

SGA2 10

Question 1

Assume that continuous random variable X has PDF that is non-zero only on segment $[a, b]$ and strictly positive on open interval (a, b) . Median of random variable X is value $m \in (a, b)$ such that $P(X < m) = P(X > m) = 1/2$.

Prove that for symmetrical PDFs median is equal to expected value.

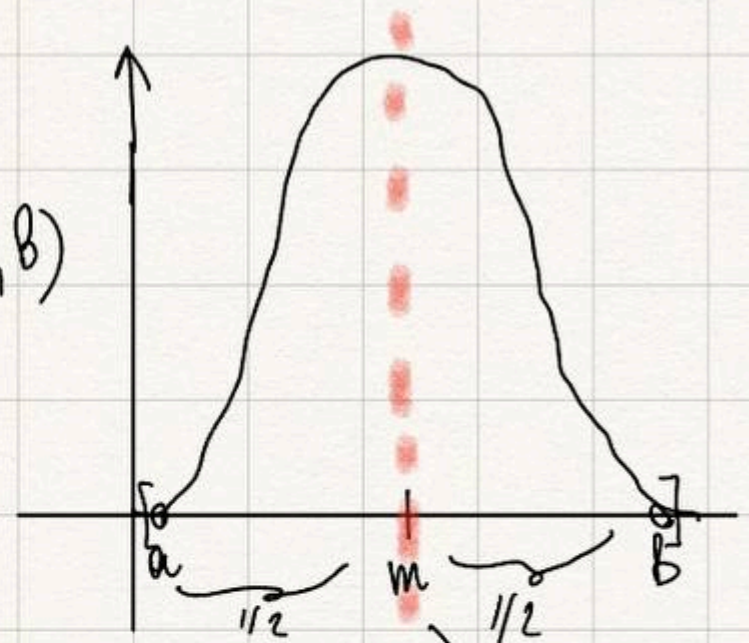
Provide an example of PDF such that median is larger than the expected value.

Let X be random variable, f be strictly increasing function and $Y = f(X)$. What can you say about medians of X and Y ?

Solution:

Let X -r.v.

$$f_x(x) = \begin{cases} 0, & x \leq a \\ f(x) > 0, & x \in (a, b), \text{ let } m \in (a, b) \\ 0, & x \geq b \end{cases}$$



$$P(X < m) = P(X > m) = 1/2 \text{ - median of } X$$

$$P(X < m) = CDF_x(m) = \int_{-\infty}^m f_x(x) dx = \int_a^m f_x(x) dx = F(m) - F(a) \quad \frac{b+a}{2}$$

$$P(X > m) = CDF_x(b) - CDF_x(m) = \int_m^b f_x(x) dx = F(b) - F(m)$$

where $F'(x) = f_x(x)$

Let assume that $f_x(x)$ is symmetrical, i.e. $f(\frac{b+a}{2} - t) = f(\frac{b+a}{2} + t)$ for $t \in [0, \frac{b-a}{2}]$, or $f(a+t) = f(b-t)$ for $t \in [0, \frac{b-a}{2}]$

Let's prove that for symmetrical $f_x(x)$, $m = \frac{a+b}{2}$ is median:

$$P(X < m) = CDF_x(m) = \int_a^m f(x) dx = \int_a^{(a+b)/2} f(x) dx = F(\frac{a+b}{2}) - F(a)$$

$$P(X > m) = F(b) - F(m) = F(b) - F(\frac{a+b}{2}) = \int_m^b f(x) dx$$

Let's replace x with $y = x - m \Rightarrow x = y + m$

$\Rightarrow \int_0^{\frac{b-a}{2}} f(y+m) dy$ ($dx = dy$). Hence $m = \frac{a+b}{2}$ and $f(x)$ is symmetrical to $x = \frac{a+b}{2}$ line then $f(y+m) = f(m-y)$ for $[0, \frac{b-a}{2}]$.

$$P(X > m) = \int_0^{(b-a)/2} f(m-y) dy$$

Let's do same replacement in $P(X < m)$ equation.

$$P(X < m) = \int_{-(b-a)/2}^0 f(m+y) dy, \text{ let } z = -y \Rightarrow P(X < m) = \int_{(b-a)/2}^0 f(m-z) (-dz) =$$

$$= \int_0^{(b-a)/2} f(m-z) dz = \int_0^{(b-a)/2} f(m+y) dy \text{ by symmetry of } f \Rightarrow$$

$\Rightarrow P(X < m) = P(X > m) = 1 - P(X < m) = P(X < m) = 1/2 \Rightarrow m = \frac{a+b}{2}$ is median.

Let's prove for symmetrical $f_X(x)$ that $E(X) = \frac{a+b}{2} = m$.

$$E(X) = \int_a^b x \cdot f(x) dx. \text{ Let } y = \frac{a+b}{2} = x - m$$

$$\int_a^b x f(x) dx = \int_{-(b-a)/2}^{+(b-a)/2} (y+m) f(y+m) dy; dx = dy$$

$$\int_{-(b-a)/2}^{+(b-a)/2} (y+m) f(y+m) dy = \int_{-(b-a)/2}^{(b-a)/2} m \cdot f(y+m) dy + \int_{-(b-a)/2}^{(b-a)/2} y f(y+m) dy = m \int_a^b f(x) dx +$$

$$+ \int_{-(b-a)/2}^0 y f(y+m) dy + \int_0^{(b-a)/2} y f(y+m) dy;$$

$\int_a^b f(x) dx = 1$ by definition of pdf x . For the 2d interval let's replace y with $z = -y$

$$EX = m \cdot 1 + \int_{-(b-a)/2}^0 (-z) f(m-z) (-dz) + \int_0^{(b-a)/2} y f(y+m) dy$$

$$EX = m + (-1) \int_0^{(b-a)/2} z f(m-z) dz + \int_0^{(b-a)/2} y f(y+m) dy \quad \text{by symmetry of } f(x).$$

$$f(m-z) = f(m+z) \text{ for } z \in [0, \frac{b-a}{2}] \Rightarrow EX = m + \int_0^{(b-a)/2} y f(y+m) dy - \int_0^{(b-a)/2} z f(m+z) dz.$$

It's easy to see that 2 integrals destroy each other \Rightarrow
 $\Rightarrow EX = m$ - median, which we wanted to prove.

Median = $m = \frac{a+b}{2} = EX$. for symmetric f_x .

Let's show that median can be larger than EX :

$$X \sim \text{pdf}_X(x) = \begin{cases} 0, & x \leq 0 \\ 2/3, & x \in (0, 1) \\ 1/3, & x \in (1, 2) \\ 0, & x \geq 2 \end{cases}$$

$$\int_{-\infty}^{\infty} f(x) dx = \int_0^1 \frac{2}{3} dx + \int_1^2 \frac{1}{3} dx = \frac{2}{3} x \Big|_0^1 + \frac{1}{3} x \Big|_1^2 = \frac{2}{3} + \frac{1}{3}(2-1) = 1 \Rightarrow$$

$\Rightarrow f(x)$ is legitimate pdf.

$$EX = \int_{-\infty}^{\infty} x f(x) dx = \int_0^1 \frac{2}{3} x dx + \int_1^2 \frac{1}{3} x dx = \frac{x^2}{3} \Big|_0^1 + \frac{x^2}{6} \Big|_1^2 = \frac{1}{3} + \frac{4}{6} - \frac{1}{6}$$

$$= 5/6$$

$m = 3/4$ - median of X .

$$P(X < m) = \int_0^{3/4} \frac{2}{3} dx = \frac{2}{3} X \Big|_0^{3/4} = \frac{2}{3} \cdot \frac{3}{4} = \frac{1}{2}$$

Then let's take $Y = -X \Rightarrow \text{median}_Y = -\text{median}_X = -3/4$

$$EY = -EX = -\frac{5}{6}; \quad -\frac{3}{4} > -\frac{5}{6} \Rightarrow \text{median}_Y > EY$$

X is r.v., $Y = f(X)$, f - strictly increasing function \Rightarrow

\Rightarrow if $P(X < m) = 1/2$ - median $_X$ for some m then

$P(f(X) < f(m)) = 1/2$ because function f preserve the order
i.e. if $a < b \Rightarrow f(a) < f(b)$.

Analogously: $P(f(X) > f(m)) = P(X > m) = 1/2 \Rightarrow f(m)$ is
median of Y .

Question 2

Let X be uniform random variable on a segment $[0, 2]$. Consider random variable $Y = X^2$. Find CDF and PDF of Y . Is PDF a bounded function?

Solution:

Let X be Uniform distribution from 0 to 2, i.e.

$X \sim \text{Uniform}[0, 2]$. Then $\text{PDF}_X(x) = \frac{1}{2-0} = \frac{1}{2}$ by the properties of uniform distributions.

$$\text{Then } \text{CDF}_X(x) = \int_{-\infty}^x \text{PDF}_X(x) dx = \int_0^x \frac{1}{2} dx = \frac{x}{2} \Big|_0^x = \frac{x}{2}, x \in [0, 2]$$

Let $Y = X^2$ then $\text{CDF}_Y(t) = P(Y \leq t) = P(X^2 \leq t) = P(X \leq \sqrt{t}) = \text{CDF}_X(\sqrt{t}) = \frac{\sqrt{t}}{2}$ where $\sqrt{t} \in [0, 2]$ because $X \in [0, 2] \Rightarrow t \in [0^2, 2^2] = [0, 4]$. When $\sqrt{t} \geq 2 \Rightarrow \text{CDF}_X(t) = 1$

$\text{PDF}_Y(t) = (\text{CDF}_Y(t))'$ by definition of PDF (properties of PDF)

$$\text{PDF}_Y(t) = \left(\frac{\sqrt{t}}{2}\right)' = \frac{1}{4\sqrt{t}}, t \in [0, 4], \text{ if } t < 0 \text{ PDF}_Y(t) = 0$$

if $t > 4$ $\text{PDF}_Y(t) = 0$

$$\text{CDF}_Y(t) = \begin{cases} 0, & t < 0 \\ \frac{\sqrt{t}}{2}, & t \in [0, 4] \\ 1, & t > 4 \end{cases}$$

$$\text{PDF}_Y(t) = \begin{cases} 0, & t < 0 \\ \frac{1}{4\sqrt{t}}, & t \in [0, 4] \\ 0, & t > 4 \end{cases}$$

Let's look at $\text{PDF}_Y(t) = \frac{1}{4\sqrt{t}}, t \in [0, 4]$

It's strictly decreasing function and continuous on $(0, 4] \Rightarrow \text{PDF}_Y(t) \geq \text{PDF}_Y(4)$ and $\text{PDF}_Y(4) \leftrightarrow \lim_{t \rightarrow 0} \text{PDF}_Y(t)$;

$\text{PDF}_Y(4) = \frac{1}{4\sqrt{4}} = \frac{1}{8}$, $\lim_{t \rightarrow 0} \text{PDF}_Y(t) = \lim_{t \rightarrow 0} \frac{1}{4\sqrt{t}} = +\infty \Rightarrow \text{PDF}_Y(t)$ can't be a bounded function. $\text{PDF}_Y(t)$ is unbounded.

Question 3

A fair coin is tossed 400 times. Let X be number of heads. Prove that

$P(X > 240) \leq 1/32$. (Hint: use Chebyshev's inequality and symmetry considerations.)

Solution:

Coin toss is Bernoulli trial with probability of success $p = 1/2$. Each of $n = 400$ coin tosses is i.i.d.

$$\text{Bern}(1/2) \Rightarrow X \sim \text{Bern}(400, 1/2)$$

$$EX = np = 400 \cdot 1/2 = 200 \quad (\text{by properties of } EX)$$

$$\text{Var } X = np(1-p) = 400 \cdot 1/2 \cdot (1-1/2) = 100$$

By Chebyshev's inequality:

$$P(|X - EX| > d) \leq \frac{\text{Var } X}{d^2}; \quad \text{Let } d = 40$$

$$P(|X - 200| > 40) \leq \frac{100}{40^2}$$

X can be from 0 to 400 because we can have all n successes at maximum and we can't have less than 0 successes.

$$P(|X - 200| > 40) = P(X < 160 \vee X > 240) \leq \frac{100}{40 \cdot 40} = \frac{1}{16}$$

$$P(X < 160) + P(X > 240) \leq \frac{1}{16}$$

By the symmetry of Bernoulli trials with $p = 1/2$ and $160 + 240 = 400$ - number of trials for each elementary outcome when X - number of success, $X > 240$ we can match exactly 1

elementary outcome where $X < 400 - 240 = 160$ by replacing all successes with failures or vice versa.

We can do this operation in reverse order so each with $X < 160$ is matched with exactly 1 outcome with $X > 240$. So we bijected all

outcomes $X < 160$ to all outcomes $X > 240 \Rightarrow$

$$\Rightarrow P(X < 160) = P(X > 240) \Rightarrow P(X < 160) + P(X > 240) =$$

$$2P(X > 240) \Rightarrow P(X > 240) \leq \frac{1}{2} \cdot \frac{1}{16} = 1/32. \text{ As}$$

we need to prove.

Question 4

Let X and Y be two independent normally distributed random variables with expected value 0 and variance 1. Find their joint PDF. Plot its level curves.

Solution:

Problem 4

Let $X \sim \mathcal{N}(0, 1)$ and $Y \sim \mathcal{N}(0, 1)$ which is independent from X .
Then $\text{PDF}_X(x) = f_X(x) = \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{x^2}{2}}$, $\text{PDF}_Y(y) = f_Y(y) = \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{y^2}{2}}$.

Because X is independent from Y the Joint $\text{PDF}_{X,Y}(x,y) = f_{X,Y}(x,y) = f_X(x) \cdot f_Y(y)$ by the properties of Joint PDF of independent events.

$$\text{So } f_{X,Y}(x,y) = f_X(x) \cdot f_Y(y) = \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{x^2}{2}} \cdot \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{y^2}{2}} =$$

$$\boxed{= \frac{1}{2\pi} \cdot e^{-\frac{(x^2+y^2)}{2}} \text{ - Joint PDF.}}$$

Let's look on Joint PDF level curves.

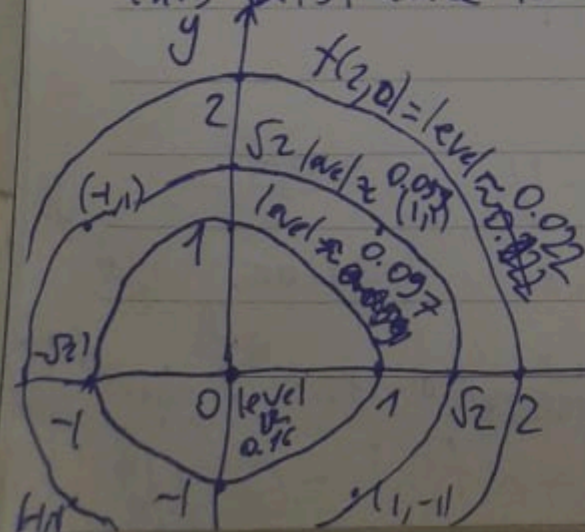
if $f_{X,Y}(x,y) = c$, c - constant, $c > 0$, then

$$c = \frac{1}{2\pi} \cdot e^{-\frac{(x^2+y^2)}{2}} \Rightarrow 2\pi c = e^{-\frac{(x^2+y^2)}{2}} \Rightarrow \ln 2\pi c = -\frac{(x^2+y^2)}{2}$$

By π, c - constants, $\ln \pi, \ln c$ also are constants as $\ln 2$. \Rightarrow
 $\Rightarrow \ln 2\pi c = -\frac{(x^2+y^2)}{2} \Rightarrow \ln 2 + \ln \pi + \ln c = (-\frac{1}{2}) \cdot (x^2+y^2) \Rightarrow$
 $\Rightarrow x^2+y^2 = (-2) \cdot \ln 2\pi c$, where $(-2) \cdot \ln 2\pi c$ is some constant d ,
 where $d \in (-\infty; +\infty)$ because $x^2+y^2 \in [0; +\infty)$ so:

$x^2+y^2 = d = \text{const}$ for all x, y which form same level curve, i.e.
 $f(x,y) = \text{const.}$

$x^2+y^2 = \text{const}$ is equation of circle with center in $(0,0)$ and
 radius $= r = \sqrt{x^2+y^2}$, because $r = \sqrt{x^2+y^2}$ is distance to point $(0,0)$ and
 this distance is constant.



$$f(1,0) = \frac{1}{2\pi} \cdot e^{-\frac{(1^2+0^2)}{2}} = \frac{1}{2\pi} \cdot e^{-\frac{1}{2}} = \frac{1}{2\pi e^{\frac{1}{2}}} \approx 0.097$$

We also have isolated dot $(0,0)$ with its own level

$$f(0,0) = \frac{1}{2\pi} \cdot e^{-\frac{(0^2+0^2)}{2}} = \frac{1}{2\pi} \cdot e^0 = \frac{1}{2\pi} \approx 0.16$$

$$f(2,0) = \frac{1}{2\pi} \cdot e^{-\frac{(2^2+0^2)}{2}} = \frac{1}{2\pi} \cdot e^{-2} = \frac{1}{2\pi e^2} \approx 0.022$$

$$f(1,1) = \frac{1}{2\pi} \cdot e^{-\frac{(1^2+1^2)}{2}} = \frac{1}{2\pi} \cdot e^{-1} = \frac{1}{2\pi e} \approx 0.059$$

Double A