

# Assignment 5 Solutions

30 marks total

1. Lightening strikes are Poisson, so numbers in non-overlapping intervals are independent.

$$X \sim \text{Poisson}(2\lambda), \quad Y \sim \text{Poisson}(4\lambda), \quad Z \sim \text{Poisson}(4\lambda).$$

$$a) p(x, y, z) = \frac{e^{-2\lambda} (2\lambda)^x}{x!} \cdot \frac{e^{-4\lambda} (4\lambda)^y}{y!} \cdot \frac{e^{-4\lambda} (4\lambda)^z}{z!}$$

for  $x, y, z \geq 0$ .

$$b) \text{ let } T = X + Y + Z \sim \text{Poisson}(10\lambda)$$

$$P(X=x, Y=y, Z=z \mid T=t) = \frac{P(T=t \mid x, y, z) \cdot P(x, y, z)}{P(t)}.$$

$$P(t \mid x, y, z) = \begin{cases} 0 & \text{if } x+y+z \neq t \\ 1 & \text{if } x+y+z = t \end{cases}$$

$$P(x, y, z \mid t) = \frac{\frac{e^{-2\lambda} (2\lambda)^x}{x!} \cdot \frac{e^{-4\lambda} (4\lambda)^y}{y!} \cdot \frac{e^{-4\lambda} (4\lambda)^z}{z!}}{\frac{e^{-10\lambda} (10\lambda)^t}{t!}}$$

$$= \frac{t!}{x! y! z!} \left(\frac{2}{10}\right)^x \left(\frac{4}{10}\right)^y \left(\frac{4}{10}\right)^z$$

So  $X, Y, Z$  is cond. by  
Multinomial

With  $n=10, p_1=\frac{1}{5}, p_2=\frac{2}{5}, p_3=\frac{2}{5}$ .

$$1. \quad X_1 \sim \text{Bin}(n, \frac{1}{8})$$

$$\boxed{6} \quad X_1 + X_4 \sim \text{Bin}(n, \frac{2}{8})$$

From the notes we have  $\text{Var}(X_1 + X_4) = n(1-p)p = n \frac{2}{8} \times \frac{6}{8} = \frac{12n}{64}$

Also have  $\text{Var}(X_1 + X_4) = \text{Var}(X_1) + \text{Var}(X_4) + 2 \text{Cov}(X_1, X_4)$

Hence  $\frac{12n}{64} = n \cdot \frac{1}{8} \times \frac{7}{8} + n \frac{1}{8} \times \frac{7}{8} + 2 \text{Cov}(X_1, X_4)$ .

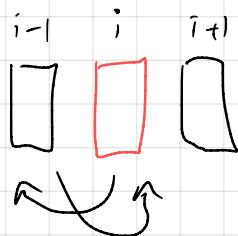
$$\Rightarrow \text{Cov}(X_1, X_4) = \frac{n}{2} \left( \frac{12}{64} - \frac{2 \times 7}{64} \right) = \underline{\frac{-n}{64}}$$

$$\begin{aligned} \text{Corr}(X_1, X_4) &= \frac{\text{Cov}(X_1, X_4)}{\sqrt{\text{Var}(X_1) \cdot \text{Var}(X_4)}} = \frac{\frac{-n}{64}}{\sqrt{\frac{7n}{64} \times \frac{7n}{64}}} \\ &= \frac{-\cancel{n}}{\cancel{64}} \cdot \frac{\cancel{7n}}{\cancel{7n}} = \underline{\frac{-1}{7}} \end{aligned}$$

3. a) That the process is Markov follows from the fact  
[8] that books are picked at random and successive choices are independent. Thus the process will be memoryless.

$$P_{i,i-1} = p$$

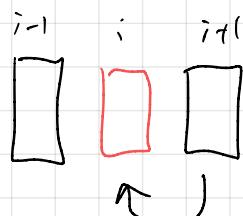
The red book is selected with probability  $p$  and moves to the right if  $i=2, 3, \dots, N$ . If  $i=1$ , it doesn't move.



$$P_{i,i+1} = \frac{1-p}{N-1}$$

A blank book is selected with probability  $1-p$ , and it is to the right of the red one, which occurs with probability  $\frac{1}{N-1}$  as there are  $N-1$

blank books and they are equally likely to be picked. If this happens the red book moves to the left.



If red book is at  $i=N$ , then there is no book to the right, hence this only works for  $i=1, \dots, N-1$ .

$$P_{i,i} = 1-p - \frac{1-p}{N-1}.$$

Probability of swapping two blank books, hence the red book doesn't move.

$$P_{1,1} = 1 - \frac{1-p}{N-1}$$

If red books is at  $i=1$ , it can only be moved by selecting the blank books to the right. This happens with probability  $\frac{1-p}{N-1}$ ,

Lence the probability it doesn't move is  $1 - \frac{1-p}{N-1}$ .

$P_{N,N} = 1-p$ . Similar to last case, the red book can only be moved by selecting it with prob  $p$ . So the probability it doesn't move is  $1-p$ .

The last three can also be obtained by 1 - the first two for the current values of  $i$ .

5). Stationary distribution satisfies  $\pi = \pi P$  where

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$$P = \begin{pmatrix} 1 - \left(\frac{1-p}{N-1}\right), & \frac{1-p}{N-1}, & 0, & 0, & 0, & \dots \\ p, & 1-p - \left(\frac{1-p}{N-1}\right), & \frac{1-p}{N-1}, & 0, & 0, & \dots \\ 0, & p, & 1-p - \left(\frac{1-p}{N-1}\right), & \frac{1-p}{N-1}, & 0, & \dots \\ \vdots & \vdots & \vdots & \vdots & \ddots & \ddots \end{pmatrix}$$

, this pattern repeats down the diagonals.

$$(\pi_1, \pi_2, \pi_3, \dots \pi_N) = (\pi_1, \pi_2, \pi_3, \dots \pi_N) P.$$

$$\text{So } \pi_1 = \pi_1 \left(1 - \frac{1-p}{N-1}\right) + p \pi_2$$

$$\pi_1 = \cancel{1} - \cancel{1} + \frac{1-p}{N-1} = p \pi_2 \Rightarrow$$

$$\boxed{\pi_2 = \frac{1-p}{(N-1)p} \pi_1}$$

Doing the next row  $\times$  col multiplication,

$$\begin{aligned}\pi_2 &= \left(\frac{1-p}{N-1}\right) \pi_1 + \left[1-p - \left(\frac{1-p}{N-1}\right)\right] \pi_2 + p \pi_3 \\ &= \cancel{\left(\frac{1-p}{N-1}\right)} \cdot \cancel{\frac{(N-1)p}{1-p}} \pi_2 + \left[1-p - \left(\frac{1-p}{N-1}\right)\right] \pi_2 + p \pi_3 \\ &= \left(1 - \left(\frac{1-p}{N-1}\right)\right) \pi_2 + p \pi_3\end{aligned}$$

This has the same form as the last eqn hence

$$\boxed{\pi_3 = \frac{1-p}{p/(N-1)} \pi_2}$$

c) Following the pattern from part (b) we see

(4)  $\boxed{\pi_{i+1} = \frac{1-p}{p/(N-1)} \pi_i \quad \text{for } i = 2, 3, \dots, N-1}$

Checking the  $i=N$  case, the lower-right part of  $P$  is,

$$\begin{pmatrix} \dots, 0, 1, \frac{1-p}{N-1}, 0 \\ \dots, p, 1-p - \left(\frac{1-p}{N-1}\right), \frac{1-p}{N-1} \\ \dots, 0, p, 1-p \end{pmatrix}$$

Hence,  $\pi_N = \left(\frac{1-p}{N-1}\right) \pi_{N-1} + (1-p) \pi_N \Rightarrow \pi_N = \frac{1-p}{(N-1)p} \pi_{N-1}$

So the pattern is the same.

Let  $A = \frac{1-p}{(N-1)p}$  then

$$\pi_2 = A \pi_1$$

$$\pi_3 = A^2 \pi_1 \text{ etc...}$$

$$\pi_i = A^{i-1} \pi_1 \text{ for } i=1, \dots, N$$

We also have  $\sum_{i=1}^N \pi_i = 1$ , so  $\pi_1 \sum_{i=1}^N A^{i-1} = 1$

$$\pi_1 \underbrace{\sum_{j=0}^{N-1} A^j}_{\text{geometric series}} = 1$$

$$\pi_1 \left( \frac{A^N - 1}{A - 1} \right) = 1$$

$$\Rightarrow \pi_i = \left( \frac{A-1}{A^N - 1} \right) \cdot A^{i-1} \text{ for } i=1, \dots, N$$