

Part IA — Probability

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§0 Introduction

Probability theory is the mathematical formulation of randomness. Examples include the modelling of random experiments like flipping a coin, throwing a die, shuffle a deck, and so on. What we want to do is to develop a mathematical framework to study randomness.

Example 0.1

Dice: outcomes $1, 2, \dots, 6$.

- $\mathbb{P}(2) = \frac{1}{6}$.
- $\mathbb{P}(\text{multiple of } 3) = \frac{2}{6} = \frac{1}{3}$.
- $\mathbb{P}(\text{not a multiple of } 3) = \frac{2}{3}$
- $\mathbb{P}(\text{prime}) = \frac{1}{2}$.
-

$$\begin{aligned}\mathbb{P}(\text{prime or multiple of } 3) &= \frac{1}{3} + \frac{1}{2} = \frac{5}{6} \\ &= \frac{4}{6} + \frac{2}{6} \\ &= \frac{6}{6} = 1.\end{aligned}$$

$$\mathbb{P}(\text{prime or multiple of } 3) = \frac{1}{3} + \frac{1}{2} - \frac{1}{6} = \frac{2}{3}$$

§1 Formal Setup

Definition 1.1 (Sample Space)

The **sample space** Ω is a set of outcomes.

Definition 1.2 (σ -algebra)

- Let \mathcal{F} a collection of subsets of Ω (called *events*).
- \mathcal{F} is a **σ -algebra** if
 - F1. $\Omega \in \mathcal{F}$.
 - F2. $A \in \mathcal{F}$ then $A^c = \Omega \setminus A \in \mathcal{F}$.
 - F3. \forall countable collections $(A_n)_{n \in \mathbb{N}} \in \mathcal{F}$, the union $\bigcup_{n \in \mathbb{N}} A_n \in \mathcal{F}$ also.

Remark 1. The motivation for F2 is so that $\mathbb{P}(A^c) = 1 - \mathbb{P}(A)$ (the probability of not A is defined as expected).

Definition 1.3 (Probability Measure)

Given σ -algebra \mathcal{F} on Ω , function $\mathbb{P} : \mathcal{F} \rightarrow [0, 1]^a$ is a **probability measure** if

- P2. $\mathbb{P}(\Omega) = 1$.
- P3. \forall countable collections $(A_n)_{n \in \mathbb{N}}$ of *disjoint* events in \mathcal{F} :

$$\mathbb{P}\left(\bigcup_{n \in \mathbb{N}} A_n\right) = \sum_{n \in \mathbb{N}} \mathbb{P}(A_n).$$

Then $(\Omega, \mathcal{F}, \mathbb{P})$ is a *probability space*.

^aP1. $\mathbb{P}(A) \geq 0$

Example 1.1

Coming back to Example 0.1. $\Omega = \{1, 2, \dots, 6\}$ so $\mathbb{P}(\Omega) = \mathbb{P}(1 \text{ or } 2 \text{ or } 3 \text{ or } 4 \text{ or } 5 \text{ or } 6) = 1$ and \mathcal{F} is all subsets of Ω .

Question

Why $\mathbb{P} : \mathcal{F} \rightarrow [0, 1]$

and not $\mathbb{P} : \Omega \rightarrow [0, 1]$?

If Ω is countable:

- In general: \mathcal{F} = all subsets of Ω , i.e. $\mathcal{P}(\Omega)$ (the power set).
- $\mathbb{P}(2)$ is shorthand for $\mathbb{P}(\{2\})$.
- \mathbb{P} is determined by $(\mathbb{P}(\{w\}), \forall w \in \Omega)$ (e.g. unfair dice).

If Ω is uncountable:

- E.g. $\Omega = [0, 1]$. Want to choose a real number, all equally likely.
- If $\mathbb{P}(\{0\}) = \alpha > 0$ then $\mathbb{P}\left(\left\{0, 1, \frac{1}{2}, \dots, \frac{1}{n}\right\}\right) = (n+1)\alpha \nless \text{if } n \text{ large as } \mathbb{P} > 1$.
- So $\mathbb{P}(\{0\}) = 0$, or $\mathbb{P}(\{0\})$ is undefined.
- What about $\mathbb{P}\left(\left\{x : x \leq \frac{1}{3}\right\}\right)$?
 - ? “Add up” all $\mathbb{P}(\{x\})$ for $x \leq \frac{1}{3}$. However this range is uncountable and we can’t take a sum of uncountably many terms.

Aside

Question

Can we choose uniformly from an infinite countable set? (E.g. $\Omega = \mathbb{N}$ or $\Omega = \mathbb{Q} \cap [0, 1]$)

Answer

No it is not possible but that’s ok there \exists lots of interesting probability measures of \mathbb{N} !

Proof. Suppose possible

- $\mathbb{P}(\{0\}) = \alpha > 0 \quad \forall \omega \in \Omega$. Then $\mathbb{P}(\Omega) = \sum_{\omega \in \Omega} \mathbb{P}(\{\omega\}) = \sum_{\omega \in \Omega} \alpha = \infty$. \nless of $\mathbb{P}2 : \mathbb{P}(\Omega) = 1$.
- $\mathbb{P}(\{0\}) = 0 \quad \forall \omega \in \Omega$. Then $\mathbb{P}(\Omega) = \sum_{\omega \in \Omega} \mathbb{P}(\{\omega\}) = \sum_{\omega \in \Omega} 0 = 0$.

□

Proposition 1.1 (From the axioms)

- $\mathbb{P}(A^c) = 1 - \mathbb{P}(A)$

Proof. A, A^c are disjoint. $A \cup A^c = \Omega$.
 $\implies \mathbb{P}(A) + \mathbb{P}(A^c) \stackrel{P_3}{=} \mathbb{P}(\Omega) \stackrel{P_2}{=} 1$

□

- $\mathbb{P}(\emptyset) = 0$
- If $A \subseteq B$ then $\mathbb{P}(A) \leq \mathbb{P}(B)$
- $\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B)$

§1.1 Examples of Probability Spaces

Example 1.2 (Uniform Choice)

Ω finite, $\Omega = \{\omega_1, \dots, \omega_n\}$, \mathcal{F} = all subsets. *uniform* choice (equally likely)

$$\mathbb{P} : \mathcal{F} \rightarrow [0, 1], \quad \mathbb{P}(A) = \frac{|A|}{|\Omega|}.$$

In particular: $\mathbb{P}(\{\omega\}) = \frac{1}{|\Omega|} \quad \forall \omega \in \Omega$.

Example 1.3 (Choosing without replacement)

n indistinguishable marbles labelled $\{1, \dots, n\}$. Pick $k \leq n$ marbles uniformly at random. Here: $\Omega = \{A \subseteq \{1, \dots, n\}, |A| = k\}$ $|\Omega| = \binom{n}{k}$

Example 1.4 (Well-shuffled deck of cards)

Uniformly chosen *permutation* of 52 cards.

$$\Omega = \{\text{all permutations of 52 cards}\}$$

$$|\Omega| = 52!$$

$$\mathbb{P}(\text{first three cards have the same suit}) = \frac{52 \times 12 \times 11 \times 49!}{52!} = \frac{22}{425}$$

$$\text{Note: } = \frac{12}{51} \times \frac{11}{50}$$

Example 1.5 (Coincident Birthdays)

There are n people; what is the probability that at least two of them share a birthday?

Assumptions:

- No leap years! (365 days)
- All birthdays are equally likely

Let $\Omega = \{1, \dots, 365\}^n$ and $\mathcal{F} = \mathcal{P}(\Omega)$.

Let $A = \{\text{at least two people share the same birthday}\}$ and so $A^c = \{\text{all } n \text{ birthdays are different}\}$.

$$\mathbb{P}(A^c) = \frac{|A^c|}{|\Omega|} = \frac{365 \times 364 \cdots \times (365 - n + 1)}{365^n}$$
$$\mathbb{P}(A) = 1 - \mathbb{P}(A^c)$$

Note that at $n = 22$, $\mathbb{P}(A) \approx 0.476$ and at $n = 23$, $\mathbb{P}(A) \approx 0.507$. So when there are at least 23 people in a room, the probability that two of them share a birthday is around 50%.

KEY IDEA: Calculating $\mathbb{P}(A^c)$ is easier than $\mathbb{P}(A)$.

$$\mathbb{P}(X_n = n) = \frac{1}{2^n}$$

$$\mathbb{P}(X_n = 0) = 0 \text{ if } n \text{ is odd}$$

What about $\mathbb{P}(X_n = 0)$ when n is even

Idea - Choose $\frac{n}{2}$ k s for $X_k = X_{k-1} + 1$ and the rest $X_k = X_{k-1} - 1$ (i.e. go up half the time and down the other half).

$$\begin{aligned}\mathbb{P}(X_n = 0) &= 2^{-n} \binom{n}{\frac{n}{2}} \\ &= \frac{n!}{2^n \left(\frac{n}{2}!\right)^2}\end{aligned}$$

Question

What happens when n is large?

§2.3 Stirling's Formula

Notation. Let $(a_n), (b_n)$ be two sequences. Say $a_n \sim b_n$ as $n \rightarrow \infty$ if $\frac{a_n}{b_n} \rightarrow 1$ as $n \rightarrow \infty$.

Example 2.1

$$n^2 + 5n + \frac{6}{n} \sim n^2$$

Example 2.2 (Non-Example)

$$\exp\left(n^2 + 5n + \frac{6}{n}\right) \not\sim \exp(n^2)$$

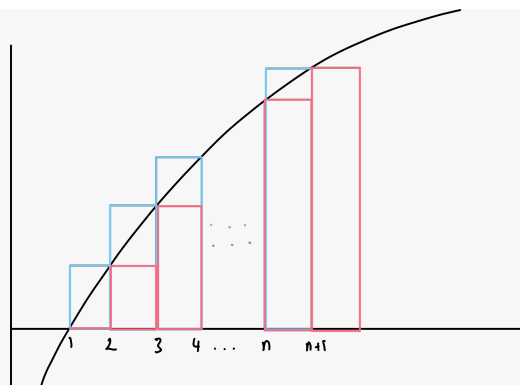
Theorem 2.1 (Stirling)

$$n! \sim \sqrt{2\pi n} n^{n+\frac{1}{2}} e^{-n} \text{ as } n \rightarrow \infty.$$

Theorem 2.2 (Weaker Version)

$$\log n! \sim n \log n.$$

Proof. $\log(n!) = \log 2 + \cdots + \log n.$



$$\begin{aligned}
 & \text{"Upper Integral"} \quad \int_1^n \log x \, dx \leq \log n! \leq \int_1^{n+1} \log x \, dx \quad \text{"Lower Integral"} \\
 & \underbrace{n \log n - n + 1}_{\sim n \log n} \leq \log n! \leq \underbrace{(n+1) \log(n+1) - n}_{\sim n \log n}
 \end{aligned}$$

Key idea: Sandwiching between lower/upper integrals. It was useful that

- $\log x$ is increasing
- $\log x$ has a nice integral!

□

§2.4 (Ordered) compositions

Definition 2.1 (Composition)

A **composition** of m with k parts is a sequence (m_1, \dots, m_k) of non-negative integers with $m_1 + \dots + m_k = m$.

Example 2.3

$$\begin{aligned}
 3 + 0 + 1 + 2 &= 6 \neq 1 + 2 + 0 + 3 = 6 \\
 \star \star \star || \star | \star \star
 \end{aligned}$$

There is a bijection between compositions *and* sequences of m stars and $(k-1)$ dividers. So the number of compositions is $\binom{m+k-1}{m}$.

Comment: Easy to mistake k with $k-1$ in no. of dividers.

§3 Properties of Probability measures

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space and $\mathbb{P} : \mathcal{F} \rightarrow [0, 1]$.

Definition 3.1 (Countable additivity)

P3 : $\mathbb{P}(\bigcup_{n \in \mathbb{N}} A_n) = \sum_{n \in \mathbb{N}} \mathbb{P}(A_n)$ for $(A_n)_{n \in \mathbb{N}}$ disjoint.

Question

What if the sets are not disjoint?

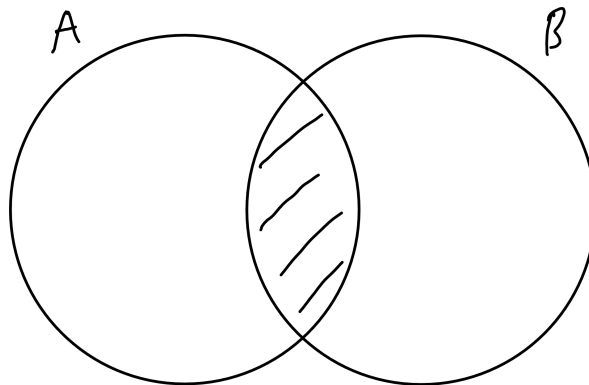
§3.1 Countable sub-additivity

Proposition 3.1 (Countable sub-additivity)

Let $(A_n)_{n \geq 1}$ be a sequence of events in \mathcal{F} . Then

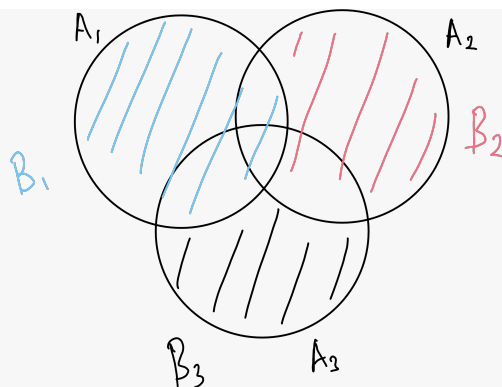
$$\mathbb{P}\left(\bigcup_{n \in \mathbb{N}} A_n\right) \leq \sum_{n \in \mathbb{N}} \mathbb{P}(A_n).$$

Intuition:



$\sum_{n \in \mathbb{N}} \mathbb{P}(A_n)$ “double counts” some sub-events.

Proof. Idea: Rewrite $\bigcup_{n \in \mathbb{N}} A_n$ as a *disjoint* union. Define $B_1 = A_1$ and $B_n = A_n \setminus (A_1 \cup \dots \cup A_{n-1}) \quad \forall n \geq 2$.
 $\underbrace{\hspace{10em}}_{\in \mathcal{F} \text{ (by Sheet 1)}}$



So

- $\bigcup_{n \in \mathbb{N}} B_n = \bigcup_{n \in \mathbb{N}} A_n$.
- $(B_n)_{n \in \mathbb{N}}$ is disjoint (by construction).
- $B_n \subseteq A_n \implies \mathbb{P}(B_n) \leq \mathbb{P}(A_n)$
Q4, Sheet 1

$$\mathbb{P}\left(\bigcup_{n \in \mathbb{N}} A_n\right) = \mathbb{P}\left(\bigcup_{n \in \mathbb{N}} B_n\right) \stackrel{P3 \text{ on } (B_n)}{=} \sum_{n \in \mathbb{N}} \mathbb{P}(B_n) \leq \sum_{n \in \mathbb{N}} \mathbb{P}(A_n)$$

□

§3.2 Continuity

Proposition 3.2 (Continuity)

Let $(A_n)_{n \in \mathbb{N}}$ be an increasing sequence of events in \mathcal{F} , i.e. $A_n \subseteq A_{n+1} \quad \forall n$. Then $\mathbb{P}(A_n) \leq \mathbb{P}(A_{n+1})$. So $\mathbb{P}(A_n)$ converges as $n \rightarrow \infty$.^a

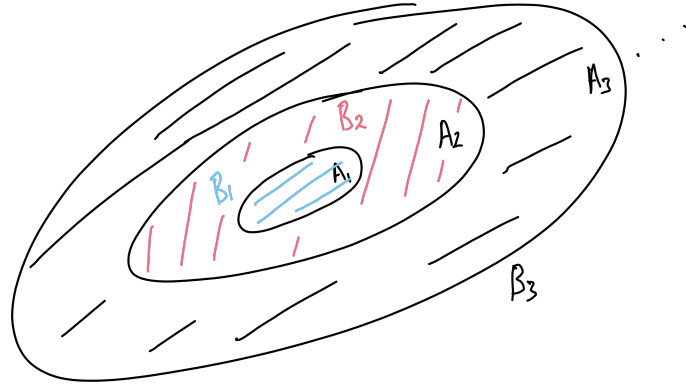
In fact: $\lim_{n \rightarrow \infty} \mathbb{P}(A_n) = \mathbb{P}(\bigcup_{n \in \mathbb{N}} A_n)$.

^aAs probabilities are bounded above by 1 and increasing.

For motivation try Q6, Sheet 1.

Proof. Let us reuse the B_n s from the previous subsection.

- $\bigcup_{k=1}^n B_k = A_n$ (disjoint union).
- $\bigcup_{n \in \mathbb{N}} B_n = \bigcup_{n \in \mathbb{N}} A_n$



$$\mathbb{P}(A_n) = \sum_{k=1}^n \mathbb{P}(B_k) \xrightarrow{n \rightarrow \infty} \sum_{k \geq 1} \mathbb{P}(B_k) = \mathbb{P}\left(\bigcup_{n \in \mathbb{N}} B_n\right) = \mathbb{P}\left(\bigcup_{n \in \mathbb{N}} A_n\right)$$

□

§3.3 Inclusion-Exclusion Principle

Background: $\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B)$.

Similarly: $A, B, C \in \mathcal{F}$

$$\mathbb{P}(A \cup B \cup C) = \mathbb{P}(A) + \mathbb{P}(B) + \mathbb{P}(C) - \mathbb{P}(A \cup B) - \mathbb{P}(B \cup C) - \mathbb{P}(C \cup A) + \mathbb{P}(A \cap B \cap C).$$

Proposition 3.3 (Inclusion-Exclusion Principle)

Let $A_1, \dots, A_n \in \mathcal{F}$, then:

$$\begin{aligned} \mathbb{P}\left(\bigcup_{i=1}^n A_i\right) &= \sum_{i=1}^n \mathbb{P}(A_i) - \sum_{1 \leq i_1 < i_2 \leq n} \mathbb{P}(A_{i_1} \cap A_{i_2}) + \sum_{1 \leq i_1 < i_2 < i_3 \leq n} \mathbb{P}(A_{i_1} \cap A_{i_2} \cap A_{i_3}) - \dots \\ &\quad + (-1)^{n+1} \mathbb{P}(A_1 \cap \dots \cap A_n) \\ &= \sum_{\substack{I \subset \{1, \dots, n\} \\ I \neq \emptyset}} (-1)^{|I|+1} \mathbb{P}\left(\bigcap_{i \in I} A_i\right) \end{aligned}$$

Note: $\sum_{1 \leq i_1 < i_2 \leq n}$ is the sum of all triples that are distinct and unordered.