
A random matrix analysis of random Fourier features: beyond the Gaussian kernel, a precise phase transition, and the corresponding double descent

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Abstract

This article characterizes the exact asymptotics of random Fourier feature (RFF) regression, in the realistic setting where the number of data samples n , their dimension p , and the dimension of feature space N are all large and comparable. In this regime, the random RFF Gram matrix no longer converges to the well-known limiting Gaussian kernel matrix (as it does when $N \rightarrow \infty$ alone), but it still has a tractable behavior that is captured by our analysis. This analysis also provides accurate estimates of training and test regression errors for large n, p, N . Based on these estimates, a precise characterization of two qualitatively different phases of learning, including the phase transition between them, is provided; and the corresponding double descent test error curve is derived from this phase transition behavior. These results do not depend on strong assumptions on the data distribution, and they perfectly match empirical results on real-world data sets.

1 Introduction

For a machine learning system having N parameters, trained on a data set of size n , asymptotic analysis as used in classical statistical learning theory typically either focuses on the (statistical) population $n \rightarrow \infty$ limit, for N fixed, or the over-parameterized $N \rightarrow \infty$ limit, for a given n . These two settings are technically more convenient to work with, yet less practical, as they essentially assume that one of the two dimensions is negligibly small compared to the other, and this is rarely the case in practice. Indeed, with a factor of 2 or 10 more data, one typically works with a more complex model. This has been highlighted perhaps most prominently in recent work on neural network models, in which the model complexity and data size increase together. For this reason, the *double asymptotic* regime where $n, N \rightarrow \infty$, with $N/n \rightarrow c$, a constant, is a particularly interesting (and likely more realistic) limit, despite being technically more challenging [48, 51, 21, 15, 37, 32, 5]. In particular, working in this regime allows for a finer quantitative assessment of machine learning systems, as

a function of their *relative* complexity N/n , as well as for a precise description of the under- to over-parameterized “phase transition” (that does not appear in the $N \rightarrow \infty$ alone analysis). This transition is largely hidden in the usual style of statistical learning theory [49], but it is well-known in the statistical mechanics approach to learning theory [48, 51, 21, 15], and empirical signatures of it have received attention recently under the name “double descent” phenomena [1, 7].

This article considers the asymptotics of random Fourier features [43], and more generally random feature maps, which may be viewed also as a single-hidden-layer neural network model, in this limit. More precisely, let $\mathbf{X} = [\mathbf{x}_1, \dots, \mathbf{x}_n] \in \mathbb{R}^{p \times n}$ denote the data matrix of size n with data vectors $\mathbf{x}_i \in \mathbb{R}^p$ as column vectors. The random feature matrix $\Sigma_{\mathbf{X}}$ of \mathbf{X} is generated by pre-multiplying some random matrix $\mathbf{W} \in \mathbb{R}^{N \times p}$ having i.i.d. entries and then passing through some *entry-wise* nonlinear function $\sigma(\cdot)$, i.e., $\Sigma_{\mathbf{X}} \equiv \sigma(\mathbf{W}\mathbf{X}) \in \mathbb{R}^{N \times n}$. Commonly used random feature techniques such as random Fourier features (RFFs) [43] and homogeneous kernel maps [50], however, rarely involve a single nonlinearity. The popular RFF maps are built with cosine and sine nonlinearities, so that $\Sigma_{\mathbf{X}} \in \mathbb{R}^{2N \times n}$ is obtained by cascading the random features of both, i.e., $\Sigma_{\mathbf{X}}^T \equiv [\cos(\mathbf{W}\mathbf{X})^T, \sin(\mathbf{W}\mathbf{X})^T]$. Note that, by combining both nonlinearities, RFFs generated from $\mathbf{W} \in \mathbb{R}^{N \times p}$ are of dimension $2N$.

The large N asymptotics of random feature maps is closely related to their limiting kernel matrices $\mathbf{K}_{\mathbf{X}}$. In the case of RFF, it was shown in [43] that *entry-wise* the Gram matrix $\Sigma_{\mathbf{X}}^T \Sigma_{\mathbf{X}}/N$ converges to the Gaussian kernel matrix $\mathbf{K}_{\mathbf{X}} \equiv \{\exp(-\|\mathbf{x}_i - \mathbf{x}_j\|^2/2)\}_{i,j=1}^n$, as $N \rightarrow \infty$. This follows from $\frac{1}{N}[\Sigma_{\mathbf{X}}^T \Sigma_{\mathbf{X}}]_{ij} = \frac{1}{N} \sum_{t=1}^N \cos(\mathbf{x}_i^T \mathbf{w}_t) \cos(\mathbf{x}_j^T \mathbf{w}_t) + \sin(\mathbf{x}_i^T \mathbf{w}_t) \sin(\mathbf{x}_j^T \mathbf{w}_t)$, for \mathbf{w}_t independent Gaussian random vectors, so that by the strong law of large numbers, for fixed n, p , $[\Sigma_{\mathbf{X}}^T \Sigma_{\mathbf{X}}/N]_{ij}$ goes to its expectation (with respect to $\mathbf{w} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_p)$) almost surely as $N \rightarrow \infty$, i.e.,

$$[\Sigma_{\mathbf{X}}^T \Sigma_{\mathbf{X}}/N]_{ij} \xrightarrow{a.s.} \mathbb{E}_{\mathbf{w}} [\cos(\mathbf{x}_i^T \mathbf{w}) \cos(\mathbf{x}_j^T \mathbf{w}) + \sin(\mathbf{x}_i^T \mathbf{w}) \sin(\mathbf{x}_j^T \mathbf{w})] \equiv \mathbf{K}_{\cos} + \mathbf{K}_{\sin}, \quad (1)$$

with

$$\mathbf{K}_{\cos} + \mathbf{K}_{\sin} \equiv e^{-\frac{1}{2}(\|\mathbf{x}_i\|^2 + \|\mathbf{x}_j\|^2)} (\cosh(\mathbf{x}_i^T \mathbf{x}_j) + \sinh(\mathbf{x}_i^T \mathbf{x}_j)) = e^{-\frac{1}{2}\|\mathbf{x}_i - \mathbf{x}_j\|^2} \equiv [\mathbf{K}_{\mathbf{X}}]_{ij}. \quad (2)$$

While this result holds in the $N \rightarrow \infty$ limit, recent advances in random matrix theory [30, 27] suggest that, in the more practical setting where N is not much larger than n, p and $n, p, N \rightarrow \infty$ at the same pace, the situation is more subtle. In particular, the above entry-wise convergence remains valid, but the convergence $\|\Sigma_{\mathbf{X}}^T \Sigma_{\mathbf{X}}/N - \mathbf{K}_{\mathbf{X}}\| \rightarrow 0$ no longer holds in spectral norm, due to the factor n , now large, in the norm inequality $\|\mathbf{A}\|_{\infty} \leq \|\mathbf{A}\| \leq n\|\mathbf{A}\|_{\infty}$ for $\mathbf{A} \in \mathbb{R}^{n \times n}$ and $\|\mathbf{A}\|_{\infty} \equiv \max_{ij} |\mathbf{A}_{ij}|$. This implies that, in the large n, p, N regime, the assessment of the behavior of $\Sigma_{\mathbf{X}}^T \Sigma_{\mathbf{X}}/N$ via $\mathbf{K}_{\mathbf{X}}$ may result in a spectral norm error that blows up. As a consequence, for various machine learning algorithms [10], the performance guarantee offered by the limiting Gaussian kernel is less likely to agree with empirical observations in real-world large-scale problems, when n, p are large.¹

1.1 Our Main Contributions

We consider the RFF model in the more realistic large n, p, N limit. While, in this setting, the RFF empirical Gram matrix does *not* converge to the Gaussian kernel matrix, we can characterize its behavior as $n, p, N \rightarrow \infty$ and provide *asymptotic performance guarantees* for RFF on large-scale problems. We also identify a phase transition as a function of the ratio N/n , including the corresponding double descent phenomenon. In more detail, our contributions are the following.

1. We provide a *precise* characterization of the asymptotics of the RFF empirical Gram matrix, in the large n, p, N limit (Theorem 1). This is accomplished by constructing a deterministic equivalent for the resolvent of the RFF Gram matrix. Based on this, the behavior of the RFF model is (asymptotically) accessible through a fixed-point equation, that can be interpreted in terms of an angle-like correction induced by the non-trivial large n, p, N limit (relative to the $N \rightarrow \infty$ alone limit).
2. We derive the asymptotic training and test mean squared errors (MSEs) of RFF ridge regression, as a function of the ratio N/n , regularization penalty λ , training as well as test sets (Theorem 2 and 3, respectively). We identify precisely the under- to over-parameterization phase transition, as a function

¹For readers not familiar with the impact of spectral norm error in learning, or with the random matrix theory techniques that we will use in our analysis, such as resolvent analysis and the use of deterministic equivalents, see Appendix A for a warm-up discussion.

of the relative model complexity N/n ; we prove the existence of a “singular” peak of test error at the $N/n = 1/2$ boundary; and we characterize the corresponding *double descent* behavior. Importantly, our results are valid *with almost no specific assumption* on the data distribution. This is a significant improvement over existing double descent analyses, which fundamentally rely on the knowledge of the data distribution (often assumed to be multivariate Gaussian for simplicity) [20, 36].

3. We provide a detailed empirical evaluation of our theoretical results, demonstrating that the theory closely matches empirical results on a range of real-world data sets (Section 3 and Section F in the supplementary material). This includes the correction due to the large n, p, N setting, sharp transitions (as a function of N/n) in angle-like quantities, and the corresponding double descent test curves. This also includes an evaluation of the impact of training-test similarity and the effect of different data sets, thus confirming, as stated in 2., that (unlike in prior work) the phase transition and double descent hold with almost no specific assumption on the data distribution.

1.2 Related Work

Here, we provide a brief review of related previous efforts.

Random features and limiting kernels. In most RFF work [44, 4, 3, 45], non-asymptotic bounds are given, on the number of random features N needed for a predefined approximation error of a given kernel matrix with fixed n, p . A more recent line of work [2, 14, 22, 9] has focused on the over-parameterized $N \rightarrow \infty$ limit of large neural networks by studying the corresponding *neural tangent kernels*. Here, we position ourselves in the more practical regime where n, p, N are all large and comparable, and provide *asymptotic performance guarantees* that better fit large-scale problems.

Random matrix theory. From a random matrix theory perspective, nonlinear Gram matrices of the type $\Sigma_X^T \Sigma_X$ have recently received an unprecedented research interests, due to their close connection to neural networks [41, 39, 8, 38], with a particular focus on the associated eigenvalue distribution. Here we propose a deterministic equivalent [11, 19] analysis for the resolvent matrix that provides access, not only to the eigenvalue distribution, but also to the regression error of central interest in this article. While most existing deterministic equivalent analyses are performed on linear models, here we focus on the *nonlinear* RFF model. From a technical perspective, the most relevant work is [30, 36]. We improve their results by considering *generic* data model on the popular RFF model.

Statistical mechanics of learning. A long history of connections between statistical mechanics and machine learning models (such as neural networks) exists, including a range of techniques to establish generalization bounds [48, 51, 21, 15], and recently there has been renewed interest [32, 34, 33, 35, 5]. Their relevance to our results lies in the use of the thermodynamic limit (akin to the large n, p, N limit), rather than the classical limits more commonly used in statistical learning theory, where uniform convergence bounds and related techniques can be applied.

Double descent in large-scale learning systems. The large n, N asymptotics of statistical models has received considerable research interests in the machine learning community [40, 20], resulting in a (somehow) counterintuitive phenomenon referred to as the “double descent.” Instead of focusing on different “phases of learning” [48, 51, 21, 15, 32], the “double descent” phenomenon focuses on an empirical manifestation of the phase boundary and refers to the empirical observations of the test error curve as a function of the model complexity, which differs from the usual textbook description of the bias-variance tradeoff [1, 26, 7, 17]. Theoretical investigation into this phenomenon mainly focuses on various regression models [13, 6, 12, 25, 20, 36]. In most cases, quite specific (and rather strong) assumptions are imposed on the input data distribution. In this respect, our work extends the analysis in [36] to handle the RFF model and its phase structure *on real-world data sets*.

1.3 Notations and Organization of the Paper

Throughout this article, we follow the convention of denoting scalars by lowercase, vectors by lowercase boldface, and matrices by uppercase boldface letters. In addition, the notation $(\cdot)^T$ denotes the transpose operator; the norm $\|\cdot\|$ is the Euclidean norm for vectors and the spectral or operator norm for matrices; and $\xrightarrow{a.s.}$ stands for almost sure convergence of random variables.

Our main results on the asymptotic training and test MSEs of RFF ridge regression are presented in Section 2, with proofs deferred to the Appendix. In Section 3, we provide detailed empirical evaluations of our main results, as well as discussions on the corresponding phase transition behavior and the double descent test curve. Concluding remarks are placed in Section 4. For more detailed discussions and empirical evaluations, we refer the readers to an extended version of this article [28].

2 Main Technical Results

In this section, we present our main theoretical results. To investigate the large n, p, N asymptotics of the RFF model, we shall technically position ourselves under the following assumption.

Assumption 1. *As $n \rightarrow \infty$, we have*

1. $0 < \liminf_n \min\{\frac{p}{n}, \frac{N}{n}\} \leq \limsup_n \max\{\frac{p}{n}, \frac{N}{n}\} < \infty$; or, practically speaking, the ratios p/n and N/n are only moderately large or moderately small.
2. $\limsup_n \|\mathbf{X}\| < \infty$ and $\limsup_n \|\mathbf{y}\|_\infty < \infty$, i.e., they are normalized with respect to n .

Under Assumption 1, we consider the RFF regression model. For training data $\mathbf{X} \in \mathbb{R}^{p \times n}$ of size n , the associated random Fourier features, $\Sigma_{\mathbf{X}} \in \mathbb{R}^{2N \times n}$, are obtained by computing $\mathbf{W}\mathbf{X} \in \mathbb{R}^{N \times n}$, for standard Gaussian random matrix $\mathbf{W} \in \mathbb{R}^{N \times p}$, and then applying entry-wise cosine and sine nonlinearities on $\mathbf{W}\mathbf{X}$, i.e., $\Sigma_{\mathbf{X}}^\top = [\cos(\mathbf{W}\mathbf{X})^\top, \sin(\mathbf{W}\mathbf{X})^\top]$ with $\mathbf{W}_{ij} \sim \mathcal{N}(0, 1)$. Given this setup, the RFF ridge regressor $\beta \in \mathbb{R}^{2N}$ is given by, for $\lambda \geq 0$,

$$\beta \equiv \frac{1}{n} \Sigma_{\mathbf{X}} \left(\frac{1}{n} \Sigma_{\mathbf{X}}^\top \Sigma_{\mathbf{X}} + \lambda \mathbf{I}_n \right)^{-1} \mathbf{y} \cdot \mathbf{1}_{2N > n} + \left(\frac{1}{n} \Sigma_{\mathbf{X}} \Sigma_{\mathbf{X}}^\top + \lambda \mathbf{I}_{2N} \right)^{-1} \frac{1}{n} \Sigma_{\mathbf{X}} \mathbf{y} \cdot \mathbf{1}_{2N < n}. \quad (3)$$

The two forms of β in (3) are equivalent for any $\lambda > 0$ and minimize the (ridge-regularized) squared loss $\frac{1}{n} \|\mathbf{y} - \Sigma_{\mathbf{X}}^\top \beta\|^2 + \lambda \|\beta\|^2$ on the training set (\mathbf{X}, \mathbf{y}) . Our objective is to characterize the large n, p, N asymptotics of both the *training MSE*, E_{train} , and the *test MSE*, E_{test} , defined as

$$E_{\text{train}} = \frac{1}{n} \|\mathbf{y} - \Sigma_{\mathbf{X}}^\top \beta\|^2, \quad E_{\text{test}} = \frac{1}{\hat{n}} \|\hat{\mathbf{y}} - \Sigma_{\hat{\mathbf{X}}}^\top \beta\|^2, \quad (4)$$

with $\Sigma_{\hat{\mathbf{X}}}^\top \equiv [\cos(\mathbf{W}\hat{\mathbf{X}})^\top, \sin(\mathbf{W}\hat{\mathbf{X}})^\top] \in \mathbb{R}^{\hat{n} \times 2N}$ on a test set $(\hat{\mathbf{X}}, \hat{\mathbf{y}})$ of size \hat{n} .

2.1 Asymptotic Deterministic Equivalent

To start, we observe that the training MSE, E_{train} , in (4), can be written as $E_{\text{train}} = \frac{\lambda^2}{n} \|\mathbf{Q}(\lambda) \mathbf{y}\|^2 = -\frac{\lambda^2}{n} \mathbf{y}^\top \partial \mathbf{Q}(\lambda) \mathbf{y} / \partial \lambda$, which depends on the quadratic form $\mathbf{y}^\top \mathbf{Q}(\lambda) \mathbf{y}$ of

$$\mathbf{Q}(\lambda) \equiv \left(\frac{1}{n} \Sigma_{\mathbf{X}}^\top \Sigma_{\mathbf{X}} + \lambda \mathbf{I}_n \right)^{-1} \in \mathbb{R}^{n \times n}, \quad (5)$$

the so-called *resolvent* of $\frac{1}{n} \Sigma_{\mathbf{X}}^\top \Sigma_{\mathbf{X}}$ (also denoted \mathbf{Q} when there is no ambiguity) with $\lambda > 0$.

In order to assess the asymptotic training MSE, it thus suffices to find a deterministic equivalent for $\mathbf{Q}(\lambda)$ (i.e., a *deterministic* matrix that captures the asymptotic behavior of the latter). One possibility is the expectation $\mathbb{E}_{\mathbf{W}}[\mathbf{Q}(\lambda)]$. Informally, if the training MSE E_{train} (that is random due to random \mathbf{W}) is “close to” some deterministic quantity \bar{E}_{train} , in the large n, p, N limit, then \bar{E}_{train} must have the same limit as $\mathbb{E}_{\mathbf{W}}[E_{\text{train}}] = -\frac{\lambda^2}{n} \partial \mathbf{y}^\top \mathbb{E}_{\mathbf{W}}[\mathbf{Q}(\lambda)] \mathbf{y} / \partial \lambda$ for $n, p, N \rightarrow \infty$. However, $\mathbb{E}_{\mathbf{W}}[\mathbf{Q}]$ involves integration (with no closed-form due to the matrix inverse), and it is not a convenient quantity with which to work. Our objective is to find an asymptotic “alternative” for $\mathbb{E}_{\mathbf{W}}[\mathbf{Q}]$ that is (i) close to $\mathbb{E}_{\mathbf{W}}[\mathbf{Q}]$ in the large $n, p, N \rightarrow \infty$ limit and (ii) numerically more accessible.

In the following theorem (proved in Appendix B), we introduce an asymptotic equivalent for $\mathbb{E}_{\mathbf{W}}[\mathbf{Q}]$. Instead of being directly related to the Gaussian kernel $\mathbf{K}_{\mathbf{X}} = \mathbf{K}_{\text{cos}} + \mathbf{K}_{\text{sin}}$ as suggested by (2) in the large- N -only limit, it depends on the two components $\mathbf{K}_{\text{cos}}, \mathbf{K}_{\text{sin}}$ in a more involved manner.

Theorem 1 (Asymptotic equivalent for $\mathbb{E}_{\mathbf{W}}[\mathbf{Q}]$). *Under Assumption 1, for \mathbf{Q} defined in (5) and $\lambda > 0$, we have, as $n \rightarrow \infty$*

$$\|\mathbb{E}_{\mathbf{W}}[\mathbf{Q}] - \bar{\mathbf{Q}}\| \rightarrow 0$$

for $\bar{\mathbf{Q}} \equiv \left(\frac{N}{n} \left(\frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} + \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}} \right) + \lambda \mathbf{I}_n \right)^{-1}$, $\mathbf{K}_{\cos} \equiv \mathbf{K}_{\cos}(\mathbf{X}, \mathbf{X})$, $\mathbf{K}_{\sin} \equiv \mathbf{K}_{\sin}(\mathbf{X}, \mathbf{X}) \in \mathbb{R}^{n \times n}$ and

$$\mathbf{K}_{\cos}(\mathbf{X}, \mathbf{X}')_{ij} = e^{-\frac{\|\mathbf{x}_i\|^2 + \|\mathbf{x}'_j\|^2}{2}} \cosh(\mathbf{x}_i^\top \mathbf{x}'_j), \quad \mathbf{K}_{\sin}(\mathbf{X}, \mathbf{X}')_{ij} = e^{-\frac{\|\mathbf{x}_i\|^2 + \|\mathbf{x}'_j\|^2}{2}} \sinh(\mathbf{x}_i^\top \mathbf{x}'_j), \quad (6)$$

where $(\delta_{\cos}, \delta_{\sin})$ is the unique positive solution to

$$\delta_{\cos} = \frac{1}{n} \text{tr}(\mathbf{K}_{\cos} \bar{\mathbf{Q}}), \quad \delta_{\sin} = \frac{1}{n} \text{tr}(\mathbf{K}_{\sin} \bar{\mathbf{Q}}). \quad (7)$$

Remark 1 (Correction to large- N behavior). Taking $N/n \rightarrow \infty$, one has $\delta_{\cos} \rightarrow 0$, $\delta_{\sin} \rightarrow 0$ so that $\frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} + \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}} \rightarrow \mathbf{K}_{\cos} + \mathbf{K}_{\sin} = \mathbf{K}_{\mathbf{X}}$ and $\bar{\mathbf{Q}} \simeq \frac{n}{N} \mathbf{K}_{\mathbf{X}}^{-1}$, for $\lambda > 0$, in accordance with the classical large- N -only prediction. In this sense, the pair $(\delta_{\cos}, \delta_{\sin})$ introduced in Theorem 1 accounts for the “correction” due to the non-trivial n/N , as opposed to the $N \rightarrow \infty$ alone analysis. Also, when the number of features N is large (i.e., as $N/n \rightarrow \infty$), the regularization effect of λ flattens out and $\bar{\mathbf{Q}}$ behaves like (a scaled version of) the inverse Gaussian kernel matrix $\mathbf{K}_{\mathbf{X}}^{-1}$ (that is well-defined if $\mathbf{x}_1, \dots, \mathbf{x}_n$ are all distinct, see [46, Theorem 2.18]).

Remark 2 (Geometric interpretation). Since $\bar{\mathbf{Q}}$ shares the same eigenspace with $\frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} + \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}}$, one can geometrically interpret $(\delta_{\cos}, \delta_{\sin})$ as a sort of “angle” between the eigenspaces of \mathbf{K}_{\cos} , \mathbf{K}_{\sin} and that of $\frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} + \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}}$. For fixed n , as $N \rightarrow \infty$, one has $\frac{1}{N} \sum_{t=1}^N \cos(\mathbf{X}^\top \mathbf{w}_t) \cos(\mathbf{w}_t^\top \mathbf{X}) \rightarrow \mathbf{K}_{\cos}$, $\frac{1}{N} \sum_{t=1}^N \sin(\mathbf{X}^\top \mathbf{w}_t) \sin(\mathbf{w}_t^\top \mathbf{X}) \rightarrow \mathbf{K}_{\sin}$, the eigenspaces of which are “orthogonal” to each other, so that $\delta_{\cos}, \delta_{\sin} \rightarrow 0$. On the other hand, as $N, n \rightarrow \infty$, the eigenspaces of \mathbf{K}_{\cos} and \mathbf{K}_{\sin} “intersect” with each other, captured by the non-trivial $(\delta_{\cos}, \delta_{\sin})$.

2.2 Asymptotic Training Performance

Theorem 1 provides an asymptotically more tractable approximation of $\mathbb{E}_{\mathbf{W}}[\mathbf{Q}]$. Together with some additional concentration arguments (e.g., from [30, Theorem 2]), this permits us to provide a complete description of the limiting behavior of the *random* bilinear form $\mathbf{a}^\top \mathbf{Q} \mathbf{b}$, for $\mathbf{a}, \mathbf{b} \in \mathbb{R}^n$ of bounded Euclidean norms, in such a way that $\mathbf{a}^\top \mathbf{Q} \mathbf{b} - \mathbf{a}^\top \bar{\mathbf{Q}} \mathbf{b} \xrightarrow{a.s.} 0$, as $n, p, N \rightarrow \infty$. This, together with the fact that $E_{\text{train}} = \frac{\lambda^2}{n} \mathbf{y}^\top \mathbf{Q}(\lambda)^2 \mathbf{y} = -\frac{\lambda^2}{n} \mathbf{y}^\top \partial \mathbf{Q}(\lambda) \mathbf{y} / \partial \lambda$, leads to the following result on the asymptotic training error, the proof of which is given in Appendix C.

Theorem 2 (Asymptotic training performance). *Under Assumption 1, for a given training set (\mathbf{X}, \mathbf{y}) and training MSE, E_{train} defined in (4), as $n \rightarrow \infty$*

$$E_{\text{train}} - \bar{E}_{\text{train}} \xrightarrow{a.s.} 0, \quad \bar{E}_{\text{train}} = \frac{\lambda^2}{n} \|\bar{\mathbf{Q}} \mathbf{y}\|^2 + \frac{N}{n} \frac{\lambda^2}{n^2} \left[\frac{\text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}})}{(1+\delta_{\cos})^2} - \frac{\text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\sin} \bar{\mathbf{Q}})}{(1+\delta_{\sin})^2} \right] \Omega \begin{bmatrix} \mathbf{y}^\top \mathbf{Q} \mathbf{K}_{\cos} \bar{\mathbf{Q}} \mathbf{y} \\ \mathbf{y}^\top \mathbf{Q} \mathbf{K}_{\sin} \bar{\mathbf{Q}} \mathbf{y} \end{bmatrix}$$

for $\bar{\mathbf{Q}}$ defined in Theorem 1 and

$$\Omega^{-1} \equiv \mathbf{I}_2 - \frac{N}{n} \begin{bmatrix} \frac{1}{n} \frac{\text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}} \mathbf{K}_{\cos})}{(1+\delta_{\cos})^2} & \frac{1}{n} \frac{\text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}} \mathbf{K}_{\sin})}{(1+\delta_{\sin})^2} \\ \frac{1}{n} \frac{\text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\sin} \bar{\mathbf{Q}} \mathbf{K}_{\cos})}{(1+\delta_{\cos})^2} & \frac{1}{n} \frac{\text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\sin} \bar{\mathbf{Q}} \mathbf{K}_{\sin})}{(1+\delta_{\sin})^2} \end{bmatrix}. \quad (8)$$

One can show that (i) for a given n and $\lambda > 0$, \bar{E}_{train} decreases as the model size N increases; and (ii) for a given ratio N/n , \bar{E}_{train} increases as the regularization penalty λ grows large, as expected.

2.3 Asymptotic Test Performance

Theorem 2 holds without any restriction on the training set, (\mathbf{X}, \mathbf{y}) , except for Assumption 1, since only the randomness of \mathbf{W} is involved, and thus one can simply treat (\mathbf{X}, \mathbf{y}) as known in this result. This is no longer the case for the test error. Intuitively, the test data $\hat{\mathbf{X}}$ cannot be chosen arbitrarily, and one must ensure that the test data “behave” statistically like the training data, in a “well-controlled” manner, so that the test MSE is asymptotically deterministic and bounded as $n, \hat{n}, p, N \rightarrow \infty$. Following this intuition, we work under the following assumption.

Assumption 2 (Data as concentrated random vectors [29]). *The training data $\mathbf{x}_i \in \mathbb{R}^p, i \in \{1, \dots, n\}$, are independently drawn (non-necessarily uniformly) from one of $K > 0$ distribution*

classes² μ_1, \dots, μ_K . There exist constants $C, \eta, q > 0$ such that for any $\mathbf{x}_i \sim \mu_k, k \in \{1, \dots, K\}$ and any 1-Lipschitz function $f : \mathbb{R}^p \rightarrow \mathbb{R}$, we have

$$\mathbb{P}(|f(\mathbf{x}_i) - \mathbb{E}[f(\mathbf{x}_i)]| \geq t) \leq Ce^{-(t/\eta)^q}, \quad t \geq 0. \quad (9)$$

The test data $\hat{\mathbf{x}}_i \sim \mu_k, i \in \{1, \dots, \hat{n}\}$ are mutually independent, but may depend on training data \mathbf{X} and $\|\mathbb{E}[\sigma(\mathbf{W}\mathbf{X}) - \sigma(\mathbf{W}\hat{\mathbf{X}})]\| = O(\sqrt{n})$ for $\sigma \in \{\cos, \sin\}$.

To facilitate the discussion of the phase transition and the double descent, we do not assume independence between training data and test data (but we do assume independence between different columns within \mathbf{X} and $\hat{\mathbf{X}}$). In this respect, Assumption 2 is weaker than the classical i.i.d. assumption, and it permits us to illustrate the impact of training-test similarity on the model performance (Section 3.3).

A first example of concentrated random vectors satisfying (9) is the random Gaussian vector $\mathcal{N}(\mathbf{0}, \mathbf{I}_p)$ [24]. Moreover, since the concentration property in (9) is stable over Lipschitz transformations [29], it holds, for any 1-Lipschitz mapping $g : \mathbb{R}^d \rightarrow \mathbb{R}^p$ and $\mathbf{z} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_d)$, that $g(\mathbf{z})$ also satisfies (9). In this respect, Assumption 2, although seemingly quite restrictive, represents a large family of “generative models”, including notably the “fake images” generated by modern generative adversarial networks (GANs) that are, by construction, Lipschitz transformations of large random Gaussian vectors [18, 47]. As such, from a practical consideration, Assumption 2 provides a more realistic and flexible statistical model for real-world data.

With Assumption 2, we have the following result on the asymptotic test error, proved in Section D.

Theorem 3 (Asymptotic test performance). *Under Assumptions 1 and 2, we have, for test MSE E_{test} defined in (4) and test data $(\hat{\mathbf{X}}, \hat{\mathbf{y}})$ satisfying $\limsup_{\hat{n}} \|\hat{\mathbf{X}}\| < \infty$, $\limsup_{\hat{n}} \|\hat{\mathbf{y}}\|_\infty < \infty$ with $\hat{n}/n \in (0, \infty)$ that, as $n \rightarrow \infty$*

$$E_{\text{test}} - \bar{E}_{\text{test}} \xrightarrow{a.s.} 0, \quad \bar{E}_{\text{test}} = \frac{1}{\hat{n}} \|\hat{\mathbf{y}} - \frac{N}{n} \hat{\Phi} \bar{\mathbf{Q}} \mathbf{y}\|^2 + \frac{N^2}{n^2 \hat{n}} \begin{bmatrix} \frac{\Theta_{\cos}}{(1+\delta_{\cos})^2} & \frac{\Theta_{\sin}}{(1+\delta_{\sin})^2} \end{bmatrix} \Omega \begin{bmatrix} \mathbf{y}^\top \bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}} \mathbf{y} \\ \mathbf{y}^\top \bar{\mathbf{Q}} \mathbf{K}_{\sin} \bar{\mathbf{Q}} \mathbf{y} \end{bmatrix}$$

for Ω defined in (8),

$$\Theta_\sigma = \frac{1}{N} \text{tr} \mathbf{K}_\sigma(\hat{\mathbf{X}}, \hat{\mathbf{X}}) + \frac{N}{n} \frac{1}{n} \text{tr} \bar{\mathbf{Q}} \hat{\Phi}^\top \hat{\Phi} \bar{\mathbf{Q}} \mathbf{K}_\sigma - \frac{2}{n} \text{tr} \bar{\mathbf{Q}} \hat{\Phi}^\top \mathbf{K}_\sigma(\hat{\mathbf{X}}, \mathbf{X}), \quad \sigma \in \{\cos, \sin\}, \quad (10)$$

and $\hat{\Phi} \equiv \frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} + \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}}$, $\hat{\Phi} \equiv \frac{\mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X})}{1+\delta_{\cos}} + \frac{\mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X})}{1+\delta_{\sin}}$, with $\mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X}), \mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X}) \in \mathbb{R}^{\hat{n} \times n}$ and $\mathbf{K}_{\cos}(\hat{\mathbf{X}}, \hat{\mathbf{X}}), \mathbf{K}_{\sin}(\hat{\mathbf{X}}, \hat{\mathbf{X}}) \in \mathbb{R}^{\hat{n} \times \hat{n}}$ defined as in (6).

Taking $(\hat{\mathbf{X}}, \hat{\mathbf{y}}) = (\mathbf{X}, \mathbf{y})$, one gets $\bar{E}_{\text{test}} = \bar{E}_{\text{train}}$, as expected. From this perspective, Theorem 3 can be seen as an extension of Theorem 2, with the “interaction” between training and test data (i.e., training-versus-test $\mathbf{K}_\sigma(\hat{\mathbf{X}}, \mathbf{X})$ and test-versus-test $\mathbf{K}_\sigma(\hat{\mathbf{X}}, \hat{\mathbf{X}})$ interaction matrices) summarized in the scalar parameter Θ_σ defined in (10), for $\sigma \in \{\cos, \sin\}$.

3 Empirical Evaluations and Practical Implications

In this section, we provide a detailed empirical evaluation, including a discussion of the behavior of the fixed-point equation in Theorem 1, and its consequences in Theorem 2 and Theorem 3. In particular, we describe the behavior of the pair $(\delta_{\cos}, \delta_{\sin})$ that characterizes the necessary correction in the large n, p, N regime, as a function of the regularization λ and the ratio N/n . This explains: (i) the mismatch between empirical regression errors from the Gaussian kernel prediction (Figure 1); and (ii) the behavior of $(\delta_{\cos}, \delta_{\sin})$ as a function of N/n , which clearly indicates two phases of learning (Figure 3) and the corresponding double descent test error curves (Figure 4).

3.1 Correction due to the Large n, p, N Regime

The RFF Gram matrix $\Sigma_{\mathbf{X}}^\top \Sigma_{\mathbf{X}}/N$ is *not* close to the classical Gaussian kernel matrix $\mathbf{K}_{\mathbf{X}}$ in the large n, p, N regime; and, as a consequence, its resolvent \mathbf{Q} , as well the training and test MSE, E_{train} and E_{test} (that are functions of \mathbf{Q}), behave quite differently from the Gaussian kernel predictions.

² $K \geq 2$ is included to cover multi-class classification problems; and K should remain fixed as $n, p \rightarrow \infty$.

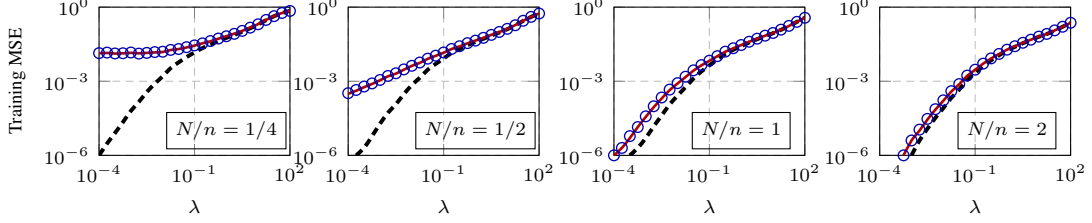


Figure 1: Training MSEs of RFF ridge regression on MNIST data (class 3 versus 7), as a function of regression penalty λ , for $p = 784$, $n = 1000$, $N = 250, 500, 1000, 2000$. Empirical results displayed in **blue** circles; Gaussian kernel predictions (assuming $N \rightarrow \infty$ alone) in **black** dashed lines; and Theorems 2 in **red** solid lines. Results obtained by averaging over 30 runs.

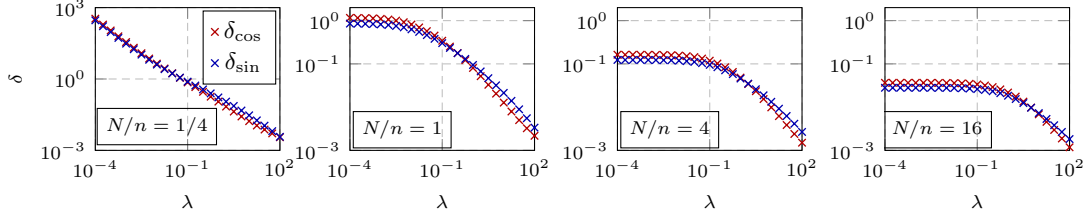


Figure 2: Behavior of $(\delta_{\cos}, \delta_{\sin})$ in (11) on MNIST data (class 3 versus 7), as a function of the regularization parameter λ , for $p = 784$, $n = 1000$, $N = 250, 1000, 4000, 16000$.

As already discussed in Remark 1 after Theorem 1, for $\lambda > 0$, the pair $(\delta_{\cos}, \delta_{\sin})$ characterizes the correction when considering n, p, N all large, compared to the large- N -only asymptotic behavior:

$$\delta_{\cos} = \frac{1}{n} \text{tr} \mathbf{K}_{\cos} \bar{\mathbf{Q}}, \quad \delta_{\sin} = \frac{1}{n} \text{tr} \mathbf{K}_{\sin} \bar{\mathbf{Q}}, \quad \bar{\mathbf{Q}} = \left(\frac{N}{n} \left(\frac{\mathbf{K}_{\cos}}{1 + \delta_{\cos}} + \frac{\mathbf{K}_{\sin}}{1 + \delta_{\sin}} \right) + \lambda \mathbf{I}_n \right)^{-1}. \quad (11)$$

To start, Figure 1 compares the training MSEs of RFF ridge regression to the predictions from Gaussian kernel regression and to the predictions from our Theorem 2, on the popular MNIST data set [23]. Observe that there is a huge gap between empirical training errors and the Gaussian kernel predictions, especially when $N/n < 1$, while our theory *consistently* fits empirical observations almost perfectly.

Next, from (11) we know that both δ_{\cos} and δ_{\sin} are decreasing functions of λ . (See Lemma 7 in Appendix E for a proof of this fact.) Figure 2 shows that: (i) over a range of different N/n , both δ_{\cos} and δ_{\sin} decrease monotonically as λ increases; (ii) the behavior for $N/n < 1$, which is decreasing from an initial value of $\delta \gg 1$, is very different from the behavior for $N/n \gtrsim 1$, where an initially flat region is observed for small values of λ and we have $\delta < 1$ for all values of λ ; and (iii) the impact of regularization λ becomes less significant as the ratio N/n becomes large. This is in accordance with the limiting behavior of $\bar{\mathbf{Q}} \simeq \frac{n}{N} \mathbf{K}_{\mathbf{X}}^{-1}$ in Remark 1 that is *independent* of λ as $N/n \rightarrow \infty$.

Note also that, while δ_{\cos} and δ_{\sin} can be geometrically interpreted as a sort of weighted “angle” between different kernel matrices (as in Remark 2), and therefore one might expect to have $\delta \in [0, 1]$, this is not the case for the leftmost plot of Figure 1 with $N/n = 1/4$. There, for small values of λ (say $\lambda \lesssim 0.1$), both δ_{\cos} and δ_{\sin} scale like λ^{-1} , while they are observed to saturate to a fixed $O(1)$ value for $N/n = 1, 4, 16$. This corresponds to two different phases of learning in the ridgeless $\lambda \rightarrow 0$ limit, as discussed in the following section.

3.2 Phase Transition and Corresponding Double Descent

Both δ_{\cos} and δ_{\sin} in (11) are decreasing functions of N , as depicted in Figure 3. (See Lemma 6 in Appendix E for a proof.) More importantly, Figure 3 also illustrates that δ_{\cos} and δ_{\sin} exhibit qualitatively different behavior, depending on the ratio N/n . For λ not too small ($\lambda = 1$ or 10), both δ_{\cos} and δ_{\sin} decrease *smoothly*, as N/n grows large. However, for λ relatively small ($\lambda = 10^{-3}$ and 10^{-7}), we observe a “phase transition” on two sides of the interpolation threshold $2N = n$. (Note

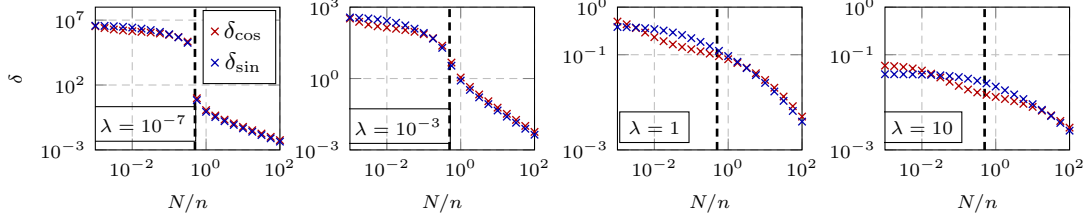


Figure 3: Behavior of $(\delta_{\cos}, \delta_{\sin})$ on MNIST data (class 3 versus 7), as a function of N/n , $p = 784$, $n = 1000$, $\lambda = 10^{-7}, 10^{-3}, 1, 10$. The **black** dashed line is the interpolation threshold $2N = n$.

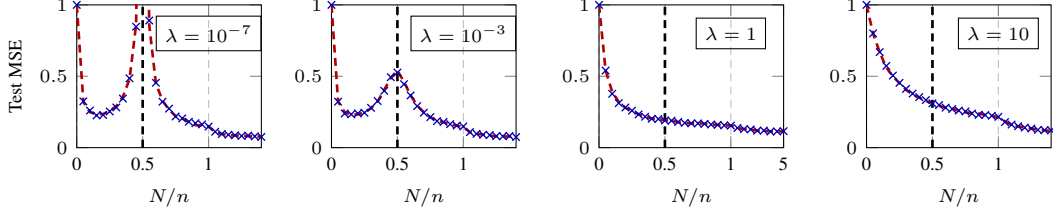


Figure 4: Empirical (**blue** crosses) and theoretical (**red** dashed lines) test errors of RFF regression as a function of the ratio N/n , on MNIST data set (class 3 versus 7), for $p = 784$, $n = 500$, $\lambda = 10^{-7}, 10^{-3}, 1, 10$. The **black** dashed line is the interpolation threshold $2N = n$.

that the scale of the y-axis is different in different subfigures.) More precisely, in the leftmost plot with $\lambda = 10^{-7}$, δ_{\cos} and δ_{\sin} “jump” from order $O(1)$ (when $2N > n$) to much higher values of the same order of λ^{-1} (when $2N < n$). A similar behavior is also observed for $\lambda = 10^{-3}$.

This phase transition can be theoretically justified by considering the *ridgeless* $\lambda \rightarrow 0$ limit in Theorem 1. First note that, for $\lambda = 0$ and $2N < n$, the (random) resolvent $\mathbf{Q}(\lambda = 0)$ in (5) is simply undefined, as it involves inverting a singular matrix $\Sigma_{\mathbf{X}}^T \Sigma_{\mathbf{X}} \in \mathbb{R}^{n \times n}$ that is of rank at most $2N < n$. As a consequence, we expect to see both \mathbf{Q} and $\bar{\mathbf{Q}}$ scale like λ^{-1} as $\lambda \rightarrow 0$ for $2N < n$, while for $2N > n$ this is no longer the case. As a consequence, we have the following two phases:

1. *Under-parameterized* with $2N < n$. Here, \mathbf{Q} is not well-defined (indeed \mathbf{Q} scales like λ^{-1}) and one must consider instead the properly scaled $\lambda \delta_{\cos}$, $\lambda \delta_{\sin}$ and $\lambda \bar{\mathbf{Q}}$ as $\lambda \rightarrow 0$.
2. *Over-parameterized* with $2N > n$, where one can take $\lambda \rightarrow 0$ in (11) to get δ_{\cos} , δ_{\sin} and $\bar{\mathbf{Q}}$.

Remark 3 (Double descent test error curves). On account of the above two phases, it is not surprising to observe a “singular” behavior at $2N = n$, when no regularization is applied. Here, we consider the (asymptotic) test MSE in Theorem 3 in the ridgeless $\lambda \rightarrow 0$ limit and focus on the situation where the test data $\tilde{\mathbf{X}}$ is sufficiently different from the training data \mathbf{X} (see more discussions on this point in Section 3.3 below). Then, the two-by-two matrix $\bar{\Omega}$ defined in (8) diverges to infinity at $2N = n$ as $\lambda \rightarrow 0$. (Indeed, the determinant $\det(\bar{\Omega}^{-1})$ scales as λ , per Lemma 5 in Appendix E.) As a consequence, we have $\bar{E}_{\text{test}} \rightarrow \infty$ as $N/n \rightarrow 1/2$, resulting in a sharp deterioration in the test performance around the interpolation threshold $2N = n$. It is also interesting to note that, while $\bar{\Omega}$ also appears in \bar{E}_{train} , we still obtain (asymptotically) zero training MSE at $2N = n$, despite the divergence of $\bar{\Omega}$ as $\lambda \rightarrow 0$, essentially due to the prefactor λ^2 in \bar{E}_{train} .

Figure 4 depicts the empirical and theoretical test MSEs with different λ . In particular, for $\lambda = 10^{-7}$ and $\lambda = 10^{-3}$, a double-descent-type behavior is observed, with a singularity at $2N = n$, while for larger values of λ ($\lambda = 1$ and 10), a smoother and monotonically decreasing test error curve is observed, as a function of N/n , in accordance with the observations in [36] on Gaussian data.

Remark 4 (Double descent as a consequence of phase transition). While the double descent phenomenon has received considerable attention recently, our analysis makes it clear that in this model (and presumably many others) it is a natural consequence of the phase transition between two qualitatively different phases of learning [32].

3.3 Impact of Training-test Similarity

We see that the (asymptotic) test error behaves entirely differently, depending on whether the test data $\hat{\mathbf{X}}$ is “close to” the training data \mathbf{X} or not. For $\hat{\mathbf{X}} = \mathbf{X}$, one has $\bar{E}_{\text{test}} = \bar{E}_{\text{train}}$ that decreases monotonically as N grows large; while for $\hat{\mathbf{X}}$ sufficiently different from \mathbf{X} (in the associated kernel space in the sense that $\mathbf{K}_\sigma(\mathbf{X}, \mathbf{X})$ is sufficiently different from $\mathbf{K}_\sigma(\hat{\mathbf{X}}, \mathbf{X})$ for $\sigma \in \{\cos, \sin\}$), \bar{E}_{test} diverges at $2N = n$ and establishes a double descent behavior. To have a more quantitative assessment of the impact of training-test similarity on the RFF model performance, we consider here the special case $\hat{\mathbf{y}} = \mathbf{y}$. Since in the ridgeless $\lambda \rightarrow 0$ limit, $\bar{\Omega}$ scales as λ^{-1} at $2N = n$ (Remark 3), one must then have $\Theta_\sigma \propto \lambda$ to “compensate” so that \bar{E}_{test} does not diverge at $2N = n$ as $\lambda \rightarrow 0$. A first example is the case where the test data is a small perturbation of the training data. In Figure 5, the test data are generated by adding Gaussian white noise of variance σ^2 to the training data, i.e.,

$$\hat{\mathbf{x}}_i = \mathbf{x}_i + \sigma \varepsilon_i \quad (12)$$

for independent $\varepsilon_i \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_p/p)$. In Figure 5, we observe that (i) below the threshold $\sigma^2 = \lambda$, the test error coincides with the training error and both are relatively small for $2N = n$; and (ii) as soon as $\sigma^2 > \lambda$, the test error diverges from the training error and grows large (but linearly in σ^2) as the noise level increases. Note also from the two rightmost plots of Figure 5 that the training-to-test “transition” at $\sigma^2 \simeq \lambda$ is *sharp* only for relatively small values of λ , as predicted by our theory.

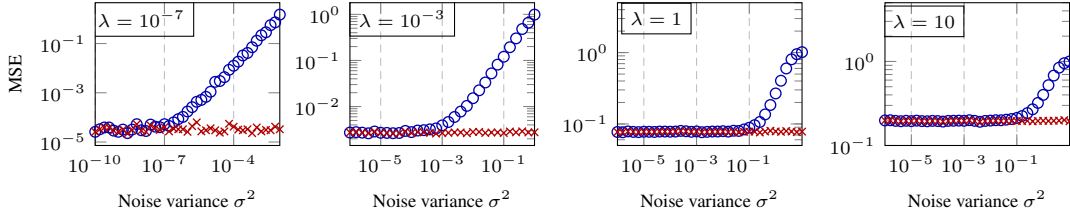


Figure 5: Empirical training (red crosses) and test (blue circles) errors of RFF ridge regression on MNIST data (class 3 versus 7), as a function of the noise level σ^2 , for $N = 512$, $p = 784$, $n = \hat{n} = 1024 = 2N$, $\lambda = 10^{-7}, 10^{-3}, 1, 10$. Results obtained by averaging over 30 runs.

4 Conclusion

We have established a precise description of the resolvent of RFF Gram matrices, and provided asymptotic training and test performance guarantees for RFF ridge regression, in the $n, p, N \rightarrow \infty$ limit. We have also discussed the under- and over-parameterized regimes, where the resolvent behaves dramatically differently. These observations involve only mild regularity assumptions on the data, yielding phase transition behavior and double descent test error curves for RFF regression that closely match experiments on real-world data. Extended to a (technically more involved) multi-layer setting in the more realistic large n, p, N regime as in [16], our analysis may shed new light on the theoretical understanding of modern deep neural nets, beyond the large- N alone neural tangent kernel limit.

Broader Impact

In this article, we provide theoretical assessment of the popular random Fourier features (RFFs), in the practical setting where n, p, N are all large and comparable. Asymptotic performance guarantees are provided for RFF ridge regression in this $n, p, N \rightarrow \infty$ limit, as an important positive impact of this work on the development of more reliable large-scale machine learning systems. The theoretical framework developed in this article presents fair and non-offensive societal consequence.

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A Warm-up Example: Sample Covariance and Marčenko-Pastur Equation

Consider the sample covariance matrix $\hat{\mathbf{C}} = \frac{1}{n}\mathbf{X}\mathbf{X}^\top$ from some data $\mathbf{X} \in \mathbb{R}^{p \times n}$ composed of n i.i.d. $\mathbf{x}_i \sim \mathcal{N}(\mathbf{0}, \mathbf{C})$ with positive definite $\mathbf{C} \in \mathbb{R}^{p \times p}$. In this zero-mean Gaussian setting, the sample covariance $\hat{\mathbf{C}}$, despite being the maximum likelihood estimator of the *population covariance* \mathbf{C} and providing *entry-wise* consistent estimate for it, is an extremely poor estimator of \mathbf{C} in a *spectral norm* sense, for n, p large. More precisely, $\|\hat{\mathbf{C}} - \mathbf{C}\| \not\rightarrow 0$ as $n, p \rightarrow \infty$ with $p/n \rightarrow c \in (0, \infty)$. Indeed, one has $\|\hat{\mathbf{C}} - \mathbf{C}\|/\|\mathbf{C}\| \approx 20\%$, even with $n = 100p$, in the simple $\mathbf{C} = \mathbf{I}_p$ setting. Figure 6 compares the eigenvalue histogram of $\hat{\mathbf{C}}$ with the population eigenvalue of \mathbf{C} , in the setting of $\mathbf{C} = \mathbf{I}_p$ and $n = 100p$. In the $\mathbf{C} = \mathbf{I}_p$ case, the limiting eigenvalue distribution of $\hat{\mathbf{C}}$ as $n, p \rightarrow \infty$ is known to be the popular Marčenko-Pastur law [31] given by

$$\mu(dx) = (1 - c^{-1}) \cdot \delta_0(x) + \frac{1}{2\pi cx} \sqrt{(x - (1 - \sqrt{c})^2)^+ ((1 + \sqrt{c})^2 - x)^+} dx \quad (13)$$

with $\delta_0(x)$ the Dirac mass at zero, $c = \lim p/n$ and $(x)^+ = \max(x, 0)$, so that the support of μ has length $(1 + \sqrt{c})^2 - (1 - \sqrt{c})^2 = 4\sqrt{c} = 0.4$ for $n = 100p$.

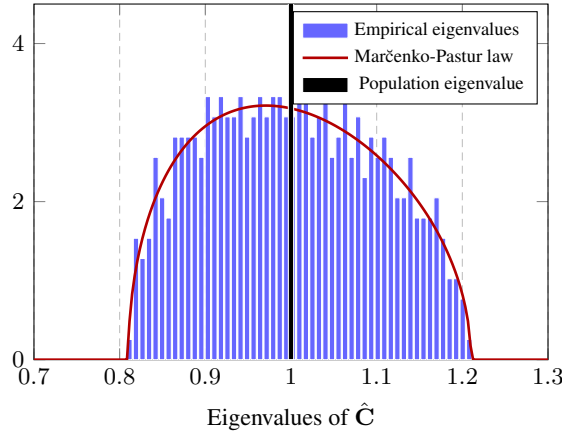


Figure 6: Eigenvalue histogram of $\hat{\mathbf{C}}$ versus the Marčenko-Pastur law, for $p = 512$ and $n = 100p$.

In the regression analysis (such as ridge regression) based on \mathbf{X} , of more immediate interest is the *resolvent* $\mathbf{Q}_{\hat{\mathbf{C}}}(\lambda) \equiv (\hat{\mathbf{C}} + \lambda \mathbf{I}_p)^{-1}$, $\lambda > 0$ of the sample covariance $\hat{\mathbf{C}}$, and more concretely, the bilinear forms of the type $\mathbf{a}^\top \mathbf{Q}_{\hat{\mathbf{C}}}(\lambda) \mathbf{b}$ for $\mathbf{a}, \mathbf{b} \in \mathbb{R}^p$. As a result of the spectral norm inconsistency $\|\hat{\mathbf{C}} - \mathbf{C}\| \not\rightarrow 0$ in the large n, p regime, it is unlikely that for most \mathbf{a}, \mathbf{b} , the convergence $\mathbf{a}^\top \mathbf{Q}_{\hat{\mathbf{C}}}(\lambda) \mathbf{b} - \mathbf{a}^\top (\mathbf{C} + \lambda \mathbf{I}_p)^{-1} \mathbf{b} \rightarrow 0$ would still hold.

While the *random* variable $\mathbf{a}^\top \mathbf{Q}_{\hat{\mathbf{C}}}(\lambda) \mathbf{b}$ is not getting close to $\mathbf{a}^\top (\mathbf{C} + \lambda \mathbf{I}_p)^{-1} \mathbf{b}$ as $n, p \rightarrow \infty$, it does exhibit a tractable asymptotically *deterministic* behavior, described by the Marčenko-Pastur equation [31] for $\mathbf{C} = \mathbf{I}_p$. Notably, for $\mathbf{a}, \mathbf{b} \in \mathbb{R}^p$ deterministic vectors of bounded Euclidean norms, we have, as $n, p \rightarrow \infty$ and $p/n \rightarrow c \in (0, \infty)$,

$$\mathbf{a}^\top \mathbf{Q}_{\hat{\mathbf{C}}}(\lambda) \mathbf{b} - m(\lambda) \cdot \mathbf{a}^\top \mathbf{b} \xrightarrow{a.s.} 0,$$

with $m(\lambda)$ the unique positive solution to the following Marčenko-Pastur equation [31]

$$c\lambda m^2(\lambda) + (1 + \lambda - c)m(\lambda) - 1 = 0. \quad (14)$$

In a sense, $\bar{\mathbf{Q}}(\lambda) \equiv m(\lambda)\mathbf{I}_p$ can be seen as a *deterministic equivalent* [19, 11] for the *random* $\mathbf{Q}_{\hat{\mathbf{C}}}(\lambda)$ that asymptotically characterizes the behavior of the latter, when bilinear forms are considered. In Figure 7 we compare the quadratic forms $\mathbf{a}^\top \mathbf{Q}_{\hat{\mathbf{C}}}(\lambda) \mathbf{a}$ as a function of λ , for $n = 10p$ and $n = 2p$. We observe that, in both cases, the RMT prediction in (14) provides a much closer match than the large- n alone asymptotic given by $\mathbf{a}^\top (\mathbf{C} + \lambda \mathbf{I}_p)^{-1} \mathbf{a}$. This, together with Figure 1 on RFF ridge regression model, conveys a strong practical motivation of this work.

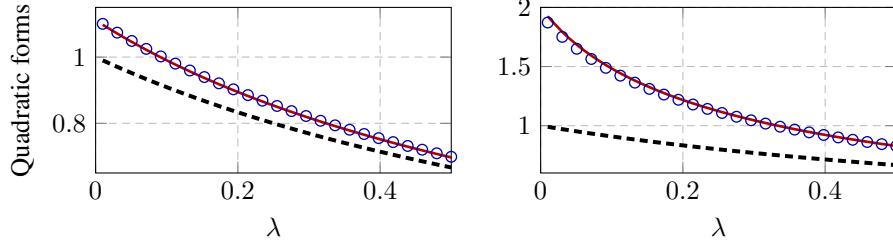


Figure 7: Quadratic forms $\mathbf{a}^\top \mathbf{Q}_{\hat{\mathbf{C}}}(\lambda) \mathbf{a}$ as a function of λ , for $p = 512$, $n = 10p$ (left) and $n = 2p$ (right). Empirical results displayed in blue circles; population predictions $\mathbf{a}^\top (\mathbf{C} + \lambda \mathbf{I}_p)^{-1} \mathbf{a}$ (assuming $n \rightarrow \infty$ alone with p fixed) in black dashed lines; and RMT prediction from (14) in red solid lines. Results obtained by averaging over 50 runs.

B Proof of Theorem 1

Our objective is to prove, under Assumption 1, the asymptotic equivalence between the expectation (with respect to \mathbf{W} , omitted from now on) $\mathbb{E}[\mathbf{Q}]$ and

$$\bar{\mathbf{Q}} \equiv \left(\frac{N}{n} \left(\frac{\mathbf{K}_{\cos}}{1 + \delta_{\cos}} + \frac{\mathbf{K}_{\sin}}{1 + \delta_{\sin}} \right) + \lambda \mathbf{I}_n \right)^{-1}$$

for $\mathbf{K}_{\cos} \equiv \mathbf{K}_{\cos}(\mathbf{X}, \mathbf{X})$, $\mathbf{K}_{\sin} \equiv \mathbf{K}_{\sin}(\mathbf{X}, \mathbf{X}) \in \mathbb{R}^{n \times n}$ defined in (6), with $(\delta_{\cos}, \delta_{\sin})$ the unique positive solution to

$$\delta_{\cos} = \frac{1}{n} \text{tr}(\mathbf{K}_{\cos} \bar{\mathbf{Q}}), \quad \delta_{\sin} = \frac{1}{n} \text{tr}(\mathbf{K}_{\sin} \bar{\mathbf{Q}}).$$

The existence and uniqueness of the above fixed-point equation is standard in random matrix literature and can be reached for instance with the standard interference function framework [54].

The asymptotic equivalence should be announced in the sense that $\|\mathbb{E}[\mathbf{Q}] - \bar{\mathbf{Q}}\| \rightarrow 0$ as $n, p, N \rightarrow \infty$ at the same pace. We shall proceed by introducing an intermediary resolvent $\tilde{\mathbf{Q}}$ (see definition in (16)) and show subsequently that

$$\|\mathbb{E}[\mathbf{Q}] - \tilde{\mathbf{Q}}\| \rightarrow 0, \quad \|\tilde{\mathbf{Q}} - \bar{\mathbf{Q}}\| \rightarrow 0.$$

In the sequel, we use $o(1)$ and $o_{\|\cdot\|}(1)$ for scalars or matrices of (almost surely if being random) vanishing absolute values or operator norms as $n, p \rightarrow \infty$.

We start by introducing the following lemma.

Lemma 1 (Expectation of $\sigma_1(\mathbf{x}_i^\top \mathbf{w}) \sigma_2(\mathbf{w}^\top \mathbf{x}_j)$). *For $\mathbf{w} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_p)$ and $\mathbf{x}_i, \mathbf{x}_j \in \mathbb{R}^p$ we have (per Definition in (6))*

$$\begin{aligned} \mathbb{E}_{\mathbf{w}}[\cos(\mathbf{x}_i^\top \mathbf{w}) \cos(\mathbf{w}^\top \mathbf{x}_j)] &= e^{-\frac{1}{2}(\|\mathbf{x}_i\|^2 + \|\mathbf{x}_j\|^2)} \cosh(\mathbf{x}_i^\top \mathbf{x}_j) \equiv [\mathbf{K}_{\cos}(\mathbf{X}, \mathbf{X})]_{ij} \equiv [\mathbf{K}_{\cos}]_{ij} \\ \mathbb{E}_{\mathbf{w}}[\sin(\mathbf{x}_i^\top \mathbf{w}) \sin(\mathbf{w}^\top \mathbf{x}_j)] &= e^{-\frac{1}{2}(\|\mathbf{x}_i\|^2 + \|\mathbf{x}_j\|^2)} \sinh(\mathbf{x}_i^\top \mathbf{x}_j) \equiv [\mathbf{K}_{\sin}(\mathbf{X}, \mathbf{X})]_{ij} \equiv [\mathbf{K}_{\sin}]_{ij} \\ \mathbb{E}_{\mathbf{w}}[\cos(\mathbf{x}_i^\top \mathbf{w}) \sin(\mathbf{w}^\top \mathbf{x}_j)] &= 0. \end{aligned}$$

Proof of Lemma 1. The proof follows the integration tricks in [52, 30]. Note in particular that the third equality holds in the case of (\cos, \sin) nonlinearity but in general not true for arbitrary Lipschitz (σ_1, σ_2) . \square

Let us focus on the resolvent $\mathbf{Q} \equiv \left(\frac{1}{n} \Sigma_{\mathbf{X}}^\top \Sigma_{\mathbf{X}} + \lambda \mathbf{I}_n \right)^{-1}$ of $\frac{1}{n} \Sigma_{\mathbf{X}}^\top \Sigma_{\mathbf{X}} \in \mathbb{R}^{n \times n}$, for random Fourier feature matrix $\Sigma_{\mathbf{X}} \equiv \begin{bmatrix} \cos(\mathbf{W}\mathbf{X}) \\ \sin(\mathbf{W}\mathbf{X}) \end{bmatrix}$ that can be rewritten as

$$\Sigma_{\mathbf{X}}^\top = [\cos(\mathbf{X}^\top \mathbf{w}_1), \dots, \cos(\mathbf{X}^\top \mathbf{w}_N), \sin(\mathbf{X}^\top \mathbf{w}_1), \dots, \sin(\mathbf{X}^\top \mathbf{w}_N)] \quad (15)$$

for \mathbf{w}_i the i -th row of $\mathbf{W} \in \mathbb{R}^{N \times p}$ with $\mathbf{w}_i \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_p)$, $i = 1, \dots, N$, that is at the core of our analysis. Note from (15) that we have

$$\Sigma_{\mathbf{X}}^T \Sigma_{\mathbf{X}} = \sum_{i=1}^N (\cos(\mathbf{X}^T \mathbf{w}_i) \cos(\mathbf{w}_i^T \mathbf{X}) + \sin(\mathbf{X}^T \mathbf{w}_i) \sin(\mathbf{w}_i^T \mathbf{X})) = \sum_{i=1}^N \mathbf{U}_i \mathbf{U}_i^T$$

with $\mathbf{U}_i = [\cos(\mathbf{X}^T \mathbf{w}_i) \quad \sin(\mathbf{X}^T \mathbf{w}_i)] \in \mathbb{R}^{n \times 2}$.

Letting

$$\tilde{\mathbf{Q}} \equiv \left(\frac{N}{n} \frac{\mathbf{K}_{\cos}}{1 + \alpha_{\cos}} + \frac{N}{n} \frac{\mathbf{K}_{\sin}}{1 + \alpha_{\sin}} + \lambda \mathbf{I}_n \right)^{-1} \quad (16)$$

with

$$\alpha_{\cos} = \frac{1}{n} \text{tr}(\mathbf{K}_{\cos} \mathbb{E}[\mathbf{Q}]), \quad \alpha_{\sin} = \frac{1}{n} \text{tr}(\mathbf{K}_{\sin} \mathbb{E}[\mathbf{Q}]) \quad (17)$$

we have, with the resolvent identity $(\mathbf{A}^{-1} - \mathbf{B}^{-1} = \mathbf{A}^{-1}(\mathbf{B} - \mathbf{A})\mathbf{B}^{-1})$ for invertible \mathbf{A}, \mathbf{B} that

$$\begin{aligned} \mathbb{E}[\mathbf{Q}] - \tilde{\mathbf{Q}} &= \mathbb{E} \left[\mathbf{Q} \left(\frac{N}{n} \frac{\mathbf{K}_{\cos}}{1 + \alpha_{\cos}} + \frac{N}{n} \frac{\mathbf{K}_{\sin}}{1 + \alpha_{\sin}} - \frac{1}{n} \Sigma_{\mathbf{X}}^T \Sigma_{\mathbf{X}} \right) \right] \tilde{\mathbf{Q}} \\ &= \mathbb{E}[\mathbf{Q}] \frac{N}{n} \left(\frac{\mathbf{K}_{\cos}}{1 + \alpha_{\cos}} + \frac{\mathbf{K}_{\sin}}{1 + \alpha_{\sin}} \right) \tilde{\mathbf{Q}} - \frac{N}{n} \frac{1}{N} \sum_{i=1}^N \mathbb{E}[\mathbf{Q} \mathbf{U}_i \mathbf{U}_i^T] \tilde{\mathbf{Q}} \\ &= \mathbb{E}[\mathbf{Q}] \frac{N}{n} \left(\frac{\mathbf{K}_{\cos}}{1 + \alpha_{\cos}} + \frac{\mathbf{K}_{\sin}}{1 + \alpha_{\sin}} \right) \tilde{\mathbf{Q}} - \frac{N}{n} \frac{1}{N} \sum_{i=1}^N \mathbb{E}[\mathbf{Q}_{-i} \mathbf{U}_i (\mathbf{I}_2 + \frac{1}{n} \mathbf{U}_i^T \mathbf{Q}_{-i} \mathbf{U}_i)^{-1} \mathbf{U}_i^T] \tilde{\mathbf{Q}}, \end{aligned}$$

for $\mathbf{Q}_{-i} \equiv \left(\frac{1}{n} \Sigma_{\mathbf{X}}^T \Sigma_{\mathbf{X}} - \frac{1}{n} \mathbf{U}_i \mathbf{U}_i^T + \lambda \mathbf{I}_n \right)^{-1}$ that is **independent** of \mathbf{U}_i (and thus \mathbf{w}_i), where we applied the following Woodbury identity.

Lemma 2 (Woodbury). *For $\mathbf{A}, \mathbf{A} + \mathbf{U}\mathbf{U}^T \in \mathbb{R}^{p \times p}$ both invertible and $\mathbf{U} \in \mathbb{R}^{p \times n}$, we have*

$$(\mathbf{A} + \mathbf{U}\mathbf{U}^T)^{-1} = \mathbf{A}^{-1} - \mathbf{A}^{-1} \mathbf{U} (\mathbf{I}_n + \mathbf{U}^T \mathbf{A}^{-1} \mathbf{U})^{-1} \mathbf{U}^T \mathbf{A}^{-1}$$

so that in particular $(\mathbf{A} + \mathbf{U}\mathbf{U}^T)^{-1} \mathbf{U} = \mathbf{A}^{-1} \mathbf{U} (\mathbf{I}_n + \mathbf{U}^T \mathbf{A}^{-1} \mathbf{U})^{-1}$.

Consider now the two-by-two matrix

$$\mathbf{I}_2 + \frac{1}{n} \mathbf{U}_i^T \mathbf{Q}_{-i} \mathbf{U}_i = \begin{bmatrix} 1 + \frac{1}{n} \cos(\mathbf{w}_i^T \mathbf{X}) \mathbf{Q}_{-i} \cos(\mathbf{X} \mathbf{w}_i) & \frac{1}{n} \cos(\mathbf{w}_i^T \mathbf{X}) \mathbf{Q}_{-i} \sin(\mathbf{X} \mathbf{w}_i) \\ \frac{1}{n} \sin(\mathbf{w}_i^T \mathbf{X}) \mathbf{Q}_{-i} \cos(\mathbf{X} \mathbf{w}_i) & 1 + \frac{1}{n} \sin(\mathbf{w}_i^T \mathbf{X}) \mathbf{Q}_{-i} \sin(\mathbf{X} \mathbf{w}_i) \end{bmatrix}$$

which, according to the following lemma, is expected to be close to $\begin{bmatrix} 1 + \alpha_{\cos} & 0 \\ 0 & 1 + \alpha_{\sin} \end{bmatrix}$ as defined in (17).

Lemma 3 (Concentration of quadratic forms). *Under Assumption 1, for $\sigma_1(\cdot), \sigma_2(\cdot)$ two real 1-Lipschitz functions, $\mathbf{w} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_p)$ and $\mathbf{A} \in \mathbb{R}^{n \times n}$ independent of \mathbf{w} with $\|\mathbf{A}\| \leq 1$, then*

$$\mathbb{P} \left(\left| \frac{1}{n} \sigma_a(\mathbf{w}^T \mathbf{X}) \mathbf{A} \sigma_b(\mathbf{X}^T \mathbf{w}) - \frac{1}{n} \text{tr}(\mathbf{A} \mathbb{E}_{\mathbf{w}}[\sigma_b(\mathbf{X}^T \mathbf{w}) \sigma_a(\mathbf{w}^T \mathbf{X})]) \right| > t \right) \leq C e^{-c n \min(t, t^2)}$$

for $a, b \in \{1, 2\}$ and some universal constants $C, c > 0$.

Proof of Lemma 3. Lemma 3 can be easily extended from [30, Lemma 1], where one observes the proof actually holds when different types of nonlinear Lipschitz functions $\sigma_1(\cdot), \sigma_2(\cdot)$ (and in particular \cos and \sin) are considered. \square

For $\mathbf{W}_{-i} \in \mathbb{R}^{(N-1) \times p}$ the random matrix $\mathbf{W} \in \mathbb{R}^{N \times p}$ with its i -th row \mathbf{w}_i removed, Lemma 3, together with the Lipschitz nature of the map $\mathbf{W}_{-i} \mapsto \frac{1}{n} \sigma_a(\mathbf{w}_i^T \mathbf{X}) \mathbf{Q}_{-i} \sigma_b(\mathbf{X} \mathbf{w}_i)$ for $\mathbf{Q}_{-i} = (\frac{1}{n} \cos(\mathbf{W}_{-i} \mathbf{X})^T \cos(\mathbf{W}_{-i} \mathbf{X}) + \frac{1}{n} \sin(\mathbf{W}_{-i} \mathbf{X})^T \sin(\mathbf{W}_{-i} \mathbf{X}) + \lambda \mathbf{I}_n)^{-1}$, leads to the following concentration result

$$\mathbb{P} \left(\left| \frac{1}{n} \sigma_a(\mathbf{w}_i^T \mathbf{X}) \mathbf{Q}_{-i} \sigma_b(\mathbf{X} \mathbf{w}_i) - \frac{1}{n} \mathbb{E}[\mathbf{Q}_{-i}] \mathbb{E}[\sigma_b(\mathbf{X} \mathbf{w}_i) \sigma_a(\mathbf{w}_i^T \mathbf{X})] \right| > t \right) \leq C' e^{-c' n \max(t^2, t)} \quad (18)$$

the proof of which follows the same line of argument of [30, Lemma 4] and is omitted here.

As a consequence, we continue to write, with again the resolvent identity, that

$$\begin{aligned}
& (\mathbf{I}_2 + \frac{1}{n} \mathbf{U}_i^\top \mathbf{Q}_{-i} \mathbf{U}_i)^{-1} - \begin{bmatrix} 1 + \alpha_{\cos} & 0 \\ 0 & 1 + \alpha_{\sin} \end{bmatrix}^{-1} \\
&= \begin{bmatrix} 1 + \frac{1}{n} \cos(\mathbf{w}_i^\top \mathbf{X}) \mathbf{Q}_{-i} \cos(\mathbf{X} \mathbf{w}_i) & \frac{1}{n} \cos(\mathbf{w}_i^\top \mathbf{X}) \mathbf{Q}_{-i} \sin(\mathbf{X} \mathbf{w}_i) \\ \frac{1}{n} \sin(\mathbf{w}_i^\top \mathbf{X}) \mathbf{Q}_{-i} \cos(\mathbf{X} \mathbf{w}_i) & 1 + \frac{1}{n} \sin(\mathbf{w}_i^\top \mathbf{X}) \mathbf{Q}_{-i} \sin(\mathbf{X} \mathbf{w}_i) \end{bmatrix}^{-1} - \begin{bmatrix} 1 + \alpha_{\cos} & 0 \\ 0 & 1 + \alpha_{\sin} \end{bmatrix}^{-1} \\
&= (\mathbf{I}_2 + \frac{1}{n} \mathbf{U}_i^\top \mathbf{Q}_{-i} \mathbf{U}_i)^{-1} \begin{bmatrix} \alpha_{\cos} - \frac{1}{n} \cos(\mathbf{w}_i^\top \mathbf{X}) \mathbf{Q}_{-i} \cos(\mathbf{X} \mathbf{w}_i) & -\frac{1}{n} \cos(\mathbf{w}_i^\top \mathbf{X}) \mathbf{Q}_{-i} \sin(\mathbf{X} \mathbf{w}_i) \\ -\frac{1}{n} \sin(\mathbf{w}_i^\top \mathbf{X}) \mathbf{Q}_{-i} \cos(\mathbf{X} \mathbf{w}_i) & \alpha_{\sin} - \frac{1}{n} \sin(\mathbf{w}_i^\top \mathbf{X}) \mathbf{Q}_{-i} \sin(\mathbf{X} \mathbf{w}_i) \end{bmatrix} \\
&\times \begin{bmatrix} \frac{1}{1+\alpha_{\cos}} & 0 \\ 0 & \frac{1}{1+\alpha_{\sin}} \end{bmatrix} \equiv (\mathbf{I}_2 + \frac{1}{n} \mathbf{U}_i^\top \mathbf{Q}_{-i} \mathbf{U}_i)^{-1} \mathbf{D}_i \begin{bmatrix} \frac{1}{1+\alpha_{\cos}} & 0 \\ 0 & \frac{1}{1+\alpha_{\sin}} \end{bmatrix},
\end{aligned}$$

where we note from (18) (and $\|\mathbf{Q}_{-i}\| \leq \lambda^{-1}$) that the matrix $\mathbb{E}[\mathbf{D}_i] = o_{\|\cdot\|}(1)$ (in fact of spectral norm of order $O(n^{-\frac{1}{2}})$). So that

$$\begin{aligned}
\mathbb{E}[\mathbf{Q}] - \tilde{\mathbf{Q}} &= \mathbb{E}[\mathbf{Q}] \frac{N}{n} \left(\frac{\mathbf{K}_{\cos}}{1 + \alpha_{\cos}} + \frac{\mathbf{K}_{\sin}}{1 + \alpha_{\sin}} \right) \tilde{\mathbf{Q}} - \frac{N}{n} \frac{1}{N} \sum_{i=1}^N \mathbb{E}[\mathbf{Q}_{-i} \mathbf{U}_i (\mathbf{I}_2 + \frac{1}{n} \mathbf{U}_i^\top \mathbf{Q}_{-i} \mathbf{U}_i)^{-1} \mathbf{U}_i^\top] \tilde{\mathbf{Q}} \\
&= \mathbb{E}[\mathbf{Q}] \frac{N}{n} \left(\frac{\mathbf{K}_{\cos}}{1 + \alpha_{\cos}} + \frac{\mathbf{K}_{\sin}}{1 + \alpha_{\sin}} \right) \tilde{\mathbf{Q}} - \frac{N}{n} \frac{1}{N} \sum_{i=1}^N \mathbb{E}[\mathbf{Q}_{-i} \mathbf{U}_i \begin{bmatrix} \frac{1}{1+\alpha_{\cos}} & 0 \\ 0 & \frac{1}{1+\alpha_{\sin}} \end{bmatrix} \mathbf{U}_i^\top] \tilde{\mathbf{Q}} \\
&\quad - \frac{N}{n} \frac{1}{N} \sum_{i=1}^N \mathbb{E}[\mathbf{Q}_{-i} \mathbf{U}_i (\mathbf{I}_2 + \frac{1}{n} \mathbf{U}_i^\top \mathbf{Q}_{-i} \mathbf{U}_i)^{-1} \mathbf{D}_i \begin{bmatrix} \frac{1}{1+\alpha_{\cos}} & 0 \\ 0 & \frac{1}{1+\alpha_{\sin}} \end{bmatrix} \mathbf{U}_i^\top] \tilde{\mathbf{Q}} \\
&= (\mathbb{E}[\mathbf{Q}] - \frac{1}{N} \sum_{i=1}^N \mathbb{E}[\mathbf{Q}_{-i}]) \frac{N}{n} \left(\frac{\mathbf{K}_{\cos}}{1 + \alpha_{\cos}} + \frac{\mathbf{K}_{\sin}}{1 + \alpha_{\sin}} \right) \tilde{\mathbf{Q}} - \frac{N}{n} \frac{1}{N} \sum_{i=1}^N \mathbb{E}[\mathbf{Q} \mathbf{U}_i \mathbf{D}_i \begin{bmatrix} \frac{1}{1+\alpha_{\cos}} & 0 \\ 0 & \frac{1}{1+\alpha_{\sin}} \end{bmatrix} \mathbf{U}_i^\top] \tilde{\mathbf{Q}},
\end{aligned}$$

where we used $\mathbb{E}_{\mathbf{w}_i}[\mathbf{U}_i \mathbf{U}_i^\top] = \mathbf{K}_{\cos} + \mathbf{K}_{\sin}$ by Lemma 1 and then Lemma 2 in reverse for the last equality. Moreover, since

$$\mathbb{E}[\mathbf{Q}] - \frac{1}{N} \sum_{i=1}^N \mathbb{E}[\mathbf{Q}_{-i}] = \frac{1}{N} \sum_{i=1}^N \mathbb{E}[\mathbf{Q} - \mathbf{Q}_{-i}] = -\frac{1}{n} \frac{1}{N} \sum_{i=1}^N \mathbb{E}[\mathbf{Q} \mathbf{U}_i (\mathbf{I}_2 + \frac{1}{n} \mathbf{U}_i^\top \mathbf{Q}_{-i} \mathbf{U}_i)^{-1} \mathbf{U}_i^\top \mathbf{Q}]$$

so that with the fact $\frac{1}{\sqrt{n}} \|\mathbf{Q} \Sigma_{\mathbf{X}}^\top\| \leq \|\sqrt{\mathbf{Q}_n^\top \Sigma_{\mathbf{X}}^\top \Sigma_{\mathbf{X}} \mathbf{Q}}\| \leq \lambda^{-\frac{1}{2}}$ we have for the first term

$$\|\mathbb{E}[\mathbf{Q}] - \frac{1}{N} \sum_{i=1}^N \mathbb{E}[\mathbf{Q}_{-i}]\| = O(n^{-1}).$$

It thus remains to treat the second term, which, with the relation $\mathbf{A} \mathbf{B}^\top + \mathbf{B} \mathbf{A}^\top \preceq \mathbf{A} \mathbf{A}^\top + \mathbf{B} \mathbf{B}^\top$ (in the sense of symmetric matrices), and the same line of arguments as above, can be shown to have vanishing spectral norm (of order $O(n^{-\frac{1}{2}})$) as $n, p, N \rightarrow \infty$.

We thus have $\|\mathbb{E}[\mathbf{Q}] - \tilde{\mathbf{Q}}\| = O(n^{-\frac{1}{2}})$, which concludes the first part of the proof of Theorem 1.

We shall show next that $\|\tilde{\mathbf{Q}} - \bar{\mathbf{Q}}\| \rightarrow 0$ as $n, p, N \rightarrow \infty$. First note from previous derivation that $\alpha_\sigma - \frac{1}{n} \text{tr} \mathbf{K}_\sigma \tilde{\mathbf{Q}} = O(n^{-\frac{1}{2}})$ for $\sigma = \cos, \sin$. To compare $\tilde{\mathbf{Q}}$ and $\bar{\mathbf{Q}}$, it follows again from the resolvent identity that

$$\tilde{\mathbf{Q}} - \bar{\mathbf{Q}} = \tilde{\mathbf{Q}} \left(\frac{N}{n} \frac{\mathbf{K}_{\cos}(\alpha_{\cos} - \delta_{\cos})}{(1 + \delta_{\cos})(1 + \alpha_{\cos})} + \frac{N}{n} \frac{\mathbf{K}_{\sin}(\alpha_{\sin} - \delta_{\sin})}{(1 + \delta_{\sin})(1 + \alpha_{\sin})} \right) \bar{\mathbf{Q}}$$

so that the control of $\|\tilde{\mathbf{Q}} - \bar{\mathbf{Q}}\|$ boils down to the control of $\max\{|\alpha_{\cos} - \delta_{\cos}|, |\alpha_{\sin} - \delta_{\sin}|\}$. To this end, it suffices to write

$$\alpha_{\cos} - \delta_{\cos} = \frac{1}{n} \text{tr} \mathbf{K}_{\cos} (\mathbb{E}[\mathbf{Q}] - \bar{\mathbf{Q}}) = \frac{1}{n} \text{tr} \mathbf{K}_{\cos} (\tilde{\mathbf{Q}} - \bar{\mathbf{Q}}) + O(n^{-\frac{1}{2}})$$

where we used $|\text{tr}(\mathbf{AB})| \leq \|\mathbf{A}\| \text{tr}(\mathbf{B})$ for nonnegative definite \mathbf{B} , together with the fact that $\frac{1}{n} \text{tr} \mathbf{K}_\sigma$ is (uniformly) bounded under Assumption 1, for $\sigma = \cos, \sin$.

As a consequence, we have

$$|\alpha_{\cos} - \delta_{\cos}| \leq |\alpha_{\cos} - \delta_{\cos}| \frac{N \frac{1}{n} \text{tr}(\mathbf{K}_{\cos} \tilde{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}})}{n (1 + \delta_{\cos})(1 + \alpha_{\cos})} + o(1).$$

It thus remains to show

$$\frac{N \frac{1}{n} \text{tr}(\mathbf{K}_{\cos} \tilde{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}})}{n (1 + \delta_{\cos})(1 + \alpha_{\cos})} < 1$$

or alternatively, by the Cauchy–Schwarz inequality, to show

$$\frac{N \frac{1}{n} \text{tr}(\mathbf{K}_{\cos} \tilde{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}})}{n (1 + \delta_{\cos})(1 + \alpha_{\cos})} \leq \sqrt{\frac{N \frac{1}{n} \text{tr}(\mathbf{K}_{\cos} \tilde{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}})}{n (1 + \delta_{\cos})^2} \cdot \frac{N \frac{1}{n} \text{tr}(\mathbf{K}_{\cos} \tilde{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}})}{n (1 + \alpha_{\cos})^2}} < 1.$$

To treat the first right-hand side term (the second can be done similarly), it unfolds from $|\text{tr}(\mathbf{AB})| \leq \|\mathbf{A}\| \cdot \text{tr}(\mathbf{B})$ for nonnegative definite \mathbf{B} that

$$\frac{N \frac{1}{n} \text{tr}(\mathbf{K}_{\cos} \tilde{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}})}{n (1 + \delta_{\cos})^2} \leq \left\| \frac{N \mathbf{K}_{\cos} \bar{\mathbf{Q}}}{n (1 + \delta_{\cos})} \right\| \frac{\frac{1}{n} \text{tr}(\mathbf{K}_{\cos} \tilde{\mathbf{Q}})}{1 + \delta_{\cos}} = \left\| \frac{N \mathbf{K}_{\cos} \bar{\mathbf{Q}}}{n (1 + \delta_{\cos})} \right\| \frac{\gamma_{\cos}}{1 + \delta_{\cos}} \leq \frac{\gamma_{\cos}}{1 + \delta_{\cos}} < 1$$

where we used the fact that $\frac{N \mathbf{K}_{\cos} \bar{\mathbf{Q}}}{n (1 + \delta_{\cos})} = \mathbf{I}_n - \frac{N \mathbf{K}_{\sin} \bar{\mathbf{Q}}}{n (1 + \delta_{\sin})} - \lambda \bar{\mathbf{Q}}$. This concludes the proof of Theorem 1. \blacksquare

C Proof of Theorem 2

To prove Theorem 2, it indeed suffices to prove the following lemma.

Lemma 4 (Asymptotic behavior of $\mathbb{E}[\mathbf{QAQ}]$). *Under Assumption 1, for \mathbf{Q} defined in (5) and symmetric nonnegative definite $\mathbf{A} \in \mathbb{R}^{n \times n}$ of bounded spectral norm, we have*

$$\left\| \mathbb{E}[\mathbf{QAQ}] - \left(\bar{\mathbf{Q}} \mathbf{A} \bar{\mathbf{Q}} + \frac{N}{n} \begin{bmatrix} \frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \mathbf{A} \mathbf{Q} \mathbf{K}_{\cos})}{(1 + \delta_{\cos})^2} & \frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \mathbf{A} \mathbf{Q} \mathbf{K}_{\sin})}{(1 + \delta_{\sin})^2} \end{bmatrix} \Omega \begin{bmatrix} \bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}} \\ \bar{\mathbf{Q}} \mathbf{K}_{\sin} \bar{\mathbf{Q}} \end{bmatrix} \right) \right\| \rightarrow 0$$

almost surely as $n \rightarrow \infty$, with $\Omega^{-1} \equiv \mathbf{I}_2 - \frac{N}{n} \begin{bmatrix} \frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}} \mathbf{K}_{\cos})}{(1 + \delta_{\cos})^2} & \frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}} \mathbf{K}_{\sin})}{(1 + \delta_{\sin})^2} \\ \frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}} \mathbf{K}_{\sin})}{(1 + \delta_{\cos})^2} & \frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\sin} \bar{\mathbf{Q}} \mathbf{K}_{\sin})}{(1 + \delta_{\sin})^2} \end{bmatrix}$. In particular, we have

$$\left\| \mathbb{E} \begin{bmatrix} \mathbf{Q} \mathbf{K}_{\cos} \mathbf{Q} \\ \mathbf{Q} \mathbf{K}_{\sin} \mathbf{Q} \end{bmatrix} - \Omega \begin{bmatrix} \bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}} \\ \bar{\mathbf{Q}} \mathbf{K}_{\sin} \bar{\mathbf{Q}} \end{bmatrix} \right\| \rightarrow 0.$$

Proof of Lemma 4. The proof of Lemma 4 essentially follows the same line of arguments as that of Theorem 1. Writing

$$\begin{aligned} \mathbb{E}[\mathbf{QAQ}] &= \mathbb{E}[\bar{\mathbf{Q}} \mathbf{A} \bar{\mathbf{Q}}] + \mathbb{E}[(\mathbf{Q} - \bar{\mathbf{Q}}) \mathbf{A} \mathbf{Q}] \\ &\simeq \bar{\mathbf{Q}} \mathbf{A} \bar{\mathbf{Q}} + \mathbb{E} \left[\mathbf{Q} \left(\frac{N}{n} \frac{\mathbf{K}_{\cos}}{1 + \delta_{\cos}} + \frac{N}{n} \frac{\mathbf{K}_{\sin}}{1 + \delta_{\sin}} - \frac{1}{n} \Sigma_{\mathbf{X}}^{\top} \Sigma_{\mathbf{X}} \right) \bar{\mathbf{Q}} \mathbf{A} \mathbf{Q} \right] \\ &= \bar{\mathbf{Q}} \mathbf{A} \bar{\mathbf{Q}} + \frac{N}{n} \mathbb{E}[\mathbf{Q} \Phi \bar{\mathbf{Q}} \mathbf{A} \mathbf{Q}] - \frac{1}{n} \sum_{i=1}^N \mathbb{E}[\mathbf{Q} \mathbf{U}_i \mathbf{U}_i^{\top} \bar{\mathbf{Q}} \mathbf{A} \mathbf{Q}] \end{aligned}$$

where we note \simeq by ignoring matrices with vanishing spectral norm (i.e., $o_{\|\cdot\|}(1)$) in the $n, p, N \rightarrow \infty$ limit and recall the shortcut $\Phi \equiv \frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} + \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}}$. Developing rightmost term with Lemma 2 as

$$\begin{aligned} \mathbb{E}[\mathbf{Q}\mathbf{U}_i\mathbf{U}_i^\top\bar{\mathbf{Q}}\mathbf{A}\mathbf{Q}] &= \mathbb{E}\left[\mathbf{Q}_{-i}\mathbf{U}_i(\mathbf{I}_2 + \frac{1}{n}\mathbf{U}_i^\top\mathbf{Q}_{-i}\mathbf{U}_i)^{-1}\mathbf{U}_i^\top\bar{\mathbf{Q}}\mathbf{A}\mathbf{Q}\right] \\ &= \mathbb{E}\left[\mathbf{Q}_{-i}\mathbf{U}_i(\mathbf{I}_2 + \frac{1}{n}\mathbf{U}_i^\top\mathbf{Q}_{-i}\mathbf{U}_i)^{-1}\mathbf{U}_i^\top\bar{\mathbf{Q}}\mathbf{A}\mathbf{Q}_{-i}\right] \\ &\quad - \frac{1}{n}\mathbb{E}\left[\mathbf{Q}_{-i}\mathbf{U}_i(\mathbf{I}_2 + \frac{1}{n}\mathbf{U}_i^\top\mathbf{Q}_{-i}\mathbf{U}_i)^{-1}\mathbf{U}_i^\top\bar{\mathbf{Q}}\mathbf{A}\mathbf{Q}_{-i}\mathbf{U}_i(\mathbf{I}_2 + \frac{1}{n}\mathbf{U}_i^\top\mathbf{Q}_{-i}\mathbf{U}_i)^{-1}\mathbf{U}_i^\top\mathbf{Q}_{-i}\right] \\ &\simeq \mathbb{E}[\mathbf{Q}_{-i}\Phi\bar{\mathbf{Q}}\mathbf{A}\mathbf{Q}_{-i}] \\ &= \mathbb{E}\left[\mathbf{Q}_{-i}\mathbf{U}_i\begin{bmatrix} \frac{1}{1+\delta_{\cos}} & 0 \\ 0 & \frac{1}{1+\delta_{\sin}} \end{bmatrix}\begin{bmatrix} \frac{1}{n}\text{tr}(\bar{\mathbf{Q}}\mathbf{A}\bar{\mathbf{Q}}\mathbf{K}_{\cos}) & 0 \\ 0 & \frac{1}{n}\text{tr}(\bar{\mathbf{Q}}\mathbf{A}\bar{\mathbf{Q}}\mathbf{K}_{\sin}) \end{bmatrix}\begin{bmatrix} \frac{1}{1+\delta_{\cos}} & 0 \\ 0 & \frac{1}{1+\delta_{\sin}} \end{bmatrix}\mathbf{U}_i^\top\mathbf{Q}_{-i}\right] \end{aligned}$$

so that

$$\begin{aligned} \mathbb{E}[\mathbf{Q}\mathbf{A}\mathbf{Q}] &\simeq \bar{\mathbf{Q}}\mathbf{A}\bar{\mathbf{Q}} + \frac{N}{n}\mathbb{E}\left[\mathbf{Q}\left(\frac{\frac{1}{n}\text{tr}(\bar{\mathbf{Q}}\mathbf{A}\bar{\mathbf{Q}}\mathbf{K}_{\cos})}{(1+\delta_{\cos})^2}\mathbf{K}_{\cos} + \frac{\frac{1}{n}\text{tr}(\bar{\mathbf{Q}}\mathbf{A}\bar{\mathbf{Q}}\mathbf{K}_{\sin})}{(1+\delta_{\sin})^2}\mathbf{K}_{\sin}\right)\mathbf{Q}\right] \\ &= \bar{\mathbf{Q}}\mathbf{A}\bar{\mathbf{Q}} + \frac{N}{n}\begin{bmatrix} \frac{\frac{1}{n}\text{tr}(\bar{\mathbf{Q}}\mathbf{A}\bar{\mathbf{Q}}\mathbf{K}_{\cos})}{(1+\delta_{\cos})^2} & \frac{\frac{1}{n}\text{tr}(\bar{\mathbf{Q}}\mathbf{A}\bar{\mathbf{Q}}\mathbf{K}_{\sin})}{(1+\delta_{\sin})^2} \end{bmatrix}\mathbb{E}\begin{bmatrix} \mathbf{Q}\mathbf{K}_{\cos}\mathbf{Q} \\ \mathbf{Q}\mathbf{K}_{\sin}\mathbf{Q} \end{bmatrix} \end{aligned} \quad (19)$$

by taking $\mathbf{A} = \mathbf{K}_{\cos}$ or \mathbf{K}_{\sin} , we result in

$$\begin{aligned} \mathbb{E}[\mathbf{Q}\mathbf{K}_{\cos}\mathbf{Q}] &\simeq \frac{c}{ac-bd}\bar{\mathbf{Q}}\mathbf{K}_{\cos}\bar{\mathbf{Q}} + \frac{b}{ac-bd}\bar{\mathbf{Q}}\mathbf{K}_{\sin}\bar{\mathbf{Q}} \\ \mathbb{E}[\mathbf{Q}\mathbf{K}_{\sin}\mathbf{Q}] &\simeq \frac{a}{ac-bd}\bar{\mathbf{Q}}\mathbf{K}_{\sin}\bar{\mathbf{Q}} + \frac{d}{ac-bd}\bar{\mathbf{Q}}\mathbf{K}_{\cos}\bar{\mathbf{Q}} \end{aligned}$$

with $a = 1 - \frac{N}{n}\frac{\frac{1}{n}\text{tr}(\bar{\mathbf{Q}}\mathbf{K}_{\cos}\bar{\mathbf{Q}}\mathbf{K}_{\cos})}{(1+\delta_{\cos})^2}$, $b = \frac{N}{n}\frac{\frac{1}{n}\text{tr}(\bar{\mathbf{Q}}\mathbf{K}_{\cos}\bar{\mathbf{Q}}\mathbf{K}_{\sin})}{(1+\delta_{\sin})^2}$, $c = 1 - \frac{N}{n}\frac{\frac{1}{n}\text{tr}(\bar{\mathbf{Q}}\mathbf{K}_{\sin}\bar{\mathbf{Q}}\mathbf{K}_{\sin})}{(1+\delta_{\sin})^2}$ and $d = \frac{N}{n}\frac{\frac{1}{n}\text{tr}(\bar{\mathbf{Q}}\mathbf{K}_{\sin}\bar{\mathbf{Q}}\mathbf{K}_{\cos})}{(1+\delta_{\cos})^2}$ such that $(1+\delta_{\sin})^2b = (1+\delta_{\cos})^2d$.

$$\mathbb{E}\begin{bmatrix} \mathbf{Q}\mathbf{K}_{\cos}\mathbf{Q} \\ \mathbf{Q}\mathbf{K}_{\sin}\mathbf{Q} \end{bmatrix} \simeq \begin{bmatrix} a & -b \\ -d & c \end{bmatrix}^{-1} \begin{bmatrix} \bar{\mathbf{Q}}\mathbf{K}_{\cos}\bar{\mathbf{Q}} \\ \bar{\mathbf{Q}}\mathbf{K}_{\sin}\bar{\mathbf{Q}} \end{bmatrix} \equiv \Omega \begin{bmatrix} \bar{\mathbf{Q}}\mathbf{K}_{\cos}\bar{\mathbf{Q}} \\ \bar{\mathbf{Q}}\mathbf{K}_{\sin}\bar{\mathbf{Q}} \end{bmatrix}$$

for $\Omega \equiv \begin{bmatrix} a & -b \\ -d & c \end{bmatrix}^{-1}$. Plugging back into (19) we conclude the proof of Lemma 4. \square

Theorem 2 can be achieved by considering the concentration of (the bilinear form) $\frac{1}{n}\mathbf{y}^\top\mathbf{Q}^2\mathbf{y}$ around its expectation $\frac{1}{n}\mathbf{y}^\top\mathbb{E}[\mathbf{Q}^2]\mathbf{y}$ (with for instance Lemma 3 in [30]), together with Lemma 4. This concludes the proof of Theorem 2. \blacksquare

D Proof of Theorem 3

Recall the definition of $E_{\text{test}} = \frac{1}{\hat{n}}\|\hat{\mathbf{y}} - \Sigma_{\hat{\mathbf{X}}}^\top\beta\|^2$ from (4) with $\Sigma_{\hat{\mathbf{X}}} = \begin{bmatrix} \cos(\mathbf{W}\hat{\mathbf{X}}) \\ \sin(\mathbf{W}\hat{\mathbf{X}}) \end{bmatrix} \in \mathbb{R}^{2N \times \hat{n}}$ on a test set $(\hat{\mathbf{X}}, \hat{\mathbf{y}})$ of size \hat{n} , and first focus on the case $2N > n$ where $\beta = \frac{1}{n}\Sigma_{\mathbf{X}}\mathbf{Q}\mathbf{y}$ as per (3). By (15), we have

$$E_{\text{test}} = \frac{1}{\hat{n}}\left\|\hat{\mathbf{y}} - \frac{1}{n}\Sigma_{\hat{\mathbf{X}}}^\top\Sigma_{\mathbf{X}}\mathbf{Q}\mathbf{y}\right\|^2 = \frac{1}{\hat{n}}\left\|\hat{\mathbf{y}} - \frac{1}{n}\sum_{i=1}^N\hat{\mathbf{U}}_i\mathbf{U}_i^\top\mathbf{Q}\mathbf{y}\right\|^2$$

where, similar to the notation $\mathbf{U}_i = [\cos(\mathbf{X}^\top\mathbf{w}_i) \quad \sin(\mathbf{X}^\top\mathbf{w}_i)] \in \mathbb{R}^{n \times 2}$ as in the proof of Theorem 1, we denote

$$\hat{\mathbf{U}}_i \equiv [\cos(\hat{\mathbf{X}}^\top\mathbf{w}_i) \quad \sin(\hat{\mathbf{X}}^\top\mathbf{w}_i)] \in \mathbb{R}^{\hat{n} \times 2}.$$

As a consequence, we further get

$$\begin{aligned}
\mathbb{E}[E_{\text{test}}] &= \frac{1}{\hat{n}} \|\hat{\mathbf{y}}\|^2 - \frac{2}{n\hat{n}} \sum_{i=1}^N \hat{\mathbf{y}}^\top \mathbb{E}[\hat{\mathbf{U}}_i \mathbf{U}_i^\top \mathbf{Q}] \mathbf{y} + \frac{1}{n^2 \hat{n}} \sum_{i,j=1}^N \mathbf{y}^\top \mathbb{E}[\mathbf{Q} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \hat{\mathbf{U}}_j \mathbf{U}_j^\top \mathbf{Q}] \mathbf{y} \\
&= \frac{1}{\hat{n}} \|\hat{\mathbf{y}}\|^2 - \frac{2}{n\hat{n}} \sum_{i=1}^N \hat{\mathbf{y}}^\top \mathbb{E} \left[\hat{\mathbf{U}}_i (\mathbf{I}_2 + \frac{1}{n} \mathbf{U}_i^\top \mathbf{Q}_{-i} \mathbf{U}_i)^{-1} \mathbf{U}_i^\top \mathbf{Q}_{-i} \right] \mathbf{y} + \frac{1}{n^2 \hat{n}} \sum_{i,j=1}^N \mathbf{y}^\top \mathbb{E}[\mathbf{Q} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \hat{\mathbf{U}}_j \mathbf{U}_j^\top \mathbf{Q}] \mathbf{y} \\
&\simeq \frac{1}{\hat{n}} \|\hat{\mathbf{y}}\|^2 - \frac{2}{n\hat{n}} \sum_{i=1}^N \hat{\mathbf{y}}^\top \mathbb{E} \left[\hat{\mathbf{U}}_i \begin{bmatrix} \frac{1}{1+\delta_{\cos}} & 0 \\ 0 & \frac{1}{1+\delta_{\sin}} \end{bmatrix} \mathbf{U}_i^\top \mathbf{Q}_{-i} \right] \mathbf{y} + \frac{1}{n^2 \hat{n}} \sum_{i,j=1}^N \mathbf{y}^\top \mathbb{E}[\mathbf{Q} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \hat{\mathbf{U}}_j \mathbf{U}_j^\top \mathbf{Q}] \mathbf{y} \\
&\simeq \frac{1}{\hat{n}} \|\hat{\mathbf{y}}\|^2 - \frac{2}{\hat{n}} \hat{\mathbf{y}}^\top \left(\frac{N}{n} \frac{\mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X})}{1 + \delta_{\cos}} + \frac{N}{n} \frac{\mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X})}{1 + \delta_{\sin}} \right) \bar{\mathbf{Q}} \mathbf{y} + \frac{1}{n^2 \hat{n}} \sum_{i,j=1}^N \mathbf{y}^\top \mathbb{E}[\mathbf{Q} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \hat{\mathbf{U}}_j \mathbf{U}_j^\top \mathbf{Q}] \mathbf{y}
\end{aligned}$$

where we similarly denote

$$\begin{aligned}
\mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X}) &\equiv \left\{ e^{-\frac{1}{2}(\|\hat{\mathbf{x}}_i\|^2 + \|\mathbf{x}_j\|^2)} \cosh(\hat{\mathbf{x}}_i^\top \mathbf{x}_j) \right\}_{i,j=1}^{\hat{n},n} \\
\mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X}) &\equiv \left\{ e^{-\frac{1}{2}(\|\hat{\mathbf{x}}_i\|^2 + \|\mathbf{x}_j\|^2)} \sinh(\hat{\mathbf{x}}_i^\top \mathbf{x}_j) \right\}_{i,j=1}^{\hat{n},n} \in \mathbb{R}^{\hat{n} \times n}.
\end{aligned}$$

Note that, different from the proof of Theorem 1 and 2 where we constantly use the fact that $\|\mathbf{Q}\| \leq \lambda^{-1}$ and

$$\frac{1}{n} \Sigma_{\mathbf{X}}^\top \Sigma_{\mathbf{X}} \mathbf{Q} = \mathbf{I}_n - \lambda \mathbf{Q}$$

so that $\|\frac{1}{n} \Sigma_{\mathbf{X}}^\top \Sigma_{\mathbf{X}} \mathbf{Q}\| \leq 1$, we do not have in general a simple control for $\|\frac{1}{n} \Sigma_{\hat{\mathbf{X}}}^\top \Sigma_{\mathbf{X}} \mathbf{Q}\|$, when arbitrary $\hat{\mathbf{X}}$ is considered. Intuitively speaking, this is due to the loss-of-control for $\|\frac{1}{n}(\Sigma_{\hat{\mathbf{X}}} - \Sigma_{\mathbf{X}})^\top \Sigma_{\mathbf{X}} \mathbf{Q}\|$ when $\hat{\mathbf{X}}$ can be chosen arbitrarily with respect to \mathbf{X} . It was remarked in [30, Remark 1] that in general only a $O(\sqrt{n})$ upper bound can be derived for $\|\frac{1}{\sqrt{n}} \Sigma_{\mathbf{X}}\|$ or $\|\frac{1}{\sqrt{n}} \Sigma_{\hat{\mathbf{X}}}\|$. Nonetheless, this problem can be resolved with the additional Assumption 2.

More precisely, note that

$$\left\| \frac{1}{n} \Sigma_{\hat{\mathbf{X}}}^\top \Sigma_{\mathbf{X}} \mathbf{Q} \right\| \leq \frac{1}{n} \|\Sigma_{\hat{\mathbf{X}}}^\top \Sigma_{\mathbf{X}} \mathbf{Q}\| + \frac{1}{n} \|(\Sigma_{\hat{\mathbf{X}}} - \Sigma_{\mathbf{X}})^\top \Sigma_{\mathbf{X}} \mathbf{Q}\| \leq 1 + \frac{1}{\sqrt{n}} \|\Sigma_{\hat{\mathbf{X}}} - \Sigma_{\mathbf{X}}\| \cdot \frac{1}{\sqrt{n}} \|\Sigma_{\mathbf{X}} \mathbf{Q}\| \quad (20)$$

it remains to show that $\|\Sigma_{\mathbf{X}} - \Sigma_{\hat{\mathbf{X}}}\| = O(\sqrt{n})$ under Assumption 2 to establish $\|\frac{1}{n} \Sigma_{\hat{\mathbf{X}}}^\top \Sigma_{\mathbf{X}} \mathbf{Q}\| = O(1)$, that is, to show that

$$\|\sigma(\mathbf{W}\mathbf{X}) - \sigma(\mathbf{W}\hat{\mathbf{X}})\| = O(\sqrt{n}) \quad (21)$$

for $\sigma \in \{\cos, \sin\}$. Note this cannot be achieved using only the Lipschitz nature of $\sigma(\cdot)$ and the fact that $\|\mathbf{X} - \hat{\mathbf{X}}\| \leq \|\mathbf{X}\| + \|\hat{\mathbf{X}}\| = O(1)$ under Assumption 1 by writing

$$\|\sigma(\mathbf{W}\mathbf{X}) - \sigma(\mathbf{W}\hat{\mathbf{X}})\| \leq \|\sigma(\mathbf{W}\mathbf{X}) - \sigma(\mathbf{W}\hat{\mathbf{X}})\|_F \leq \|\mathbf{W}\|_F \cdot \|\mathbf{X} - \hat{\mathbf{X}}\| = O(n). \quad (22)$$

where we recall that $\|\mathbf{W}\| = O(\sqrt{n})$ and $\|\mathbf{W}\|_F = O(n)$. Nonetheless, from [29, Proposition B.1] we have that the product $\mathbf{W}\mathbf{X}$, and thus $\sigma(\mathbf{W}\mathbf{X})$, strongly concentrates around its expectation in the sense of (9), so that

$$\begin{aligned}
\|\sigma(\mathbf{W}\mathbf{X}) - \sigma(\mathbf{W}\hat{\mathbf{X}})\| &\leq \|\sigma(\mathbf{W}\mathbf{X}) - \mathbb{E}[\sigma(\mathbf{W}\mathbf{X})]\| + \|\mathbb{E}[\sigma(\mathbf{W}\mathbf{X}) - \sigma(\mathbf{W}\hat{\mathbf{X}})]\| \\
&\quad + \|\sigma(\mathbf{W}\hat{\mathbf{X}}) - \mathbb{E}[\sigma(\mathbf{W}\hat{\mathbf{X}})]\| = O(\sqrt{n})
\end{aligned}$$

under Assumption 2. As a results, we are allowed to control $\frac{1}{n} \Sigma_{\hat{\mathbf{X}}}^\top \Sigma_{\mathbf{X}} \mathbf{Q}$ and similarly $\frac{1}{n} \Sigma_{\mathbf{X}}^\top \Sigma_{\hat{\mathbf{X}}} \mathbf{Q}$ in the same vein as $\frac{1}{n} \Sigma_{\mathbf{X}}^\top \Sigma_{\mathbf{X}} \mathbf{Q}$ in the proof of Theorem 1 and 2 in Appendix B and C, respectively.

It thus remains to handle the last term (noted \mathbf{Z}) as follows

$$\begin{aligned}
\mathbf{Z} &\equiv \frac{1}{n^2 \hat{n}} \sum_{i,j=1}^N \mathbf{y}^\top \mathbb{E}[\mathbf{Q} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \hat{\mathbf{U}}_j \mathbf{U}_j^\top \mathbf{Q}] \mathbf{y} \\
&= \frac{1}{n^2 \hat{n}} \sum_{i=1}^N \mathbf{y}^\top \mathbb{E}[\mathbf{Q} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \hat{\mathbf{U}}_i \mathbf{U}_i^\top \mathbf{Q}] \mathbf{y} + \frac{1}{n^2 \hat{n}} \sum_{i=1}^N \sum_{j \neq i} \mathbf{y}^\top \mathbb{E}[\mathbf{Q} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \hat{\mathbf{U}}_j \mathbf{U}_j^\top \mathbf{Q}] \mathbf{y} = \mathbf{Z}_1 + \mathbf{Z}_2
\end{aligned}$$

where \mathbf{Z}_1 term can be treated as

$$\begin{aligned}
\mathbf{Z}_1 &\equiv \frac{1}{n^2 \hat{n}} \sum_{i=1}^N \mathbf{y}^\top \mathbb{E}[\mathbf{Q} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \hat{\mathbf{U}}_i \mathbf{U}_i^\top \mathbf{Q}] \mathbf{y} \\
&= \frac{1}{n \hat{n}} \sum_{i=1}^N \mathbf{y}^\top \mathbb{E}[\mathbf{Q}_{-i} \mathbf{U}_i (\mathbf{I}_2 + \frac{1}{n} \mathbf{U}_i^\top \mathbf{Q}_{-i} \mathbf{U}_i)^{-1} \frac{1}{n} \hat{\mathbf{U}}_i^\top \hat{\mathbf{U}}_i (\mathbf{I}_2 + \frac{1}{n} \mathbf{U}_i^\top \mathbf{Q}_{-i} \mathbf{U}_i)^{-1} \mathbf{U}_i^\top \mathbf{Q}_{-i}] \mathbf{y} \\
&\simeq \frac{1}{n \hat{n}} \sum_{i=1}^N \mathbf{y}^\top \mathbb{E}[\mathbf{Q}_{-i} \mathbf{U}_i \begin{bmatrix} \frac{1}{1+\delta_{\cos}} & 0 \\ 0 & \frac{1}{1+\delta_{\sin}} \end{bmatrix} \begin{bmatrix} \frac{1}{n} \text{tr} \hat{\mathbf{K}}_{\cos} & 0 \\ 0 & \frac{1}{n} \text{tr} \hat{\mathbf{K}}_{\sin} \end{bmatrix} \begin{bmatrix} \frac{1}{1+\delta_{\cos}} & 0 \\ 0 & \frac{1}{1+\delta_{\sin}} \end{bmatrix} \mathbf{U}_i^\top \mathbf{Q}_{-i}] \mathbf{y} \\
&\simeq \frac{N}{n} \frac{1}{\hat{n}} \mathbf{y}^\top \mathbb{E} \left[\mathbf{Q} \left(\frac{\frac{1}{n} \text{tr} \mathbf{K}_{\cos}(\hat{\mathbf{X}}, \hat{\mathbf{X}})}{(1+\delta_{\cos})^2} \mathbf{K}_{\cos} + \frac{\frac{1}{n} \text{tr} \mathbf{K}_{\sin}(\hat{\mathbf{X}}, \hat{\mathbf{X}})}{(1+\delta_{\sin})^2} \mathbf{K}_{\sin} \right) \mathbf{Q} \right] \mathbf{y} \\
&\simeq \frac{N}{n} \frac{1}{\hat{n}} \begin{bmatrix} \frac{1}{n} \text{tr} \mathbf{K}_{\cos}(\hat{\mathbf{X}}, \hat{\mathbf{X}}) & \frac{1}{n} \text{tr} \frac{1}{n} \text{tr} \mathbf{K}_{\sin}(\hat{\mathbf{X}}, \hat{\mathbf{X}})}{(1+\delta_{\sin})^2} \end{bmatrix} \Omega \begin{bmatrix} \mathbf{y}^\top \bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}} \mathbf{y} \\ \mathbf{y}^\top \bar{\mathbf{Q}} \mathbf{K}_{\sin} \bar{\mathbf{Q}} \mathbf{y} \end{bmatrix}
\end{aligned}$$

where we apply Lemma 4 and recall

$$\mathbf{K}_{\cos}(\hat{\mathbf{X}}, \hat{\mathbf{X}}) \equiv \left\{ e^{-\frac{1}{2}(\|\hat{\mathbf{x}}_i\|^2 + \|\hat{\mathbf{x}}_j\|^2)} \cosh(\hat{\mathbf{x}}_i^\top \hat{\mathbf{x}}_j) \right\}_{i,j=1}^{\hat{n}}, \quad \mathbf{K}_{\sin}(\hat{\mathbf{X}}, \hat{\mathbf{X}}) \equiv \left\{ e^{-\frac{1}{2}(\|\hat{\mathbf{x}}_i\|^2 + \|\hat{\mathbf{x}}_j\|^2)} \sinh(\hat{\mathbf{x}}_i^\top \hat{\mathbf{x}}_j) \right\}_{i,j=1}^{\hat{n}}$$

Moving on to \mathbf{Z}_2 and we write

$$\begin{aligned}
\mathbf{Z}_2 &\equiv \frac{1}{n^2 \hat{n}} \mathbb{E} \sum_{i=1}^N \sum_{j \neq i} \mathbf{y}^\top \mathbf{Q} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \hat{\mathbf{U}}_j \mathbf{U}_j^\top \mathbf{Q} \mathbf{y} \\
&= \frac{1}{n^2 \hat{n}} \mathbb{E} \sum_{i=1}^N \sum_{j \neq i} \mathbf{y}^\top \mathbf{Q}_{-j} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \hat{\mathbf{U}}_j (\mathbf{I}_2 + \frac{1}{n} \mathbf{U}_j^\top \mathbf{Q}_{-j} \mathbf{U}_j)^{-1} \mathbf{U}_j^\top \mathbf{Q}_{-j} \mathbf{y} \\
&\quad - \frac{1}{n^2 \hat{n}} \mathbb{E} \sum_{i=1}^N \sum_{j \neq i} \mathbf{y}^\top \mathbf{Q}_{-j} \mathbf{U}_j (\mathbf{I}_2 + \frac{1}{n} \mathbf{U}_j^\top \mathbf{Q}_{-j} \mathbf{U}_j)^{-1} \mathbf{U}_j^\top \mathbf{Q}_{-j} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \hat{\mathbf{U}}_j (\mathbf{I}_2 + \frac{1}{n} \mathbf{U}_j^\top \mathbf{Q}_{-j} \mathbf{U}_j)^{-1} \mathbf{U}_j^\top \mathbf{Q}_{-j} \mathbf{y} \\
&\simeq \frac{1}{n \hat{n}} \mathbb{E} \sum_{i=1}^N \sum_{j \neq i} \mathbf{y}^\top \mathbf{Q}_{-j} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \left(\frac{\mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X})}{1+\delta_{\cos}} + \frac{\mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X})}{1+\delta_{\sin}} \right) \mathbf{Q}_{-j} \mathbf{y} \\
&\quad - \frac{1}{n^2 \hat{n}} \mathbb{E} \sum_{i=1}^N \sum_{j \neq i} \mathbf{y}^\top \mathbf{Q}_{-j} \mathbf{U}_j \begin{bmatrix} \frac{1}{1+\delta_{\cos}} & 0 \\ 0 & \frac{1}{1+\delta_{\sin}} \end{bmatrix} \begin{bmatrix} \frac{1}{n} \text{tr}(\mathbf{Q}_{-j} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X})) & 0 \\ 0 & \frac{1}{n} \text{tr}(\mathbf{Q}_{-j} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X})) \end{bmatrix} \\
&\quad \begin{bmatrix} \frac{1}{1+\delta_{\cos}} & 0 \\ 0 & \frac{1}{1+\delta_{\sin}} \end{bmatrix} \mathbf{U}_j^\top \mathbf{Q}_{-j} \mathbf{y} \equiv \mathbf{Z}_{21} - \mathbf{Z}_{22}.
\end{aligned}$$

For the term \mathbf{Z}_{21} , note that $\mathbf{Q}_{-j} \simeq \mathbf{Q}$ and **depends** on \mathbf{U}_i (and $\hat{\mathbf{U}}_i$), such that

$$\begin{aligned}
\mathbf{Z}_{21} &\equiv \frac{1}{n^2 \hat{n}} \mathbb{E} \sum_{i=1}^N \sum_{j \neq i} \mathbf{y}^\top \mathbf{Q}_{-j} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \left(\frac{\mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X})}{1 + \delta_{\cos}} + \frac{\mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X})}{1 + \delta_{\sin}} \right) \mathbf{Q}_{-j} \mathbf{y} \\
&\simeq \frac{N}{n} \frac{1}{n \hat{n}} \mathbb{E} \sum_{i=1}^N \mathbf{y}^\top \mathbf{Q} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \left(\frac{\mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X})}{1 + \delta_{\cos}} + \frac{\mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X})}{1 + \delta_{\sin}} \right) \mathbf{Q} \mathbf{y} \\
&= \frac{N}{n} \frac{1}{n \hat{n}} \mathbb{E} \sum_{i=1}^N \mathbf{y}^\top \mathbf{Q}_{-i} \mathbf{U}_i (\mathbf{I}_2 + \frac{1}{n} \mathbf{U}_i^\top \mathbf{Q}_{-i} \mathbf{U}_i)^{-1} \hat{\mathbf{U}}_i^\top \hat{\Phi} \mathbf{Q}_{-i} \mathbf{y} \\
&\quad - \frac{N}{n} \frac{1}{n \hat{n}} \mathbb{E} \sum_{i=1}^N \mathbf{y}^\top \mathbf{Q}_{-i} \mathbf{U}_i (\mathbf{I}_2 + \frac{1}{n} \mathbf{U}_i^\top \mathbf{Q}_{-i} \mathbf{U}_i)^{-1} \hat{\mathbf{U}}_i^\top \hat{\Phi} \mathbf{Q}_{-i} \mathbf{U}_i (\mathbf{I}_2 + \frac{1}{n} \mathbf{U}_i^\top \mathbf{Q}_{-i} \mathbf{U}_i)^{-1} \mathbf{U}_i^\top \mathbf{Q}_{-i} \mathbf{y} \\
&\simeq \frac{N}{n} \frac{1}{n \hat{n}} \mathbb{E} \sum_{i=1}^N \mathbf{y}^\top \mathbf{Q}_{-i} \left(\frac{\mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X})}{1 + \delta_{\cos}} + \frac{\mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X})}{1 + \delta_{\sin}} \right)^\top \hat{\Phi} \mathbf{Q}_{-i} \mathbf{y} \\
&\quad - \frac{N}{n} \frac{1}{n \hat{n}} \mathbb{E} \sum_{i=1}^N \mathbf{y}^\top \mathbf{Q}_{-i} \mathbf{U}_i \begin{bmatrix} \frac{1}{1+\delta_{\cos}} & 0 \\ 0 & \frac{1}{1+\delta_{\sin}} \end{bmatrix} \frac{1}{n} \hat{\mathbf{U}}_i^\top \hat{\Phi} \mathbf{Q}_{-i} \mathbf{U}_i \begin{bmatrix} \frac{1}{1+\delta_{\cos}} & 0 \\ 0 & \frac{1}{1+\delta_{\sin}} \end{bmatrix} \mathbf{U}_i^\top \mathbf{Q}_{-i} \mathbf{y}
\end{aligned}$$

where we recall the shortcut $\Phi \equiv \frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} + \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}}$ and similarly $\hat{\Phi} \equiv \frac{\mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X})}{1+\delta_{\cos}} + \frac{\mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X})}{1+\delta_{\sin}} \in \mathbb{R}^{\hat{n} \times n}$. As a consequence, we further have, with Lemma 4 that

$$\begin{aligned}
\mathbf{Z}_{21} &\simeq \left(\frac{N}{n} \right)^2 \frac{1}{\hat{n}} \mathbf{y}^\top \mathbb{E} \left[\mathbf{Q} \hat{\Phi}^\top \hat{\Phi} \mathbf{Q} \right] \mathbf{y} \\
&\quad - \frac{N}{n} \frac{1}{n \hat{n}} \mathbb{E} \sum_{i=1}^N \mathbf{y}^\top \mathbf{Q}_{-i} \mathbf{U}_i \begin{bmatrix} \frac{1}{1+\delta_{\cos}} & 0 \\ 0 & \frac{1}{1+\delta_{\sin}} \end{bmatrix} \begin{bmatrix} \frac{1}{n} \text{tr}(\hat{\Phi} \bar{\mathbf{Q}} \mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X})^\top) & 0 \\ 0 & \frac{1}{n} \text{tr}(\hat{\Phi} \bar{\mathbf{Q}} \mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X})^\top) \end{bmatrix} \\
&\quad \times \begin{bmatrix} \frac{1}{1+\delta_{\cos}} & 0 \\ 0 & \frac{1}{1+\delta_{\sin}} \end{bmatrix} \mathbf{U}_i^\top \mathbf{Q}_{-i} \mathbf{y} \\
&\simeq \left(\frac{N}{n} \right)^2 \frac{1}{\hat{n}} \mathbf{y}^\top \mathbb{E} \left[\mathbf{Q} \hat{\Phi}^\top \hat{\Phi} \mathbf{Q} \right] \mathbf{y} \\
&\quad - \left(\frac{N}{n} \right)^2 \frac{1}{\hat{n}} \mathbb{E} \mathbf{y}^\top \mathbf{Q} \left(\frac{1}{n} \text{tr}(\hat{\Phi} \bar{\mathbf{Q}} \mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X})^\top) \frac{\mathbf{K}_{\cos}}{(1+\delta_{\cos})^2} + \frac{1}{n} \text{tr}(\hat{\Phi} \bar{\mathbf{Q}} \mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X})^\top) \frac{\mathbf{K}_{\sin}}{(1+\delta_{\sin})^2} \right) \mathbf{Q} \mathbf{y} \\
&\simeq \left(\frac{N}{n} \right)^2 \frac{1}{\hat{n}} \mathbf{y}^\top \mathbb{E} \left[\mathbf{Q} \hat{\Phi}^\top \hat{\Phi} \mathbf{Q} \right] \mathbf{y} - \left(\frac{N}{n} \right)^2 \frac{1}{\hat{n}} \mathbf{y}^\top \left(\begin{bmatrix} \frac{1}{n} \text{tr}(\hat{\Phi} \bar{\mathbf{Q}} \mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X})^\top) & \frac{1}{n} \text{tr}(\hat{\Phi} \bar{\mathbf{Q}} \mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X})^\top) \\ \frac{1}{(1+\delta_{\cos})^2} & \frac{1}{(1+\delta_{\sin})^2} \end{bmatrix} \mathbb{E} \begin{bmatrix} \mathbf{Q} \mathbf{K}_{\cos} \mathbf{Q} \\ \mathbf{Q} \mathbf{K}_{\sin} \mathbf{Q} \end{bmatrix} \right) \mathbf{y} \\
&\simeq \left(\frac{N}{n} \right)^2 \frac{1}{\hat{n}} \mathbf{y}^\top \bar{\mathbf{Q}} \Phi^\top \hat{\Phi} \bar{\mathbf{Q}} \mathbf{y} \\
&\quad + \left(\frac{N}{n} \right)^2 \frac{1}{\hat{n}} \begin{bmatrix} \frac{1}{n} \text{tr} \bar{\mathbf{Q}} \frac{N}{n} \hat{\Phi}^\top \hat{\Phi} \bar{\mathbf{Q}} \mathbf{K}_{\cos} - \frac{1}{n} \text{tr} \bar{\mathbf{Q}} \hat{\Phi} \mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X}) & \frac{1}{n} \text{tr} \bar{\mathbf{Q}} \frac{N}{n} \hat{\Phi}^\top \hat{\Phi} \bar{\mathbf{Q}} \mathbf{K}_{\sin} - \frac{1}{n} \text{tr} \bar{\mathbf{Q}} \hat{\Phi}^\top \mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X}) \\ \frac{1}{(1+\delta_{\cos})^2} & \frac{1}{(1+\delta_{\sin})^2} \end{bmatrix} \Omega \begin{bmatrix} \mathbf{y}^\top \bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}} \mathbf{y} \\ \mathbf{y}^\top \bar{\mathbf{Q}} \mathbf{K}_{\sin} \bar{\mathbf{Q}} \mathbf{y} \end{bmatrix}
\end{aligned}$$

The last term \mathbf{Z}_{22} can be similarly treated as

$$\mathbf{Z}_{22} \simeq \frac{1}{n^2 \hat{n}} \mathbb{E} \sum_{i=1}^N \sum_{j \neq i} \mathbf{y}^\top \mathbf{Q}_{-j} \mathbf{U}_j \begin{bmatrix} \frac{1}{n} \text{tr}(\mathbf{Q} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X})) & 0 \\ 0 & \frac{1}{n} \text{tr}(\mathbf{Q} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X})) \end{bmatrix} \mathbf{U}_j^\top \mathbf{Q}_{-j} \mathbf{y}$$

where by Lemma 2 we deduce

$$\begin{aligned}
\frac{1}{n} \text{tr}(\mathbf{Q} \mathbf{U}_i \hat{\mathbf{U}}_i^\top \mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X})) &\simeq \frac{1}{n} \text{tr} \left(\mathbf{Q}_{-i} \mathbf{U}_i (\mathbf{I}_2 + \mathbf{U}_i^\top \mathbf{Q}_{-i} \mathbf{U}_i)^{-1} \hat{\mathbf{U}}_i^\top \mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X}) \right) \\
&\simeq \frac{1}{n} \text{tr} \left(\mathbf{Q}_{-i} \mathbf{U}_i \begin{bmatrix} \frac{1}{1+\delta_{\cos}} & 0 \\ 0 & \frac{1}{1+\delta_{\sin}} \end{bmatrix} \hat{\mathbf{U}}_i^\top \mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X}) \right) \simeq \frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \hat{\Phi}^\top \mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X}))
\end{aligned}$$

so that by again Lemma 4

$$\begin{aligned}
\mathbf{Z}_{22} &\simeq \frac{N}{n} \frac{1}{n\hat{n}} \mathbb{E} \sum_{j=1}^N \mathbf{y}^\top \mathbf{Q}_{-j} \mathbf{U}_j \begin{bmatrix} \frac{\frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \hat{\Phi}^\top \mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X}))}{(1+\delta_{\cos})^2} & 0 \\ 0 & \frac{\frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \hat{\Phi}^\top \mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X}))}{(1+\delta_{\sin})^2} \end{bmatrix} \mathbf{U}_j^\top \mathbf{Q}_{-j} \mathbf{y} \\
&\simeq \left(\frac{N}{n} \right)^2 \frac{1}{\hat{n}} \mathbf{y}^\top \mathbb{E} \left[\mathbf{Q} \left(\frac{\frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \hat{\Phi}^\top \mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X}))}{(1+\delta_{\cos})^2} \mathbf{K}_{\cos} + \frac{\frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \hat{\Phi}^\top \mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X}))}{(1+\delta_{\sin})^2} \mathbf{K}_{\sin} \right) \mathbf{Q} \right] \mathbf{y} \\
&\simeq \left(\frac{N}{n} \right)^2 \frac{1}{\hat{n}} \mathbf{y}^\top \left(\bar{\mathbf{Q}} \Xi \bar{\mathbf{Q}} + \frac{N}{n} \begin{bmatrix} \frac{\frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \Xi \bar{\mathbf{Q}} \mathbf{K}_{\cos})}{(1+\delta_{\cos})^2} & \frac{\frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \Xi \bar{\mathbf{Q}} \mathbf{K}_{\sin})}{(1+\delta_{\sin})^2} \end{bmatrix} \Omega \begin{bmatrix} \bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}} \\ \bar{\mathbf{Q}} \mathbf{K}_{\sin} \bar{\mathbf{Q}} \end{bmatrix} \right) \mathbf{y} \\
&\simeq \left(\frac{N}{n} \right)^2 \frac{1}{\hat{n}} \begin{bmatrix} \frac{\frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \hat{\Phi}^\top \mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X}))}{(1+\delta_{\cos})^2} & \frac{\frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \hat{\Phi}^\top \mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X}))}{(1+\delta_{\sin})^2} \end{bmatrix} \Omega \begin{bmatrix} \mathbf{y}^\top \bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}} \mathbf{y} \\ \mathbf{y}^\top \bar{\mathbf{Q}} \mathbf{K}_{\sin} \bar{\mathbf{Q}} \mathbf{y} \end{bmatrix}.
\end{aligned}$$

Assembling the estimates for \mathbf{Z}_1 , \mathbf{Z}_{21} and \mathbf{Z}_{22} , we get

$$\begin{aligned}
\mathbb{E}[E_{\text{test}}] &\simeq \frac{1}{\hat{n}} \|\hat{\mathbf{y}}\|^2 - \frac{2}{\hat{n}} \hat{\mathbf{y}}^\top \frac{N}{n} \hat{\Phi} \bar{\mathbf{Q}} \mathbf{y} + \frac{1}{\hat{n}} \mathbf{y}^\top \left(\frac{N^2}{n^2} \bar{\mathbf{Q}} \hat{\Phi}^\top \hat{\Phi} \bar{\mathbf{Q}} \right) \mathbf{y} + \left(\frac{N}{n} \right)^2 \frac{1}{n\hat{n}} \times \\
&\quad \left[\frac{\frac{n}{N} \text{tr} \mathbf{K}_{\cos}(\hat{\mathbf{X}}, \hat{\mathbf{X}}) + \frac{N}{n} \text{tr} \bar{\mathbf{Q}} \hat{\Phi}^\top \hat{\Phi} \bar{\mathbf{Q}} \mathbf{K}_{\cos} - 2 \text{tr} \bar{\mathbf{Q}} \hat{\Phi}^\top \mathbf{K}_{\cos}(\hat{\mathbf{X}}, \mathbf{X})}{(1+\delta_{\cos})^2} - \frac{\frac{n}{N} \text{tr} \mathbf{K}_{\sin}(\hat{\mathbf{X}}, \hat{\mathbf{X}}) + \frac{N}{n} \text{tr} \bar{\mathbf{Q}} \hat{\Phi}^\top \hat{\Phi} \bar{\mathbf{Q}} \mathbf{K}_{\sin} - 2 \text{tr} \bar{\mathbf{Q}} \hat{\Phi}^\top \mathbf{K}_{\sin}(\hat{\mathbf{X}}, \mathbf{X})}{(1+\delta_{\sin})^2} \right] \\
&\quad \times \Omega \begin{bmatrix} \mathbf{y}^\top \bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}} \mathbf{y} \\ \mathbf{y}^\top \bar{\mathbf{Q}} \mathbf{K}_{\sin} \bar{\mathbf{Q}} \mathbf{y} \end{bmatrix}
\end{aligned}$$

which, up to further simplifications, concludes the proof of Theorem 3.

E Several Useful Lemmas

Lemma 5 (Some useful properties of Ω). *For any $\lambda > 0$ and Ω defined in (8), we have*

1. *all entries of Ω are positive;*
2. *for $2N = n$, $\det(\Omega^{-1})$, as well as the entries of Ω , scales like λ as $\lambda \rightarrow 0$;*

Proof. Developing the inverse we obtain

$$\Omega = \begin{bmatrix} 1 - \frac{N}{n} \frac{\frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}} \mathbf{K}_{\cos})}{(1+\delta_{\cos})^2} & -\frac{N}{n} \frac{\frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}} \mathbf{K}_{\sin})}{(1+\delta_{\sin})^2} \\ -\frac{N}{n} \frac{\frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\sin} \bar{\mathbf{Q}} \mathbf{K}_{\cos})}{(1+\delta_{\cos})^2} & 1 - \frac{N}{n} \frac{\frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\sin} \bar{\mathbf{Q}} \mathbf{K}_{\sin})}{(1+\delta_{\sin})^2} \end{bmatrix}^{-1}$$

we have $[\Omega^{-1}]_{11} = \frac{1}{1+\delta_{\cos}} + \frac{\lambda}{n} \text{tr} \bar{\mathbf{Q}} \frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} \bar{\mathbf{Q}} + \frac{N}{n} \frac{1}{n} \text{tr} \bar{\mathbf{Q}} \frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} \bar{\mathbf{Q}} \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}} > 0$, $[\Omega^{-1}]_{12} < 0$, and similarly $[\Omega^{-1}]_{21} < 0$, $[\Omega^{-1}]_{22} > 0$. Furthermore, the determinant writes

$$\begin{aligned}
\det(\Omega^{-1}) &= \left(1 - \frac{1}{n} \text{tr} \bar{\mathbf{Q}} \frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} + \frac{\lambda}{n} \text{tr} \bar{\mathbf{Q}} \frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} \bar{\mathbf{Q}} \right) \left(1 - \frac{1}{n} \text{tr} \bar{\mathbf{Q}} \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}} + \frac{\lambda}{n} \text{tr} \bar{\mathbf{Q}} \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}} \bar{\mathbf{Q}} \right) \\
&+ \left(1 - \frac{1}{n} \text{tr} \bar{\mathbf{Q}} \frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} + 1 - \frac{1}{n} \text{tr} \bar{\mathbf{Q}} \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}} + \frac{\lambda}{n} \text{tr} \bar{\mathbf{Q}} \left(\frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} + \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}} \right) \bar{\mathbf{Q}} \right) \\
&\times \frac{N}{n} \frac{1}{n} \text{tr} \bar{\mathbf{Q}} \frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} \bar{\mathbf{Q}} \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}}
\end{aligned}$$

where we constantly use the fact that $\bar{\mathbf{Q}} \frac{N}{n} \left(\frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} + \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}} \right) = \mathbf{I}_n - \lambda \bar{\mathbf{Q}}$. Note that

$$\begin{aligned}
1 - \frac{1}{n} \text{tr} \bar{\mathbf{Q}} \frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} &= \frac{1}{1+\delta_{\cos}} > 0, \quad 1 - \frac{1}{n} \text{tr} \bar{\mathbf{Q}} \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}} = \frac{1}{1+\delta_{\sin}} > 0 \\
\frac{1}{1+\delta_{\cos}} + \frac{1}{1+\delta_{\sin}} &= 2 - \frac{n}{N} + \frac{\lambda}{N} \text{tr} \bar{\mathbf{Q}} > 0
\end{aligned}$$

so that 1) $\det(\Omega^{-1}) > 0$ and 2) for $2N = n$, $\det(\Omega^{-1})$ scales like λ as $\lambda \rightarrow 0$. \square

Lemma 6 (Derivatives with respect to N). *Let Assumption 1 holds, for any $\lambda > 0$ and*

$$\begin{cases} \delta_{\cos} = \frac{1}{n} \text{tr}(\mathbf{K}_{\cos} \bar{\mathbf{Q}}) = \frac{1}{n} \text{tr} \mathbf{K}_{\cos} \left(\frac{N}{n} \left(\frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} + \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}} \right) + \lambda \mathbf{I}_n \right)^{-1} \\ \delta_{\sin} = \frac{1}{n} \text{tr}(\mathbf{K}_{\sin} \bar{\mathbf{Q}}) = \frac{1}{n} \text{tr} \mathbf{K}_{\sin} \left(\frac{N}{n} \left(\frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} + \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}} \right) + \lambda \mathbf{I}_n \right)^{-1} \end{cases}$$

defined in Theorem 1, we have that $(\delta_{\cos}, \delta_{\sin})$ and $\|\bar{\mathbf{Q}}\|$ are all decreasing functions of N . Note in particular that the same conclusion holds for $2N > n$ as $\lambda \rightarrow 0$.

Proof. We write

$$\begin{bmatrix} \frac{\partial \delta_{\cos}}{\partial N} \\ \frac{\partial \delta_{\sin}}{\partial N} \end{bmatrix} = -\frac{1}{n} \Omega \begin{bmatrix} \frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \Phi \bar{\mathbf{Q}} \mathbf{K}_{\cos}) \\ \frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \Phi \bar{\mathbf{Q}} \mathbf{K}_{\sin}) \end{bmatrix} = -\frac{n}{N} \frac{1}{n} \Omega \begin{bmatrix} \delta_{\cos} - \frac{\lambda}{n} \text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}}) \\ \delta_{\sin} - \frac{\lambda}{n} \text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\sin} \bar{\mathbf{Q}}) \end{bmatrix} \quad (23)$$

for Ω defined in (8) and $\Phi = \frac{\mathbf{K}_{\cos}}{1+\delta_{\cos}} + \frac{\mathbf{K}_{\sin}}{1+\delta_{\sin}}$, which, together with Lemma 5, allows us to conclude that $\frac{\partial \delta_{\cos}}{\partial N}, \frac{\partial \delta_{\sin}}{\partial N} < 0$. Further note that

$$\frac{\partial \bar{\mathbf{Q}}}{\partial N} = -\frac{1}{n} \bar{\mathbf{Q}} \left(\Phi - \frac{\mathbf{K}_{\cos}}{(1+\delta_{\cos})^2} N \frac{\partial \delta_{\cos}}{\partial N} - \frac{\mathbf{K}_{\sin}}{(1+\delta_{\sin})^2} N \frac{\partial \delta_{\sin}}{\partial N} \right) \bar{\mathbf{Q}}$$

which concludes the proof. \square

Lemma 7 (Derivative with respect to λ). *For any $\lambda > 0$, $(\delta_{\cos}, \delta_{\sin})$ and $\|\bar{\mathbf{Q}}\|$ defined in Theorem 1 decrease as λ grows large.*

Proof. Taking the derivative of $(\delta_{\cos}, \delta_{\sin})$ with respect to $\lambda > 0$, we have explicitly

$$\begin{bmatrix} \frac{\partial \delta_{\cos}}{\partial \lambda} \\ \frac{\partial \delta_{\sin}}{\partial \lambda} \end{bmatrix} = -\Omega \begin{bmatrix} \frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\cos} \bar{\mathbf{Q}}) \\ \frac{1}{n} \text{tr}(\bar{\mathbf{Q}} \mathbf{K}_{\sin} \bar{\mathbf{Q}}) \end{bmatrix} \quad (24)$$

which, together with the fact that all entries of Ω are positive (Lemma 5), allows us to conclude that $\frac{\partial \delta_{\cos}}{\partial \lambda}, \frac{\partial \delta_{\sin}}{\partial \lambda} < 0$. Further considering

$$\frac{\partial \bar{\mathbf{Q}}}{\partial \lambda} = \bar{\mathbf{Q}} \left(\frac{N}{n} \frac{\mathbf{K}_{\cos}}{(1+\delta_{\cos})^2} \frac{\partial \delta_{\cos}}{\partial \lambda} + \frac{N}{n} \frac{\mathbf{K}_{\sin}}{(1+\delta_{\sin})^2} \frac{\partial \delta_{\sin}}{\partial \lambda} - \mathbf{I}_n \right) \bar{\mathbf{Q}}$$

and thus the conclusion for $\bar{\mathbf{Q}}$. \square

F Additional Real-world Data sets

We have presented results in detail for one particular real-world data set, the MNIST data set, but we have extensive empirical results demonstrating that similar conclusions hold more broadly. As an example of this, here we present numerical evaluations of our results on several other real-world image data sets. We consider the classification task on another two MNIST-like data sets composed of 28×28 grayscale images: the Fashion-MNIST [53] and the Kannada-MNIST [42] data sets. Each image is represented as a $p = 784$ -dimensional vector and the output targets $\mathbf{y}, \hat{\mathbf{y}}$ are taken to have $-1, +1$ entries depending on the image class. As a consequence, both the training and test MSEs defined in (4) are approximately 1 for $N = 0$ and not-too-large regularization λ , as observed in Figure 4 and Figure 10 below. For each data set, images were jointly centered and scaled so to fall close to the setting of Assumption 1 on \mathbf{X} and $\hat{\mathbf{X}}$.

In Figure 8, we compare the empirical training and test NSEs with their limiting behaviors derived from Theorem 2 and 3, as a function of the penalty parameter λ , on a training set of size $n = 1024$ (512 images from class 5 and 512 images from class 6) with feature dimension $N = 256$, on both data sets. A close fit between theory and practice is observed, for moderately large values of n, p, N , demonstrating thus a wide practical applicability of the proposed asymptotic analyses, particularly compared to the (limiting) Gaussian kernel predictions per Figure 1.

In Figure 9, we report the behavior of the pair $(\delta_{\cos}, \delta_{\sin})$ for small values of $\lambda = 10^{-7}$ and 10^{-3} . Similar to the two leftmost plots in Figure 3 for MNIST, a jump from the under- to over-parameterized

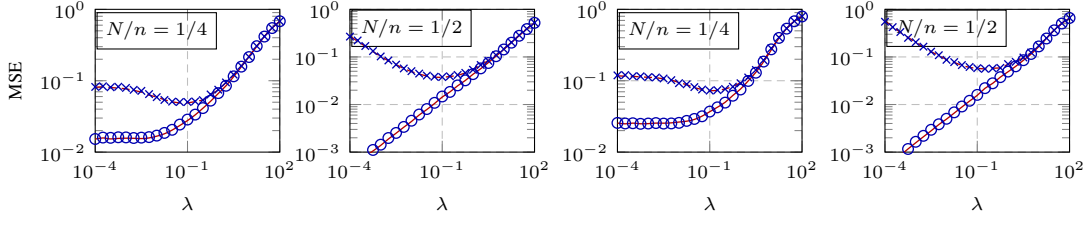


Figure 8: MSEs of RFF regression on Fashion-MNIST (**left two**) and Kannada-MNIST (**right two**) data (class 5 versus 6), as a function of regression parameter λ , for $p = 784$, $n = \hat{n} = 1024$, $N = 256$ and 512 . Empirical results displayed in **blue** (circles for training and crosses for test); and the asymptotics from Theorem 2 and 3 displayed in **red** (solid lines for training and dashed for test). Results obtained by averaging over 30 runs.

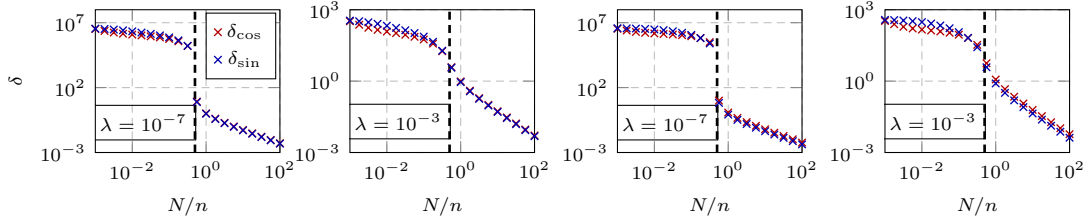


Figure 9: Behavior of $(\delta_{\cos}, \delta_{\sin})$ in (11), on Fashion-MNIST (**left two**) and Kannada-MNIST (**right two**) data (class 8 versus 9), for $p = 784$, $n = 1000$, $\lambda = 10^{-7}$ and 10^{-3} . The **black** dashed line is the interpolation threshold $2N = n$.

regime occurs at the interpolation threshold $2N = n$, in both Fashion- and Kannada-MNIST data sets, clearly indicating the two phases of learning and the phase transition between them.

In Figure 10, we report the empirical and theoretical test errors as a function of the ratio N/n , on a training test of size $n = 500$ (250 images from class 8 and 250 images from class 9), by varying feature dimension N . An exceedingly small regularization $\lambda = 10^{-7}$ is applied to mimic the “ridgeless” limiting behavior as $\lambda \rightarrow 0$. On both data sets, double-descent-type test curves are observed where the test errors goes down and up, with a singular peak around $2N = n$, and then goes down monotonically as N continues to increase when $2N > n$.

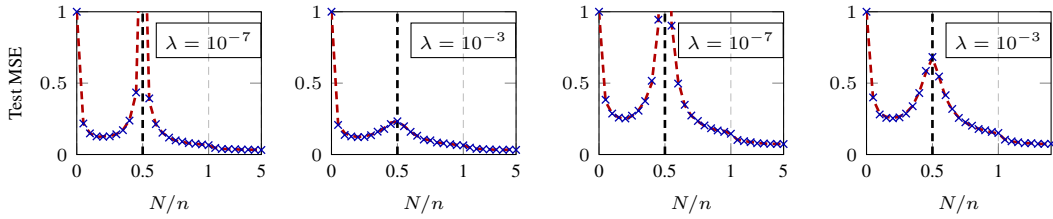


Figure 10: Empirical (crosses) and theoretical (dashed lines) test error of RFF regression, as a function of the ratio N/n , on Fashion-MNIST (**left two**) and Kannada-MNIST (**right two**) data (class 8 versus 9), for $p = 784$, $n = 500$, $\lambda = 10^{-7}$ and 10^{-3} . The **black** dashed line is the interpolation threshold $2N = n$. Results obtained by averaging over 30 runs.