# Analysis of riboflavin

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# **Data Description**



Figure 1.1: Data

The data **riboflavin** comes from the package **hdi** of R, with 71 observations and 4088 continuous explanatory variables. So this is a  $n \ll p$  problem.

# **PCA**

#### **Cumulative Variance Explained by Principal Components**

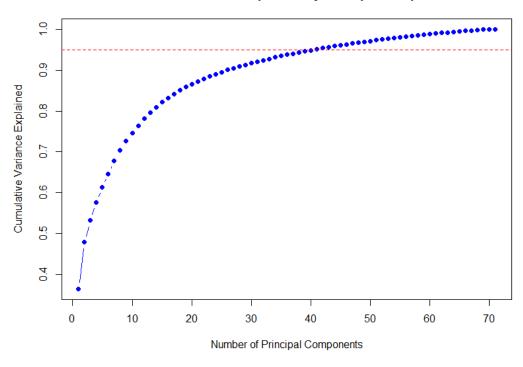


Figure 2.1: PCA Accumulated Variance

First, I try PCA to reduce the dimensionality. Considering that PCA is an unsupervised learning method and there is no clear standard how much cumulative variance should I choose, I choose 95% to keep more information. In this way, there are 41 variables left.

CHAPTER 2. PCA

4

```
riboflavin <- read.csv('riboflavin.csv')</pre>
_{2} X <- riboflavin [-1]
з y <- riboflavin[1]
4 pca_result <- prcomp(X, scale. = TRUE)
  explained_variance <- pca_result $sdev^2 / sum(pca_result
      $sdev^2)
  cumulative_variance <- cumsum(explained_variance)</pre>
  num components <- 41
 cumulative variance [num components]
plot (cumulative_variance, type = "b", pch = 19, col = "
      blue", xlab = "Number⊔of⊔Principal⊔Components", ylab
     = "Cumulative Variance Explained", main = "Cumulative
      ⊔Variance ⊔Explained ⊔by ⊔Principal ⊔Components")
abline(h = 0.95, col = "red", lty = 2)
PC_variables <- list()
  for (i in 1:41) {
    PC_index <- which.max(abs(pca_result$rotation[, i]))
    PC_variable <- rownames(pca_result$rotation)[PC_index]
    PC_variables [[paste0("PC", i)]] <- PC_variable
15
16
  selected columns <- riboflavin [, unlist (PC variables)]
riboflavin2 <- cbind(y, selected_columns)</pre>
```

# Kernel Smoothing

#### Histogram of y with Density Curve

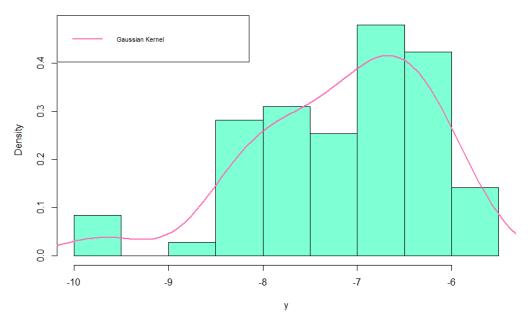


Figure 3.1: Kernel Smoothing of y

Next, I want to explore the distribution of y, so I make a histogram of y, and use Gaussian kernel with optimal bandwidth based on the cross validation to smooth. The plot shows that y is unimodal and left-skewed, so it might not be normal.

```
hcv_result <- sm.density(y, hcv = TRUE)
optimal_bandwidth <- hcv_result $h
hist(riboflavin $y, breaks = "FD", freq = FALSE, main = "
HistogramuofuyuwithuDensityuCurve", xlab = "y", col =
"#7FFFD4", border = "black")</pre>
```

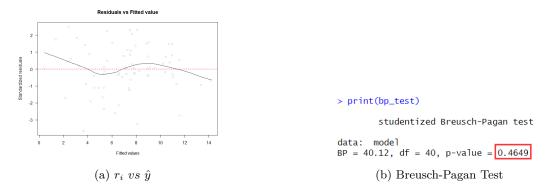
```
density estimate <- sm. density (y, h = optimal bandwidth,
       display = "none")
5 lines (density_estimate $ eval.points, density_estimate $
     estimate, col = "\#FF69B4", lwd = 2)
6 legend ("topleft", legend = c("Gaussian ⊔ Kernel"), col = c
      ("\#FF69B4"), lwd = 2, cex = 0.7)
  ##### Box-Cox #####
  y \leftarrow as.matrix(y)
  \min_{y} < -\min_{y}
  if (\min_{y} \le 0)
    y \leftarrow y - \min_{y} + 1
15
  boxcox result \leftarrow boxcox(y \sim 1, lambda = seq(-2, 2, by =
      0.1))
  best lambda <- boxcox result $x [which.max(boxcox result $y]
  print (paste ("Best lambda:", best_lambda))
  if (best lambda = 0) {
    y boxcox < log(y)
  } else {
21
    y_boxcox \leftarrow (y^best_lambda - 1) / best_lambda
22
  }
23
24
  ### New Kernel Estimation ###
  hcv result <- sm. density (y, hcv = TRUE)
  optimal_bandwidth <- hcv_result$h
  hist (riboflavin2_boxcox$y, breaks = "FD", freq = FALSE,
     main = "Histogram of y with Density Curve",
       xlab = "y", col = "#7FFFD4", border = "black")
  density estimate <- sm. density (y, h = optimal bandwidth,
       display = "none")
  lines (density_estimate $eval.points, density_estimate $
     estimate, col = "\#FF69B4", lwd = 2)
  legend ("topleft", legend = c("Gaussian Kernel"), col = c
     ("\#FF69B4"), lwd = 2, cex = 0.7)
```

### Linear Model

### 4.1 Hypothesis Testing

Fitting a linear model and making inferences requires 5 assumptions, and I will test each one.

#### 4.1.1 Linearity and Homoscedasticity



The residual plot shows that the  $r'_i$ s are randomly distribute around 0, so linearity is satisfied. And the p-value of Breusch-Pagan Test is very large, which means that we cannot reject  $H_0$ : Variance of the error term is a constant, i.e. homoscedasticity is satisfied.

```
6 bp_test <- bptest(model)
7 print(bp_test)
```

#### 4.1.2 Independence

There's no information of the time when the data is collected, so we just assume that the data is independent on the time dimension.

#### 4.1.3 Multicollinearity

```
> vif(model)
                                                                          x.UREA_at
   x.YNDJ_at
                 x.YACD_at
                               x.YXDJ_at
                                              x.YWAE_at
                                                            x.XYLB_at
                                                                                      x.SPOVID_at
                                                                                                       x.PHRK_at
                                                                                                                     x.ADHA_at
   12.969803
                  5.343119
                                4.195356
                                              5.353218
                                                             6.829044
                                                                            7.810606
                                                                                         3.949662
                                                                                                        3.819092
                                                                                                                      3.705367
   x.RTBA at
                 x.BOFA at
                               x.YKPC at
                                              x.YCNK at
                                                            x.PABB at
                                                                       x.SPOTTGA at
                                                                                        x.CSBA at
                                                                                                       x.HUTP at
                                                                                                                     x.YCLK at
                  5.023135
                                4.572193
                                               6.114210
                                                             2.973912
                                                                           3.164068
                                                                                         9.635928
                                                                                                        2.753163
                                                                                                                      5.293606
    3.351136
   x.YUSJ at
                 x.YKRP at
                               x.XKDO at
                                             x.CHER at
                                                            x.NADB at
                                                                          x.YURK at
                                                                                        x.YTAB_at
                                                                                                       x.YERO at
                                                                                                                     x.YRHA at
   2.617433
                  2.958422
                                3.326167
                                               4.163400
                                                             2.050740
                                                                          4.249037
x.YVFK_at
                                                                                        7.776301
x.DPPD_at
                                                                                                        4.762744
                                                                                                                      3.834362
   x.ALST_at
                 x.PYRF_at
                               x.CYSE_at
                                              x.FFH_at
                                                          x.YHDS_r_at
                                                                                                       x.YEFA_at
                                                                                                                     x.PRFA_at
    3.399612
                  2.926549
                                4.095913
                                             13.977456
x.GLNM_at
                                                             2.172475
                                                                          62.333463
                                                                                                        5.829215
                                                                                        19.462113
                                                                                                                      2.841525
                               x.YVFO at
   x.CSAA at
                 x.YEZC at
    4.162516
                  5.737347
                               55.285061
                                              4.669691
```

Figure 4.2: VIF of the Original Model

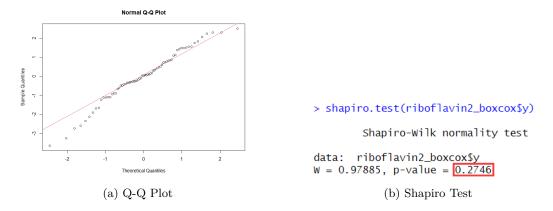
```
any(vif(model11)>10)
                                                                                                                                                                                                                                                                                        [1] TRUE
> any(vif(model12)>10)
\verb|mode|| 11 <- || lm(y \sim .-x.NADB\_at.1-x.YVFO\_at, || data = || ribof|| lavin2\_boxcox)|
                                                                                                                                                                                                                                                                                        [1] TRUE
modell2 \leftarrow lm(y \sim .-x.NADB_at.1-x.FFH_at, data = riboflavin2\_boxcox)
                                                                                                                                                                                                                                                                                             any(vif(modell3)>10)
                                                                                                                                                                                                                                                                                        [1] TRUE
model13 <- lm(y ~ .-x.NADB_at.1-x.DPPD_at, data = riboflavin2_boxcox)
                                                                                                                                                                                                                                                                                       > any(vif(model14)>10)
[1] TRUE
model14 <- lm(y ~ .-x.NADB_at.1-x.YVFK_at, data = riboflavin2_boxcox)
                                                                                                                                                                                                                                                                                                            (b) VIF
                                                                           (a) Deleting One Variable
                                                                                                                                                                                                                                                                                             any(vif(model21)>10)
                                                                                                                                                                                                                                                                                        [1] FALSE
                                                                                                                                                                                                                                                                                              any(vif(model22)>10)
                                                                                                                                                                                                                                                                                        [1] FALSE
                                                                                                                                                                                                                                                                                             any(vif(model23)>10)
                                                                                                                                                                                                                                                                                        [1] TRUE
> any(vif(model24)>10)
model21 \leftarrow lm(y \sim .-x.NADB_at.1-x.YVFO_at-x.FFH_at, data = riboflavin2\_boxcox)
\label{eq:model22} $$ \mbox{model22} < - \mbox{lm}(y \sim .-x.NADB_at.1-x.DPPD_at-x.YVFO_at, data = riboflavin2_boxcox) $$ \mbox{model23} < - \mbox{lm}(y \sim .-x.NADB_at.1-x.FFH_at-x.DPPD_at, data = riboflavin2_boxcox) $$ \mbox{model23} < - \mbox{lm}(y \sim .-x.NADB_at.1-x.FFH_at-x.DPPD_at, data = riboflavin2_boxcox) $$ \mbox{model23} < - \mbox{lm}(y \sim .-x.NADB_at.1-x.FFH_at-x.DPPD_at, data = riboflavin2_boxcox) $$ \mbox{model23} < - \mbox{lm}(y \sim .-x.NADB_at.1-x.FFH_at-x.DPPD_at, data = riboflavin2_boxcox) $$ \mbox{model23} < - \mbox{lm}(y \sim .-x.NADB_at.1-x.FFH_at-x.DPPD_at, data = riboflavin2_boxcox) $$ \mbox{model23} < - \mbox{lm}(y \sim .-x.NADB_at.1-x.FFH_at-x.DPPD_at, data = riboflavin2_boxcox) $$ \mbox{model23} < - \mbox{lm}(y \sim .-x.NADB_at.1-x.FFH_at-x.DPPD_at, data = riboflavin2_boxcox) $$ \mbox{model23} < - \mbox{lm}(y \sim .-x.NADB_at.1-x.FFH_at-x.DPPD_at, data = riboflavin2_boxcox) $$ \mbox{model23} < - \mbox{lm}(y \sim .-x.NADB_at.1-x.FFH_at-x.DPPD_at, data = riboflavin2_boxcox) $$ \mbox{model23} < - \mbox{lm}(y \sim .-x.NADB_at.1-x.FFH_at-x.DPPD_at, data = riboflavin2_boxcox) $$ \mbox{model23} < - \mbox{lm}(y \sim .-x.NADB_at.1-x.FFH_at-x.DPPD_at, data = riboflavin2_boxcox) $$ \mbox{model24} < - \mbox{lm}(y \sim .-x.NADB_at.1-x.FFH_at-x.DPPD_at, data = riboflavin2_boxcox) $$ \mbox{lm}(y \sim .-x.NADB_at.1-x.DPD_at, data = riboflavin2_boxcox) $$ \mbox{lm}(y \sim .-x.NADB_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD_at.1-x.DPD
                                                                                                                                                                                                                                                                                        [1] FALSE
                                                                                                                                                                                                                                                                                             any(vif(model25)>10)
\verb|mode|| 24 <- || \verb|m(y \sim .-x.NADB_at.1-x.YVFK_at-x.DPPD_at, | data = riboflavin2\_boxcox)|
                                                                                                                                                                                                                                                                                        [1] FALSE
model25 <- lm(y ~ .-x.NADB_at.1-x.YVFK_at-x.FFH_at, data = riboflavin2_boxcox)
                                                                                                                                                                                                                                                                                       > any(vif(model26)>10)
[1] TRUE
\verb|mode|| 26 <- | \verb|m(y \sim .-x.NADB_at.1-x.YVFK_at-x.YVFO_at, data = riboflavin2\_boxcox)| \\
                                                                          (a) Deleting Two Variables
                                                                                                                                                                                                                                                                                                            (b) VIF
```

When I first fit the model, it shows that the  $\mathbf{x.NADB\_at.1}$  term can be completely represented by other variables, so I delete it and fit a new one. Then, I compute the VIF of each variable and check whether any of them is more than 10, and I find four of them with multicollinearity problem and none of them is significant. When I try to delete one, there still exists multicollinearty. So I test with two of the four deleted and choose one with the highest  $R^2$ , which turn out to be  $\mathbf{x.DPPD\_at}$  and  $\mathbf{x.YVFO\_at}$  are removed, and get the  $\mathbf{model25}$ .

```
\begin{array}{lll} & \text{vif(model)} \\ & \text{2 model11} <- \text{lm(y} \sim .-x.\text{NADB\_at.1-x.YVFO\_at, data} = \\ & & \text{riboflavin2\_boxcox)} \end{array}
```

```
model12 \leftarrow lm(y \sim .-x.NADB_at.1-x.FFH_at, data =
    riboflavin2 boxcox)
model13 \leftarrow lm(y \sim .-x.NADB_at.1-x.DPPD_at, data =
    riboflavin2_boxcox)
model14 \leftarrow lm(y \sim .-x.NADB_at.1-x.YVFK_at, data =
    riboflavin2 boxcox)
model21 \leftarrow lm(y \sim .-x.NADB_at.1-x.YVFO_at-x.FFH_at, data
     = riboflavin2 boxcox)
model22 <- lm(y ~ .-x.NADB at.1-x.DPPD at-x.YVFO at,
    data = riboflavin2_boxcox)
model23 <- lm(y ~ .-x.NADB_at.1-x.FFH_at-x.DPPD_at, data
     = riboflavin2 boxcox)
model24 \leftarrow lm(y \sim .-x.NADB_at.1-x.YVFK_at-x.DPPD_at,
    data = riboflavin2_boxcox)
model25 <- lm(y ~ .-x.NADB_at.1-x.YVFK_at-x.FFH_at, data
     = riboflavin2 boxcox)
model26 <- lm(y ~ .-x.NADB at.1-x.YVFK at-x.YVFO at,
    data = riboflavin 2 boxcox)
```

#### 4.1.4 Normality



We can learn from the kernel smoothing that y is unlikely normally distributed. So I make a Box-Cox transformation to y and then make the Shapiro test and draw the Q-Q plot of y. Since p-value of the test is large, we cannot reject  $H_0$ : y is normally distributed. Hence, normality is satisfied.

```
shapiro.test(riboflavin25_boxcox$y)
qqnorm(resid(model25))
qqline(resid(model25), col = 2)
```

### 4.2 Model Comparison

```
Residual standard error: 1.975 on 32 degrees of freedom
Multiple R-squared: 0.8267, Adjusted R-squared: 0.6209
F-statistic: 4.017 on 38 and 32 DF, p-value: 6.055e-05
                   (a) R^2 and Adjusted R^2 of Full Model
> # prediction
> predicted <- predict(model2, x)</pre>
> lm_mse <- mean((riboflavin2_boxcox[,1]-predicted)^2)</pre>
> 1m mse
[1] 1.75762
                     (b) Prediction MSE of Full Model
Residual standard error: 1.61 on 55 degrees of freedom
Multiple R-squared: 0.802, Adjusted R-squared: 0.748
F-statistic: 14.85 on 15 and 55 DF, p-value: 3.155e-14
                   (a) R^2 and Adjusted R^2 of AIC Model
> prediction_AIC <- predict(modelAIC, newx = x)</pre>
> AIC_mse <- mean((prediction_AIC - riboflavin2_boxcox$y)^2)</pre>
> AIC_mse
[1] 2.007884
                     (b) Prediction MSE of AIC Model
Residual standard error: 1.667 on 58 degrees of freedom
Multiple R-squared: 0.7761, Adjusted R-squared: 0.7298
F-statistic: 16.76 on 12 and 58 DF, p-value: 1.163e-14
                   (a) R^2 and Adjusted R^2 of BIC Model
> prediction_BIC <- predict(modelBIC, newdata = riboflavin2_boxcox)</pre>
> BIC_mse <- mean((prediction_BIC - y)^2)</pre>
> BIC_mse
[1] 2.270461
                     (b) Prediction MSE of BIC Model
```

```
> ss_res <- sum((y - y_pred)^2)
> ss_tot <- sum((y - mean(y))^2)
> r_squared <- 1 - ss_res / ss_tot</pre>
> n <- length(y)
> lasso_coef <- coef(final_lasso_model, s = best_lambda)</pre>
> p <- sum(lasso_coef != 0) - 1
> adjusted_r_squared <- 1 - ((1 - r_squared) * (n - 1) / (n - p - 1))
> r_squared
[1] 0.7219365
> adjusted_r_squared
[1] 0.6256838
                     (a) R^2 and Adjusted R^2 of Lasso Model
> y_pred <- predict(final_lasso_model, newx = as.matrix(x))</pre>
> LASSO_mse <- mean((y - y_pred)^2)
> LASSO_mse
[1] 2.819975
```

(b) Prediction MSE of Lasso Model

Here I try AIC, BIC and Lasso to choose variables and the whole dataset is taken as both train data and test data. Although the MSE of the full model seems to be the smallest, it only means that it performs the best in the original data. Thus, in general, consider both to minimize MSE and maximize adjusted  $R^2$ , the best model should be the AIC model.

```
### full model ###
  predicted <- predict(model2, x)</pre>
  lm_mse \leftarrow mean((riboflavin2\_boxcox[,1]-predicted)^2)
  lm mse
  summary( model2)
  ### AIC ###
  modelAIC <- stepAIC (model2, direction = 'both')
  prediction AIC <- predict(modelAIC, newx = x)</pre>
  AIC_mse <- mean((prediction_AIC - riboflavin2_boxcox$y)
      ^2)
10 AIC mse
  summary (modelAIC)
12 ### BIC ###
modelBIC <- stepAIC (model2, direction = 'both', k = log(
      nrow (model.frame (model2))))
  prediction_BIC <- predict(modelBIC, newdata =</pre>
      riboflavin2 boxcox)
  BIC_mse \leftarrow mean((prediction_BIC - y)^2)
  BIC mse
17 summary (modelBIC)
18 ### LASSO ###
  lasso_model <- cv.glmnet(as.matrix(x), y, alpha = 1)
```

```
best\_lambda <- lasso\_model\$lambda.min
  final_lasso_model <- glmnet(x, y, alpha = 1, lambda =
      best_lambda)
  lasso_coefficients <- coef(final_lasso_model)
y_pred <- predict(final_lasso_model, newx = as.matrix(x)
      )
LASSO_mse \leftarrow mean((y - y_pred)^2)
25 LASSO_mse
ss_res <-sum((y - y_pred)^2)
  ss\_tot \leftarrow sum((y - mean(y))^2)
  r_squared <- 1 - ss_res / ss_tot
_{29} n \leftarrow length (y)
  lasso_coef <- coef(final_lasso_model, s = best_lambda)
_{31} p <- sum(lasso_coef != 0) - 1
_{32} adjusted_r_squared <- 1 - ((1 - r_squared) * (n - 1) / (
      n - p - 1)
33 r_squared
34 adjusted_r_squared
```

### Other Methods

### 5.1 Multivariate Adaptive Regression Splines (MARS)

```
> ss_res <- sum((y - predicted)^2)</pre>
> ss_tot <- sum((y - mean(y))^2)
> r_squared <- 1 - ss_res / ss_tot</pre>
> n <- length(y)
> p <- length(mars_model$selected.terms) - 1</pre>
> adjusted_r_squared <- 1 - ((1 - r_squared) * (n - 1) / (n - p - 1))
> r_squared
[1] 0.8547289
> adjusted_r_squared
[1] 0.8246728
                    (a) R^2 and Adjusted R^2 of MARS Model
> predicted <- predict(mars_model, x)</pre>
> mars_mse <- mean((riboflavin2_boxcox[,1]-predicted)^2)</pre>
> mars_mse
[1] 1.473264
                      (b) Prediction MSE of MARS Model
```

I use the MARS method to fit model, the  $\mathbb{R}^2$  is 0.85, adjusted  $\mathbb{R}^2$  is 0.82 and prediction MSE is 1.47.

```
1 mars_model <- mars(x, y, degree = 2)
2 summary(mars_model)
3 predicted <- predict(mars_model, x)
4 mars_mse <- mean((riboflavin2_boxcox[,1]-predicted)^2)
5 
6 ss_res <- sum((y - predicted)^2)
7 ss_tot <- sum((y - mean(y))^2)
8 r_squared <- 1 - ss_res / ss_tot
9 n <- length(y)
10 p <- length(mars_model$selected.terms) - 1</pre>
```

### 5.2 Classification And Regression Tree(CART)

(b) Prediction MSE of CART Model

Then I use the CART method to fit model, the  $R^2$  is 0.83, adjusted  $R^2$  is 0.81 and prediction MSE is 1.7. But I think CART method is not highly explanatory for continuous variables, since there might be little difference between different classes.

```
1 tree_model <- tree(y ~ .,data = riboflavin2_boxcox)
2 plot(tree_model)
3 text(tree_model)
4
5 predicted <- predict(tree_model, as.data.frame(x))
6 tree_mse <- mean((riboflavin2_boxcox[,1]-predicted)^2)
7
8 ss_res <- sum((y - predicted)^2)
9 ss_tot <- sum((y - mean(y))^2)
10 r_squared <- 1 - ss_res / ss_tot
11 n <- length(y)
12 p <- length(unique(tree_model$frame$var[tree_model$frame$var != "<leaf>"]))
13 adjusted_r_squared <- 1 - ((1 - r_squared) * (n - 1) / (n - p - 1))
14 r_squared
15 adjusted_r_squared</pre>
```

## Elastic Net

Considering that PCA is not y oriented, I try Lasso and Elastic Net directly.

### 6.1 Lasso & Elastic Net

```
> ss_res <- sum((y - prediction_lasso)^2)
> ss_tot <- sum((y - mean(y))^2)
> r_squared <- 1 - ss_res / ss_tot</pre>
> p <- length(which(coef(model_lasso, s = "lambda.min") != 0)) - 1</pre>
> adjusted_r_squared <- 1 - ((1 - r_squared) * (n - 1) / (n - p - 1))
> r_squared
[1] 0.6046345
> adjusted_r_squared
[1] 0.4677773
                    (a) R^2 and Adjusted R^2 of Lasso Model
> mse_lasso <- mean((prediction_lasso - riboflavin$y)^2)</pre>
> mse_lasso
[1] 4.009591
                      (b) Prediction MSE of Lasso Model
> ss_res <- sum((y - prediction_elastic)^2)</pre>
> ss_tot <- sum((y - mean(y))^2)
> r_squared <- 1 - ss_res / ss_tot
> n <- length(y)
> p <- length(which(coef(model_elastic, s = "lambda.min") != 0)) - 1</pre>
> adjusted_r_squared <- 1 - ((1 - r_squared) * (n - 1) / (n - p - 1))
> r_squared
[1] 0.7361478
> adjusted_r_squared
[1] 0.6070286
                 (a) R^2 and Adjusted R^2 of Elastic Net Model
> mse_elastic <- mean((prediction_elastic - riboflavin$y)^2)</pre>
> mse_elastic
[1] 2.675852
                   (b) Prediction MSE of Elastic Net Model
```

First, I use the whole data to fit the models. Comparing the two model, obviously

that Elastic Net is better than Lasso, since Lasso is a special case of Elastic Net.

```
### Lasso ###
2 model_lasso <- cv.glmnet(as.matrix(x), y, alpha = 1)</pre>
  coef(model lasso, s = "lambda.min")
  sum(coef(model lasso, s = "lambda.min") != 0)
  prediction_lasso <- predict(model_lasso, newx = as.</pre>
     matrix(riboflavin[, -1]))
  print(prediction lasso)
  mse_lasso <- mean((prediction_lasso - riboflavin$y)^2)</pre>
  mse lasso
  ss_res \leftarrow sum((y - prediction_lasso)^2)
  ss\_tot \leftarrow sum((y - mean(y))^2)
  r_squared <- 1 - ss_res / ss_tot
  n \leftarrow length(y)
 p <- length (which (coef (model lasso, s = "lambda.min") !=
       0)) - 1
  adjusted_r\_squared \leftarrow 1 - ((1 - r\_squared) * (n - 1) / (
     n - p - 1)
  r squared
  adjusted_r_squared
  ### Elastic Net ###
  alpha_values \leftarrow seq(0, 1, by = 0.1)
  cv_results <- lapply(alpha_values, function(alpha) {
    cv.glmnet(as.matrix(x), y, alpha = alpha)
  cv mse <- sapply(cv results, function(cv) min(cv$cvm))
  best_alpha_index <- which.min(cv_mse)
  best_alpha <- alpha_values[best_alpha_index]
  best_lambda <- cv_results [[best_alpha_index]] $lambda.min
  model_elastic <- glmnet(x, y, alpha = best_alpha, lambda
      = best lambda)
  elastic net coefficients <- coef(model elastic)
  sum(elastic net coefficients != 0)
  elastic_net_coefficients [elastic_net_coefficients != 0]
  prediction_elastic <- predict(model_elastic, newx = as.</pre>
     matrix(riboflavin[, -1]))
  print(prediction_elastic)
```

### 6.2 Elastic Net & Average Elastic Net

```
> sst <- sum((y_test - mean(y_test))^2)
> sse <- sum((y_test - y_test_pred)^2)
> r_squared <- 1 - (sse / sst)
>
> n <- length(y_train)
> p <- length(selected_variables)
> adjusted_r_squared <- 1 - ((1 - r_squared) * (n - 1) / (n - p - 1))
> r_squared
[1] 0.8570204
> adjusted_r_squared
[1] 0.6387884

(a) R² and Adjusted R² of Average Elastic Net Model
> mse <- mean((y_test - y_test_pred)^2)
> mse
[1] 0.08749318
```

(b) Prediction MSE of Average Elastic Net Model

Next, I want to fit a model that can perform better, so I randomly divide 70% of the data into train data and the rest as test data to fit a best elastic net model. After repeat the procedure for 100 times, I choose the variables selected for more than 50 times and take their average coefficients as the new coefficients. In this term, the adjusted  $R^2$  increases from 0.61 to 0.64 and prediction MSE decreases from 2.68 to 0.09. It shows that multiple biased classifiers can be combined to form one robust classifier.

```
num_repeats <- 100
variable_selection_frequency <- matrix(0, nrow = num_
repeats, ncol = ncol(X))
alpha_values <- seq(0, 1, by = 0.1)
for (i in 1:num_repeats) {</pre>
```

```
set.seed(i)
5
     train indices <- sample(1:nrow(X), size = 0.7 * nrow(X)
6
        ))
     X_train <- X[train_indices,]
     y_train <- y[train_indices]</pre>
     X \text{ test} \leftarrow X[-\text{train indices},]
     y_test <- y[-train_indices]
10
     cv_results <- lapply(alpha_values, function(alpha) {
11
       cv.glmnet(as.matrix(X train), y train, alpha = alpha
12
          )
     })
13
     cv_mse <- sapply(cv_results, function(cv) min(cv$cvm))
14
     best alpha index <- which.min(cv mse)
15
     best_alpha <- alpha_values[best_alpha_index]
     best_lambda <- cv_results [[best_alpha_index]] $lambda.
^{17}
        min
     final model <- glmnet(X train, y train, alpha = best
18
        alpha, lambda = best lambda)
     coefs <- coef(final_model, s = best_lambda)
19
     selected\_vars \leftarrow as.numeric(coefs[-1] != 0)
20
     variable selection frequency [i, ] <- selected vars
21
22
   selection_frequency <- colMeans(variable_selection_
      frequency)
   selected_variables <- which (selection_frequency > 0.5)
25 X_train_selected <- X_train[, selected_variables]
26 X_test_selected <- X_test[, selected_variables]
  final model <- glmnet(X train selected, y train, alpha =
       best alpha, lambda = best lambda)
  y test pred <- predict (final model, newx = X test
      selected)
  mean((y_test - y_test_pred)^2)
   sst \leftarrow sum((y_test - mean(y_test))^2)
   sse \leftarrow sum((y test - y test pred)^2)
   r_{squared} \leftarrow 1 - (sse / sst)
  n <- length (y_train)
  p <- length (selected_variables)
  adjusted_r\_squared \leftarrow 1 - ((1 - r\_squared) * (n - 1) / (
      n - p - 1)
  r squared
   adjusted_r_squared
```

# Conlusion

I think the best model is PCA + MARS, because considering that there is a high possibility of the existence of correlations between gene expressions and MARS takes the interactions into account, therefore is better than simple linear model. And compared to directly Elastic Net Method, PCA can deal with multicollinearity better. So this model may have better predictions. Also, it is more interpretive than the Regression Tree.