MATRIX DAN VEKTOR ACAK

- Random Vector vector whose individual elements are random variables
- Random Matrix matrix whose individual elements are random variables

The expected value of a random vector or matrix is the matrix containing the expected values of the individual elements, i.e.,

$$\mathbf{E}[\mathbf{X}] = \begin{bmatrix} \mathbf{E}[\mathbf{x}_{11}] & \mathbf{E}[\mathbf{x}_{12}] & \cdots & \mathbf{E}[\mathbf{x}_{1p}] \\ \mathbf{E}[\mathbf{x}_{21}] & \mathbf{E}[\mathbf{x}_{22}] & \cdots & \mathbf{E}[\mathbf{x}_{2p}] \\ \vdots & \vdots & \ddots & \vdots \\ \mathbf{E}[\mathbf{x}_{n1}] & \mathbf{E}[\mathbf{x}_{n2}] & \cdots & \mathbf{E}[\mathbf{x}_{np}] \end{bmatrix}$$

where

$$\mathsf{E}\!\left[x_{ij}\right] = \begin{cases} \sum_{\mathsf{all}\ x_{ij}} x_{ij} p_{ij}\left(x_{ij}\right) \\ \int\limits_{-\infty}^{\infty} x_{ij} f_{ij}\left(x_{ij}\right) dx_{ij} \end{cases}$$

Note that for random matrices **X** and **Y** of the same dimension, conformable matrices of constants **A** and **B**, and scalar c

- $\blacksquare E(cX) = cE(X)$
- $\blacksquare E(X+Y) = E(X) + E(Y)$
- \blacksquare E(AXB)= AE(X)B

Mean Vector – random vector whose elements are the means of the corresponding random variables, i.e.,

$$\mathsf{E}\!\left[x_i\right] = \mu_i = \begin{cases} \sum_{\mathsf{all}\ x_i} x_i p_i\left(x_i\right) \\ \int\limits_{-\infty}^{\infty} x_i f_i\left(x_i\right) dx_i \end{cases}$$

In matrix notation we can write the mean vector as

$$\mathbf{E}[\mathbf{X}] = \begin{bmatrix} \mathbf{E}[\mathbf{X}_1] \\ \mathbf{E}[\mathbf{X}_2] \\ \vdots \\ \mathbf{E}[\mathbf{X}_p] \end{bmatrix} = \begin{bmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_p \end{bmatrix} = \mathbf{\mu}$$

For the bivariate probability distribution

X_1 X_2	-4	1	3	p ₁ (x ₁)
1	0.15	0.20	0.05	0.40
2	0.10	0.25	0.25	0.60
$p_2(x_2)$	0.25	0.45	0.30	1.00

the mean vector is

$$\mu = E[\mathbf{X}] = \begin{bmatrix} E[x_1] \\ E[x_2] \end{bmatrix} = \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix} = \begin{bmatrix} \sum_{\text{all } x_1} x_1 p_1(x_1) \\ \sum_{\text{all } x_2} x_2 p_2(x_2) \end{bmatrix}$$
$$= \begin{bmatrix} 0.4(1) + 0.6(2) \\ 0.1(-4) + 0.25(1) + 0.25(3) \end{bmatrix} = \begin{bmatrix} 1.60 \\ 0.35 \end{bmatrix}$$

Covariance Matrix – random symmetric vector whose diagonal elements are variances of the corresponding random variables, i.e.,

$$E\left[\left(x_{i}-\mu_{i}\right)^{2}\right]=\sigma_{i}^{2}=\begin{cases} \sum_{\text{all }x_{i}}\left(x_{i}-\mu_{i}\right)^{2}p_{i}\left(x_{i}\right)\\ \int_{-\infty}^{\infty}\left(x_{i}-\mu_{i}\right)^{2}f_{i}\left(x_{i}\right)dx_{i} \end{cases}$$

and whose off-diagonal elements are covariances of the corresponding random variable pairs, i.e.,

$$E\Big[\Big(x_{i}-\mu_{i}\Big)\Big(x_{k}-\mu_{k}\Big)\Big] = \sigma_{ik} = \begin{cases} \sum\limits_{\text{all }x_{i}}\sum\limits_{\text{all }x_{k}}\Big(x_{i}-\mu_{i}\Big)\Big(x_{k}-\mu_{k}\Big)p_{ik}\left(x_{i},x_{k}\right) \\ \int\limits_{-\infty}^{\infty}\int\limits_{-\infty}^{\infty}\Big(x_{i}-\mu_{i}\Big)\Big(x_{k}-\mu_{k}\Big)f_{ik}\left(x_{i},x_{k}\right)dx_{i}dx_{k} \end{cases}$$

notice that if we this expression, when i = k, returns the variance, i.e.,

$$\sigma_{ii} = \sigma_i^2$$

In matrix notation we can write the covariance matrix as

$$\begin{split} E\Big[\left(\boldsymbol{X} - \boldsymbol{\mu} \right) \left(\boldsymbol{X} - \boldsymbol{\mu} \right)^{'} \Big] &= E\Big[\begin{pmatrix} (x_{1} - \mu_{1}) \\ (x_{2} - \mu_{2}) \\ \vdots \\ (x_{p} - \mu_{p}) \end{pmatrix} \Big[\left(x_{1} - \mu_{1} \right) \quad \left(x_{2} - \mu_{2} \right) \quad \cdots \quad \left(x_{p} - \mu_{p} \right) \Big] \\ &= \begin{bmatrix} E\Big[\left(x_{1} - \mu_{1} \right)^{2} \Big] & E\Big[\left(x_{1} - \mu_{1} \right) \left(x_{2} - \mu_{2} \right) \Big] \quad \cdots \quad E\Big[\left(x_{1} - \mu_{1} \right) \left(x_{p} - \mu_{p} \right) \Big] \\ E\Big[\left(x_{2} - \mu_{2} \right) \left(x_{1} - \mu_{1} \right) \Big] & E\Big[\left(x_{2} - \mu_{2} \right)^{2} \Big] \quad \cdots \quad E\Big[\left(x_{2} - \mu_{2} \right) \left(x_{p} - \mu_{p} \right) \Big] \\ \vdots & \vdots & \ddots & \vdots \\ E\Big[\left(x_{p} - \mu_{p} \right) \left(x_{1} - \mu_{1} \right) \Big] & E\Big[\left(x_{p} - \mu_{p} \right) \left(x_{2} - \mu_{2} \right) \Big] \quad \cdots \quad E\Big[\left(x_{p} - \mu_{p} \right)^{2} \Big] \end{aligned}$$

For the bivariate probability distribution we used earlier

X_1 X_2	-4	1	3	p ₁ (x ₁)
1	0.15	0.20	0.05	0.40
2	0.10	0.25	0.25	0.60
$p_2(x_2)$	0.25	0.45	0.30	1.00

the covariance matrix is

$$\begin{split} E\Big[\big(\boldsymbol{X} - \boldsymbol{\mu} \big) \big(\boldsymbol{X} - \boldsymbol{\mu} \big)^{'} \Big] &= \begin{bmatrix} E\Big[\big(x_{1} - \mu_{1} \big)^{2} \Big] & E\Big[\big(x_{1} - \mu_{1} \big) \big(x_{2} - \mu_{2} \big) \Big] \\ E\Big[\big(x_{2} - \mu_{2} \big) \big(x_{1} - \mu_{1} \big) \Big] & E\Big[\big(x_{2} - \mu_{2} \big)^{2} \Big] \\ &= \begin{bmatrix} \sum_{\text{all } x_{1}} \big(x_{1} - \mu_{1} \big)^{2} p_{1} \big(x_{1} \big) & \sum_{\text{all } x_{1}} \sum_{\text{all } x_{2}} \big(x_{1} - \mu_{1} \big) \big(x_{2} - \mu_{2} \big) p_{12} \big(x_{1}, x_{2} \big) \\ \sum_{\text{all } x_{2}} \sum_{\text{all } x_{1}} \big(x_{2} - \mu_{2} \big) \big(x_{1} - \mu_{1} \big) p_{21} \big(x_{2}, x_{1} \big) & \sum_{\text{all } x_{2}} \big(x_{2} - \mu_{2} \big)^{2} p_{2} \big(x_{2} \big) \end{bmatrix} = \begin{bmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{21} & \sigma_{22} \end{bmatrix} = \Sigma \end{split}$$

which can be computed as

$$\begin{split} E\Big[\big(\boldsymbol{X} - \boldsymbol{\mu} \big) \big(\boldsymbol{X} - \boldsymbol{\mu} \big)^{\cdot} \Big] &= \begin{bmatrix} \sum_{\text{all } x_1} \left(x_1 - \mu_1 \right)^2 p_1 \big(x_1 \big) & \sum_{\text{all } x_2} \sum_{n_1} \left(x_2 - \mu_2 \right) p_{12} \big(x_1, x_2 \big) \\ \sum_{\text{all } x_2} \sum_{\text{all } x_1} \left(x_2 - \mu_2 \right) \big(x_1 - \mu_1 \big) p_{21} \big(x_2, x_1 \big) & \sum_{\text{all } x_2} \left(x_2 - \mu_2 \right)^2 p_2 \big(x_2 \big) \end{bmatrix} \\ &= \begin{bmatrix} (1 - 1.6)^2 \big(0.4 \big) + \big(2 - 1.6 \big)^2 \big(0.6 \big) & + \big(1 - 1.6 \big) \big(0.4 - 0.35 \big) \big(0.15 \big) + \big(1 - 1.6 \big) \big(1 - 0.35 \big) \big(0.2 \big) \\ & + \big(1 - 1.6 \big) \big(1 - 0.35 \big) \big(0.25 \big) + \big(2 - 1.6 \big) \big(0.4 \big) - 0.35 \big) \big(0.25 \big) \\ &+ \big(1 - 0.35 \big) \big(1 - 1.6 \big) \big(0.15 \big) + \big(1 - 0.35 \big) \big(2 - 1.6 \big) \big(0.25 \big) \\ &+ \big(1 - 0.35 \big) \big(1 - 1.6 \big) \big(0.02 \big) + \big(1 - 0.35 \big) \big(2 - 1.6 \big) \big(0.25 \big) \\ &+ \big(3 - 0.35 \big) \big(1 - 1.6 \big) \big(0.05 \big) + \big(3 - 0.35 \big) \big(2 - 1.6 \big) \big(0.25 \big) \\ &= \begin{bmatrix} 0.24 & 0.39 \\ 0.39 & 7.0275 \end{bmatrix} = \begin{bmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{21} & \sigma_{22} \end{bmatrix} = \Sigma \end{split}$$

Thus the population represented by the bivariate probability distribution

X_1 X_2	-4	1	3	p ₁ (x ₁)
1	0.15	0.20	0.05	0.40
2	0.10	0.25	0.25	0.60
$p_2(x_2)$	0.25	0.45	0.30	1.00

Have population mean vector and variancecovariance matrix

$$\mu = \begin{bmatrix} 1.60 \\ 0.35 \end{bmatrix}$$
 and $\Sigma = \begin{bmatrix} 0.24 & 0.39 \\ 0.39 & 7.0275 \end{bmatrix}$

We can use the population variance-covariance matrix Σ to calculate the population correlation matrix ρ . Individual population correlation coefficients are defined as

$$\rho_{ik} = \frac{\sigma_{ik}}{\sqrt{\sigma_{ii}}\sqrt{\sigma_{kk}}}$$

In matrix notation we can write the correlation matrix as

$$\rho = \begin{bmatrix} \frac{\sigma_{11}}{\sqrt{\sigma_{11}}\sqrt{\sigma_{11}}} & \frac{\sigma_{12}}{\sqrt{\sigma_{11}}\sqrt{\sigma_{22}}} & \cdots & \frac{\sigma_{1p}}{\sqrt{\sigma_{11}}\sqrt{\sigma_{pp}}} \\ \frac{\sigma_{21}}{\sqrt{\sigma_{22}}\sqrt{\sigma_{11}}} & \frac{\sigma_{22}}{\sqrt{\sigma_{22}}\sqrt{\sigma_{22}}} & \cdots & \frac{\sigma_{2p}}{\sqrt{\sigma_{22}}\sqrt{\sigma_{pp}}} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\sigma_{p1}}{\sqrt{\sigma_{pp}}\sqrt{\sigma_{11}}} & \frac{\sigma_{p2}}{\sqrt{\sigma_{pp}}\sqrt{\sigma_{22}}} & \cdots & \frac{\sigma_{pp}}{\sqrt{\sigma_{pp}}\sqrt{\sigma_{pp}}} \end{bmatrix} = \begin{bmatrix} \rho_{11} & \rho_{12} & \cdots & \rho_{1p} \\ \rho_{1k} & \rho_{1k} & \cdots & \rho_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ \rho_{p1} & \rho_{p2} & \cdots & \rho_{pp} \end{bmatrix}$$

We can easily show that

$$\sum = \mathbf{V}^{1/2} \rho \mathbf{V}^{1/2}$$

where

$$\mathbf{V}^{1/2} = egin{bmatrix} \sqrt{\sigma_{11}} & 0 & \cdots & 0 \\ 0 & \sqrt{\sigma_{22}} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \sqrt{\sigma_{pp}} \end{bmatrix}$$

which implies

$$\boldsymbol{\rho} = \left(\mathbf{V}^{1/2}\right)^{-1} \boldsymbol{\Sigma} \left(\mathbf{V}^{1/2}\right)^{-1}$$

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$p_2(x_2)$	0.25	0.45	0.30	1.00

the variance matrix is

$$\mathbf{V}^{1/2} = \begin{bmatrix} \sqrt{\sigma_{11}} & 0 \\ 0 & \sqrt{\sigma_{22}} \end{bmatrix} = \begin{bmatrix} \sqrt{0.24} & 0 \\ 0 & \sqrt{7.0275} \end{bmatrix} = \begin{bmatrix} 0.48989 & 0 \\ 0 & 2.65094 \end{bmatrix}$$

so the population correlation matrix is

$$\rho = \begin{pmatrix} \mathbf{V}^{1/2} \end{pmatrix}^{-1} \Sigma \begin{pmatrix} \mathbf{V}^{1/2} \end{pmatrix}^{-1} \\
= \begin{bmatrix} 2.04124 & 0 \\ 0 & 0.37739 \end{bmatrix} \begin{bmatrix} 0.24 & 0.39 \\ 0.39 & 7.0275 \end{bmatrix} \begin{bmatrix} 2.04124 & 0 \\ 0 & 0.37739 \end{bmatrix} \\
= \begin{bmatrix} 1.00 & 0.30 \\ 0.30 & 1.00 \end{bmatrix}$$

We often deal with variables that naturally fall into groups. In the simplest case, we have two groups of size q and p – q of variables.

Under such circumstances, it may be convenient to partition the matrices and vectors.

Here we have a mean vector and variancecovariance matrix:

$$\mu\begin{bmatrix} \mu_1 \\ \vdots \\ \mu_q \\ \vdots \\ \mu_p \end{bmatrix} = \begin{pmatrix} \mu^{(1)} \\ \mu^{(2)} \\ \end{pmatrix}, \quad \Sigma = \begin{bmatrix} \sigma_{11} & \cdots & \sigma_{1p} & \sigma_{1,q+1} & \cdots & \sigma_{1p} \\ \vdots & \ddots & \vdots & \vdots & \ddots & \vdots \\ \sigma_{q1} & \cdots & \sigma_{qq} & \sigma_{q,q+1} & \cdots & \sigma_{qp} \\ & & & & & & \\ \hline \sigma_{q+1,1} & \cdots & \sigma_{q+1,q} & \sigma_{q+1,q+1} & \cdots & \sigma_{q+1,p} \\ \vdots & \ddots & \vdots & & \vdots & \ddots & \vdots \\ \sigma_{p1} & \cdots & \sigma_{pq} & \sigma_{p,q+1} & \cdots & \sigma_{pp} \end{bmatrix} = \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ & & & \\ \hline \Sigma_{21} & & \Sigma_{22} \end{bmatrix}$$

Rules for Mean Vectors and Covariance Matrices for Linear Combinations of Random Variables: The linear combination of real constants **c** and random variables **X**

$$\mathbf{c'X} = \begin{bmatrix} c_1 & c_2 & \cdots & c_p \end{bmatrix} = \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_p \end{bmatrix} = \sum_{j=1}^p c_j X_j$$

has mean vector

$$\mathsf{E} \big[\mathbf{c}' \mathbf{X} \big] = \mathbf{c}' \mathsf{E} \big[\mathbf{X} \big] = \mathbf{c}' \boldsymbol{\mu}$$

The linear combination of real constants **c** and random variables **X**

$$\mathbf{c'X} = \begin{bmatrix} \mathbf{c}_1 & \mathbf{c}_2 & \cdots & \mathbf{c}_p \end{bmatrix} = \begin{bmatrix} \mathbf{x}_1 \\ \mathbf{x}_2 \\ \vdots \\ \mathbf{x}_p \end{bmatrix} = \sum_{j=1}^{p} \mathbf{c}_j \mathbf{x}_j$$

also has variance

$$Var[\mathbf{c}'\mathbf{X}] = \mathbf{c}' \sum \mathbf{c}$$

Suppose, for example, we have random vector

$$\mathbf{X} = \begin{bmatrix} \mathbf{X}_1 \\ \mathbf{X}_2 \end{bmatrix}$$

and we want to find the mean vector and covariance matrix for the linear combinations:

$$Z_1 = X_1 - X_2$$
, i.e., $Z = \begin{bmatrix} Z_1 \\ Z_2 \end{bmatrix} = \begin{bmatrix} 1 & -1 \\ 1 & 1 \end{bmatrix} \begin{bmatrix} X_1 \\ X_2 \end{bmatrix} = CX$