

Expectation Maximization [3] for Fitting Mixture Models

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1. Introduction

To learn the MoG parameters [4] $\theta = \{\lambda_k, \mu_k, \Sigma_k\}_{k=1}^K$ from training data $\{x_i\}_{i=1}^I$ it applies the EM algorithm [5]. And it initializes the parameters randomly and alternate between performing the E- and M-steps.

In the E-step, it maximizes the bound with respect to the distributions $q_i(h_i)$ by finding the posterior probability distribution $\Pr(h_i | x_i)$ of each hidden variable h_i given the observation x_i and the current parameter settings in Equation 1.

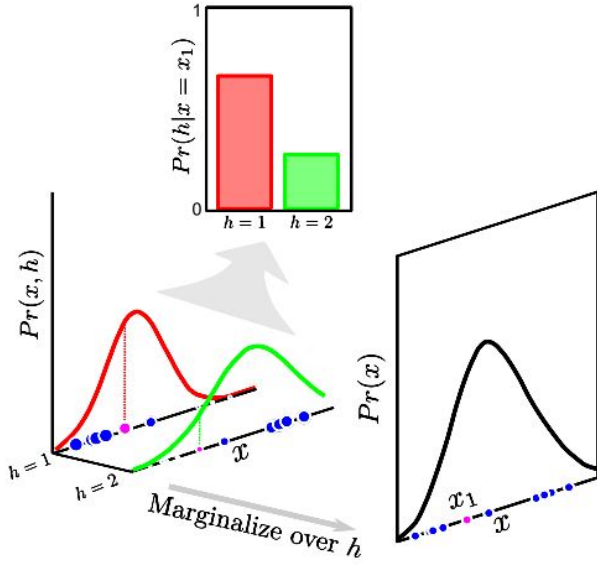


Figure 1. E-step for fitting the mixture of Gaussians model. For each of the I data points $x_1 \dots x_I$, it calculates the posterior distribution $\Pr(h_i | x_i)$ over the hidden variable h_i . The posterior probability [2] $\Pr(h_i = k | x_i)$ that h_i takes value k can be understood as the responsibility of normal distribution k for data point x_i . For example, for data point x_1 (magenta circle), component 1 (red curve) is more than twice as likely to be responsible than component 2 (green curve). Note that in the joint distribution (left), the size of the projected data point indicates the responsibility.

$$\begin{aligned} q_i(h_i) &= \Pr(h_i = k | x_i, \theta^{[t]}) \\ &= \frac{\Pr(x_i | h_i = k, \theta^{[t]}) \Pr(h_i = k, \theta^{[t]})}{\sum_{j=1}^K \Pr(x_i | h_i = j, \theta^{[t]}) \Pr(h_i = j, \theta^{[t]})} \\ &= \frac{\lambda_k \text{Norm}_{x_i}[\mu_k, \Sigma_k]}{\sum_{j=1}^K \lambda_j \text{Norm}_{x_i}[\mu_j, \Sigma_j]} \\ &= r_{ik} \end{aligned} \quad (1)$$

In other words it computes the probability $\Pr(h_i = k | x_i, \theta^{[t]})$ that the k^{th} normal distribution was responsible for the i^{th} data point (Figure 1). It denotes this responsibility by r_{ik} for short. In the M-step, it maximizes the bound with respect to the parameters $\theta = \{\lambda_k, \mu_k, \Sigma_k\}_{k=1}^K$ in Equation 2.

$$\begin{aligned} \hat{\theta}^{[t+1]} &= \underset{\theta}{\operatorname{argmax}} \left[\sum_{i=1}^I \sum_{k=1}^K \hat{q}_i(h_i = k) \log[\Pr(x_i, h_i = k | \theta)] \right] \\ &= \underset{\theta}{\operatorname{argmax}} \left[\sum_{i=1}^I \sum_{k=1}^K r_{ik} \log[\lambda_k \text{Norm}_{x_i}[\mu_k, \Sigma_k]] \right] \end{aligned} \quad (2)$$

2. Conclusions

This maximization can be performed by taking the derivative of the expression with respect to the parameters, equating the result to zero and rearranging, taking care to enforce the constraint $\sum_k \lambda_k = 1$ using Lagrange multipliers [1].

References

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