

Three

Joint Distributions

1 JOINT DISTRIBUTIONS FOR MULTIPLE RANDOM VARIABLES

Very often, we are interested in more than one random variables *simultaneously*.

- For example, an investigator might be interested in both the height (H) and the weight (W) of individuals from a certain population.
- Another investigator could be interested in both the hardness (H) and the tensile strength (T) of a piece of cold-drawn copper.

DEFINITION 1 (TWO-DIMENSIONAL RANDOM VECTOR)

Let E be an experiment and S be a corresponding sample space. Suppose X and Y are two functions each assigning a real number to each $s \in S$.

*We call (X, Y) a **two-dimensional random vector**, or a **two-dimensional random variable**.*

DEFINITION 2 (RANGE SPACE)

*Similar to the one-dimensional situation, we can denote the **range space** of (X, Y) by*

$$R_{X,Y} = \{(x, y) \mid x = X(s), y = Y(s), s \in S\}.$$

The definitions above can be extended to more than two random variables.

DEFINITION 3 (*n*-DIMENSIONAL RANDOM VECTOR)

Let X_1, X_2, \dots, X_n be n functions each assigning a real number to every outcome $s \in S$.

We call (X_1, X_2, \dots, X_n) a *n -dimensional random vector*, or a *n -dimensional random variable*.

We define the discrete and continuous two-dimensional random variables as follows.

DEFINITION 4

(X, Y) is a **discrete two-dimensional random variable** if the number of possible values of $(X(s), Y(s))$ are finite or countable. That is, the possible values of $(X(s), Y(s))$ may be represented by

$$(x_i, y_j), \quad i = 1, 2, 3, \dots; j = 1, 2, 3, \dots$$

(X, Y) is a **continuous two-dimensional random variable** if the possible values of $(X(s), Y(s))$ can assume any value in some region of the Euclidean space \mathbb{R}^2 .

REMARK

We can view X and Y separately to judge whether (X, Y) is discrete or continuous.

- If both X and Y are discrete random variables, then (X, Y) is discrete.
- If both X and Y are continuous random variables, then (X, Y) is continuous.
- Clearly, there are other cases. For example, X is discrete, but Y is continuous. These are not the focus of this course.

EXAMPLE 3.1

Consider a TV set that needs to be serviced.

Let X be the age of the set, rounded to the nearest year, and Y be the numbers of defective components in the set.

Then (X, Y) is a discrete 2-dimensional random variable and its range space is given as

$$R_{X,Y} = \{(x, y) \mid x = 0, 1, 2, \dots; y = 0, 1, 2, \dots, n\},$$

where n is the total number of components in the TV.

For example, $(X, Y) = (5, 3)$ means that the TV is 5 years old and has 3 defective components.

Joint Probability Function

We will now introduce the probability functions for discrete and continuous random vectors.

For the discrete random vector, similar to the one-dimensional case, we define its probability function by associating a number with each possible value of the random variable.

DEFINITION 5 (DISCRETE JOINT PROBABILITY FUNCTION)

Let (X, Y) be a 2-dimensional **discrete** random variable. Its **joint probability (mass) function** is defined by

$$f_{X,Y}(x, y) = P(X = x, Y = y),$$

for $(x, y) \in R_{X,Y}$.

PROPERTIES OF THE DISCRETE JOINT PROBABILITY FUNCTION

The joint probability mass function has the following properties:

(1) $f_{X,Y}(x,y) \geq 0$ for any $(x,y) \in R_{X,Y}$.

(2) $f_{X,Y}(x,y) = 0$ for any $(x,y) \notin R_{X,Y}$.

(3)
$$\sum_{i=1}^{\infty} \sum_{j=1}^{\infty} f_{X,Y}(x_i, y_j) = \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} P(X = x_i, Y = y_j) = 1.$$

Equivalently,
$$\sum \sum_{(x,y) \in R_{X,Y}} f(x,y) = 1.$$

(4) Let A be any subset of $R_{X,Y}$, then

$$P((X,Y) \in A) = \sum \sum_{(x,y) \in A} f_{X,Y}(x,y).$$

EXAMPLE 3.2

Find the value of k such that

$$f(x,y) = kxy, \quad \text{for } x = 1, 2, 3 \text{ and } y = 1, 2, 3,$$

can serve as a joint probability function.

Solution:

Note that $R_{X,Y} = \{(x,y) \mid x = 1,2,3; y = 1,2,3\}$, and

$$\begin{array}{lll} f(1,1) = k, & f(1,2) = 2k, & f(1,3) = 3k, \\ f(2,1) = 2k, & f(2,2) = 4k, & f(2,3) = 6k, \\ f(3,1) = 3k, & f(3,2) = 6k, & f(3,3) = 9k. \end{array}$$

Using Property (3), we have

$$\begin{aligned} 1 &= \sum \sum_{(x,y) \in R_{X,Y}} f(x,y) \\ &= 1k + 2k + 3k + 2k + 4k + 6k + 3k + 6k + 9k. \end{aligned}$$

This results in $k = 1/36$.

DEFINITION 6 (CONTINUOUS JOINT PROBABILITY FUNCTION)

Let (X, Y) be a 2-dimensional *continuous* random variable. Its *joint probability (density) function* is a function $f_{X,Y}(x, y)$ such that

$$P((X, Y) \in D) = \iint_{(x,y) \in D} f_{X,Y}(x, y) \, dy \, dx,$$

for any $D \subset \mathbb{R}^2$. More specifically,

$$P(a \leq X \leq b, c \leq Y \leq d) = \int_a^b \int_c^d f_{X,Y}(x, y) \, dy \, dx.$$

PROPERTIES OF THE CONTINUOUS JOINT PROBABILITY FUNCTION

The joint probability density function has the following properties:

(1) $f_{X,Y}(x,y) \geq 0$, for any $(x,y) \in R_{X,Y}$.

(2) $f_{X,Y}(x,y) = 0$, for any $(x,y) \notin R_{X,Y}$.

(3) $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x,y) \, dx \, dy = 1$.

Equivalently, $\iint_{(x,y) \in R_{X,Y}} f_{X,Y}(x,y) \, dx \, dy = 1$.

EXAMPLE 3.3

Find the value c such that $f(x,y)$ below can serve as a joint probability density function for a random variable (X,Y) :

$$f(x,y) = \begin{cases} cx(x+y), & 0 \leq x \leq 1; 1 \leq y \leq 2, \\ 0, & \text{elsewhere.} \end{cases}$$

Solution:

In order for $f(x, y)$ to be a probability density function, we need

$$\begin{aligned} 1 &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) \, dy \, dx = \int_0^1 \int_1^2 cx(x + y) \, dy \, dx \\ &= c \int_0^1 x \left[x + \frac{1}{2}y^2 \right]_1^2 \, dx = c \int_0^1 x(x + 1.5) \, dx \\ &= c \left[\frac{1}{3}x^3 + 1.5 \cdot \frac{1}{2}x^2 \right]_0^1 = c \cdot \frac{13}{12}. \end{aligned}$$

This implies that $c = 12/13$.

2 MARGINAL AND CONDITIONAL DISTRIBUTIONS

We now consider the marginal distributions.

Put simply, the marginal distribution of X is the individual distribution of X , ignoring the value of Y .

DEFINITION 7 (MARGINAL PROBABILITY DISTRIBUTION)

*Let (X, Y) be a two-dimensional random variable with joint probability function $f_{X,Y}(x, y)$. We define the **marginal distribution** of X as follows.*

If Y is a discrete random variable, then for any x ,

$$f_X(x) = \sum_y f_{X,Y}(x, y).$$

If Y is a continuous random variable, then for any x ,

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) \, dy.$$

REMARK

- $f_Y(y)$ for Y is defined in the same way as that of X .
- We can view the marginal distribution as the “projection” of the 2D function $f_{X,Y}(x,y)$ to the 1D function.
- Intuitively, it is the distribution of X by ignoring the presence of Y .

For example, consider a person from a certain community.

- Suppose $X =$ body weight, $Y =$ height, and (X,Y) has joint distribution $f_{X,Y}(x,y)$.
- The marginal distribution $f_X(x)$ of X is the **distribution of body weights for all people in the community**.

- $f_X(x)$ should not involve the variable y . This can be viewed from its definition: y is either summed out or integrated over.
- $f_X(x)$ is a **probability function**; so it satisfies all the properties of the probability function.

EXAMPLE 3.4

We revisit Example 3.2. The joint probability function is given by

$$f(x,y) = \frac{1}{36}xy, \quad \text{for } x = 1, 2, 3 \text{ and } y = 1, 2, 3.$$

Note that X has three possible values: 1, 2, and 3. The marginal distribution for X is given by

- for $x = 1$, $f_X(1) = f(1,1) + f(1,2) + f(1,3) = 6/36 = 1/6$.
- for $x = 2$, $f_X(2) = f(2,1) + f(2,2) + f(2,3) = 12/36 = 1/3$.
- for $x = 3$, $f_X(3) = f(3,1) + f(3,2) + f(3,3) = 18/36 = 1/2$.

For other values of x , $f_X(x) = 0$.

Alternatively, for each $x \in \{1, 2, 3\}$,

$$f_X(x) = \sum_y f(x, y) = \sum_{y=1}^3 \frac{1}{36} xy = \frac{1}{36} x \sum_{y=1}^3 y = \frac{1}{6} x.$$

DEFINITION 8 (CONDITIONAL DISTRIBUTION)

Let (X, Y) be a random variable with joint probability function $f_{X,Y}(x, y)$. Let $f_X(x)$ be the marginal probability function for X . Then for any x such that $f_X(x) > 0$, the **conditional probability function of Y given $X = x$** is defined to be

$$f_{Y|X}(y|x) = \frac{f_{X,Y}(x, y)}{f_X(x)}.$$

REMARK

- For any y such that $f_Y(y) > 0$, we can similarly define the **conditional distribution of X given $Y = y$** as

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)}.$$

- $f_{Y|X}(y|x)$ is defined only for x such that $f_X(x) > 0$; likewise $f_{X|Y}(x|y)$ is defined only for y such that $f_Y(y) > 0$.
- The intuitive meaning of $f_{Y|X}(y|x)$: the distribution of Y given that the random variable X is observed to take the value x .

- Considering y as the variable (and x as a fixed value), $f_{Y|X}(y|x)$ is a probability function, so it must satisfy all the properties of a probability function.
- However, $f_{Y|X}(y|x)$ is not a probability function for x . This means that there is **NO** requirement that

- $\int_{-\infty}^{\infty} f_{Y|X}(y|x) dx = 1$, for X continuous; or

- $\sum_x f_{Y|X}(y|x) = 1$, for X discrete.

- With this definition, we immediately have
 - If $f_X(x) > 0$, $f_{X,Y}(x,y) = f_X(x)f_{Y|X}(y|x)$.
 - If $f_Y(y) > 0$, $f_{X,Y}(x,y) = f_Y(y)f_{X|Y}(x|y)$.

- One immediate application of the conditional distribution is to compute, for continuous random variable,

$$P(Y \leq y|X = x) = \int_{-\infty}^y f_{Y|X}(y|x) dy;$$
$$E(Y|X = x) = \int_{-\infty}^{\infty} y f_{Y|X}(y|x) dy.$$

Their interpretations are clear: the former is the probability that $Y \leq y$, given $X = x$; the latter is the average value of Y given $X = x$.

For the discrete case, the results can be similarly established, based on the definition of $f_{Y|X}(y|x)$.

EXAMPLE 3.5

We revisit Examples 3.2 and 3.4. The joint probability function for (X, Y) is given by

$$f(x, y) = xy/36, \quad \text{for } x = 1, 2, 3 \text{ and } y = 1, 2, 3.$$

The marginal probability function for X is

$$f_X(x) = x/6, \quad \text{for } x = 1, 2, 3.$$

Therefore $f_{Y|X}(y|x)$ is defined for any $x = 1, 2, 3$:

$$f_{Y|X}(y|x) = \frac{f(x, y)}{f_X(x)} = \frac{(xy/36)}{(x/6)} = y/6, \quad \text{for } y = 1, 2, 3.$$

We can also compute

$$P(Y = 2|X = 1) = f_{Y|X}(2|1) = \frac{1}{6} \times 2 = 1/3;$$

$$\begin{aligned} P(Y \leq 2|X = 1) &= P(Y = 1|X = 1) + P(Y = 2|X = 1) \\ &= f_{Y|X}(1|1) + f_{Y|X}(2|1) = 1/6 + 1/3 = 1/2; \end{aligned}$$

$$\begin{aligned} E(Y|X = 2) &= 1 \cdot f_{Y|X}(1|2) + 2 \cdot f_{Y|X}(2|2) + 3 \cdot f_{Y|X}(3|2) \\ &= 1 \cdot (1/6) + 2 \cdot (2/6) + 3 \cdot (3/6) = 7/3. \end{aligned}$$

3 INDEPENDENT RANDOM VARIABLES

We next discuss independence for random variables.

DEFINITION 9 (INDEPENDENT RANDOM VARIABLES)

Random variables X and Y are *independent* if and only if for *any* x and y ,

$$f_{X,Y}(x,y) = f_X(x)f_Y(y).$$

Random variables X_1, X_2, \dots, X_n are *independent* if and only if for *any* x_1, x_2, \dots, x_n ,

$$f_{X_1, X_2, \dots, X_n}(x_1, x_2, \dots, x_n) = f_{X_1}(x_1)f_{X_2}(x_2) \cdots f_{X_n}(x_n).$$

REMARK

- The above definition is applicable whether (X, Y) is continuous or discrete.
- The “product feature” in the definition implies one necessary condition for independence: $R_{X,Y}$ needs to be a product space. In the sense that if X and Y are independent, for any $x \in R_X$ and any $y \in R_Y$, we have

$$f_{X,Y}(x,y) = f_X(x)f_Y(y) > 0,$$

implying $R_{X,Y} = \{(x,y) | x \in R_X; y \in R_Y\} = R_X \times R_Y$.

Conclusion:

If $R_{X,Y}$ is not a product space, then X and Y are not independent!

PROPERTIES OF INDEPENDENT RANDOM VARIABLES

Suppose X, Y are independent random variables.

- (1) If A and B are arbitrary subsets of \mathbb{R} , the events $X \in A$ and $Y \in B$ are independent events in S . Thus

$$P(X \in A; Y \in B) = P(X \in A)P(Y \in B).$$

In particular, for any real numbers x, y ,

$$P(X \leq x; Y \leq y) = P(X \leq x)P(Y \leq y).$$

(2) For arbitrary functions $g_1(\cdot)$ and $g_2(\cdot)$, $g_1(X)$ and $g_2(Y)$ are independent. For example,

- X^2 and Y are independent.
- $\sin(X)$ and $\cos(Y)$ are independent.
- e^X and $\log(Y)$ are independent.

(3) Independence is connected with conditional distribution.

- If $f_X(x) > 0$, then $f_{Y|X}(y|x) = f_Y(y)$.
- If $f_Y(y) > 0$, then $f_{X|Y}(x|y) = f_X(x)$.

EXAMPLE 3.6

The joint probability function of (X, Y) is given below.

x	y			$f_X(x)$
	1	3	5	
2	0.1	0.2	0.1	0.4
4	0.15	0.3	0.15	0.6
$f_Y(y)$	0.25	0.5	0.25	1

Are X and Y independent?

Solution:

We need to check that for every x and y combination, whether we have

$$f_{X,Y}(x,y) = f_X(x)f_Y(y).$$

For example, from the table, we have $f_{X,Y}(2,1) = 0.1$; $f_X(2) = 0.4$, $f_Y(1) = 0.25$. Therefore

$$f_{X,Y}(2,1) = 0.1 = 0.4 \times 0.25 = f_X(2)f_Y(1).$$

In fact, we can check for each $x \in \{2,4\}$ and $y \in \{1,3,5\}$ combination, the equality holds. Therefore X and Y are independent.

4 EXPECTATION AND COVARIANCE

Similar to one dimensional random variable, we can talk about the expectation of a random vector.

DEFINITION 10 (EXPECTATION)

Consider any two variable function $g(x, y)$.

If (X, Y) is a discrete random variable,

$$E(g(X, Y)) = \sum_x \sum_y g(x, y) f_{X, Y}(x, y).$$

If (X, Y) is a continuous random variable,

$$E(g(X, Y)) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f_{X, Y}(x, y) dy dx.$$

If we let

$$g(X, Y) = (X - E(X))(Y - E(Y)) = (X - \mu_X)(Y - \mu_Y),$$

the expectation $E[g(X, Y)]$ leads to the covariance of X and Y .

DEFINITION 11 (COVARIANCE)

The *covariance* of X and Y is defined to be

$$\text{cov}(X, Y) = E[(X - E(X))(Y - E(Y))].$$

REMARK

If X and Y are discrete random variables,

$$\text{cov}(X, Y) = \sum_x \sum_y (x - \mu_X)(y - \mu_Y) f_{X,Y}(x, y).$$

If X and Y are continuous random variables,

$$\text{cov}(X, Y) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (x - \mu_X)(y - \mu_Y) f_{X,Y}(x, y) \, dx \, dy.$$

PROPERTIES OF THE COVARIANCE

The covariance has the following properties.

(1) $\text{cov}(X, Y) = E(XY) - E(X)E(Y).$

This is true because

$$\begin{aligned}\text{cov}(X, Y) &= E[(X - \mu_X)(Y - \mu_Y)] = E[XY - Y\mu_X - X\mu_Y + \mu_X\mu_Y] \\ &= E(XY) - \mu_X E(Y) - \mu_Y E(X) + \mu_X\mu_Y \\ &= E(XY) - \mu_X\mu_Y - \mu_Y\mu_X + \mu_X\mu_Y = E(XY) - \mu_X\mu_Y.\end{aligned}$$

(2) If X and Y are independent, then $\text{cov}(X, Y) = 0$. However, $\text{cov}(X, Y) = 0$ does not imply that X and Y are independent.

Take note that the two statements can be summarised as:

(i) $X \perp Y \Rightarrow \text{cov}(X, Y) = 0$;

(ii) $X \perp Y \not\Leftarrow \text{cov}(X, Y) = 0$.

For (i), note that if X and Y are independent, then $f_{X,Y}(x, y) = f_X(x)f_Y(y)$. So

$$\begin{aligned} E(XY) &= \sum_i \sum_j x_i y_j f_{X,Y}(x_i, y_j) = \sum_i \sum_j x_i y_j f_X(x_i) f_Y(y_j) \\ &= \sum_i x_i f_X(x_i) \sum_j y_j f_Y(y_j) = E(X)E(Y). \end{aligned}$$

$$(3) \operatorname{cov}(aX + b, cY + d) = ac \cdot \operatorname{cov}(X, Y).$$

This can be derived using the following 3 formulas:

$$(i) \operatorname{cov}(X, Y) = \operatorname{cov}(Y, X);$$

$$(ii) \operatorname{cov}(X + b, Y) = \operatorname{cov}(X, Y);$$

$$(iii) \operatorname{cov}(aX, Y) = a \operatorname{cov}(X, Y).$$

$$(4) V(aX + bY) = a^2V(X) + b^2V(Y) + 2ab \cdot \operatorname{cov}(X, Y).$$

This can be derived using the following 2 formulas:

$$(i) V(aX) = a^2V(X);$$

$$(ii) V(X + Y) = V(X) + V(Y) + 2\operatorname{cov}(X, Y).$$

EXAMPLE 3.7

We are given the joint distribution for (X, Y) :

x	y				$f_X(x)$
	0	1	2	3	
0	1/8	1/4	1/8	0	1/2
1	0	1/8	1/4	1/8	1/2
$f_Y(y)$	1/8	3/8	3/8	1/8	1

- (i) Find $E(Y - X)$.
- (ii) Find $\text{cov}(X, Y)$.

Solution:

(i) Method 1:

$$\begin{aligned} E(Y - X) &= (0 - 0)(1/8) + (1 - 0)(1/4) + (2 - 0)(1/8) \\ &\quad + \dots + (3 - 1)(1/8) = 1. \end{aligned}$$

Method 2:

$$E(Y - X) = E(Y) - E(X) = 1.5 - 0.5 = 1,$$

where

$$E(Y) = 0 \cdot (1/8) + 1 \cdot (3/8) + 2 \cdot (3/8) + 3 \cdot (1/8) = 1.5$$

$$E(X) = 0 \cdot (1/2) + 1 \cdot (1/2) = 0.5.$$

(ii) We use $\text{cov}(X, Y) = E(XY) - E(X)E(Y)$ to compute. Note that we have computed $E(X)$ and $E(Y)$ in Part (i).

$$\begin{aligned} E(XY) &= (0)(0)(1/8) + (0)(1)(1/4) + (0)(2)(1/8) \\ &\quad + \dots + (1)(3)(1/8) = 1. \end{aligned}$$

Therefore

$$\text{cov}(X, Y) = E(XY) - E(X)E(Y) = 1 - (0.5)(1.5) = 0.25.$$