



模式识别 Pattern Recognition

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1 高斯混合模型

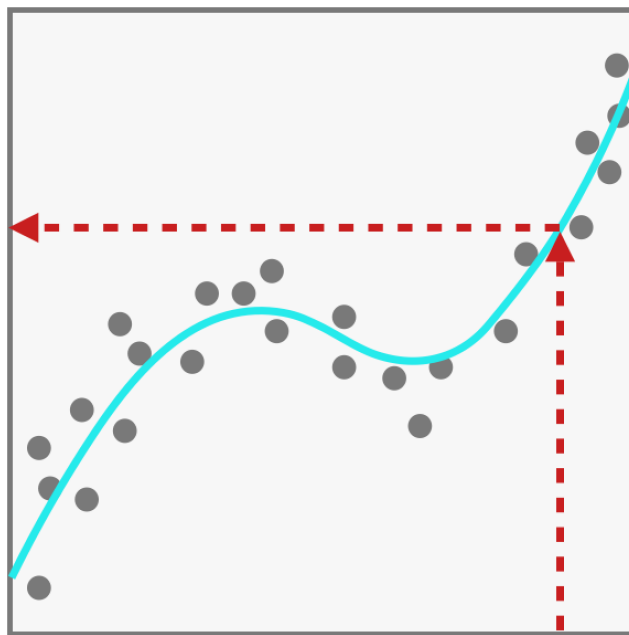
2 模型优化

3 医学图像应用

Recap: 线性回归



- 通过已知的输入数据（特征）预测一个连续的数值输出。



Recap: 线性回归



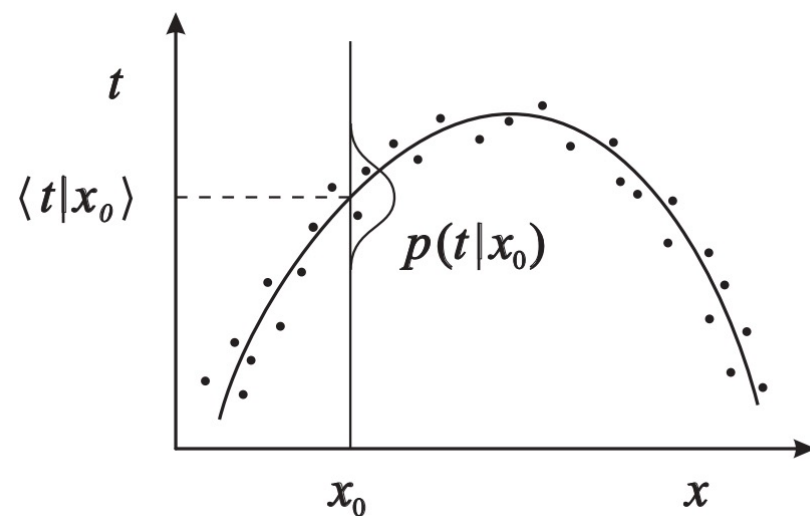
- 对于回归任务的MLE估计本质就是高斯假设下的最小二乘法

Example (contd.)

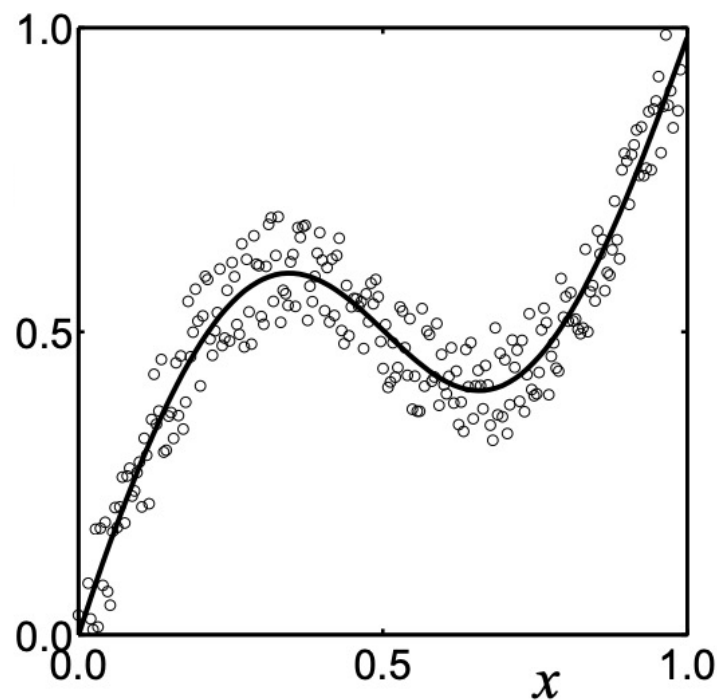
$$\begin{aligned}\mathcal{L}(\theta) &= -\sum_{i=1}^N \log p(y_i | \mathbf{x}_i, \theta) = -\sum_{i=1}^N \log \mathcal{N}(y_i | \mathbf{x}_i^\top \theta, \sigma^2) \\ &= -\sum_{i=1}^N \log \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y_i - \mathbf{x}_i^\top \theta)^2}{2\sigma^2}\right) \\ &= -\sum_{i=1}^N \log \exp\left(-\frac{(y_i - \mathbf{x}_i^\top \theta)^2}{2\sigma^2}\right) - \sum_{i=1}^N \log \frac{1}{\sqrt{2\pi\sigma^2}} \\ &= \frac{1}{2\sigma^2} \sum_{i=1}^N (y_i - \mathbf{x}_i^\top \theta)^2 - \sum_{i=1}^N \log \frac{1}{\sqrt{2\pi\sigma^2}}.\end{aligned}$$

The second term is **constant**.

\Rightarrow minimizing $\mathcal{L}(\theta) \Rightarrow$ solving the least-squares problem.



- 建立一个模型，使其能够根据输入的特征预测一个数值型的目标变量

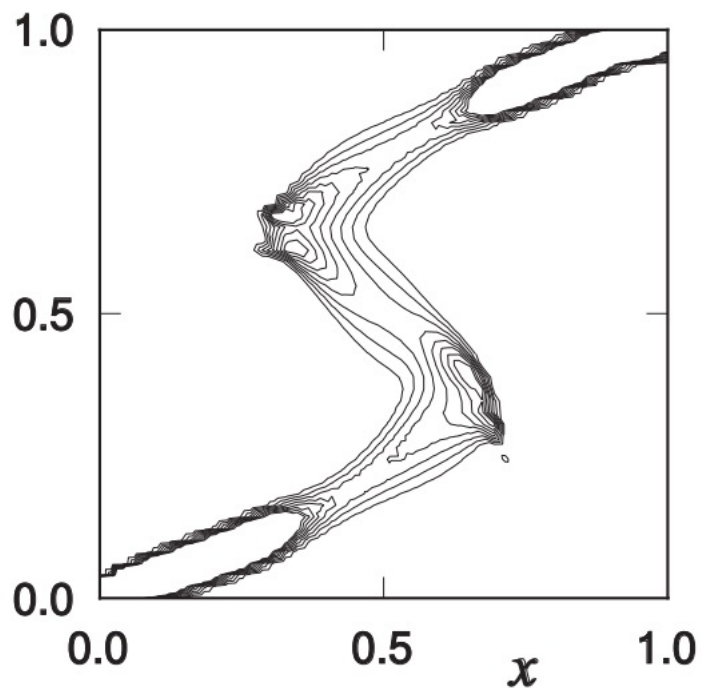


- But wait, 类似的输入特征一定是同样的解吗？

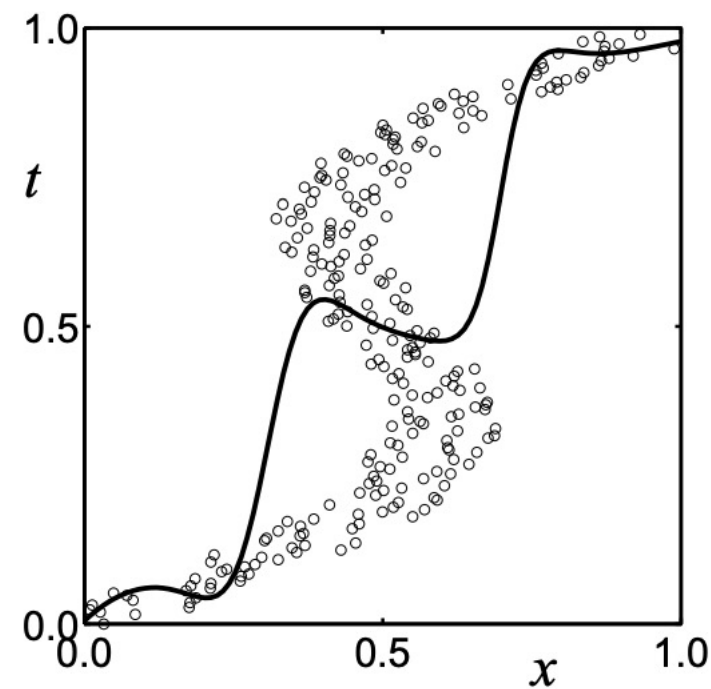
线性回归的局限



- How about this?
- 同一个输入特征，有多个可能的输出



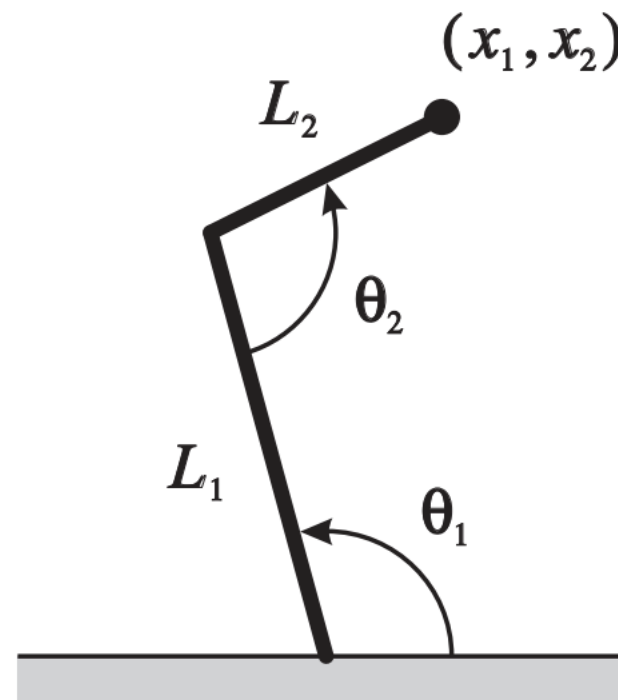
线性回归



- 现实中的例子：机械臂
- 需要从观测得到的 (x_1, x_2) 推断出 (θ_1, θ_2)

$$x_1 = L_1 \cos(\theta_1) - L_2 \cos(\theta_1 + \theta_2)$$

$$x_2 = L_1 \sin(\theta_1) - L_2 \sin(\theta_1 + \theta_2)$$

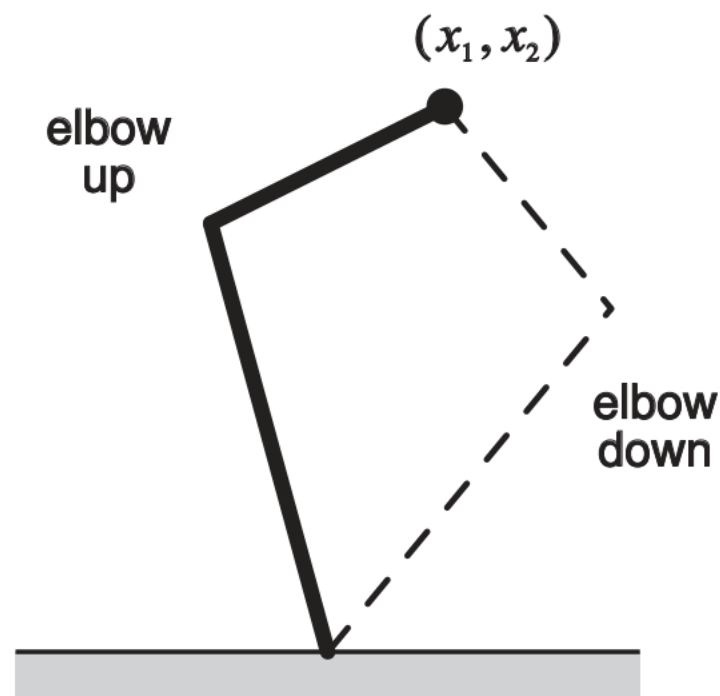


- 现实中的例子：机械臂
- 需要从观测得到的 (x_1, x_2) 推断出 (θ_1, θ_2)

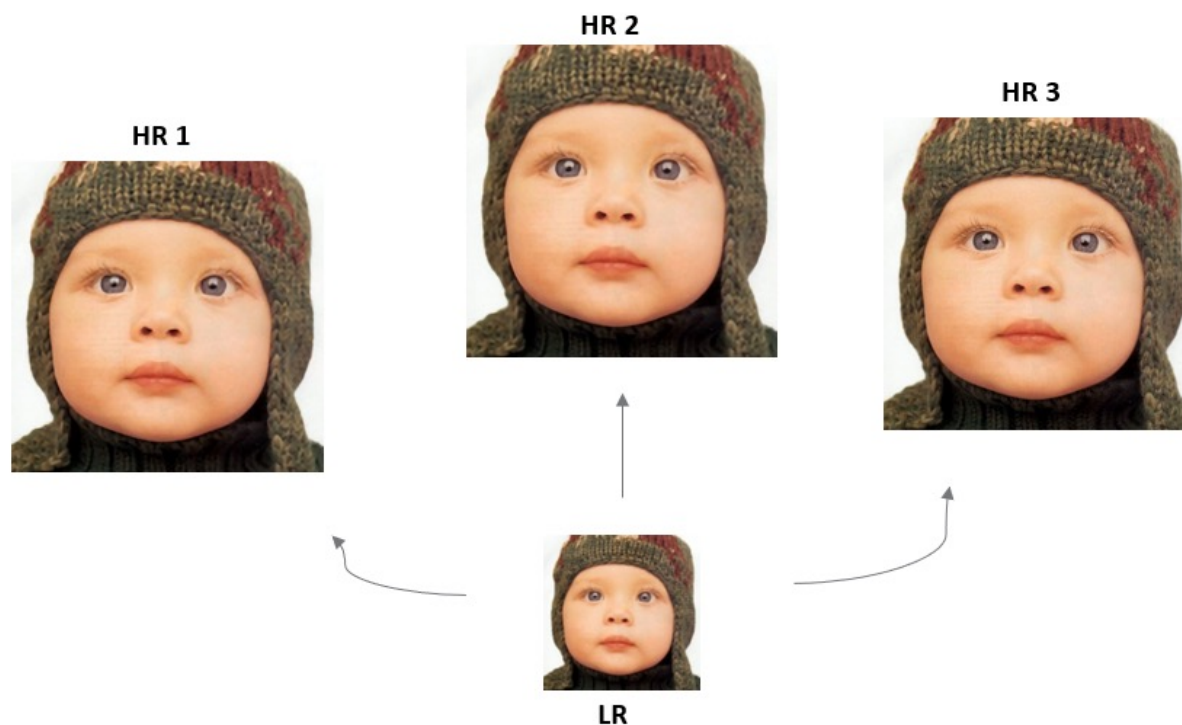
$$x_1 = L_1 \cos(\theta_1) - L_2 \cos(\theta_1 + \theta_2)$$

$$x_2 = L_1 \sin(\theta_1) - L_2 \sin(\theta_1 + \theta_2)$$

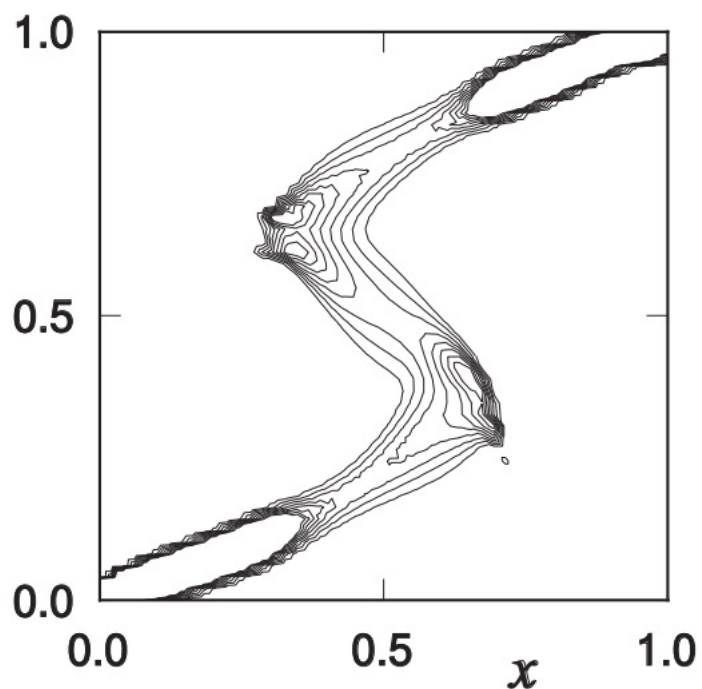
- 但是问题也是有两个解的



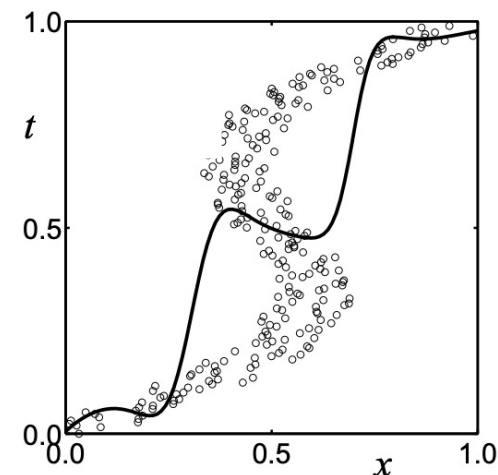
- 现实中的例子：图像重建
- 一个多分辨率，可以对应多个高分辨率
- “ill-posed” problem



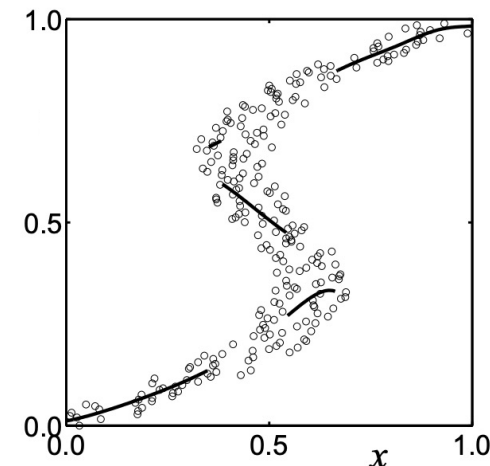
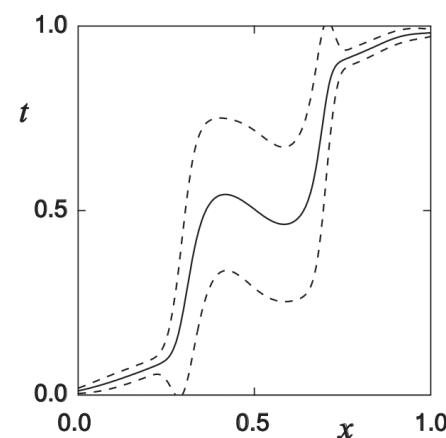
- 混合模型可以对多输出进行建模
- 对于同一个输入特征，取最大可能输出



线性回归



混合模型



- 利用K的简单分布的凸集合来表示一个复杂的分布

$$p(\mathbf{x}) = \sum_{k=1}^K \pi_k p_k(\mathbf{x}),$$

$$0 \leq \pi_k \leq 1, \sum_{k=1}^K \pi_k = 1.$$

π_k : *mixture weights*.

- 为了简单，假设每一个子分布都服从高斯分布

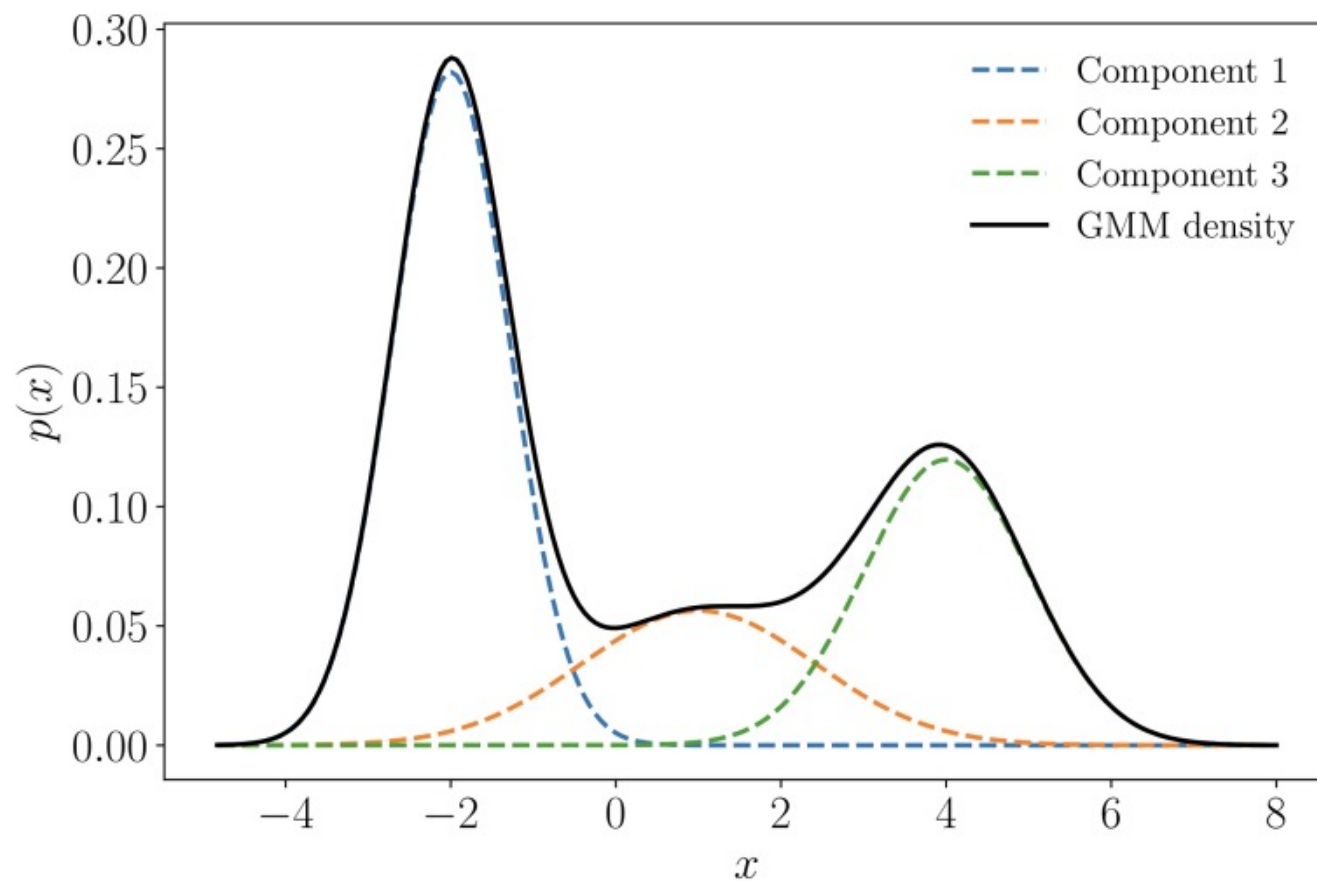
Gaussian Mixture Model

A Gaussian mixture model is a density model where we combine a finite number of K Gaussian distributions $\mathcal{N}(\mathbf{x} \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$ such that

$$p(\mathbf{x} \mid \boldsymbol{\theta}) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x} \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

$$0 \leq \pi_k \leq 1, \sum_{k=1}^K \pi_k = 1,$$

where $\boldsymbol{\theta} := \{\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k, \pi_k \mid k = 1, \dots, K\}$.



$$p(x | \theta) = 0.5\mathcal{N}(x | -2, 0.5) + 0.2\mathcal{N}(x | 1, 2) + 0.3\mathcal{N}(x | 4, 1).$$

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- 问题假设

A dataset $\mathcal{X} = \{\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_N\}$, where each \mathbf{x}_i is drawn i.i.d. from an unknown distribution $p(\mathbf{x})$.

Parameters: $\boldsymbol{\theta} := \{\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k, \pi_k \mid k = 1, \dots, K\}$.

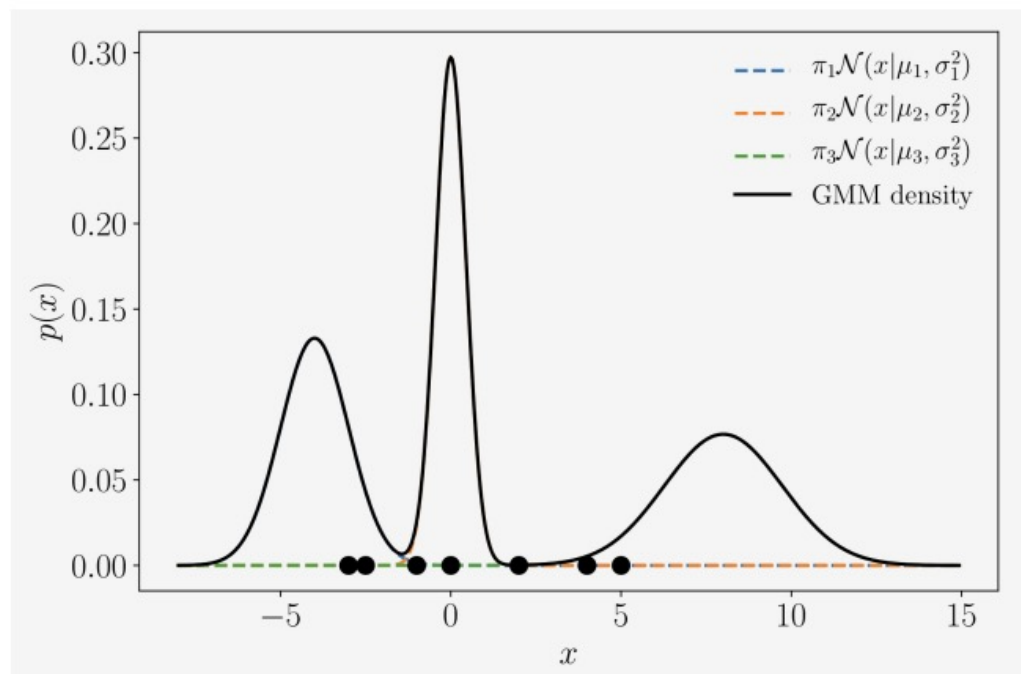
- 初始状态

$$\mathcal{X} = \{-3, -2.5, -1, 0, 2, 4, 5\}.$$

$$K = 3.$$

$$p_1(x) = \mathcal{N}(x \mid -4, 1), \quad p_2(x) = \mathcal{N}(x \mid 0, 0.2), \quad p_3(x) = \mathcal{N}(x \mid 8, 3).$$

$$\pi_1 = \pi_2 = \pi_3 = 1/3.$$



- 利用最大似然估计（MLE）来计算参数

By the i.i.d. assumption, we have the factorized likelihood

$$p(\mathcal{X} | \boldsymbol{\theta}) = \prod_{i=1}^N p(\mathbf{x}_i | \boldsymbol{\theta}), \quad p(\mathbf{x}_i | \boldsymbol{\theta}) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k).$$

Then the log-likelihood is

$$\mathcal{L} := \log p(\mathcal{X} | \boldsymbol{\theta}) = \sum_{i=1}^N \log p(\mathbf{x}_i | \boldsymbol{\theta}) = \sum_{i=1}^N \log \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k).$$

- 问题是，需要优化的参数太多
- 所以用EM算法来优化参数
- 本质上就是一个时间只优化一个参数，固定其他参数

- 对三个变量分别分析

Necessary conditions for a local optimum of \mathcal{L} :

$$\begin{aligned}\frac{\partial \mathcal{L}}{\partial \boldsymbol{\mu}_k} = \mathbf{0}^\top &\iff \sum_{i=1}^N \frac{\partial \log p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \boldsymbol{\mu}_k} = \mathbf{0}^\top \\ \frac{\partial \mathcal{L}}{\partial \boldsymbol{\Sigma}_k} = \mathbf{0}^\top &\iff \sum_{i=1}^N \frac{\partial \log p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \boldsymbol{\Sigma}_k} = \mathbf{0}^\top \\ \frac{\partial \mathcal{L}}{\partial \pi_k} = 0 &\iff \sum_{i=1}^N \frac{\partial \log p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \pi_k} = 0.\end{aligned}$$

Applying the chain rule:

$$\frac{\partial \log p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = \frac{1}{p(\mathbf{x}_i | \boldsymbol{\theta})} \frac{\partial p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \boldsymbol{\theta}}$$

and

$$\frac{1}{p(\mathbf{x}_i | \boldsymbol{\theta})} = \frac{1}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}.$$

- 对三个变量分别分析
- 分别进行推导可以得到

Theorem [Update of the Means]

The update of the mean parameters μ_k , $k = 1, \dots, K$, of the GMM is given by

$$\mu_k^{new} = \frac{\sum_{i=1}^N r_{ik} \mathbf{x}_i}{\sum_{i=1}^N r_{ik}}.$$

Theorem [Update of the Covariances]

The update of the covariance parameters Σ_k , $k = 1, \dots, K$, of the GMM is given by

$$\Sigma_k^{new} = \frac{1}{N_k} \sum_{i=1}^N r_{ik} (\mathbf{x}_i - \mu_k)(\mathbf{x}_i - \mu_k)^\top,$$

where

$$r_{ik} := \frac{\pi_k \mathcal{N}(\mathbf{x}_i | \mu_k, \Sigma_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_i | \mu_j, \Sigma_j)}.$$

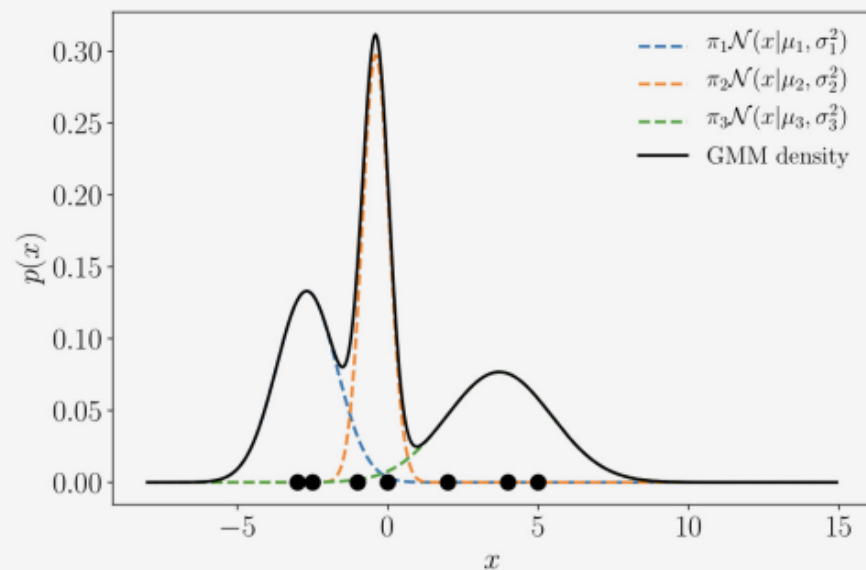
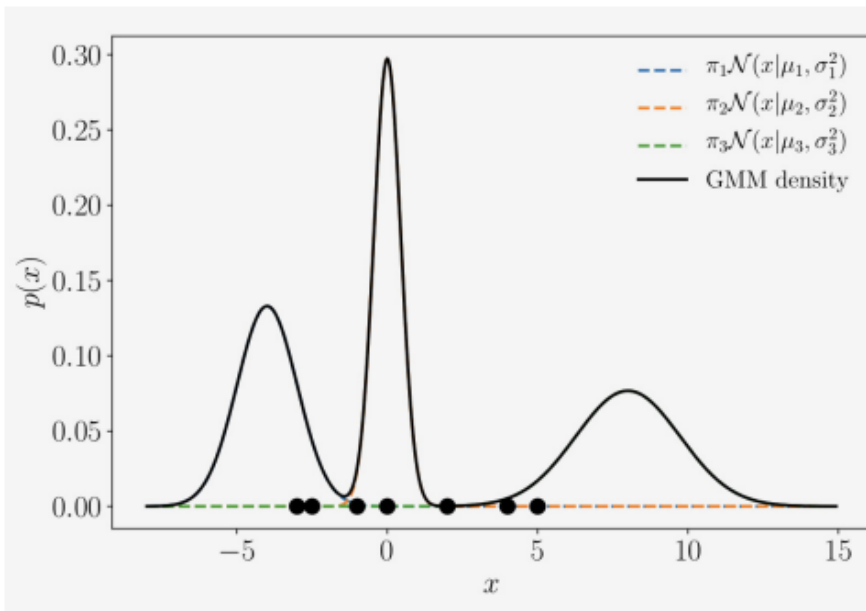
and $N_k := \sum_{i=1}^N r_{ik}$.

Theorem [Update of the Mixture Weights]

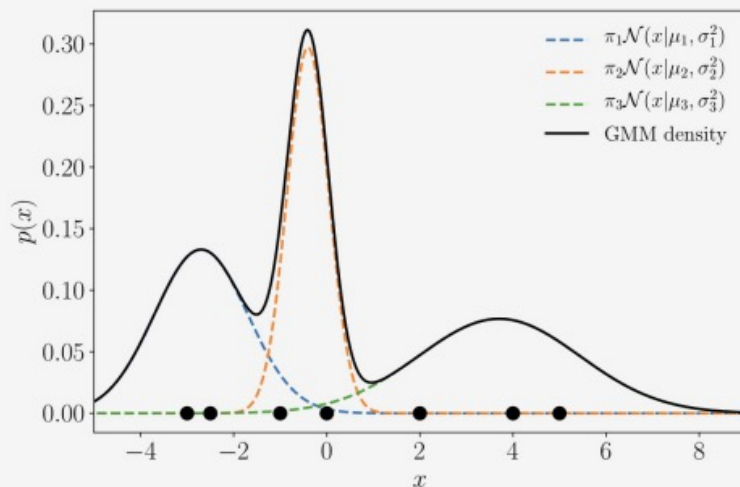
The update of the mixture weights of the GMM is given by

$$\pi_k^{new} = \frac{N_k}{N}, \quad k = 1, \dots, K.$$

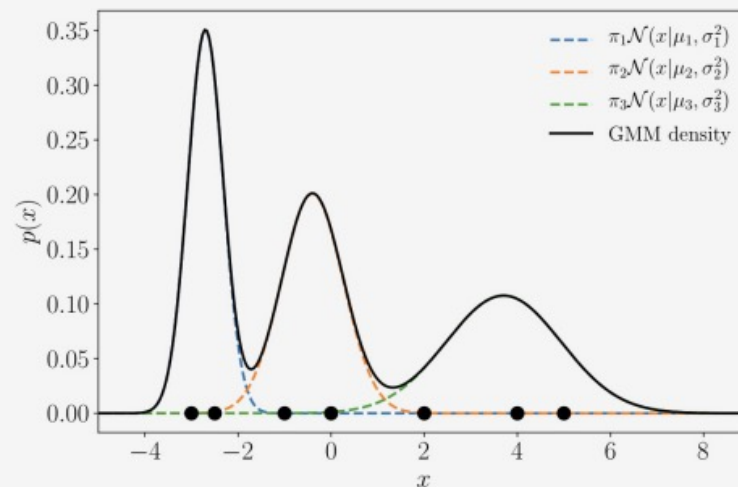
- N : the number of data points.
- $N_k := \sum_{i=1}^N r_{ik}$.



- $\mu_1 : -4 \rightarrow -2.7.$
- $\mu_2 : 0 \rightarrow -0.4.$
- $\mu_3 : 8 \rightarrow 3.7.$

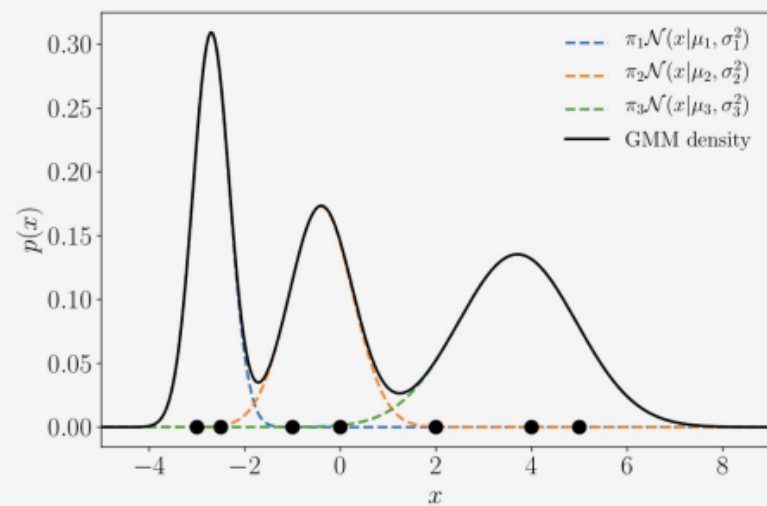
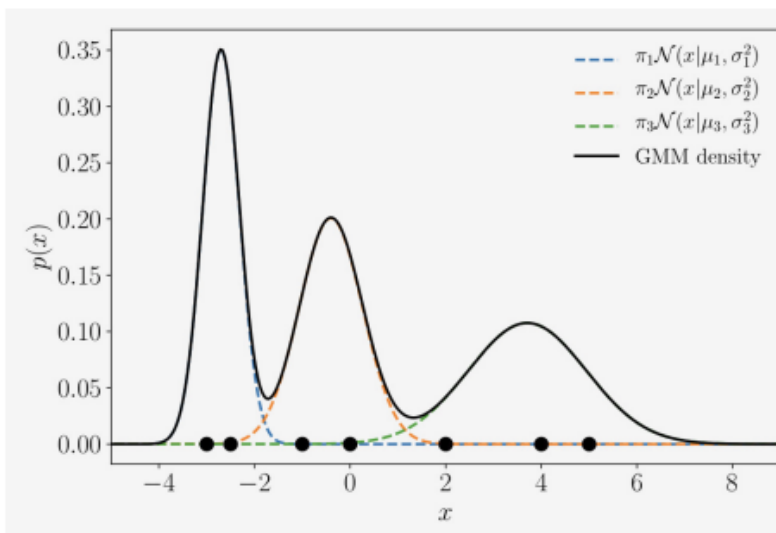


(a) GMM density and individual components prior to updating the variances.



(b) GMM density and individual components after updating the variances.

- $\sigma_1^2 : 1 \rightarrow 0.14$.
- $\sigma_2^2 : 0.2 \rightarrow 0.44$.
- $\sigma_3^2 : 3 \rightarrow 1.53$.



- $\pi_1 : \frac{1}{3} \rightarrow 0.29.$
- $\pi_2 : \frac{1}{3} \rightarrow 0.29.$
- $\pi_3 : \frac{1}{3} \rightarrow 0.42.$

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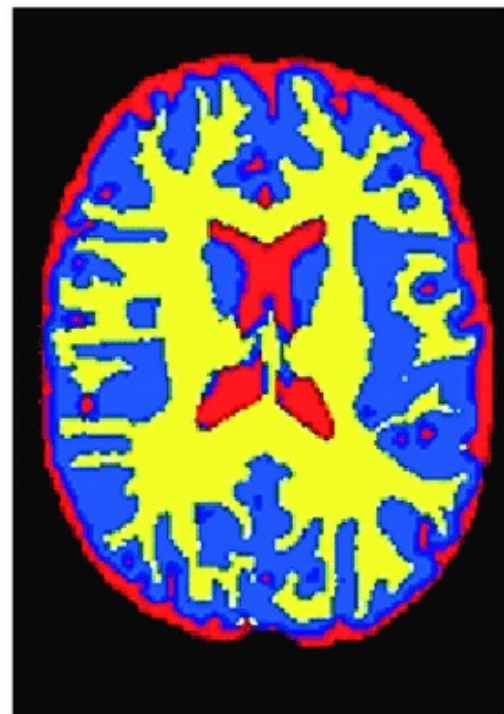
3 医学图像应用

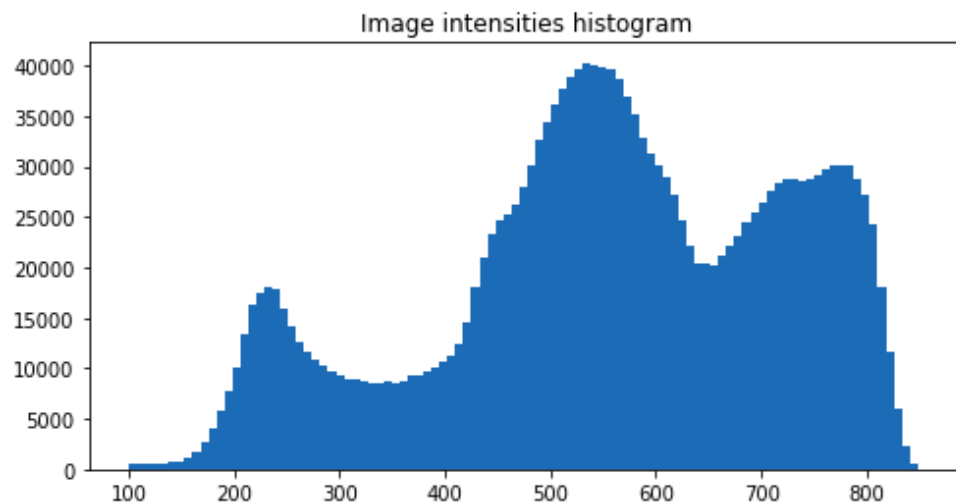
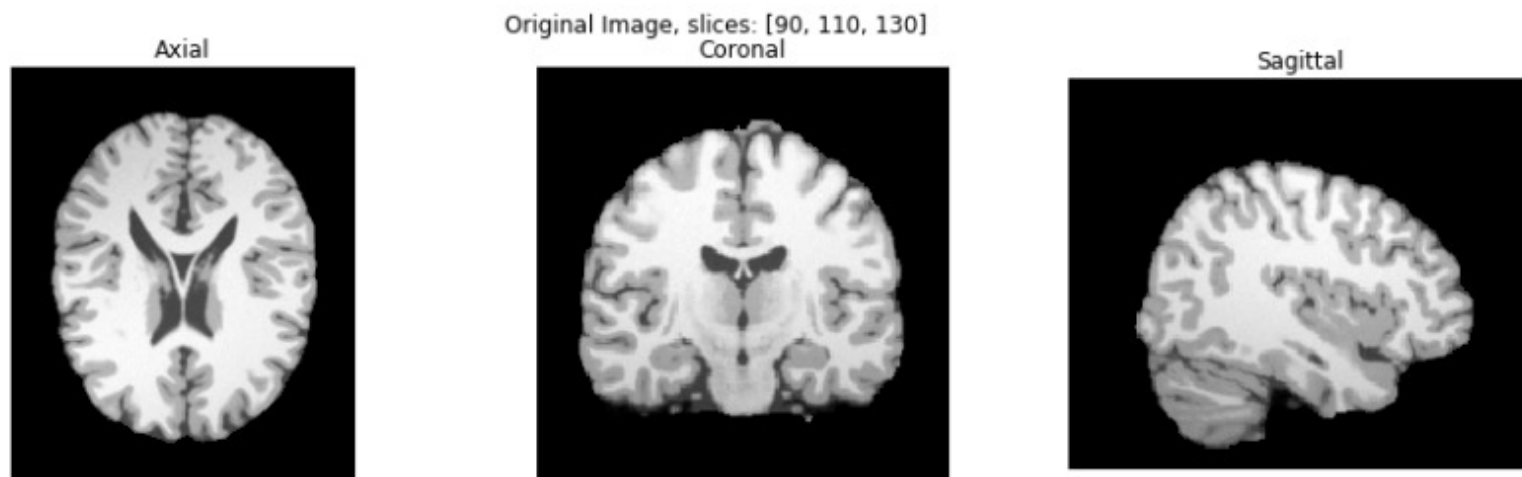
- 将MRI图像分割成为GM, WM和CSF

A



B



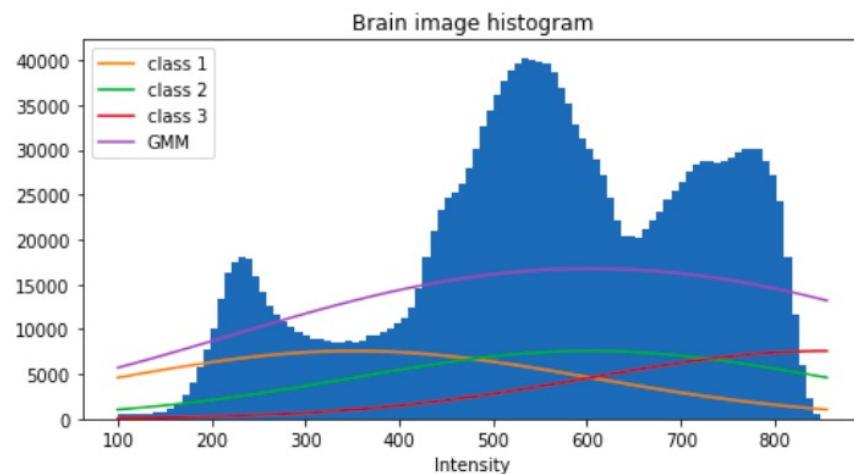


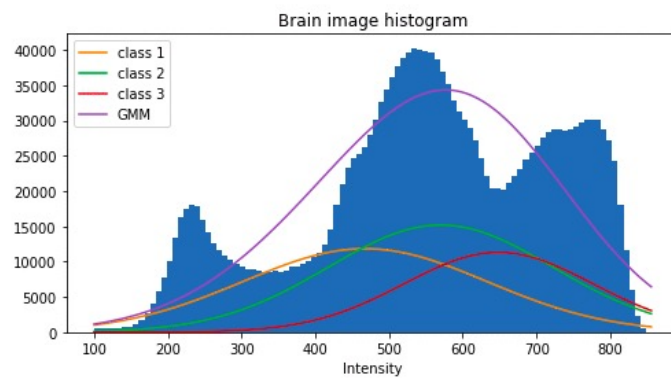
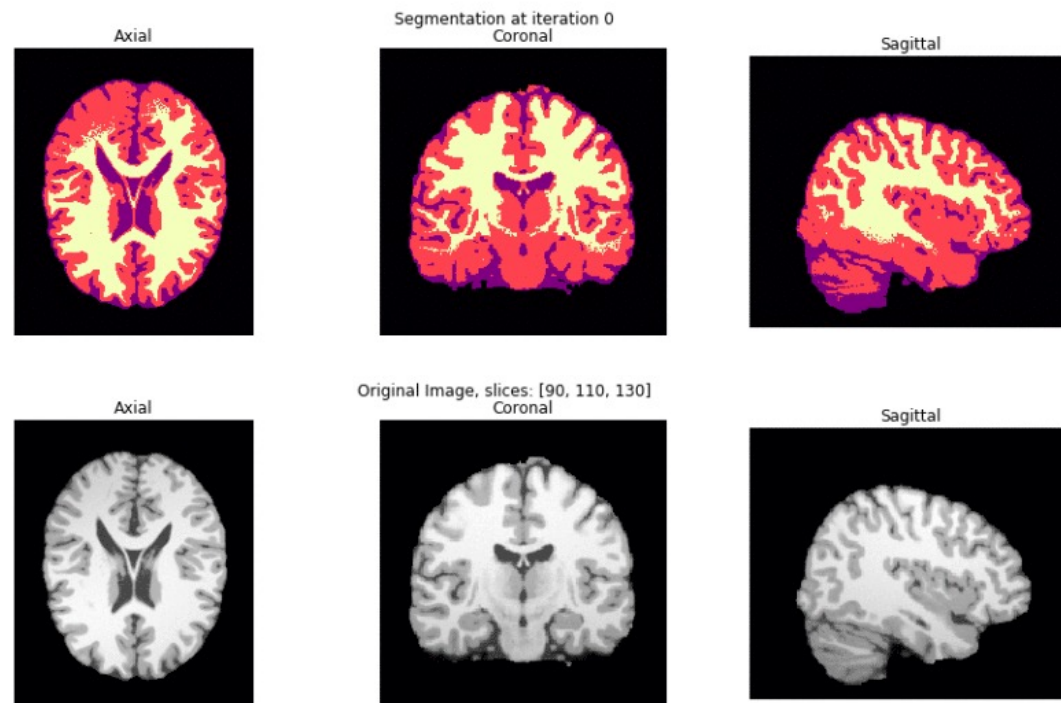
- 假设 D 是图像, I 是期望得到的标注
- 我们希望最大化 $p(I|D)$
- 直接优化就是如果有K-means, 只能得到hard label, 那么这个决定是武断的。我们丢失了“这个点很有可能是白质”的重要信息。

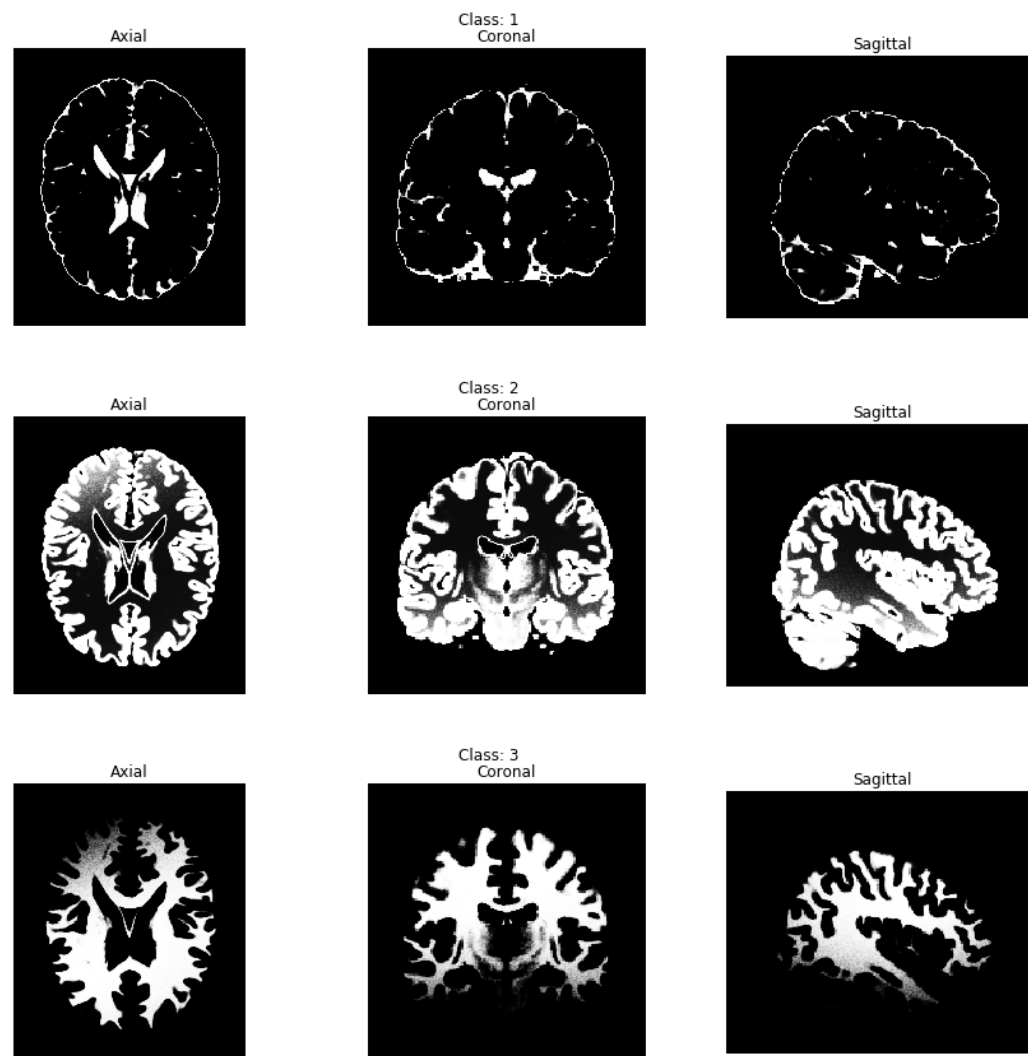
- 根据贝叶斯法则，我们有 $p(I|D) \propto p(D|I)p(I)$
- 分割->不同组织下的后验概率加权之和最大

- 根据贝叶斯法则，我们有 $p(I|D) \propto p(D|I) p(I)$
- 分割 \rightarrow 不同组织下的后验概率加权之和最大

$$\ln p(\mathbf{D}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \sum_i^I \ln \left\{ \sum_k^K \pi_k \mathcal{N}(\mathbf{d}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\}.$$







- 回归模型不能对多输出情况进行建模
- 混合模型可以通过生成模型的思想建模多输出问题
- 高斯混合模型本质是用EM算法来进行优化的
- 高斯混合模型是k-means聚类的一个概率化形式
- EM算法不保证全局最优，但往往总是有效