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## 《机器学习基础》第7次作业

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T<sub>1</sub>

解. 对于二项 logistic 回归模型, 有

$$\log L(\theta) = \sum_{i=1}^N \left[ y_i (w \cdot x_i + b) - \log(1 + e^{w \cdot x_i + b}) \right]$$

参数  $\theta$  的极大似然估计

$$\hat{\theta} = (\hat{w}, \hat{b}) = \arg \max_{\theta=(w,b)} \log L(\theta) = \arg \min_{\theta=(w,b)} [-\log L(\theta)]$$

利用拟牛顿法求解, 算法框架如下:

(BFGS 算法)

输入: 目标函数  $f(\theta) = -\log L(\theta)$ 

$$g(\theta) = \nabla f(\theta)$$

精度要求  $\varepsilon$ 输出: 极小点  $\hat{\theta} = \arg \min_{\theta=(w,b)} f(\theta)$ i) 任意选定初始  $\theta^{(0)}$ , 取正定对称矩阵  $B_0$ , 令  $k=0$ ii) 计算  $g_k = g(\theta^{(k)})$ 若  $\|g_k\| < \varepsilon$ , 输出  $\hat{\theta} = \theta^{(k)}$ , 停止;  
否则转到 iii)

iii) 由

$$B_k p_k = -g_k$$

求出  $p_k$ iv) 一维搜索: 求  $\lambda_k$  s.t.

$$f(\theta^{(k)} + \lambda_k p_k) = \min_{\lambda \geq 0} f(\theta^{(k)} + \lambda p_k)$$

$$v) \text{ 令 } \theta^{(k+1)} = \theta^{(k)} + \lambda_k p_k$$

$$vi) \text{ 计算 } g_{k+1} = g(\theta^{(k+1)})$$

若  $\|g_{k+1}\| < \varepsilon$ , 输出  $\hat{\theta} = \theta^{(k+1)}$ , 停止;

否则, 计算

$$B_{k+1} = B_k + \frac{y_k y_k^T}{y_k^T \delta_k} - \frac{B_k \delta_k \delta_k^T B_k}{\delta_k^T B_k \delta_k}$$

其中

$$y_k = g_{k+1} - g_k$$

$$\delta_k = \theta^{(k+1)} - \theta^{(k)}$$

vii)  $\Delta k = k+1$ , 转到 iii)T<sub>2</sub>

证: 1)

极大似然估计:

$$\text{先验概率 } \hat{P}(Y=c_k) = \frac{\sum_{j=1}^N I(Y_j=c_k)}{N} \quad \textcircled{1}$$

$$k=1, 2, \dots, K$$

条件概率

$$\hat{P}(X^{(i)}=a_i^{(i)} | Y=c_k) = \frac{\sum_{j=1}^N I(X_j^{(i)}=a_i^{(i)}, Y_j=c_k)}{\sum_{j=1}^N I(Y_j=c_k)} \quad \textcircled{2}$$

$$l=1, 2, \dots, m_i, \quad i=1, 2, \dots, n$$

先证①. 记

$$p = P(Y=c_k)$$

$$q = \sum_{j=1}^N I(Y_j=c_k)$$

似然函数

$$L(p) = C_N^q p^q (1-p)^{N-q}$$

$$\text{令 } \frac{\partial L(p)}{\partial p} = 0, \quad \forall p$$

$$q p^{N-1} (1-p)^{N-1} = (N-1) (1-p)^{N-2} p^2$$

$$\text{非平凡解为 } \hat{p} = \frac{q}{N} = \frac{\sum_{j=1}^N I(Y_j = C_k)}{N}$$

$$\text{即 } \hat{p}(Y = C_k) = \frac{\sum_{j=1}^N I(Y_j = C_k)}{N};$$

类似地, 对  $\theta$ , 记

$$p_l = P(X^{(i)} = a_l^{(i)} | Y = C_k)$$

$$q_l = \sum_{j=1}^N I(X_j^{(i)} = a_l^{(i)}, Y_j = C_k)$$

似然函数

$$L(p) = C_N^q p_1^{q_1} (1-p_1)^{N-q_1}$$

经过相似计算过程, 即得

$$\hat{p}(X^{(i)} = a_l^{(i)} | Y = C_k) = \hat{p}_l = \frac{\sum_{j=1}^N I(X_j^{(i)} = a_l^{(i)}, Y_j = C_k)}{\sum_{j=1}^N I(Y_j = C_k)};$$

2) 贝叶斯估计:

先验概率

$$\hat{p}_\lambda(Y = C_k) = \frac{\sum_{j=1}^N I(Y_j = C_k) + \lambda}{N + K\lambda} \quad (3)$$

$$k = 1, 2, \dots, K$$

条件概率

$$\hat{p}_\lambda(X^{(i)} = a_l^{(i)} | Y = C_k) = \frac{\sum_{j=1}^N I(X_j^{(i)} = a_l^{(i)}, Y_j = C_k) + \lambda}{\sum_{j=1}^N I(Y_j = C_k) + m_l \lambda} \quad (4)$$

$$l = 1, 2, \dots, m_i$$

$$i = 1, 2, \dots, n$$

先证 (4).

考虑  $p_\lambda(Y = C_i) = \theta_i, i = 1, \dots, K$   
为随机变量

$C$  分布为多项分布. 故其先验分布为 Dirichlet 分布. 记参数为  $\lambda$ , 则有

$$P(\theta_1, \dots, \theta_K; \lambda) = \frac{1}{B(\lambda)} \prod_{i=1}^K \theta_i^{\lambda_i - 1} \quad (1)$$

$$\text{训练集 } D = \{(x_i, y_i)\}_{i=1}^N$$

$$\text{记 } M_i = \sum_{j=1}^N I(Y_j = C_i) \\ i = 1, 2, \dots, K$$

$$\theta = (\theta_1, \dots, \theta_K)$$

$$M = (M_1, \dots, M_K)$$

则后验分布

$$P(\theta | M) = \frac{P(M | \theta) P(\theta)}{\int P(M | \theta) P(\theta) d\theta}$$

考虑后验概率  $P(M | \theta)$  服从多项分布.

$$P(M | \theta) = \frac{K!}{M_1! \dots M_K!} \theta_1^{M_1} \dots \theta_K^{M_K} \quad (2)$$

由 (2), 又由  $\int P(M | \theta) P(\theta) d\theta$  为定值, 知

$$P(\theta | M) \propto \prod_{i=1}^K \theta_i^{M_i + \lambda - 1}$$

也服从 Dirichlet 分布

因此

$$\hat{p}_\lambda(Y = C_k) = E(\theta_k) = \frac{M_k + \lambda}{N + K\lambda} \\ = \frac{\sum_{j=1}^N I(Y_j = C_k) + \lambda}{N + K\lambda};$$

(4) 的证明是完全类似的, 只需考虑

$$p_\lambda(Y = C_l, X^{(i)} = a_l^{(i)} | Y = C_k) = \theta_l, \\ l = 1, 2, \dots, m_i$$

$$M_i = \sum_{j=1}^N I(x_j^{(i)} = a_i^{(i)}, y_j = c_k)$$

使用完全相同的证明过程，可得

$$\begin{aligned} \hat{p}_\lambda(X^{(i)} = a_i^{(i)} | Y = c_k) &= E(\theta_i) \\ &= \frac{\sum_{j=1}^N I(x_j^{(i)} = a_i^{(i)}, y_j = c_k) + \lambda}{\sum_{j=1}^N I(y_j = c_k) + m_i \lambda} \end{aligned}$$

证毕！