

#### 第四周第二次作业:

P65-66 2. 1) Find  $n$  r.v.s s.t. every  $(n-1)$  of them are independent but not all of them.

Consider the following  $n$  r.v.s:

$$Z_1, Z_2, \dots, Z_{n-1}, Z_n \text{ s.t. } \sum_{i=1}^n Z_i = 1, Z_i \in \mathbb{R}.$$

For every  $(n-1)$  of them are independent, but their sum is static, so these  $n$  r.v.s are not independent.

2) Find incidents  $\Lambda_1, \Lambda_2$  and  $\Lambda_3$  s.t.  $\Lambda_1$  and  $\Lambda_2$  are independent

$\Lambda_1$  and  $\Lambda_3$  are independent, but  $\Lambda_1$  and  $\Lambda_2 \cup \Lambda_3$  are not independent

Consider a family with two kids, one is younger and the other one older.

Let  $\Lambda_1 = \{ \text{The older kid is a girl} \}$

$\Lambda_2 = \{ \text{The younger kid is a girl} \}$

$\Lambda_3 = \{ \text{Two kids have the same sex} \}$

$\Rightarrow \Lambda_1$  and  $\Lambda_2$  independent

$\Lambda_1$  and  $\Lambda_3$  independent

$$P(\Lambda_1) = \frac{1}{2} \quad P(\Lambda_2 \cup \Lambda_3) = \frac{3}{4}, \quad P(\Lambda_1 \cap (\Lambda_2 \cup \Lambda_3)) = \frac{1}{4} \neq P(\Lambda_1)P(\Lambda_2 \cup \Lambda_3)$$

$\Rightarrow \Lambda_1$  and  $\Lambda_2 \cup \Lambda_3$  not independent

3.  $\{E_\alpha: \alpha \in A\}$  are independent  $\Leftrightarrow \{F_\alpha: \alpha \in A, F_\alpha = E_\alpha \text{ or } E_\alpha^c\}$  are independent

To prove the events  $\bigcup_{\alpha \in A} E_\alpha$  ( $\beta \in B$ ) are independent.

We just need to prove that  $\forall B' \subseteq B, B' = \{\beta_1, \beta_2, \dots, \beta_n\}$  finite

events  $\bigcup_{\alpha \in A_{\beta_i}} E_\alpha$  ( $i = 1, 2, \dots, n$ ) are independent.

$$\Leftrightarrow \prod_{i=1}^n P\left\{\bigcup_{\alpha \in A_{\beta_i}} E_\alpha\right\} = P\left\{\bigcap_{i=1}^n \bigcup_{\alpha \in A_{\beta_i}} E_\alpha\right\}$$

By Exclusion Exp: 
$$P\left\{\bigcup_{\alpha \in A_{\beta_i}} E_\alpha\right\} = \sum_{\alpha \in A_{\beta_i}} P(E_\alpha) - \sum_{\substack{\{\alpha_1, \alpha_2\} \in A_{\beta_i}^2 \\ \alpha_1 \neq \alpha_2}} P(E_{\alpha_1} \cap E_{\alpha_2}) + \dots$$

$$= \sum_{\alpha \in A_{\beta_i}} P(E_\alpha) - \sum_{\substack{\{\alpha_1, \alpha_2\} \in A_{\beta_i}^2 \\ \alpha_1 \neq \alpha_2}} P(E_{\alpha_1})P(E_{\alpha_2}) + \dots$$

$$\text{Thus } \prod_{i=1}^n P\left\{\bigcup_{\alpha \in A_{\beta_i}} E_\alpha\right\} = \prod_{i=1}^n \left[ \sum_{\alpha \in A_{\beta_i}} P(E_\alpha) - \sum_{\substack{\{\alpha_1, \alpha_2\} \in A_{\beta_i}^2 \\ \alpha_1 \neq \alpha_2}} P(E_{\alpha_1})P(E_{\alpha_2}) + \dots \right]$$

On the other hand 
$$P\left\{\bigcap_{i=1}^n \bigcup_{\alpha \in A_{\beta_i}} E_\alpha\right\} = P\left\{\bigcup_{(\alpha_1, \alpha_2, \dots, \alpha_n) \in \prod_{i=1}^n A_{\beta_i}} (E_{\alpha_1} \cap E_{\alpha_2} \cap \dots \cap E_{\alpha_n})\right\}$$

$$= \sum_{\alpha \in \vec{B}} P\{E_{\alpha_1} \cap E_{\alpha_2} \cap \dots \cap E_{\alpha_n}\} - \dots$$

Compare RHS of the two equations, they are the same.

6.  $X$  and  $f(X)$  independent  $\Leftrightarrow F_1(x_0)F_2[f(x_0)] = F_3[X \circ f(x_0)] \quad \forall x_0 \in \mathbb{R}$

where  $F_1, F_2, F_3$  are relatively d.f. of  $X, f(X)$  and  $(X, f(X))$

But we know that  $F_1(x_0) = F_2[f(x_0)] = F_3[X \circ f(x_0)]$  a.s. for  $x_0$

Thus  $F(x) \in \{0, 1\} \Rightarrow X$  is constant with probability one

and when so  $X$  and  $f(x)$  are independent

$$7. \bigcap_{j=1}^{\infty} E_j = \lim_{n \rightarrow \infty} \bigcap_{j=1}^n E_j$$

$$\text{So } P(\bigcap_{j=1}^{\infty} E_j) = P(\lim_{n \rightarrow \infty} \bigcap_{j=1}^n E_j) = \lim_{n \rightarrow \infty} P(\bigcap_{j=1}^n E_j)$$

$$\{E_j : 1 \leq j < \infty\} \text{ independent} \Rightarrow \forall \{i_1, i_2, \dots, i_k\} \subseteq \{1, 2, \dots\}.$$

$$\{E_{i_1}, E_{i_2}, \dots, E_{i_k}\} \text{ independent} \Rightarrow P(\bigcap_{j=1}^n E_j) = \prod_{j=1}^n P(E_j)$$

$$\Rightarrow P(\bigcap_{j=1}^{\infty} E_j) = \lim_{n \rightarrow \infty} \prod_{j=1}^n P(E_j) = \prod_{j=1}^{\infty} P(E_j)$$

$$\bigcup_{j=1}^{\infty} E_j = \lim_{n \rightarrow \infty} \bigcup_{j=1}^n E_j$$

$$\text{So } P(\bigcup_{j=1}^{\infty} E_j) = P(\lim_{n \rightarrow \infty} \bigcup_{j=1}^n E_j) = \lim_{n \rightarrow \infty} P(\bigcup_{j=1}^n E_j)$$

Since  $E_1, E_2, \dots$  independent, we have

$$\begin{aligned} P(\bigcup_{j=1}^n E_j) &= \sum_{j=1}^n P(E_j) - \sum_{1 \leq i < j \leq n} P(E_i)P(E_j) + \dots + (-1)^{n-1} \prod_{j=1}^n P(E_j) \\ &= 1 - \prod_{j=1}^n (1 - P(E_j)) \end{aligned}$$

$$\Rightarrow P(\bigcup_{j=1}^{\infty} E_j) = \lim_{n \rightarrow \infty} [1 - \prod_{j=1}^n (1 - P(E_j))] = 1 - \prod_{j=1}^{\infty} (1 - P(E_j))$$

$$8. \text{ Denote } M = \max\{x_1, x_2, \dots, x_n\}, m = \min\{x_1, x_2, \dots, x_n\}$$

$$\Rightarrow F_M(x) = P\{M \leq x\} = P\{x_1 \leq x, x_2 \leq x, \dots, x_n \leq x\}$$

Since  $x_1, x_2, \dots, x_n$  are independent, this becomes

$$F_M(x) = \prod_{i=1}^n P(x_i \leq x) = F_1(x) F_2(x) \dots F_n(x)$$

The same way we can get the minimum

$$F_m(x) = 1 - \prod_{i=1}^n [1 - F_i(x)]$$

$$17. \quad P\left(\bigcup_{j=1}^n E_j^{(n)}\right) \rightarrow 0 \text{ as } n \rightarrow +\infty$$

$$\text{By } P\left(\bigcup_{j=1}^n E_j^{(n)}\right) \geq \sum_{j=1}^n P(E_j^{(n)}) - \sum_{1 \leq j < k \leq n} P(E_j^{(n)} E_k^{(n)})$$

$$\Rightarrow \lim_{n \rightarrow \infty} \frac{\sum_{j=1}^n P(E_j^{(n)}) - \sum_{1 \leq j < k \leq n} P(E_j^{(n)}) P(E_k^{(n)})}{P\left(\bigcup_{j=1}^n E_j^{(n)}\right)} = 1$$

$$\text{On the other hand } \sum_{1 \leq j < k \leq n} P(E_j^{(n)}) P(E_k^{(n)}) = \left(\sum_{j=1}^n P(E_j^{(n)})\right)^2 - \sum_{j=1}^n P^2(E_j^{(n)}) \sim \left[\sum_{j=1}^n P(E_j^{(n)})\right]^2$$

$$\text{So } \lim_{n \rightarrow \infty} \frac{\sum_{j=1}^n P(E_j^{(n)})}{P\left(\bigcup_{j=1}^n E_j^{(n)}\right)} = 1 \text{ and } P\left(\bigcup_{j=1}^n E_j\right) \sim \sum_{j=1}^n P(E_j^{(n)})$$

Pb Thm 3.3.4. finite or infinite sequence  $\{x_j\}$  on  $(\mathbb{R}^1, \mathcal{B}^1)$

and d.f.s.  $\Rightarrow$  There exists a finite or infinite probability space

$(\Omega, \mathcal{F}, P)$  and a sequence of independent r.v.s  $\{x_j\}$  defined on it

s.t.  $\forall j, x_j$  is the p.m. of  $x_j$ .

Proof: Without loss of generality we suppose the given sequence is infinite.

$\forall n$ , suppose  $(\Omega_n, \mathcal{F}_n, P_n)$  be a probability space in which there exists

an r.v.  $x_n$  with  $\mu_n$  as its p.m.

Define the infinite measure space  $\Omega = \Omega_1 \times \Omega_2 \times \dots$

on the collection of all points  $\omega = \{\omega_1, \omega_2, \dots\}$

A subset  $E$  of  $\Omega$  will be called a "finite-product set" iff

it's of the form  $E = \prod_{n=1}^{\infty} F_n$  where each  $F_n \in \mathcal{F}_n$  and all but a finite number of these  $F_n$ 's are equal to the corresponding  $\Omega_n$ 's.

Thus  $W \in E \Leftrightarrow W_n \in F_n, n \geq 1$

Define a set function  $P$  on  $\mathcal{F}_0$  as follows.

(1) First.  $\forall E$  as finite-product set, let:

$$P(E) = \prod_{n=1}^{\infty} P_n(F_n)$$

where all but a finite number of the factors

on the right side are equal to one.

(2) Next, if  $E \in \mathcal{F}_0$  and  $E = \bigcup_{k=1}^n E^{(k)}$

where the  $E^{(k)}$ 's are disjoint finite-product sets, we put

$$P(E) = \sum_{k=1}^n P(E^{(k)})$$

If a given set  $E$  in  $\mathcal{F}_0$  has two representations of the form above.

Then it is not difficult to see that the two definitions of

$P(E)$  agree. Hence the set function  $P$  is uniquely defined on  $\mathcal{F}_0$

and it is clearly positive with  $P(\Omega) = 1$ ,

and additive on  $\mathcal{F}_0$  by definition.

$\Rightarrow \mathcal{P}$  as defined on  $\mathcal{F}_0$  is a p.m.