

Problem 1

(As I have been struggling with typing align* (so many formulas) in R markdown, so my question 1 is here typed in LaTeX, and the remaining questions are typed in rmd.)

Solution:

1.

(a) Cumulative distribution function is

$$F(y; \theta) = 1 - e^{-\frac{y^2}{2\theta}},$$

Therefore, p.d.f. is given by $f(y; \theta) = F'(y; \theta) = \frac{1}{\theta} e^{-\frac{y^2}{2\theta}}$

Survival function $S(y) = P(Y > y) = \int_y^\infty f(y; \theta) = 1 - F(y; \theta) = e^{-\frac{y^2}{2\theta}}$

Therefore, the likelihood function is

$$\begin{aligned} L(\theta) &= \prod_{i=1}^n \{ [f(y_i; \theta)]^{r_i} [S(C; \theta)]^{1-r_i} \} \\ &= \prod_{i=1}^n \left\{ \left[\frac{1}{\theta} e^{-\frac{y_i^2}{2\theta}} \right]^{r_i} \left[e^{-\frac{C^2}{2\theta}} \right]^{1-r_i} \right\} \\ &= \frac{1}{\theta^{\sum_{i=1}^n r_i}} e^{-\frac{1}{2\theta} \sum_{i=1}^n y_i^2 r_i} e^{-\frac{1}{2\theta} \sum_{i=1}^n C^2 (1-r_i)} \\ &= \frac{1}{\theta^{\sum_{i=1}^n r_i}} e^{-\frac{1}{2\theta} \sum_{i=1}^n x_i^2}, \end{aligned}$$

The corresponding log likelihood is

$$l(\theta) = \log L(\theta) = -\sum_{i=1}^n R_i \log \theta - \frac{\sum_{i=1}^n X_i^2}{2\theta}$$

We thus have

$$U = \frac{d}{d\theta} \log L(\theta) = -\frac{\sum_{i=1}^n R_i}{\theta} + \frac{\sum_{i=1}^n X_i^2}{2\theta^2},$$

let $U = 0$, we have

$$\hat{\theta}_{MLE} = \frac{\sum_{i=1}^n X_i^2}{2 \sum_{i=1}^n R_i}.$$

(b) We have

$$\frac{d^2}{d\theta^2} \log L(\theta) = \frac{\sum_{i=1}^n r_i}{\theta^2} - \frac{\sum_{i=1}^n X_i^2}{\theta^3}.$$

The fisher information is

$$\begin{aligned} I(\theta) &= -E \left(\frac{\sum_{i=1}^n r_i}{\theta^2} - \frac{\sum_{i=1}^n X_i^2}{\theta^3} \right) \\ &= -\frac{1}{\theta^2} n E(R) + \frac{1}{\theta^3} n E(X^2) \end{aligned}$$

Note that R is a binary random variable and so

$$\begin{aligned}
E(R) &= 1 \times \Pr(R = 1) + 0 \times \Pr(R = 0) \\
&= \Pr(R = 1) \\
&= \Pr(Y \leq C) \\
&= F(C; \theta) \\
&= 1 - e^{-\frac{y^2}{2\theta}}.
\end{aligned}$$

$$\begin{aligned}
E(X^2) &= \int_0^\infty X_i^2 f(x; \theta) dx \\
&= \int_0^C X_i^2 f(x; \theta) dx + \int_C^\infty X_i^2 f(x; \theta) dx \\
&= \int_0^C y^2 f(y; \theta) dy + \int_C^\infty C^2 f(x; \theta) dx \\
&= -C^2 e^{-\frac{C^2}{(2\theta)}} + 2\theta \left(1 - e^{-\frac{C^2}{(2\theta)}} \right) + C^2 S(C) \\
&= 2\theta \left(1 - e^{-\frac{C^2}{(2\theta)}} \right)
\end{aligned}$$

Therefore,

$$\begin{aligned}
I(\theta) &= -\frac{1}{\theta^2} n E(R) + \frac{1}{\theta^3} n E(X^2) \\
&= -\frac{n}{\theta^2} \left(1 - e^{-\frac{C^2}{2\theta}} \right) + \frac{2n}{\theta^2} \left(1 - e^{-\frac{C^2}{(2\theta)}} \right) \\
&= \frac{n}{\theta^2} \left(1 - e^{-\frac{C^2}{2\theta}} \right)
\end{aligned}$$

(c) As mle is asymptotic normal distributed, we need to find the mean and variance of mle.

From 1(a), we know that $\hat{\theta}_{\text{MLE}} = \frac{\sum_{i=1}^n X_i^2}{2 \sum_{i=1}^n R_i}$,

Therefore,

$$\begin{aligned}
\text{Var}(\theta) &= \frac{1}{\sqrt{J(\theta)}} \\
&= \frac{1}{\sqrt{-\frac{\sum_{i=1}^n R_i}{\theta^2} + \frac{\sum_{i=1}^n X_i^2}{\theta^3}}} \\
&= \frac{\theta}{\sqrt{-\sum_{i=1}^n R_i + \frac{\sum_{i=1}^n X_i^2}{\theta}}} \\
&= \frac{\sum_{i=1}^n X_i^2}{(\sum_{i=1}^n R_i)^{\frac{3}{2}}}
\end{aligned}$$

therefore, the 95% confidence interval for θ is $[\hat{\theta} - 1.96 \frac{1}{\sqrt{J(\theta)}}, \hat{\theta} + 1.96 \frac{1}{\sqrt{J(\theta)}}]$,

i.e., $[\frac{\sum_{i=1}^n X_i^2}{2 \sum_{i=1}^n R_i} - 0.98 \frac{\sum_{i=1}^n X_i^2}{(\sum_{i=1}^n R_i)^{\frac{3}{2}}}, \frac{\sum_{i=1}^n X_i^2}{2 \sum_{i=1}^n R_i} + 0.98 \frac{\sum_{i=1}^n X_i^2}{(\sum_{i=1}^n R_i)^{\frac{3}{2}}}]$