

31. Properties of Expectations

Covariance, Variance of Sums [Ross S7.4]

Proposition 31.1 *If X and Y are independent, then for any functions $g(x)$ and $h(y)$:*

- i) $E[g(X)h(Y)] = E[g(X)]E[h(Y)]$
- ii) $g(X)$ and $h(Y)$ are independent.

Why?

$$\begin{aligned} \text{i) } E[g(X)h(Y)] &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x)h(y)f_{XY}(x,y)dxdy \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x)h(y)f_X(x)f_Y(y)dxdy \\ &= \int_{-\infty}^{\infty} g(x)f_X(x)dx \int_{-\infty}^{\infty} h(y)f_Y(y)dy \\ &= E[h(Y)]E[g(X)] \end{aligned}$$

ii) Let $A = \{x \in \mathbb{R} \mid g(x) \leq a\}$ and $B = \{y \in \mathbb{R} \mid h(y) \leq b\}$. Then:

$$\begin{aligned} P[g(X) \leq a, h(Y) \leq b] &= P[X \in A, Y \in B] \\ &= P[X \in A] P[Y \in B] && \text{since } X \text{ and } Y \text{ are independent} \\ &= P[g(X) \leq a] P[h(Y) \leq b] \end{aligned}$$

For a single random variable X , its mean and variance give us some information about X .

For two random variables X and Y , **covariance** (and **correlation**) will give us information about the relationship between the pair X and Y .

Definition 31.1: The **covariance** between X and Y , denoted $Cov[X, Y]$, is defined as

$$Cov[X, Y] = E [(X - E[X])(Y - E[Y])]$$

Just as $Var[X] = E[X^2] - (E[X])^2$, we also have:

$$\begin{aligned} Cov[X, Y] &= E [(X - E[X])(Y - E[Y])] \\ &= E [XY - E[X]Y - E[Y]X + E[X]E[Y]] \\ &= E[XY] + E[-E[X]Y] + E[-E[Y]X] + E[E[X]E[Y]] \\ &= E[XY] - E[X]E[Y] - E[Y]E[X] + E[X]E[Y] \\ &= E[XY] - E[X]E[Y] \end{aligned}$$

Note: If X and Y are independent, then $E[XY] = E[X]E[Y]$
so $Cov[X, Y] = 0$.

Example 31.1: Does $Cov[X, Y] = 0$ imply X and Y are independent?

Solution:

Proposition 31.2

- i) $Cov[X, Y] = E[(X - \mu_X)(Y - \mu_Y)] = Cov[Y, X]$
- ii) $Cov[X, X] = E[(X - \mu_X)(X - \mu_X)] = Var[X]$
- iii) $Cov[aX, Y] = E[(aX - a\mu_X)(Y - \mu_Y)] = aCov[X, Y] = Cov[X, aY]$
- iv) $Cov\left[\sum_{i=1}^n X_i, \sum_{j=1}^m Y_j\right] = \sum_{i=1}^n \sum_{j=1}^m Cov[X_i, Y_j]$

Why?

For iv), let $U = \sum_{i=1}^n X_i$ $V = \sum_{j=1}^m Y_j$

$$E[X_i] = \mu_i \qquad E[Y_j] = \nu_j$$

Then: $E[U] = \sum_{i=1}^n \mu_i$ $E[V] = \sum_{j=1}^m \nu_j$

$$\begin{aligned} \text{So, } Cov[U, V] &= E\left[\left(\sum_{i=1}^n X_i - \sum_{i=1}^n \mu_i\right)\left(\sum_{j=1}^m Y_j - \sum_{j=1}^m \nu_j\right)\right] \\ &= E\left[\sum_{i=1}^n (X_i - \mu_i) \sum_{j=1}^m (Y_j - \nu_j)\right] \end{aligned}$$

$$\begin{aligned}
&= E \left[\sum_{i=1}^n \sum_{j=1}^m (X_i - \mu_i)(Y_j - \nu_j) \right] \\
&= \sum_{i=1}^n \sum_{j=1}^m E [(X_i - \mu_i)(Y_j - \nu_j)] \\
&= \sum_{i=1}^n \sum_{j=1}^m Cov[X_i, Y_j]
\end{aligned}$$

Now,

$$\begin{aligned}
Var \left[\sum_{i=1}^n X_i \right] &= Cov \left[\sum_{i=1}^n X_i, \sum_{j=1}^n X_j \right] && \text{by ii) of Prop. 31.2} \\
&= \sum_{i=1}^n \sum_{j=1}^n Cov[X_i, X_j] && \text{by iv) of Prop. 31.2} \\
&= \sum_{\substack{i,j \\ j=i}} Cov[X_i, X_j] + \sum_{\substack{i,j \\ j \neq i}} Cov[X_i, X_j] \\
&= \sum_{i=1}^n Var[X_i] + \sum_{\substack{i,j \\ j \neq i}} Cov[X_i, X_j] && \text{by ii) of Prop. 31.2} \\
&= \sum_{i=1}^n Var[X_i] + 2 \sum_{\substack{i,j \\ i < j}} Cov[X_i, X_j] && \text{by i) of Prop. 31.2}
\end{aligned}$$

If for $i \neq j$, each pair X_i, X_j are independent, then:

$$Var \left[\sum_{i=1}^n X_i \right] = \sum_{i=1}^n Var[X_i]$$

Example 31.2: Compute the variance of $X \sim \text{Binomial}(n, p)$.

Solution:

Example 31.3: Recall (from Example 30.3) that $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$ is called the **sample mean**. Let

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$$

be the **sample variance**.

Let X_1, \dots, X_n be iid with (common) mean μ and variance σ^2 .

Find a) $\text{Var}[\bar{X}]$ and b) $E[S^2]$. [b) is hard]

Solution:

Note: Since $E[S^2] = \sigma^2$ then S^2 is an unbiased estimator of variance.

Example 31.4: Let X_1, \dots, X_n be iid with variance σ^2 . Recall $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$ is the sample mean. $X_i - \bar{X}$ is called the ***i*th deviation**.

Compute $Cov[X_i - \bar{X}, \bar{X}]$.

Solution: