# Introduction to the rstpm2 package

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#### Abstract

This vignette outlines the methods and provides some examples for link-based survival models as implemented in the R rstpm2 package.

Keywords: survival, splines.

# 1. Background and theory

Link-based survival models provide a flexible and general approach to modelling survival or time-to-event data. The survival function S(t|x) to time t for covariates x is defined in terms of a link function G and a linear prediction  $\eta(t,x)$ , such that

$$S(t|x) = G(\eta(t,x))$$

where  $\eta$  is a function of both time t and covariates x. The linear predictor can be constructed in a flexible manner. Royston and Parmar (2003) focused on time being modelled using natural splines for log-time, including left truncation and relative survival. We have implemented the Royston-Parmar model class and extended it in several ways, allowing for: (i) general parametric models for  $\eta(t,x)$ , including B-splines and natural splines for different transformations of time; (ii) general semi-parametric models for  $\eta(t,x)$  including penalised smoothers together with unpenalised parametric functions; (iii) interval censoring; and (iv) frailties using Gamma and log-Normal distributions. Fully parametric models are estimated using maximum penalised likelihood, while the semi-parametric models are estimated using maximum penalised likelihood with smoothing parameters selected using A more detailed theoretical development is available from the paper by Liu, Pawitan and Clements (available on request).

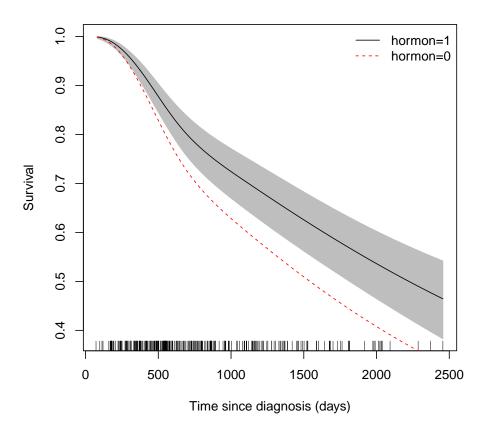
# 2. Independent survival analysis

We begin with some simple proportional hazard models using the breancer dataset. We can fit the models using very similar syntax to coxph, except that we need to specify the degrees of freedom for the baseline smoother. Typical values for df are 3-6. For this model the model parameters include an intercept term, time-invariant log-hazard ratios, and parameters for the baseline smoother. The default for the baseline smoother is to use the nsx function, which is a limited extension to the splines::ns function, with log of the time effect.

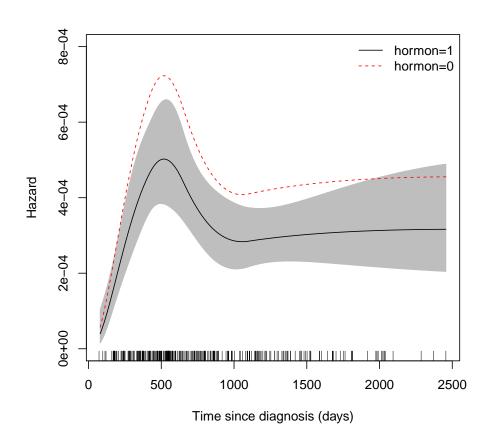
```
> fit <- stpm2(Surv(rectime,censrec==1)~hormon,</pre>
               data=brcancer, df=4)
> summary(fit)
Maximum likelihood estimation
Call:
mle2(minuslog1 = neg11, start = coef, eval.only = TRUE, vecpar = TRUE,
    gr = function (beta)
        localargs <- args
        localargs$init <- beta</pre>
        localargs$return_type <- "gradient"</pre>
        return(.Call("model_output", localargs, PACKAGE = "rstpm2"))
    \}, control = list(parscale = c(1, 1, 1, 1, 1, 1), maxit = 300),
    lower = -Inf, upper = Inf)
Coefficients:
                           Estimate Std. Error z value
                                                            Pr(z)
                                       0.72643 -9.3578 < 2.2e-16 ***
(Intercept)
                           -6.79773
hormon
                           -0.36406
                                       0.12491 -2.9144 0.003563 **
                                       0.71677 7.9523 1.831e-15 ***
nsx(log(rectime), df = 4)1 5.69995
nsx(log(rectime), df = 4)2 	 4.85614
                                       0.48002 10.1166 < 2.2e-16 ***
                                       1.41268 7.1731 7.331e-13 ***
nsx(log(rectime), df = 4)3 10.13328
nsx(log(rectime), df = 4)4 	 4.70626
                                       0.33016 14.2545 < 2.2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
-2 log L: 5212.943
> ## utility to exponentiate the ith components
> expi <- function(x,i=1:length(x)) { x[i] <- exp(x[i]); x }
> fit.cox <- coxph(Surv(rectime,censrec==1)~hormon, data=brcancer)</pre>
> rbind(coxph=coef(summary(fit.cox)),
        stpm2=expi(coef(summary(fit))["hormon",c(1,1,2:4)],2))
             coef exp(coef) se(coef)
                                                    Pr(>|z|)
hormon -0.3640099 0.6948843 0.1250446 -2.911041 0.003602266
stpm2 -0.3640574 0.6948513 0.1249147 -2.914449 0.003563175
>
```

We see that the hazard ratios are very similar to the coxph model. The model fit can also be used to estimate a variety of parameters. For example, we can easily estimate survival for a given parameter.

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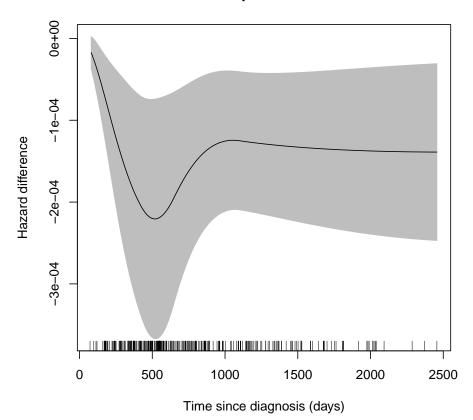
We can also calculate the hazards.



Usefully, we can also estimate survival differences and hazard differences. We define the survival differences using a reference covariate pattern using the newdata argument, and then define an exposed function which takes the newdata and transforms for the 'exposed' covariate pattern. As an example

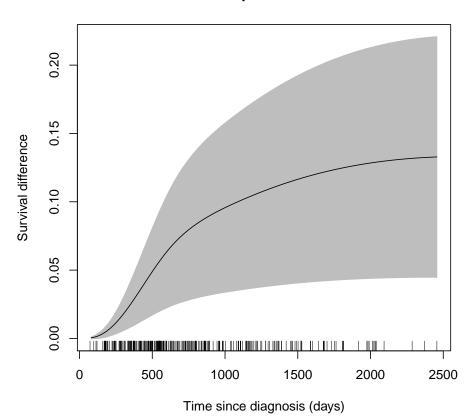
```
> plot(fit,newdata=data.frame(hormon=0), type="hdiff",
+ exposed=function(data) transform(data, hormon=1),
+ main="hormon=1 compared with hormon=0",
+ xlab="Time since diagnosis (days)")
>
```

## hormon=1 compared with hormon=0



```
> plot(fit,newdata=data.frame(hormon=0), type="sdiff",
+ exposed=function(data) transform(data, hormon=1),
+ main="hormon=1 compared with hormon=0",
+ xlab="Time since diagnosis (days)")
```

### hormon=1 compared with hormon=0



## 3. Mean survival

This has a useful interpretation for causal inference.

$$E_Z(S(t|Z, X = 1)) - E_Z(S(t|Z, X = 0))$$

fit <- stpm2(...)
predict(fit,type="meansurv",newdata=data)</pre>

## 4. Cure models

For cure, we use the melanoma dataset used by Andersson and colleagues for cure models with Stata's stpm2 (see http://www.pauldickman.com/survival/).

Initially, we merge the patient data with the all cause mortality rates.

```
> popmort2 <- transform(rstpm2::popmort,exitage=age,exityear=year,age=NULL,year=NULL)
> colon2 <- within(rstpm2::colon, {
+    status <- ifelse(surv_mm>120.5,1,status)
```

```
+ tm <- pmin(surv_mm, 120.5)/12
    exit <- dx+tm*365.25
   sex <- as.numeric(sex)</pre>
    exitage <- pmin(floor(age+tm),99)</pre>
    exityear <- floor(yydx+tm)</pre>
    ##year8594 <- (year8594=="Diagnosed 85-94")
+ })
> colon2 <- merge(colon2,popmort2)</pre>
For comparisons, we fit the relative survival model without and with cure.
> fit0 <- stpm2(Surv(tm, status %in% 2:3)~I(year8594=="Diagnosed 85-94"),
                data=colon2,
+
                bhazard=colon2$rate, df=5)
>
> summary(fit <- stpm2(Surv(tm, status %in% 2:3)~I(year8594=="Diagnosed 85-94"),
                       data=colon2,
                       bhazard=colon2$rate,
                       df=5,cure=TRUE))
Maximum likelihood estimation
Call:
mle2(minuslog1 = neg11, start = coef, eval.only = TRUE, vecpar = TRUE,
    gr = function (beta)
        localargs <- args
        localargs$init <- beta</pre>
        localargs$return_type <- "gradient"</pre>
        return(.Call("model_output", localargs, PACKAGE = "rstpm2"))
    \}, control = list(parscale = c(1, 1, 1, 1, 1, 1, 1), maxit = 300),
    lower = -Inf, upper = Inf)
Coefficients:
                                       Estimate Std. Error z value
                                                                         Pr(z)
(Intercept)
                                      -3.977663 0.054783 -72.6075 < 2.2e-16
I(year8594 == "Diagnosed 85-94")TRUE -0.155511 0.025089 -6.1984 5.704e-10
                                    3.323382 0.053170 62.5043 < 2.2e-16
nsx(log(tm), df = 5, cure = TRUE)1
nsx(log(tm), df = 5, cure = TRUE)2 3.628899 0.053164 68.2580 < 2.2e-16
nsx(log(tm), df = 5, cure = TRUE)3 1.634974 0.022466 72.7744 < 2.2e-16
nsx(log(tm), df = 5, cure = TRUE)4 6.592489 0.111515 59.1177 < 2.2e-16
nsx(log(tm), df = 5, cure = TRUE)5 3.371954 0.042790 78.8021 < 2.2e-16
(Intercept)
I(year8594 == "Diagnosed 85-94")TRUE ***
```

```
nsx(log(tm), df = 5, cure = TRUE)1
nsx(log(tm), df = 5, cure = TRUE)2
nsx(log(tm), df = 5, cure = TRUE)3
nsx(log(tm), df = 5, cure = TRUE)4
nsx(log(tm), df = 5, cure = TRUE)5
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
-2 log L: 42190.77
> predict(fit,head(colon2),se.fit=TRUE)
   Estimate
                lower
                          upper
1 0.8611043 0.8543118 0.8676051
2 0.7934962 0.7850418 0.8016614
3 0.6967834 0.6863627 0.7069356
4 0.8611043 0.8543118 0.8676051
5 0.8221508 0.8143496 0.8296594
6 0.8611043 0.8543118 0.8676051
```

The estimate for the year parameter from the model without cure is within three significant figures with that in Stata. For the predictions, the Stata model gives:

```
surv surv_lci surv_uci |
  |-----|
1. | .86108264 .8542898 .8675839 |
2. | .79346526 .7850106 .8016309 |
3. | .69674037 .6863196 .7068927 |
4. | .86108264
           .8542898
                    .8675839 |
5. | .82212425
            .8143227
                    .8296332 |
  |-----|
6. | .86108264
            .8542898
                    .8675839 |
  +----+
```

We can estimate the proportion of failures prior to the last event time:

```
> 1-predict(fit, newdata.eof, type="surv", se.fit=TRUE)
```

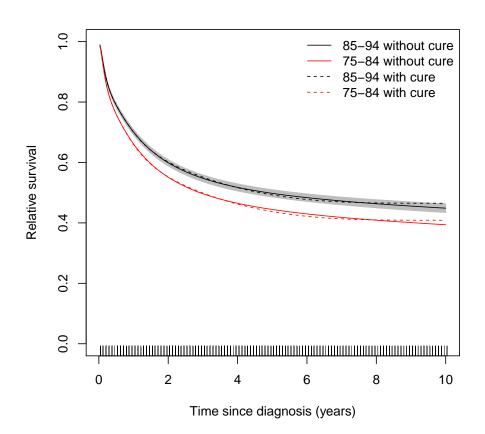
```
Estimate lower upper
1 0.5912976 0.6054691 0.5771835
2 0.5350852 0.5485412 0.5217471

> predict(fit, newdata.eof, type="haz", se.fit=TRUE)

Estimate lower upper
1 1.253896e-06 1.092818e-06 1.438717e-06
2 1.073307e-06 9.334233e-07 1.234153e-06
```

>

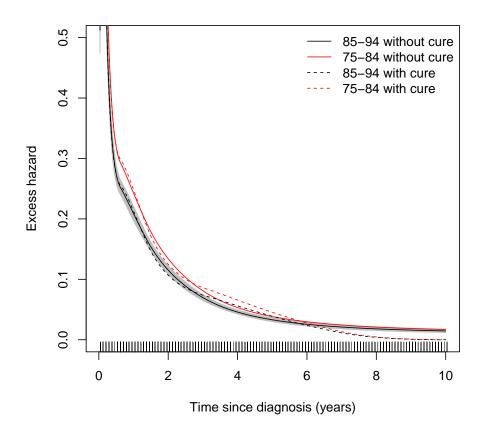
We can plot the predicted survival estimates:



#### And the hazard curves:

```
> plot(fit0, newdata=data.frame(year8594 = "Diagnosed 85-94", tm=tms),
       ylim=c(0,0.5), type="hazard",
       xlab="Time since diagnosis (years)", ylab="Excess hazard")
 plot(fit0,newdata=data.frame(year8594 = "Diagnosed 75-84", tm=tms),
       type="hazard",
       add=TRUE,line.col="red",rug=FALSE)
> plot(fit,newdata=data.frame(year8594 = "Diagnosed 85-94", tm=tms),
       type="hazard",
       add=TRUE,ci=FALSE,lty=2,rug=FALSE)
> plot(fit,newdata=data.frame(year8594="Diagnosed 75-84", tm=tms),
       type="hazard",
       add=TRUE,rug=FALSE,line.col="red",ci=FALSE,lty=2)
> legend("topright",c("85-94 without cure", "75-84 without cure",
                      "85-94 with cure", "75-84 with cure"),
         col=c(1,2,1,2), lty=c(1,1,2,2), bty="n")
+
>
```

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