

Inferential Statistics

L2 - Descriptive statistics and statistical models

Erlis Ruli (erlis.ruli@unipd.it)

Department of Statistics, University of Padova

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The samples

The data collected in an experiment consist of observations x_1, x_2, \dots, x_n on a variable of interest, which are then used to learn about the data-generating mechanism.

The list x_1, \dots, x_n is called the observed sample and n is called the sample size.

We assume that x_1, \dots, x_n is a realisation of the random sample X_1, \dots, X_n , with X_i assumed mutually independent with equal marginal pdf f .

The distinction between observed and random sample is much like the difference between a measured voltage (observed) and the voltmeter.

By the definition of independence, the joint pdf of the sample is

$$\underline{f(x_1, \dots, x_n; \theta) = \prod_{i=1}^n f(x_i; \theta),}$$

where $f(x_i; \theta)$ is the density for X_i which depends on some unknown parameter θ .

Example 1

Let X_1, \dots, X_n be a random sample from the population $\text{Exp}(1/\beta)$. X_i may be time (years) until failure for n identical circuit boards put to test.

The joint pdf is

$$\underline{f(x_1, \dots, x_n; \beta) = \prod_{i=1}^n f(x_i; \beta) = \prod_{i=1}^n \frac{1}{\beta} e^{-x_i/\beta} = \frac{1}{\beta^n} e^{-(x_1 + \dots + x_n)/\beta}}$$

We could use this to compute, say the probability that all boards last at least 5 years:

$$\underline{P(X_1 \geq 5, \dots, X_n \geq 5) = \int_5^\infty \cdots \int_5^\infty \prod_{i=1}^n \frac{1}{\beta} e^{(x_i/\beta)} dx_1 \cdots dx_n = e^{-5n/\beta}.}$$

Summary statistics

Typically we are interested at some function of the sample. These are called descriptive or summary statistics.

Some examples are moment-based statistics:

- sample average $\bar{X} = \frac{1}{n} \sum_i X_i$ and the observed counterpart $\bar{x} = \frac{1}{n} \sum_i x_i$
- sample variance $S^2 = \frac{1}{n-1} \sum_i (X_i - \bar{X})^2$ and the observed counterpart $s^2 = \frac{1}{n-1} \sum_i (x_i - \bar{x})^2$; s is commonly called standard deviation.
- sample k th moment $\bar{X}^k = \frac{1}{n} \sum_i X_i^k$ and the observed counterpart.

Order statistics

Let $X_{(1)} = \min_{1 \leq i \leq n} X_i$ be the smallest observation, $X_{(2)}$ be the second smallest and so on $X_{(n)} = \max_{1 \leq i \leq n} X_i$.

The list $X_{(1)}, \dots, X_{(n)}$ is called order statistics, and are the basis of the following summary statistics

- the median Q_2 =
$$\begin{cases} X_{(\frac{n+1}{2})} & \text{if } n \text{ is odd} \\ (X_{(n/2)} + X_{(n/2+1)})/2 & \text{if } n \text{ is even} \end{cases}$$
- the first and third quartile, $Q_1 = X_{[0.25(n+1)]}$ and $Q_3 = X_{[0.75(n+1)]}$, resp
- the p th sample quantile, $p \in (0, 1)$ is $X_{[p(n+1)]}$
- inter quartile range $IQR = Q_3 - Q_1$
- median absolute deviation from the median (MAD) =
 $\text{median}(|X_1 - Q_2|, \dots, |X_n - Q_2|)$

and their observed counterparts; $[x]$ is the greatest integer $\leq x$.

Uses and relations

The above summary statistics serve different purposes:

\bar{X} , Q_1 , Q_2 , Q_3 , $X_{[p(n+1)]}$ are measures of location and are used when we want to provide a single typical value of the sample

S^2 , S , MAD , IQR are measures of spread, useful when we want to describe the variability of the sample.

→ "non-simmetry"
→ length of the "tail"

You might heard about skewness, kurtosis. These are additional features of the shape of distribution/sample.

Sample measures target their population counterparts, e.g. \bar{X} for μ_X , Q_2 for $\xi_{0.5}$, S^2 , MAD for σ^2 , etc.

Example 2

Suppose the sample of size is $n = 12$ and the 0.65th quantile is wanted. Then $[0.65 \cdot (12 + 1)] = 8$, so the 0.65th quantile is $X_{(8)}$. The answer would have been the same if wanted the 0.69th quantile.

(Two different examples)

↳ nice

Consider now the observed sample 1.1, 0.5, 0.4, 3, 2.2, so $x_1 = 1.1$, $x_2 = 0.5$ and so on. The observed order statistics are

$$\underline{x_{(1)} = 0.4, x_{(2)} = 0.5, x_{(3)} = 1.1, x_{(4)} = 2.2, x_{(5)} = 3.}$$

We find that $\bar{x} = 1.44$, $s^2 = 1.273$, $q_1 = 0.4$, $q_2 = 1.1$, $q_3 = 2.2$, and $\text{mad} = 1.03782$.

Histogram

Useful when we want to get an idea of the pdf of a sample.

Let x_1, \dots, x_n be the observed sample and consider a partition in intervals $(a_{j-1}, a_j]$, $j = 1, \dots, m$, with $m < n$ covering the sample.

Defined by the piecewise function

$$\underline{h_n(x) = \frac{1}{n(a_j - a_{j-1})} \sum_{i=1}^n \mathbf{1}_{(a_{j-1}, a_j]}(x_i), \quad \text{for all } x \in (a_{j-1}, a_j].}$$

Typically, $(a_{j-1}, a_j]$ are equal-length intervals and $m = 2 \text{ iqr} / n^{1/3}$
(Friedman-Diaconis rule). The $h_n(x)$ thus targets $f(x)$, the population
pdf, i.e. the pdf from which the observations come from.

Empirical distribution function

converges faster
→ than the histogram
to the true pdf

Given the random sample (rs) X_1, \dots, X_n , the edf is defined by

$$\underline{F_n(x) = \sum_{i=1}^n I_{X_i}(x)}, \quad \text{for all } x \in \mathbb{R},$$

where $I_{X_i}(x)$ is Bernoulli rv with success probability $P(X_i \leq x)$. For each x , F_n is thus a random variable.

The corresponding observed version is

$$\underline{\hat{F}_n(x) = n^{-1} \sum_{i=1}^n \mathbf{1}_{x_i}(x)}, \quad \text{for all } x \in \mathbb{R},$$

$\mathbf{1}_{x_i}(x)$ takes value 1 if $x_i \leq x$ and 0 otherwise.

F_n and its observed version \hat{F}_n target $F(x)$, the population df.

Example 3

Compute \hat{F}_n from the observed sample 1.1, 0.5, 0.3, 1.1, 5.

First, we have to get the sorted list, which is 0.4, 0.5, 1.1, 1.1, 5. Then we observe that

- for $-\infty < x < 0.4$ there are no observations, so $\sum_i \mathbf{1}_{x_i}(x) = 0$
- for $0.4 \leq x < 0.5$ there is only one observation, so $\sum_i \mathbf{1}_{x_i}(x) = 1$
- and so on,
- for $1.1 \leq x < 5$ there are two observations, so $\sum_i \mathbf{1}_{x_i}(x) = 2$.

Hence

$$\hat{F}_n(x) = \begin{cases} 0 & \text{if } x < 0.4 \\ 1/5 & \text{if } 0.4 \leq x < 0.5 \\ 2/5 & \text{if } 0.5 \leq x < 1.1 \\ 4/5 & \text{if } 1.1 \leq x < 5 \\ 1 & \text{if } 5 \leq x. \end{cases}$$