Sample problems for the second lecture (Mar. 8)

1. Let
$$a_n = (-1)^n \cdot (1 + \frac{1}{n})$$
 and $\epsilon = 0.1$.

(a) Find

$$b_n := \sup_{k > n} a_k = \sup\{a_n, a_{n+1}, a_{n+2}, \ldots\}.$$

Is $\{b_n\}$ a **monotonic** sequence?

- (b) Find the value of $\limsup_{n\to\infty} a_n$.
- (c) Show that $a_n < \limsup_{m \to \infty} a_m + \epsilon$ for sufficiently large n.
- (d) Show that $a_n > \limsup_{m \to \infty} a_m \epsilon$ for **infinitely many** n.
- (e) Find

$$c_n := \inf_{k>n} a_k = \inf\{a_n, a_{n+1}, a_{n+2}, \ldots\}.$$

Is $\{c_n\}$ a monotonic sequence?

- (f) Find the value of $\liminf_{n\to\infty} a_n$.
- (g) Show that $a_n > \liminf_{m \to \infty} a_m \epsilon$ for sufficiently large n.
- (h) Show that $a_n < \liminf_{m \to \infty} a_m + \epsilon$ for **infinitely many** n.

Solution.

(a) $\{b_n\} = \{\sup_{k \ge n} a_k\}$ is a monotonically non-increasing sequence by its definition. With $a_n = (-1)^n \cdot (1 + \frac{1}{n})$, we obtain

$$b_n := \sup_{k \ge n} a_k = \sup\{a_n, a_{n+1}, a_{n+2}, \ldots\}$$

$$= \begin{cases} 1 + \frac{1}{n}, & n \text{ even} \\ 1 + \frac{1}{n+1}, & n \text{ odd} \end{cases}$$

$$= 1 + \frac{1}{2\lceil \frac{n}{2} \rceil}.$$

This confirms that $\{b_n\}$ is monotonically non-increasing (not strictly decreasing).

(b)

$$\limsup_{n \to \infty} a_n = \lim_{n \to \infty} \left(\sup_{k \ge n} a_n \right) \text{ (By definition of limsup)}$$

$$= \lim_{n \to \infty} b_n$$
(The limit always exists by the monotone convergence theorem over $\mathbb{R} \cup \{\pm \infty\}$.)
$$= 1.$$

(c)

$$a_n < \limsup_{m \to \infty} a_m + \epsilon \Leftrightarrow (-1)^n \cdot \left(1 + \frac{1}{n}\right) < 1 + 0.1$$

which is true for all $n \ge 11$, i.e., for sufficiently large n.

(d)

$$a_n > \limsup_{m \to \infty} a_m - \epsilon \Leftrightarrow (-1)^n \cdot \left(1 + \frac{1}{n}\right) > 1 - 0.1 = 0.9.$$

Hence, for a given (positive) integer K, there exists **one** N > K such that $a_n > 0.9$ (e.g., $N = 2\lceil \frac{K+1}{2} \rceil$).

(e) $\{b_n\} = \{\inf_{k \ge n} a_k\}$ is a monotonically non-decreasing sequence by its definition. With $a_n = (-1)^n \cdot (1 + \frac{1}{n})$, we obtain

$$c_n := \sup_{k \ge n} a_k = \inf\{a_n, a_{n+1}, a_{n+2}, \dots\}$$

$$= \begin{cases} -\left(1 + \frac{1}{n}\right), & n \text{ odd} \\ -\left(1 + \frac{1}{n+1}\right), & n \text{ even} \end{cases}$$

$$= -\left(1 + \frac{1}{2\lfloor \frac{n}{2} \rfloor + 1}\right).$$

This confirms that $\{c_n\}$ is monotonically non-decreasing (not strictly increasing).

(f)

$$\lim_{n \to \infty} \inf a_n = \lim_{n \to \infty} \left(\inf_{k \ge n} a_n \right) \quad \text{(By definition of limsup)}$$

$$= \lim_{n \to \infty} c_n$$
(The limit always exists by the monotone convergence theorem over $\mathbb{R} \cup \{\pm \infty\}$.)
$$= -1.$$

(g)

$$a_n > \liminf_{m \to \infty} a_m - \epsilon \Leftrightarrow (-1)^n \cdot \left(1 + \frac{1}{n}\right) > -1 - 0.1$$

which is true for all $n \geq 10$, i.e., for sufficiently large n.

(h)

$$a_n < \liminf_{m \to \infty} a_m + \epsilon \Leftrightarrow (-1)^n \cdot \left(1 + \frac{1}{n}\right) < -1 + 0.1 = -0.9.$$

Hence, for a given (positive) integer K, there exists **one** N > K such that $a_n < -0.9$ (e.g., $N = 2 \lceil \frac{K+1}{2} \rceil + 1$).

2. Give an example of $\{a_n\}$ and $\{b_n\}$ such that the following strict inequalities hold:

$$\lim_{n \to \infty} \inf a_n + \lim_{n \to \infty} \inf b_n < \lim_{n \to \infty} \inf (a_n + b_n)
< \lim_{n \to \infty} \sup a_n + \lim_{n \to \infty} \inf b_n
< \lim_{n \to \infty} \sup (a_n + b_n)
< \lim_{n \to \infty} \sup a_n + \lim_{n \to \infty} \sup b_n.$$

Solution. Let $a_n = (-1)^n$. Then, the above inequalities become:

$$-1 + \liminf_{n \to \infty} b_n < \liminf_{n \to \infty} ((-1)^n + b_n)$$

$$< 1 + \liminf_{n \to \infty} b_n$$

$$< \limsup_{n \to \infty} ((-1)^n + b_n)$$

$$< 1 + \limsup_{n \to \infty} b_n,$$

which are equivalent to:

$$-1 + \liminf_{n \to \infty} b_n < \min \left\{ \liminf_{m \to \infty} (-1 + b_{2m+1}), \liminf_{m \to \infty} (1 + b_{2m}) \right\}$$

$$< 1 + \liminf_{n \to \infty} b_n$$

$$< \max \left\{ \limsup_{m \to \infty} (-1 + b_{2m+1}), \limsup_{m \to \infty} (1 + b_{2m}) \right\}$$

$$< 1 + \limsup_{n \to \infty} b_n.$$

These can be equivalently rewritten as:

$$-1 + \liminf_{n \to \infty} b_n < \min \left\{ -1 + \liminf_{m \to \infty} b_{2m+1}, 1 + \liminf_{m \to \infty} b_{2m} \right\}$$
(1)
$$< 1 + \liminf_{n \to \infty} b_n$$
(2)
$$< \max \left\{ -1 + \limsup_{m \to \infty} b_{2m+1}, 1 + \limsup_{m \to \infty} b_{2m} \right\}$$
(3)
$$< 1 + \limsup_{n \to \infty} b_n.$$
(4)

Since

$$\liminf_{n \to \infty} b_n = \min \Big\{ \liminf_{m \to \infty} b_{2m+1}, \liminf_{m \to \infty} b_{2m} \Big\},$$

the validations of (1) and (2) require

$$\underbrace{\liminf_{n \to \infty} b_n}_{=A} = \liminf_{m \to \infty} b_{2m} < \underbrace{\liminf_{m \to \infty} b_{2m+1}}_{=B}$$

and

$$-1 + A < \min \{-1 + B, 1 + A\} = -1 + B$$

< 1 + A.

A solution is to let B = 1 + A.

Similarly, we note that

$$\limsup_{n \to \infty} b_n = \max \Big\{ \limsup_{m \to \infty} b_{2m+1}, \limsup_{m \to \infty} b_{2m} \Big\},\,$$

for which the validation of (4) requires

$$\underbrace{\limsup_{n \to \infty} b_n}_{=C} = \limsup_{m \to \infty} b_{2m+1} > \underbrace{\limsup_{m \to \infty} b_{2m}}_{=D}$$

and

$$\max\{-1 + C, 1 + D\} < 1 + C.$$

A solution is to let C = 1 + D.

As a consequence, any $\{b_n\}$ satisfying

$$\begin{cases} \liminf_{m \to \infty} b_{2m} = A < \limsup_{m \to \infty} b_{2m} = D \\ \liminf_{m \to \infty} b_{2m+1} = 1 + A < \limsup_{m \to \infty} b_{2m+1} = 1 + D \end{cases}$$

is the solution.

3. Suppose the probability space is given as

$$\Omega = \{ \blacktriangle, \blacktriangledown, \square, \blacksquare, \diamondsuit, \blacklozenge \}
\mathcal{F} = \{ \text{all subsets of } \Omega \} = \text{powerset of } \Omega
P = \begin{cases} P(\blacktriangle) = 0.05, P(\blacktriangledown) = 0.1, P(\square) = 0.15, \\ P(\blacksquare) = 0.2, P(\diamondsuit) = 0.25, P(\spadesuit) = 0.25 \end{cases}$$

Give a random process X_1, X_2, X_3, \ldots , each of which is defined over the above probability space. Let

$$X_3(\blacktriangle) = 1;$$
 $X_3(\blacksquare) = 2;$ $X_7(\blacktriangle) = 1;$ $X_7(\blacksquare) = 2$
 $X_3(\blacktriangledown) = 2;$ $X_3(\lozenge) = 1;$ $X_7(\blacktriangledown) = 1;$ $X_7(\lozenge) = 2$
 $X_3(\square) = 1;$ $X_3(\spadesuit) = 2;$ $X_7(\square) = 1;$ $X_7(\spadesuit) = 2$

Then the distributions of X_3 and X_7 , as well as the joint distribution between X_3 and X_7 , are well defined (by simply providing the mappings sperately) as indicated in the following.

- (a) Find $Pr[X_3 = 1]$.
- (b) Find $\Pr[X_7 = 2]$.
- (c) Find $Pr[X_3 = 1 \text{ and } X_7 = 2].$

Solution.

(a)

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\Pr[X_3 = 1] = P\left(\{\omega \in \Omega : X_3(\omega) = 1\}\right) \quad \text{(By the function mapping of } X_3\text{)}
\text{(We can deal with "=" because } X_3 \text{ is a real-valued function mapping.)}
= P\left(\{\blacktriangle, \square, \blacklozenge\}\right) \quad \text{(Back to the probability measure } P\text{)}
= P\left(\{\blacktriangle\} \cup \{\square\} \cup \{\lozenge\}\right)
= P\left(\{\blacktriangle\}\right) + P\left(\{\square\}\right) + P\left(\{\lozenge\}\right) \quad \text{(Axiom 3 of probability measure: Countable additivity for disjoint sets)}
= 0.05 + 0.15 + 0.25
= 0.45.
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(b)

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\Pr[X_7 = 2] = P(\{\omega \in \Omega : X_7(\omega) = 2\}) \quad \text{(By the function mapping of } X_7)
(\text{We can deal with "=" because } X_7 \text{ is a real-valued function mapping.)}
= P(\{\blacksquare, \lozenge, \spadesuit\}) \quad \text{(Back to the probability measure } P)
= P(\{\blacksquare\} \cup \{\lozenge\} \cup \{\spadesuit\}\})
= P(\{\blacksquare\}) + P(\{\lozenge\}) + P(\{\diamondsuit\}\}) \quad \text{(Axiom 3 of probability measure: Countable additivity for disjoint sets)}
= 0.2 + 0.25 + 0.25
= 0.7.
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(c)

$$\Pr[X_3 = 1 \text{ and } X_7 = 2] = P\left(\{\omega \in \Omega : X_3(\omega) = 1 \text{ and } X_7(\omega) = 2\}\right)$$

$$= P\left(\{\omega \in \Omega : X_3(\omega) = 1\} \cap \{\omega \in \Omega : X_7(\omega) = 2\}\right)$$

$$= P\left(\{\blacktriangle, \square, \spadesuit\} \cap \{\blacksquare, \lozenge, \spadesuit\}\right)$$

$$= P\left(\{\clubsuit\}\right)$$

$$= 0.25.$$

- 4. An event E is **ergodic** if $\mathbb{T}^{-1}(E) = E$. A random process is ergodic if all ergodic events are either with probability one or with probability zero.
 - (a) For events over two-sided sequences, the inverse mapping \mathbb{T}^{-1} can be defined as:

$$\mathbb{T}^{-1}(\boldsymbol{x}) = \mathbb{T}^{-1}(\dots, x_{-2}, x_{-1}, x_0, x_1, x_2, \dots) = \dots, y_{-2}, y_{-1}, y_0, y_1, y_2, \dots,$$

with $y_n = x_{n-1}$. Give the ergodic set E of two-sided sequences over $\{0,1\}^{\infty}$, which contain

$$\dots, 0, 1, 0, 1, 0, 1, \dots$$

List all possible values of

$$\lim_{n \to \infty} \frac{x_{-n} + \dots + x_{-2} + x_{-1} + x_0 + x_1 + x_2 + \dots + x_n}{2n + 1}$$

for $\boldsymbol{x} \in E$. Is

$$\cdots = \mathbb{T}^{-2}(E) = \mathbb{T}^{-1}(E) = E = \mathbb{T}(E) = \mathbb{T}^{2}(E) = \cdots$$

i.e., the set stays the same as time varies.

(b) For events over one-sided sequences as considered in most communication systems, the inverse (time-shift) mapping \mathbb{T}^{-1} in general does not exist. As an extension, Shields [P. C. Shields 1991, p. 3] adopted the definition as the following:

$$\mathbb{T}^{-1}E := \{ \boldsymbol{x} \in \mathcal{X}^{\infty} \colon \mathbb{T}\boldsymbol{x} \in E \}$$

which includes all right-shift counterparts of "transient" elements into E (and is consistent with the definition of \mathbb{T}^{-1} for sets over

two-sided sequences). Give the ergodic set E of one-sided sequences over $\{0,1\}^{\infty}$, which contain

$$0, 1, 0, 1, 0, 1, \ldots$$

List all possible values of

$$\lim_{n\to\infty} \frac{x_1 + x_2 + \dots + x_n}{n}$$

for
$$\mathbf{x} = (x_1, x_2, x_3, \ldots) \in E$$
. Is

$$\cdots = \mathbb{T}^{-2}(E) = \mathbb{T}^{-1}(E) = E = \mathbb{T}(E) = \mathbb{T}^{2}(E) = \cdots$$

i.e., the set stays the same as time varies.

Solution.

(a) $E = \{ \boldsymbol{x} \in \{0, 1\}^{\infty} : (\forall n) x_n = x_{n+2} \text{ and } x_n \neq x_{n+1} \}$ and the limit is unique for all elements in E and is equal to $\frac{1}{2}$.

The set definitely stays the same as time varies.

(b)

$$E = \bigcup_{k=1}^{\infty} E_k,$$

where

$$E_k = \{ \boldsymbol{x} \in \{0, 1\}^{\infty} : (\forall \ n \ge k) \ x_n = x_{n+2} \text{ and } x_n \ne x_{n+1} \}.$$

For example,

$$E_1 = \{010101..., 101010...\},\$$

$$E_2 = \{ {\color{red}0010101 \dots, {\color{red}1010101 \dots, {\color{red}0101010 \dots, {\color{red}1101010 \dots}}} \},$$

and

$$E_3 = \{00010101..., 01010101..., 10010101..., 11010101..., 01010101..., 01101010..., 10101010..., 11101010...\}$$

For $x \in E_k$ (with k fixed), we obtain

$$\lim_{n \to \infty} \frac{x_1 + \dots + x_{k-1} + x_k + \dots + x_n}{n}$$

$$= \lim_{n \to \infty} \left(\underbrace{\frac{x_1 + \dots + x_{k-1}}{n}}_{\text{go to zero as } n \to \infty} + \underbrace{\frac{x_k + \dots + x_n}{n}}_{n} \right)$$

$$= \lim_{n \to \infty} \left(\underbrace{\frac{n - k + 1}{2}}_{n} \right) = \frac{1}{2}.$$

Thus, the limit is unique for all elements in E and is equal to $\frac{1}{2}$. This set definitely stays the same as time varies (either by "chopping" as \mathbb{T} or by "enlarging" as \mathbb{T}^{-1}).

- 5. Suppose Jack and Mary separately make consecutive observations on the same phenomenon in a one-observation-at-a-time manner, and then compute the time averages based on their own observations. Answer the following two questions for each of the subproblems.
 - Does it guarantee that Jack and Mary will conclude almost the same average value as the number of observations they make grows very large?
 - Does it guarantee that their average value equals the ensemble average (i.e., expected value of the phenomenon) as the number of observations tends to infinity?
 - (a) The observations on the phenomenon constitute an ergodic process.
 - (b) The observations on the phenomenon constitute a stationary process.
 - (c) The observations on the phenomenon constitute a stationary-ergodic process.
 - (d) The observations on the phenomenon follow the law of large numbers.
 - (e) The observations on the phenomenon constitute an independent and identically distributed (i.i.d.) random process.
 - (f) The observations on the phenomenon constitute a memoryless random process.
 - (g) The observations on the phenomenon constitute a Markov process.

Solution.

- (a) Yes for the first question but no for the second; in other words, the time average may not equal to the ensemble average.
- (b) No for both questions; in other words, Jack and Mary may obtain very different averages even if both accurately and professionally do their measurements.
- (c) Yes for both questions.

- (d) Yes for both questions. This is exactly the behavior dictated by the law of large numbers, i.e., the time average converges to the ensemble average.
- (e) Yes for both questions as an i.i.d. random process is stationary ergodic.
- (f) Yes for both questions as "memoryless" is an alternative name for "i.i.d."
- (g) Not necessarily for both questions since a Markov process can be neither stationary nor ergodic. Note that an irreducible, homogeneous, aperiodic, finite-state Markov process with stationary initial probability is stationary ergodic.