In [1]: %matplotlib inline import numpy as np import pandas as pd import matplotlib.pyplot as plt import statsmodels.formula.api as sm import statsmodels.api as sm2 import statistics as stat from numpy.random import normal from numba import njit import seaborn as sns import pylab from scipy import optimize as opt from scipy import stats as st from statsmodels.tsa.arima process import ArmaProcess from statsmodels.graphics.tsaplots import plot pacf,plot acf Probelm 1 **OLS Implementation:** In [2]: #load data df=pd.read_csv("https://raw.githubusercontent.com/dompazz/FinTech590-RiskManagement/main/Week02/Project/problem In [3]: Out[3]: У -1.166289 1.014680 **1** -0.426878 0.262715 -1.477892 -1.044772 3.049119 0.804363 -2.123732 -0.689514 **95** -0.588599 0.652704 -0.218138 0.067676 0.342822 1.214472 0.337376 0.608974 1.153817 -0.683444 100 rows × 2 columns In [4]: x=df.xy=df.y In [5]: # fit ols model (Y respect to X) result = sm.ols(formula="y ~ x", data=df).fit() In [6]: print(result.params) 0.037877 Intercept 0.428004 dtype: float64 In [7]: # print the output from fitting OLS of Y respect to X print(result.summary()) Dep. Variable: y R-squared: 0.268
Model: OLS Adj. R-squared: 0.261
Method: Least Squares F-statistic: 35.89
Date: Sat, 15 Jan 2022 Prob (F-statistic): 3.47e-08
Time: 06:34:38 Log-Likelihood: -120.46
No. Observations: 100 AIC: 244.9
Df Residuals: 98 BIC:
Df Model: OLS Regression Results 98 BIC: 1 Df Model: Covariance Type: nonrobust coef std err t P>|t| [0.025 0.975] Intercept 0.0379 0.082 0.461 0.646 -0.125 0.201 x 0.4280 0.071 5.990 0.000 0.286 0.570 _____ Omnibus: 5.101 Durbin-Watson:
0.078 Jarque-Bera (JB): Prob(Omnibus): 0.145 Prob(JB): 2.246 Cond. No. Kurtosis: 1.20 ______ Warnings: [1] Standard Errors assume that the covariance matrix of the errors is correctly specified. In [8]: # calculate the expected y, given x=5y h=0.037877+5* 0.428004 var = np.sum(result.resid**2)/result.df resid print("E[y|x=5]=", y_h ,"\n") print("Var= ",var,"\n") E[y|x=5] = 2.1778969999999999Var= 0.664670142845936 Expected Y ,given X=5, is 2.1778969999999997 from OLS, Variance = 0.664670142845936 **Conditional Distribution of Multivariate Normal Implementation:** In [9]: @njit **def** $f(z, \mu, \Sigma)$: The density function of multivariate normal distribution. Parameters z: ndarray(float, dim=2) random vector, N by 1 μ: ndarray(float, dim=1 or 2) the mean of z, N by 1Σ: ndarray(float, dim=2) the covarianece matrix of z, N by 1 z = np.atleast 2d(z) $\mu = np.atleast 2d(\mu)$ Σ = np.atleast 2d(Σ) N = z.sizetemp1 = np.linalg.det(Σ) ** (-1/2) temp2 = np.exp(-.5 * (z - μ).T @ np.linalg.inv(Σ) @ (z - μ)) return (2 * np.pi) ** (-N/2) * temp1 * temp2 In [10]: class MultivariateNormal: Class of multivariate normal distribution. Parameters μ: ndarray(float, dim=1) the mean of z, N by 1 Σ: ndarray(float, dim=2) the covarianece matrix of z, N by 1 Arguments μ, Σ: see parameters μs: list(ndarray(float, dim=1)) list of mean vectors $\mu 1$ and $\mu 2$ in order Σs: list(list(ndarray(float, dim=2))) 2 dimensional list of covariance matrices $\Sigma 11$, $\Sigma 12$, $\Sigma 21$, $\Sigma 22$ in order βs: list(ndarray(float, dim=1)) list of regression coefficients $\beta1$ and $\beta2$ in order **def** __init__(self, μ , Σ): "initialization" $self.\mu = np.array(\mu)$ $self.\Sigma = np.atleast 2d(\Sigma)$ def partition(self, k): 11 11 11 Given k, partition the random vector z into a size k vector z1 and a size N-k vector z2. Partition the mean vector $\boldsymbol{\mu}$ into $\mu 1$ and $\mu 2$, and the covariance matrix Σ into $\Sigma 11$, $\Sigma 12$, $\Sigma 21$, $\Sigma 22$ correspondingly. Compute the regression coefficients $\beta1$ and $\beta2$ using the partitioned arrays. $\mu = self.\mu$ $\Sigma = \text{self.}\Sigma$ self. μ s = [μ [:k], μ [k:]] self. Σ s = [[Σ [:k, :k], Σ [:k, k:]], $[\Sigma[k:, :k], \Sigma[k:, k:]]$ self. β s = [self. Σ s[0][1] @ np.linalg.inv(self. Σ s[1][1]), $self.\Sigma s[1][0]$ @ np.linalg.inv(self. $\Sigma s[0][0]$)] def cond_dist(self, ind, z): Compute the conditional distribution of z1 given z2, or reversely. Argument ind determines whether we compute the conditional distribution of z1 (ind=0) or z2 (ind=1). Returns μ_hat: ndarray(float, ndim=1) The conditional mean of z1 or z2. Σ hat: ndarray(float, ndim=2) The conditional covariance matrix of z1 or z2. $\beta = self.\beta s[ind]$ $\mu s = self.\mu s$ $\Sigma s = self.\Sigma s$ $\mu_{hat} = \mu s[ind] + \beta @ (z - \mu s[1-ind])$ Σ hat = Σ s[ind][ind] - β @ Σ s[1-ind][1-ind] @ β .T **return** μ _hat, Σ _hat In [11]: # covariance matrix cov=df.cov() Out[11]: У 1.315195 0.562908 **y** 0.562908 0.898883 In [12]: xx = cov.x[0]xy=cov.x[1]yx = cov.y[0]yy=cov.y[1]In [13]: #set up mu and sigma $\mu = \text{np.array}([\text{stat.mean}(x), \text{stat.mean}(y)])$ $\Sigma = \text{np.array}([[xx, yx], [xy,yy]])$ In [14]: #set object of MultivariateNormal class $multi normal = MultivariateNormal(\mu, \Sigma)$ In [15]: k = 1 # choose partition # partition and compute regression coefficients multi normal.partition(k) multi normal.βs[1] array([[0.42800371]]) Out[15]: In [16]: # compute the cond. dist. of y ind = 1 $y_1 = np.array([5]) # given x=5$ μ1_hat, Σ1_hat = multi_normal.cond_dist(ind, y_1) print(' μ 1_hat, Σ 1_hat = ', μ 1_hat, Σ 1_hat) $\mu1_{hat}$, $\Sigma1_{hat} = [2.17789539] [[0.6579563]]$ Conditional expectation of Y, given X=5, is 2.17789539 from Conditional Distribution of Multivariate Normal, the variance is 0.6579563. Conclusion: Since 2.17789539 ≈ 2.177896999999997, we can conclude that the OLS equation and the Conditional distribution equation are in fact the same thing. Although their variance is not exactly the same(but very close): 0.6579563 vs.0.664670142845936 Problem 2 (1) In [17]: #load data from problem2.csv df2=pd.read_csv("https://raw.githubusercontent.com/dompazz/FinTech590-RiskManagement/main/Week02/Project/proble In [18]: df2 Out[18]: Х У **o** -1.614399 -1.695691 **1** -0.900999 0.409843 **2** -0.170662 1.043979 **3** 2.097252 2.708814 4 0.140208 0.052374 **95** -1.115215 -2.145361 **96** -0.564690 -1.916765 **97** -1.098674 -0.110209 **98** -0.562357 0.181756 99 1.044383 -0.170417 100 rows × 2 columns In [19]: x=df2.xy=df2.y In [20]: # fit ols model result = sm.ols(formula="y ~ x", data=df2).fit() In [21]: print(result.summary()) OLS Regression Results ______ Dep. Variable: y R-squared: OLS Adj. R-squared: Model: 0.186 Least Squares F-statistic: Method: Date: 23.68 Sat, 15 Jan 2022 Prob (F-statistic): 06:34:47 Log-Likelihood: 4.34e-06 -159.99No. Observations: 100 AIC: 324.0 Df Residuals: 98 BIC: 329.2 Df Model: 1 Covariance Type: nonrobust ______ coef std err t P>|t| [0.025 0.975] Intercept 0.1198 0.121 0.990 0.325 -0.120 x 0.6052 0.124 4.867 0.000 0.358 0.852 _____ 14.146 Durbin-Watson: 1.885 Omnibus: 43.673 0.001 Jarque-Bera (JB): Prob(Omnibus): 3.28e-10 -0.267 Prob(JB): Skew: 6.193 Cond. No. Kurtosis: 1.03 ______ [1] Standard Errors assume that the covariance matrix of the errors is correctly specified. In [22]: # list the error vector result.resid Out[22]: 0 -0.838485 1 0.835296 2 1.027428 3 1.319711 4 -0.152317 . . . 95 -1.590264 96 -1.694848 97 0.434878 98 0.402261 -0.922319 Length: 100, dtype: float64 In [23]: # plot the density curve of error vector sns.distplot(result.resid, hist=True, kde=True, bins=int(180/5), color = 'darkblue', hist kws={'edgecolor':'black'}, kde kws={'linewidth': 4}) <matplotlib.axes. subplots.AxesSubplot at 0x7fbea80d5550> Out[23]: 0.6 0.5 0.4 0.3 0.2 0.1 Implication from the result above: The density curve above is clearly not a normal distribution, especially on the tails, and the top of the bell shaped curve is on the right hand side of the mean. In [24]: # test the 68-95-99.7 rule for the density curve of error vector def normal percent(mean, std): count=0; for r in result.resid : if r>(mean-std) and r<(mean+std):</pre> count=count+1 return count std=np.std(result.resid) mean=np.mean(result.resid) print("68-95-99.7 rule:\n") print(normal percent(mean,std), normal percent(mean,std*2),normal percent(mean,std*3)) 68-95-99.7 rule: 76 97 98 Implication from the result above: from the block above, we see that the density curve of the error vector did not match the 68-95-99.7 rule well. Too much data in the center and too little at the end. In [25]: # plot the qq plot for error vector sm2.qqplot(x, line='45')pylab.show() 2 1 nple Quantiles San -2Theoretical Quantiles Implication from the result above: From the qq plot above, we see that the error vector is not normally distributed, especially at the tails In [26]: # shapiro wilk test shapiro test = st.shapiro(result.resid) shapiro_test ShapiroResult(statistic=0.938385546207428, pvalue=0.00015389148029498756) Out[26]: Implication from the result above: Null Hypothesis: Error vectors are normally distributed. Let Alpha = 0.05. Since the p value from the shaprio wilk test is 0.00015389148029498756, which is smaller than 0.05, the null hypothesis is rejected and the error vector is probably ot normally distributed. **Final Conclusion:** From the density curve, the 68-95-99.7 rule, the qq plot, and the shaprio wilk test that are described above, we can see that all of these indicates that the error vector is not normally distributed. However, it is roughly close to normal. (2) & (3)In [27]: # Fit the mle with normal assup def mle_norm(param_vec): $b_0 = param_vec[0]$ $b_1 = param_vec[1]$ yhat=b_0+b_1*x l= -np.sum(np.log(st.norm.pdf(y - yhat))) return 1 # Fit the mle with t assup def mle t(param vec): $b_0 = param_vec[0]$ $b_1 = param_{vec}[1]$ yhat=b_0+b_1*x l = -np.sum(np.log(st.t.pdf(y - yhat, len(x)-2)))return 1 #optimization(minimization) model_norm = opt.minimize(mle_norm, np.array([1, 1])) model_t= opt.minimize(mle_t, np.array([1, 1])) # calculate the sse for mle_norm and mle_t e_norm=y-model_norm.x[0]-model_norm.x[1]*x e t= y-model t.x[0]-model t.x[1]*x sse_norm=sum(e_norm*e_norm) sse_t=sum(e_t*e_t) print("The optimized parameters generated from MIE with assumption of normality:") print(" $\beta 0 = \text{",model_norm.x[0],"} \beta 1 = \text{",model_norm.x[1],"}$ "," sse= ",sse norm,"\n") print("The optimized parameters generated from MIE with assumption of T distributin:") $\beta 0 = \text{",model_t.x[0],"} \beta 1 = \text{",model_t.x[1],"}n","$ The optimized parameters generated from MlE with assumption of normality: $\beta 0 = 0.11983615555152172$ $\beta 1 = 0.6052048075913198$ 143.61484854062627 The optimized parameters generated from MIE with assumption of T distributin: $\beta 0 = 0.12325309647619272$ $\beta 1 = 0.5951244996627715$ 143.6256458563721 Conclusion 2 and 3 combined: Since the sse from MLE_t is 143.6256458563721, which is larger than the sse of 143.61484854062627 from MLE_normal, we know that the using MLE given the assumption of normality has the best fit. The fitted parameters of each are shown in the block above, the parameters from the mle_norm model and mle_t model is close to each other but not the same. We notice that the parameters generated from MLE with assumpotion of normality is the same with those from OLS. The differences of the parameters from two models and the fact that the parameters generated from MLE with assumpotion of normality is the same with those from OLS indicates that if we break the normality assumption, the value of the estimated parameters would be different and the expected value of y would also be different **Problem 3** (1) In [28]: #fit AR process and draw the simulation, ACF and PACF def ar (p, titles): fig, axes=plt.subplots(3,3,figsize=(20,15)) for i in range(3): process = ArmaProcess(ar = p[i]) data = process.generate sample(nsample=1000) axes[i][0].plot(data) axes[i][0].set title(titles[i]) fig = plot acf(data, lags=25, ax=axes[i][1]) fig = plot_pacf(data, lags=25, ax=axes[i][2]) ar([[1,-0.3],[1,-0.3,-0.3],[1,-0.3,0.3,0.3]),["AR(1)","AR(2)","AR(3)"])Autocorrelation Partial Autocorrelation 1.0 1.0 0.8 0.8 0.6 0.6 0.4 0.4 0.2 0.2 -2 0.0 0.0 -3 Autocorrelation Partial Autocorrelation 0.8 0.8 0.6 0.6 0.4 0.4 0.2 0.2 0.0 200 600 Autocorrelation AR(3) Partial Autocorrelation 1.0 1.0 0.6 0.6 0.4 0.4 0.2 0.2 0.0 0.0 -0.2 -0.2-0.4-0.4 Conclusion 1: From the graph above we find that only the first lag of the PACF of AR(1) is significantly different from 0; the first and the second lag of the PACF of AR(2) are significantly different from 0; the first, second, and the third lag of AR(3) are significantly different from 0.Therefore the PACF of AR process can be used to determine the order. On the other hand, there is no clear pattern for ACF. This feature can also help us to identify if the process is AR or not. (2)In [29]: #fit MA process and draw the simulation, ACF and PACF def ma (p, titles): fig, axes=plt.subplots(3,3,figsize=(20,15)) for i in range(3): process = ArmaProcess(ma = p[i]) data = process.generate sample(nsample=1000) axes[i][0].plot(data) axes[i][0].set_title(titles[i]) fig = plot_acf(data, lags=25, ax=axes[i][1]) fig = plot pacf(data, lags=25, ax=axes[i][2]) ma([[1,-0.3],[1,-0.3,-0.3],[1,-0.3,0.3,0.3]),["MA(1)","MA(2)","MA(3)"])Partial Autocorrelation Autocorrelation 1.0 1.0 0.8 0.8 0.6 0.4 0.4 0.2 0.2 -2 0.0 0.0 -3 -0.2 -0.2 600 1000 15 20 25 25 MA(2) Autocorrelation Partial Autocorrelation 1.0 1.0 0.8 0.8 0.6 0.6 0.4 0.2 0.2 -10.0 0.0 -0.2 -0.2 -3 15 Partial Autocorrelation Autocorrelation 1.0 1.0 0.8 0.8 0.6 0.6 0.4 Conclusion 2: From the graph above we find that only first lag of the PACF of MA(1) is significantly different from 0; the first and the second lag of the ACF of MA(2) are significantly different from 0; the first, second, and the third lag of MA(3) are significantly different from 0. Therefore the ACF of MA process can be used to determine the order. On the other hand, there is no clear pattern for PACF of MA process. Just like from conclusion 1, this feature can also help us to identify if the process is MA or not. Final Conclusion(Identify type and order): So when there is no clear pattern of the ACF but the first n lags of the PACF is significantly different from 0 and the significance of the rest of the lags decrease intensely, then this might be an AR(n) process. On the other hand, when there is no clear pattern of the PACF but the first n lags of the ACF is significantly different from 0 and the significance of the rest of the lags decrease intensely, then this might be a MA(n) process.