Econometrics 720

Logistic Regression

The logistic regression model arises from the desire to model posterior probabilities of K classes via linear functions in x, while at the same time ensuring that they sum to one and remain in the interval [0,1].

A model that complies to the above is:

$$log \frac{Pr(G=1|X=x)}{Pr(G=K|X=x)} = \beta_{10} + \beta_1^T x$$

$$log(odds_1) = \beta_{10} + \beta_1^T x \text{ where } odds_1 = \frac{Pr(G=1|X=x)}{Pr(G=K|X=x)}$$

$$odds_1 = e^{\beta_{10} + \beta_1^T x}$$

$$log \frac{Pr(G=2|X=x)}{Pr(G=K|X=x)} = \beta_{20} + \beta_2^T x$$

$$log(odds_2) = \beta_{20} + \beta_2^T x \text{ where } odds_2 = \frac{Pr(G=1|X=x)}{Pr(G=K|X=x)}$$

$$odds_2 = e^{\beta_{20} + \beta_2^T x}$$

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$$log \frac{Pr(G = K - 1|X = x)}{Pr(G = K|X = x)} = \beta_{(K-1)0} + \beta_{K-1}^{T} x$$

$$log(odds_{K-1}) = \beta_{(K-1)0} + \beta_{K-1}^{T} x \text{ where } odds_{K-1} = \frac{Pr(G = K - 1|X = x)}{Pr(G = K|X = x)}$$

$$odds_{K-1} = e^{\beta_{(K-1)0} + \beta_{K-1}^{T} x}$$

The above imply the use of linear decision boundaries, based on the CDF of a logistic distribution.

Consider the following:

$$\begin{split} \Sigma_{l=1}^{K-1}log(odds_{l}) &= \Sigma_{l=1}^{K-1}\beta_{l0} + \beta_{l}^{T}x \\ \frac{1}{Pr(G=K|X=x)}\Sigma_{l=1}^{K-1}Pr(G=l|X=x) &= \Sigma_{l=1}^{K-1}e^{\beta_{l0}+\beta_{l}^{T}x} \\ \frac{1}{Pr(G=K|X=x)}\left(1 - Pr(G=K|X=x)\right) &= \Sigma_{l=1}^{K-1}e^{\beta_{l0}+\beta_{l}^{T}x} \\ \frac{1}{Pr(G=K|X=x)} &= 1 + \Sigma_{l=1}^{K-1}e^{\beta_{l0}+\beta_{l}^{T}x} \\ Pr(G=K|X=x) &= \frac{1}{1 + \Sigma_{l=1}^{K-1}e^{\beta_{l0}+\beta_{l}^{T}x}} \\ Pr(G=k|X=x) &= e^{\beta_{k0}+\beta_{k}^{T}x}Pr(G=K|X=x) \\ Pr(G=k|X=x) &= e^{\beta_{k0}+\beta_{k}^{T}x}\frac{1}{1 + \Sigma_{l=1}^{K-1}e^{\beta_{l0}+\beta_{l}^{T}x}} \end{split}$$

Clearly $\Sigma_{l=1}^K Pr(G=l|X=x)$ is equal to 1, and all probabilities depend on the full parameter set $\theta = \{\beta_{10}, \beta_1, \beta_{20}, \beta_2, \dots, \beta_{(K-1)0}, \beta_{K-1}\}$. These probabilities are denoted by $Pr(G=k|x=x) = p_k(x;\theta)$

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Estimation of Logistic Regression models using Newton-Raphson

Objective: Estimate the parameters that maximize the conditional likelihood of G given X, using the modelling/training data.

The conditional log-likelihood function

- Denote $p_k(x_i; \theta) = Pr(G = k | X = X_i; \theta)$.
- Given the first input x_1 , the posterior probability if its class being g_1 is $Pr(G = g_1|X = x_1)$.
- Assuming independence of observations, the posterior probability for the N observations each having class $g_i = 1, 2, ..., N$, given their inputs $x_1, x_2, ..., x_N$ is:

$$\prod_{i=1}^{N} Pr(G = g_i | X = x_i)$$

and therefore the conditional log likelihood is given as

$$\begin{split} l(\theta) &= & \Sigma_{i=1}^{N} log Pr(G=g_{i}|X=x_{i}) \\ &= & \Sigma_{i=1}^{N} log p_{g_{i}}(x_{i};\theta) \end{split}$$

where $p_{g_i}(x_i;\theta)$ is the probability of being in the group g that is associated with the i'th observation.

• Consider only cases where K = 2, binary classification. The log-likelihood function can be simplified by using the following:

$$y_i$$
 = 1 when g_i = 1
 y_i = 0 when g_i = 2
 $p_1(x;\theta)$ = $p(x;\theta)$
 $p_2(x;\theta)$ = 1 - $p(x;\theta)$ for two groups

- Since K = 2, the parameters $\theta = \beta_{10}, \beta_1$. Denote $\beta = (\beta_{10}, \beta_1)'$
- If $y_i = 1$, i.e, $g_i = 1$,

$$log p_{g_i}(x; \beta) = log p_1(x; \beta)$$
$$= 1.log p_1(x; \beta)$$
$$= y_i log p_1(x; \beta)$$

If $y_i = 0$, i.e, $g_i = 0$,

$$log p_{g_i}(x; \beta) = log p_2(x; \beta)$$

$$= 1.log(1 - p_1(x; \beta))$$

$$= (1 - y_i)log(1 - p_1(x; \beta))$$

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Since
$$y_i = 0$$
 or $(1 - y_i) = 0$ we have
$$log p_{g_i}(x; \beta) = y_i log p_1(x; \beta) + (1 - y_i) log (1 - p_1(x; \beta))$$

• The conditional log-likelihood function then is:

$$\begin{split} l(\beta) &= & \sum_{i=1}^{N} \left[y_i \log p(x_i, \beta) + (1 - y_i) \log (1 - p(x_i; \beta)) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \log p(x_i, \beta) + \log (1 - p(x_i; \beta)) - y_i \log (1 - p(x_i; \beta)) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \log \frac{p(x_i, \beta)}{(1 - p(x_i; \beta)} + \log (1 - p(x_i; \beta)) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i + \log (1 - \frac{e^{\beta^T x_i}}{1 + e^{\beta^T x_i}}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i + \log \left(\frac{1}{1 + e^{\beta^T x_i}} \right) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= & \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right] \\ &= \sum_{i=1}^{N} \left[y_i \beta^T x_i - \log (1 + e^{\beta^T x_i}) \right]$$

Maximum Likelihood Estimation

In order to maximise the log-likelihood function we set its derivatives equal to zero.

$$\frac{\partial l(\beta)}{\partial \beta} = \sum_{i=1}^{N} x_i (y_i - p(x_i; \beta)) = 0$$

which are p + 1 score equations non-linear in β .

In order to solve the score equations , we use the Newton-Raphson algorithm, which requires the second derivative or Hessian matrix

$$\frac{\partial^2 l(\beta)}{\partial \beta \partial \beta^t} = -\sum_{i=1}^N x_i x_i^T p(x_i; \beta) (1 - p(x_i; \beta))$$

The Newton-Raphson algorithm relates to the following update formula:

$$\beta^{new} = \beta^{old} - \left(\frac{\partial^2 l(\beta)}{\partial \beta \partial \beta^t}\right)^{-1} \frac{\partial l(\beta)}{\partial \beta}$$

where the derivatives are evaluated at β^{old} .

In matrix notation we get:

$$\begin{array}{lcl} \frac{\partial l(\beta)}{\partial \beta} & = & X^T(y-p) \\ \frac{\partial^2 l(\beta)}{\partial \beta \partial \beta^t} & = & -X^T \, W \, X \end{array}$$

y the vector of y_i values

X the $N \times (p+1)$ matrix of x_i values with

p the vector of fitted probabilities with i'th element $p(x_i; \beta^{old})$

W a $N \times N$ diagonal matrix of weights with i'th diagonal element $p(x_i; \beta^{old})(1 - p(x_i; \beta^{old}))$

The Newton-Raphson update formula is:

$$\begin{split} \beta^{new} & = & \beta^{old} + (X^T W X)^{-1} X^T (y-p) \\ & = & (X^T W X)^{-1} (X^T W X) \beta^{old} + (X^T W X)^{-1} X^T (y-p) \\ & = & (X^T W X)^{-1} X^T W (X \beta^{old} + W^{-1} (y-p)) \\ & = & (X^T W X)^{-1} X^T W z \end{split}$$

where

$$z = (X\beta^{old} + W^{-1}(y-p))$$

The vector z is sometimes called the adjusted response, and the equations are solved repeatedly.