w) dw = \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-iwt} \frac{\phi_S(w)}{\phi_U(w)} dw \quad (5)$$

This inspired Stefinski and Carroll[1990] to estimate $f_T(t)$ by $\hat{f}_{T,h}(t) = (Nh)^{-1} \sum_{i=1}^N L_U\{(t - S_i)/h\}$ where,

$$L_U(v) = \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-i w v} \frac{\phi_L(w)}{\phi_U(w/h)} dw \quad (6)$$

- Based on this idea Fan and Troung [1993] proposed a consistent errors-in-variable regression estimator replacing the $L\{(t - T_i)/h\}$ with $L_U\{(t - S_i)/h\}$ i.e

$$\tilde{\mu}(t) = \frac{\sum_{i=1}^N \pi_0(t, X_i) Y_i L_U\{h^{-1}(t - S_i)\}}{\sum_{i=1}^N L_U\{h^{-1}(t - S_i)\}} \quad (7)$$

- In fact we have,

$$\mathbb{E}[L_U(\{h^{-1}(t - S)\}) | T, X, Y] = L\{h^{-1}(t - T)/h\} \quad (8)$$

Estimation of $\pi_0(t, X)$

- **Bad News?** The straightforward estimate of π_0 by plug in is subject to sensitivity on low values of $f_{T|X}$ since small errors in estimating $f_{T|X}$ lead to large error. So we will treat π_0 as a whole
- if T_i are fully observable from Ai et al[2021]

$$\mathbb{E}\{\pi_0(t, X)u(X)v(T)\} = \mathbb{E}\{v(T)\}\mathbb{E}\{u(X)\} \quad (9)$$

holds for any integrable function $u(X)$ and $v(T)$

- **Bad News?** non parametric estimation of $\mathbb{E}\{v(T)\}$ from contaminated data $\{S_i\}_{i=1}^N$ is difficult if not impossible.
- **Solution** We will estimate the projection $\pi_0(t, \cdot) : \mathcal{X} \rightarrow \mathbb{R}$ for fixed $t \in \mathcal{T}$. We see

$$\mathbb{E}\{\pi_0(t, X)u(X)|T = t\} = \int_{\mathcal{X}} \frac{f_T(t)}{f_{T|X}(t|x)} u(x) f_{X|T}(x|t) dx = \mathbb{E}\{u(X)\} \quad (1)$$

holds for any integrable $u(X)$

Estimation of $\pi_0(t, X)$

Theorem 3.1 for fixed $t \in \mathcal{T}$ and any integrable function $u(X)$,

$$\lim_{h_0 \rightarrow 0} \frac{\mathbb{E}[\pi(t, X)u(X)L_U\{h_0^{-1}(t - S)\}]}{\mathbb{E}[L_U\{h_0^{-1}(t - S)\}]} = \mathbb{E}[u(X)] \quad (11)$$

if and only if $\pi(t, X) = \pi_0(t, X)$ a.s

- We can solve a sample analogue of (11) for any integrable function $u(X)$ where $h_0 \rightarrow 0$ as $N \rightarrow \infty$ but this need solving infinite number of equations which is impractical
- **Solution** We can approximate infinite dimensional space of $u(x)$ using a finite known basis function with dimension K (power series, B-splines etc) i.e $u_K(X) = (u_{K,1}(X), \dots, u_{K,K}(X))^T$ which approximates any $u(X)$ as $K \rightarrow \infty$. Now,

$$\lim_{h_0 \rightarrow 0} \frac{\mathbb{E}[\pi(t, X)u_K(X)L_U\{h_0^{-1}(t - S)\}]}{\mathbb{E}[L_U\{h_0^{-1}(t - S)\}]} = \mathbb{E}[u_K(X)] \quad (12)$$

Estimation of $\pi_0(t, X)$

Let $\rho(\cdot)$ be a globally concave and increasing function. Define a strictly concave function

$$G_t^*(\lambda) = \lim_{h_0 \rightarrow 0} \frac{\mathbb{E}[\rho\{\lambda^T u_K(X)\} L_U\{h_0^{-1}(t - S)\}]}{\mathbb{E}[L_U\{h_0^{-1}(t - S)\}]} - \lambda^T \mathbb{E}[u_K(X)]$$

and $\lambda_t^* = \operatorname{argmax}_{\lambda \in \mathbb{R}^k} G_t^*(\lambda)$ Hence $\nabla G_t^*(\lambda_t^*) = 0$ which implies (12) holds if

$$\pi^*(t, X) = \rho'\{\lambda_t^{*T} u_K(X)\} \quad (13)$$

Hence this paper proposes

$$\hat{\pi}(t, X) = \rho'\{\hat{\lambda}_t^T u_K(X)\} \quad (14)$$

with $\hat{\lambda}_t = \operatorname{argmax}_{\lambda \in \mathbb{R}^k} \hat{G}_t(\lambda)$

$$\hat{G}_t(\lambda) = \lim_{h_0 \rightarrow 0} \frac{\sum_{i=1}^N \rho\{\lambda^T u_K(X)\} L_U\{h_0^{-1}(t - S_i)\}}{\sum_{i=1}^N L_U\{h_0^{-1}(t - S_i)\}} - \lambda^T \{N^{-1} \sum_{i=1}^N u_K(X)\} \quad (15)$$

Equivalent Duel Solution interpretation

The proposed estimator $\hat{\pi}(t, X)$ is also a duel solution to a local generalised empirical likelihood maximization problem: for fixed $t \in \mathcal{T}$,

$$\left\{ \begin{array}{l} \max_{\{\pi_i\}_{i=1}^N} - \frac{\sum_{i=1}^N D(\pi_i) L_U(\{t - S_i\}/h_0)}{\sum_{i=1}^N L_U(\{t - S_i\}/h_0)} \\ \text{subject to } \frac{\sum_{i=1}^N \pi_i u_K(X_i) L_U(\{t - S_i\}/h_0)}{\sum_{i=1}^N L_U(\{t - S_i\}/h_0)} = \frac{1}{N} \sum_{i=1}^N u_K(X_i), \end{array} \right.$$

where $D(v)$ is a distance measure from v to 1 for $v \in \mathbb{R}$, which is continuously differentiable and satisfies that $D(1) = 0$ and

$$\rho(-v) = D\{(D')^{-1}(v)\} - v \cdot (D')^{-1}(v).$$

Now once we get $\hat{\pi}(t, X)$ we obtain estimator of $\mu(t)$:

$$\hat{\mu}(t) = \frac{\sum_{i=1}^N \hat{\pi}(t, X_i) Y_i L_U\{(t - S_i)/h_0\}}{\sum_{i=1}^N L_U\{(t - S_i)/h\}} = \frac{1}{N} \sum_{i=1}^N u_K(X_i) \quad (17)$$

Proof of Theorem 3.1

We first prove that for every fixed $t \in \mathcal{T}$ and any integrable function $u(X)$, $\mathbb{E}\{\pi(t, X)u(X) \mid T = t\} = \mathbb{E}\{u(X)\}$ holds if and only if $\pi(t, X) = \pi_0(t, X)$ a.s.. The sufficient part is obvious and we here show the necessary part. Since for all $t \in \mathcal{T}$ and any integrable function $u(X)$, we have $\mathbb{E}\{\pi(t, X)u(X) \mid T = t\} = \mathbb{E}\{u(X)\}$, comparing to (11), we see that

$$\mathbb{E} [\{\pi(t, X) - \pi_0(t, X)\} u(X) \mid T = t] = 0$$

for all $t \in \mathcal{T}$ and any integrable function $u(X)$. Taking $u(X) = \exp(a^T X)$ for $a \in \mathbb{R}^r$, we have

$$\mathbb{E} [\{\pi(t, X) - \pi_0(t, X)\} \exp(a^T X) \mid T = t] = 0$$

for all $a \in \mathbb{R}^r$. Thus, according to the uniqueness of Laplace transform, we have that $\pi(t, \cdot) = \pi_0(t, \cdot)$ a.s.. Next, we show that

$$\lim_{h_0 \rightarrow 0} \frac{\mathbb{E} [\pi(t, X)u(X)L_U \{(t - S)/h_0\}]}{\mathbb{E} [L_U \{(t - S)/h_0\}]} = \mathbb{E}\{\pi(t, X)u(X) \mid T = t\}$$

Proof of Theorem 3.1

Note,

$$\begin{aligned} & \lim_{h_0 \rightarrow 0} \frac{\mathbb{E} [\pi(t, X) u(X) L_U \{(t - S)/h_0\}]}{\mathbb{E} [L_U \{(t - S)/h_0\}]} \\ &= \lim_{h_0 \rightarrow 0} \frac{\mathbb{E} (\pi(t, X) u(X) \mathbb{E} [L_U \{(t - S)/h_0\} \mid T, X])}{\mathbb{E} (\mathbb{E} [L_U \{(t - S)/h_0\} \mid T, X])} \\ &= \lim_{h_0 \rightarrow 0} \frac{\mathbb{E} (\pi(t, X) u(X) \mathbb{E} [L_U \{(t - S)/h_0\} \mid T])}{\mathbb{E} (\mathbb{E} [L_U \{(t - S)/h_0\} \mid T])} \quad (S \perp X \mid T) \\ &= \lim_{h_0 \rightarrow 0} \frac{h_0^{-1} \mathbb{E} [\pi(t, X) u(X) L \{(t - T)/h_0\}]}{h_0^{-1} \mathbb{E} [L \{(t - T)/h_0\}]} \quad (\text{by (8)}). \end{aligned}$$

Proof of Theorem 3.1

For the numerator, we have

$$\begin{aligned} & \lim_{h_0 \rightarrow 0} h_0^{-1} \mathbb{E} [\pi(t, X) u(X) L \{(t - T)/h_0\}] \\ &= \lim_{h_0 \rightarrow 0} h_0^{-1} \iint \pi(t, x) u(x) L \{(t - t')/h_0\} f_{T,X}(t', x) dt' dx \\ &= - \lim_{h_0 \rightarrow 0} \iint \pi(t, x) u(x) L(z) f_{T,X}(t - zh_0, x) dz dx \\ &= - \iint \pi(t, x) u(x) L(z) f_{T,X}(t, x) dz dx \\ &= \mathbb{E} \{ \pi(t, X) u(X) \mid T = t \} \cdot f_T(t). \end{aligned}$$

Similarly, we have

$$\lim_{h_0 \rightarrow 0} h_0^{-1} \mathbb{E} [L \{(t - T)/h_0\}] = f_T(t).$$

The results then follows.

Assumptions

- **Assumption 2** The kernel function $L(\cdot)$ is an even function such that $\int_{-\infty}^{\infty} L(u)du = 1$ and has finite moments of order 3 .
- **Assumption 3** Assume
 - (i) the support \mathcal{X} of \mathbf{X} is a compact subset of \mathbb{R}^r . The support \mathcal{T} of the treatment variable T is a compact subset of \mathbb{R} .
 - (ii) (Strict Positivity) there exist a positive constant η_{\min} such that $f_{T|X}(t | x) \geq \eta_{\min} > 0$, for all $x \in \mathcal{X}$.
- **Assumption 4**
 - (i) The densities $f_T(t)$, $f_{T|X}(t | \mathbf{X})$, and $f_{T|Y,X}(t | Y, X)$ are third-order continuously differentiable w.r.t. t almost surely.
 - (ii) The derivatives of $f_{T|X}(t | X)$ and $f_{T|Y,X}(t | Y, X)$, denoted by $\{\partial_t^d f_{T|X}(t | X), \partial_t^d f_{T|Y,X}(t | Y, X)$ for $d = 0, 1, 2, 3\}$, are integrable almost surely in t .
- **Assumption 5** For every $t \in \mathcal{T}$,
 - (i) the function $\pi_0(t, x)$ is s -times continuously differentiable w.r.t. $x \in \mathcal{X}$, where $s > r/2$ is an integer;
 - (ii) there exist $\lambda_t \in \mathbb{R}^K$ and a positive constant $\alpha > 0$ such that
$$\sup_{x \in \mathcal{X}} \left| (\rho')^{-1} \{ \pi_0(t, x) \} - \lambda_t^\top u_K(x) \right| = O(K^{-\alpha}).$$

Assumptions

- **Assumption 6**

- (i) For every K , the eigenvalues of $\mathbb{E} [u_K(\mathbf{X})u_K(\mathbf{X})^\top \mid T = t]$ are bounded away from zero and infinity, and twice differentiable w.r.t. t for $t \in \mathcal{T}$.
- (ii) There is a sequence of constants $\zeta(K)$ satisfying $\sup_{x \in \mathcal{X}} \|u_K(x)\| \leq \zeta(K)$, such that $\zeta(K) \{K^{-\alpha} + h_0^2 + h^2\} \rightarrow 0$ as $N \rightarrow \infty$, where $\|\cdot\|$ denotes the Euclidean norm.

- **Assumption 7**

- For every $t \in \mathcal{T}$, there exist $\gamma_t \in \mathbb{R}^K$ and a positive constant $\ell > 0$ such that $\sup_{x \in \mathcal{X}} |m(t, x) - \gamma_t^\top u_K(x)| = O(K^{-\ell})$, where $m(t, x) = \mathbb{E}[Y \mid T = t, X = x]$.

- **Assumption 8** Following are bounded for some $\delta > 0$, for all $t \in \mathcal{T}$.

- $R_1^{2+\delta}(t) = \mathbb{E}[|\pi_0(t, X)Y - \mu(t)|^{2+\delta} \mid T = t]$
- $R_2^{2+\delta}(t) = \mathbb{E}[|\pi_0(t, X)m(t, X) - \mu(t)|^{2+\delta} \mid T = t]$ and
- $R_3^{2+\delta}(t) := \mathbb{E}[|\pi_0(t, X)\{Y - m(t, X)\}|^{2+\delta} \mid T = t]$

Ordinary Smooth and Super Smooth error

- An ordinary smooth error of order $\beta \geq 1$ satisfies,

$$\lim_{t \rightarrow \infty} t^\beta \phi_U(t) = c \text{ and } \lim_{t \rightarrow \infty} t^{\beta+1} \phi_U^{(1)}(t) = -c\beta \quad (18)$$

for some $c > 0$

- A super smooth error of order $\beta \geq 1$ satisfies,

$$d_0 |t|^{\beta_0} e^{-|t|^\beta/\gamma} \leq \phi_U(t) \leq d_1 |t|^{\beta_1} e^{-|t|^\beta/\gamma} \text{ as } |t| \rightarrow \infty \quad (19)$$

for positive constants d_0, d_1, γ and constants β_0 and β_1

- **Assumption O** (Ordinary Smooth Case)

- $\|\phi_L\|_\infty < \infty$, $\int_{-\infty}^\infty |t|^{\beta+1} \{|\phi_L(t)| + |\partial_t \phi_L(t)|\} dt < \infty$ and $\int_{-\infty}^\infty |t^\beta \phi_L(t)|^2 dt < \infty$.

- **Assumption S** (Super Smooth Case)

- $\phi_L(t)$ is support on $[-1, 1]$ and bounded.

Asymptotics for estimated weight function

Theorem 4.1 Suppose that the error U is ordinary smooth of order β satisfying (18) and that Assumption O holds. Under Assumptions 2-6 and $\zeta(K)\{K/(Nb_0^{1+2\beta})\}^{1/2} \rightarrow 0$ as $N \rightarrow \infty$, for every fixed $t \in \mathcal{T}$, then

$$\sup_{x \in \mathcal{X}} |\hat{\pi}(t, x) - \pi_0(t, x)| = O_p \left(\zeta(K)\{K^{-\alpha} + b_0^2\} + \zeta(K) \left\{ \frac{K}{Nb_0^{1+2\beta}} \right\}^{1/2} \right),$$

$$\int_{\mathcal{X}} |\hat{\pi}(t, x) - \pi_0(t, x)|^2 dF_X(x) = O_p \left(\{K^{-2\alpha} + b_0^4\} + \frac{K}{Nb_0^{1+2\beta}} \right)$$

$$\frac{1}{N} \sum_{i=1}^N |\hat{\pi}(t, X_i) - \pi_0(t, X_i)|^2 = O_p \left(\{K^{-2\alpha} + b_0^4\} + \frac{K}{Nh_0^{1+2\beta}} \right).$$

Asymptotics for estimated weight function

Theorem 4.3 Suppose that the error U is supersmooth of order β satisfying (19) and Assumption S holds. Under Assumptions 2 – 6 and $\zeta^2(K)K \cdot (Nh_0)^{-1} \cdot e^{(2h_0^{-\beta}/\gamma)} \rightarrow 0$ as $N \rightarrow \infty$, for every fixed $t \in \mathcal{T}$, then

$$\begin{aligned} \sup_{x \in \mathcal{X}} |\hat{\pi}(t, x) - \pi_0(t, x)| &= O_p \left(\zeta(K) \left[\{K^{-\alpha} + b_0^2\} + \frac{e^{(h_0^{-\beta}/\gamma)}}{\sqrt{h_0}} \cdot \sqrt{\frac{K}{N}} \right] \right), \\ \int_{\mathcal{X}} |\hat{\pi}(t, x) - \pi_0(t, x)|^2 \, dF_X(x) &= O_p \left(\{K^{-2\alpha} + b_0^4\} + \frac{e^{(2h_0^{-\beta}/\gamma)}}{h_0} \cdot \frac{K}{N} \right), \\ \frac{1}{N} \sum_{i=1}^N |\hat{\pi}(t, X_i) - \pi_0(t, X_i)|^2 &= O_p \left(\{K^{-2\alpha} + b_0^4\} + \frac{\exp(2h_0^{-\beta}/\gamma)}{h_0} \cdot \frac{K}{N} \right) \end{aligned}$$

Asymptotics for unconditional ADRF

Let us define for $i = 1, \dots, N$,

$\eta_{h,h_0}(S_i, X_i, Y_i; t) = \phi_h(S_i, X_i, Y_i; t) + \psi_{h_0}(S_i, X_i, Y_i; t)$ where,

- $\phi_h(S_i, X_i, Y_i; t) = [\pi_0(t, X_i)Y_i L_{U,h}(t - S_i) - \mathbb{E}\{\pi_0(t, X)Y L_{U,h}(t - S)\}] - \mu(t)[L_{U,h}(t - S_i) - \mathbb{E}\{L_{U,h}(t - S)\}]$
- $\psi_{h_0}(S_i, X_i, Y_i; t) = \mu(t)[L_{U,h_0}(t - S_i) - \mathbb{E}\{L_{U,h_0}(t - S)\}] - [m(t, X_i)\pi_0(t, X_i)L_{U,h_0}(t - S_i) - \mathbb{E}\{m(t, X_i)\pi_0(t, X)L_{U,h}(t - S)\}]$
- $L_{U,h}(v) = h^{-1}L_U(v/h)$
- $V_j = f_T^{-2}(t)(R_j^2 f_T) * f_U(t).C$ for $j = 1, 2$
- $C = \int_{-\infty}^{\infty} J^2(v)dv = (2\pi c^2)^{-1} \int |w|^{2\beta} \phi_L^2(w)dw$
- R_1^2, R_2^2 are defined as in Assumption 8 also,
 $(R_1 R_2)(t) = \mathbb{E}[\{\pi_0(t, X)Y - \mu(t)\}\{\mu(t) - \pi_0(t, X)m(t, X)\}|T = t]$
and $v_h(t) = \mathbb{E}\{L_{U,h}^2(t - S)\}$

Asymptotics for unconditional ADRF

Theorem 4.2 Suppose that the error U is ordinary smooth of order β satisfying (18) and Assumption O holds. Under Assumptions 1 – 8 and $\frac{(K^{-l}+h_0^2)(K^{-\alpha}+h_0^2)}{h^2} + \frac{(h \wedge h_0)^{1+2\beta}}{h_0^{1+2\beta}} \frac{K}{\sqrt{N}} \rightarrow 0$, then where

$$\begin{aligned}\hat{\mu}_t - \mu_t &= \frac{\kappa_{21}}{2} \left[\frac{f_T(t)\Phi_1(t) - \mu(t)\partial_t^2 f_T(t)}{f_T(t)} \right] h^2 + o(h^2) \\ &+ \frac{\kappa_{21}}{2} \left[\frac{\mu(t)\partial_t^2 f_T(t) - f_T(t)\Phi_2(t)}{f_T(t)} \right] h_0^2 + o(h_0^2) \\ &+ \sum_{i=1}^N \frac{\eta_{h,h_0}(S_i, X_i, Y_i; t)}{N f_T(t)} + o_P\left(\frac{1}{\sqrt{N(h \wedge h_0)^{1+2\beta}}}\right)\end{aligned}$$

where, $\kappa_{21} = \int u^2 L(u) du$, $\Phi_1(t) = \mathbb{E}[\{Y \partial_t^2 f_{T|X}(t|x)\} / \{f_{T|X}(t|X)\}]$ and $\Phi_2(t) = \mathbb{E}[\{m(t, X) \partial_t^2 f_{T|X}(t|X)\} / \{f_{T|X}(t|X)\}]$

Asymptotics for unconditional ADRF

In fact,

- if $h = o(h_0)$ then $\sqrt{h^{1+2\beta}/N} \sum_{i=1}^N \frac{\eta_{h,h_0}(S_i, X_i, Y_i; t)}{f_T(t)} \rightarrow N(0, V_1)$
- if $h_0 = o(h)$ then $\sqrt{h_0^{1+2\beta}/N} \sum_{i=1}^N \frac{\eta_{h,h_0}(S_i, X_i, Y_i; t)}{f_T(t)} \rightarrow N(0, V_2)$
- if $h_0 = \tilde{c}h$ then $\sqrt{h^{1+2\beta}/N} \sum_{i=1}^N \frac{\eta_{h,h_0}(S_i, X_i, Y_i; t)}{f_T(t)} \rightarrow N(0, V_3)$ where,

$$V_3 = \frac{(R_1^2 f_T) * f_U(t)}{f_T^2(t)} \int_{-\infty}^{\infty} J^2(v) dv + \frac{(R_2^2 f_T) * f_U(t)}{\tilde{c}^{2+2\beta} f_T^2(t)} \int_{-\infty}^{\infty} J^2(v/\tilde{c}) dv \\ + \frac{2(R_1 R_2 f_T) * f_U(t)}{\tilde{c}^{1+\beta} f_T^2(t)} \int_{-\infty}^{\infty} J(v) J(v/\tilde{c}) dv$$

As long as $(K^{-l} + h_0)(K^{-\alpha} + h_0) = o(h^2)$, the error arising from the sieve approximation is asymptotically negligible. For example, $\hat{\mu}(t) - \mu(t)$ achieves the optimal convergence rate, $N^{-2/(2\beta+5)}$, if $h_0 \sim h \sim N^{-1/(2\beta+5)}$. Also we need $K = o(h^{-2})$, $\alpha + l > 1$ and

- $\alpha > 1/2$ if spline basis is used
- $\alpha > 1$ if a power series is used

Asymptotics for unconditional ADRF

Theorem 4.4 Suppose that the error U is super smooth of order β satisfying (19) and Assumption S holds. Letting $e(h) = h^{1/2}e^{-\frac{h^{-\beta}}{\gamma}}$ we have $v_h(t) = O(e^{-2}(h))$. If Assumptions 1 – 8 holds and for $h \rightarrow \infty, v_h(t) \rightarrow \infty$ and $\frac{(K^{-l}+h_0^2)(K^{-\alpha}+h_0^2)}{h^2} + \frac{1}{\{e(h)\wedge e(h_0)\}} \frac{K}{\sqrt{N}} \rightarrow 0$ as $N \rightarrow \infty$, then for fixed t

$$\begin{aligned}\hat{\mu}_t - \mu_t &= \frac{\kappa_{21}}{2} \left[\frac{f_T(t)\Phi_1(t) - \mu(t)\partial_t^2 f_T(t)}{f_T(t)} \right] h^2 + o(h^2) \\ &+ \frac{\kappa_{21}}{2} \left[\frac{\mu(t)\partial_t^2 f_T(t) - f_T(t)\Phi_2(t)}{f_T(t)} \right] h_0^2 + o(h_0^2) \\ &+ \sum_{i=1}^N \frac{\eta_{h,h_0}(S_i, X_i, Y_i; t)}{N f_T(t)} \{1 + o_P(1)\}\end{aligned}$$

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- $\sum_{i=1}^N \frac{\eta_{h,h_0}(S_i, X_i, Y_i; t)}{N f_T(t)} = O_P(N^{-1/2} \{e(h) \wedge e(h_0)\}^{-1})$

Moreover, if $v_h(t) \geq d_1 f_S(t) h^{d_3} e^{(2h^{-\beta}/\gamma - d_2 h^{-d_4\beta})}$ for constants $d_2 > 0$, $0 < d_4 < 1$ and d_1, d_3 we have

- $[\text{var}\{\eta_{h,h_0}(S_i, X_i, Y_i; t)\}]^{-1/2} \frac{1}{\sqrt{N}} \sum_{i=1}^N \{\eta_{h,h_0}(S_i, X_i, Y_i; t)\} \rightarrow N(0, 1)$

As in the ordinary smooth case, as long as $\pi_0(t, \cdot)$ is sufficiently smooth or K grows sufficiently fast, and h_0 decays fast enough, the sieve approximation error of our estimator is asymptotically negligible and the dominating bias term is the same as that in the ordinary smooth case.

- From the theorem, when $h \sim h_0$ and $h \wedge h_0 = d(\log N)^{-1/\beta}$ for a constant $d > (2/\gamma)^{1/\beta}$, one finds that the rate of variance, $N^{-1} \{e(h) \wedge e(h_0)\}^{-2} = o(h^4 + h_0^4)$, is negligible compared to the asymptotic bias, and the convergence rate of $\hat{\mu}(t) - \mu(t)$ is $(\log N)^{-2/\beta}$.

Preliminaries

Lets assume K and h_0 are given, then CV criteria to this context to choose h , we see $CV(h) = \sum_{i=1}^N \{\hat{\pi}(T_i, X_i)Y_i - \hat{\mu}^{-i}(T_i)\}^2 w(T_i)$ is difficult to compute since we do not observe T . To tackle this,

- generate two additional sets of contaminated data, $S_{i,d}^* = S_i + U_{id}^*$ and $S_{i,d}^{**} = S_{i,d}^* + U_{i,d}^{**}$, with $d = 1(1)D$ large and i.i.d $U_{i,d}^{**}, U_{i,d}^{**} \sim U$
- Based on this we first obtain $(\hat{\pi}_d^*, \hat{\mu}_d^*)$ and then $(\hat{\pi}_d^{**}, \hat{\mu}_d^{**})$
- compute \hat{h}^* and \hat{h}^{**} which minimises $\sum_{i=1}^D CV_d^*(h)/D$ and $\sum_{i=1}^D CV_d^{**}(h)/D$ respectively
- Intuitively, we then expect the relationship between \hat{h}^* and our target bandwidth h to be similar to that between \hat{h}^* and \hat{h}^{**} . Thus, the authors considered that $h/\hat{h}^* \approx \hat{h}^*/\hat{h}^{**}$ and used a linear back-extrapolation procedure that would give the bandwidth $\hat{h}_{DH} = (\hat{h}^*)^2/\hat{h}^{**}$

Two-step procedure and local constant extrapolation

Extending the SIMEX method to choose three parameters simultaneously is computationally unstable thus the paper proposes

- Plug-in bandwidth h_{PI} for the kernel de convolution estimator with bandwidth h of the density of T proposed by Delaigle and Gijbels (2002) minimises the asymptotic MSE of the estimator, whose bias is of rate of h^2 and standard deviation $\sqrt{v_h(t)/N}$
- set $h_0 = h_{PI}$ and choose $K = \lfloor \tilde{c} h_{PI}^{-2} \log(h_{PI} + 1) \rfloor$ such that $K \geq 2$, where \tilde{c} minimises the generalised CV criterion (Craven and Wahba, 1978)
- The linear back-extrapolation sometimes gave highly unstable results as which extrapolant function should be used in practice is unknown hence they used local constant estimator (Fan and Gijbels, 1996), with $h_d^* = c_d^* h_{PI}$ and $h_d^{**} = c_d^{**} h_{PI}$ where c_d^* , c_d^{**} minimise $CV_d^*(h)$ and $CV_d^{**}(h)$, respectively and we choose $\hat{h} = \sum_{d=1}^D h_d^* \phi\{(\hat{h}^* - h_d^{**})/b\} / \sum_{d=1}^D \phi\{(\hat{h}^* - h_d^{**})/b\}$

Thank You