

Great Likelihood Estimation of the Multivariate Normal Distribution

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Multivariate normal distribution

Give the probability density function of the normal distribution in multivariate form($\mathbf{x} \sim N_p(\mathbf{u}, \mathbf{\Sigma})$):

$$f(\mathbf{x}) = \frac{1}{(2\pi)^{\frac{p}{2}} |\mathbf{\Sigma}|^{\frac{1}{2}}} \exp\left[-\frac{1}{2}(\mathbf{x} - \mathbf{u})^T \mathbf{\Sigma}^{-1}(\mathbf{x} - \mathbf{u})\right]$$

s.t. $\mathbf{x} \in \mathbb{R}^p, \mathbf{u} \in \mathbb{R}^p, \mathbf{\Sigma} \in \mathcal{S}_+^p$.

Then for the sample set $\mathcal{D} = \{(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_n, y_n)\}$, s.t. $\mathbf{x}_i \in \mathbb{R}^n, y_i \in \mathbb{R}, |\mathcal{D}| = n$. With reference to the definition, the great likelihood function is given:

$$\begin{aligned} L(\mathbf{u}, \mathbf{\Sigma}) &= \prod_{i=1}^n f(\mathbf{x}_i) \\ &= \frac{1}{(2\pi)^{\frac{pn}{2}} |\mathbf{\Sigma}|^{\frac{n}{2}}} \exp\left[-\frac{1}{2} \sum_{i=1}^n (\mathbf{x}_i - \mathbf{u})^T \mathbf{\Sigma}^{-1}(\mathbf{x}_i - \mathbf{u})\right] \end{aligned}$$

Take the logarithm of it and simplify it:

$$\ln L(\mathbf{u}, \mathbf{\Sigma}) = -\frac{pn}{2} \ln(2\pi) - \frac{n}{2} \ln |\mathbf{\Sigma}| - \frac{1}{2} \sum_{i=1}^n (\mathbf{x}_i - \mathbf{u})^T \mathbf{\Sigma}^{-1}(\mathbf{x}_i - \mathbf{u})$$

Matrix derivation formula

Give the derivation formulas that may be used:

$$\frac{\partial \ln |\mathbf{X}|}{\partial \mathbf{X}} = \mathbf{X}^{-T} \quad (1)$$

$$\frac{\partial \ln(\boldsymbol{\lambda}^T \mathbf{X}^{-1} \boldsymbol{\lambda})}{\partial \mathbf{X}} = -(\mathbf{X}^{-1} \boldsymbol{\lambda} \boldsymbol{\lambda}^T \mathbf{X}^{-1})^T \quad (2)$$

$$\frac{\partial (\boldsymbol{\lambda} - \mathbf{x})^T \boldsymbol{\Sigma}^{-1} (\boldsymbol{\lambda} - \mathbf{x})}{\partial \mathbf{x}} = [(\boldsymbol{\lambda} - \mathbf{x})^T (\boldsymbol{\Sigma}^{-T} + \boldsymbol{\Sigma}^{-1})]^T \quad (3)$$

Then write $\frac{\partial f(\mathbf{X})}{\partial \mathbf{X}} = \mathbf{A}$, which $\mathbf{X} \in \mathbb{R}^{n \times n} : \mathbb{R}^{n \times n} \rightarrow \mathbb{R}$. For a symmetric matrix $\mathbf{X} \in S^n$, We have the following form:

$$\frac{\partial f(\mathbf{X})}{\partial \mathbf{X}} = \mathbf{A}^T + \mathbf{A} - \mathbf{A} \circ \mathbf{E} \quad (4)$$

\circ is Hadamard product, proof is not given here.

It is known that $\mathbf{X} \in \mathcal{S}^n$, then (1)、(2) can transfer to (5)、(6):

$$\frac{\partial \ln |\mathbf{X}|}{\partial \mathbf{X}} = \mathbf{X}^{-1} + \mathbf{X}^{-T} - \mathbf{X}^{-T} \circ \mathbf{E} \quad (5)$$

$$\begin{aligned} \frac{\partial \ln(\boldsymbol{\lambda}^T \mathbf{X}^{-1} \boldsymbol{\lambda})}{\partial \mathbf{X}} &= -(\mathbf{X}^{-1} \boldsymbol{\lambda} \boldsymbol{\lambda}^T \mathbf{X}^{-1}) \\ &\quad - (\mathbf{X}^{-1} \boldsymbol{\lambda} \boldsymbol{\lambda}^T \mathbf{X}^{-1})^T \\ &\quad + (\mathbf{X}^{-1} \boldsymbol{\lambda} \boldsymbol{\lambda}^T \mathbf{X}^{-1})^T \circ \mathbf{E} \end{aligned} \quad (6)$$

From equation (3):

$$\begin{aligned} \frac{\partial \ln L(\mathbf{u}, \boldsymbol{\Sigma})}{\partial \mathbf{u}} &= -\frac{1}{2} \sum_{i=0}^n \frac{\partial (\mathbf{x}_i - \mathbf{u})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x}_i - \mathbf{u})}{\partial \mathbf{u}} \\ &= \sum_{i=1}^n \boldsymbol{\Sigma}^{-1} (\mathbf{x}_i - \mathbf{u}) \\ &= \boldsymbol{\Sigma}^{-1} \sum_{i=1}^n (\mathbf{x}_i - \mathbf{u}) \end{aligned} \quad (7)$$

From equation (5)、(6):

$$\begin{aligned}
\frac{\partial \ln L(\mathbf{u}, \Sigma)}{\partial \Sigma} &= -\frac{n}{2} \frac{\partial \ln |\Sigma|}{\partial \Sigma} - \frac{1}{2} \sum_{i=0}^n \frac{\partial (\mathbf{x}_i - \mathbf{u})^T \Sigma^{-1} (\mathbf{x}_i - \mathbf{u})}{\partial \Sigma} \\
&= \frac{1}{2} \left[\sum_{i=0}^n (\Sigma^{-1} (\mathbf{x}_i - \mathbf{u}) (\mathbf{x}_i - \mathbf{u})^T \Sigma^{-1} - n \Sigma^{-1})^T \right. \\
&\quad + \sum_{i=0}^n (\Sigma^{-1} (\mathbf{x}_i - \mathbf{u}) (\mathbf{x}_i - \mathbf{u})^T \Sigma^{-1} - n \Sigma^{-1}) \\
&\quad \left. - \sum_{i=0}^n (\Sigma^{-1} (\mathbf{x}_i - \mathbf{u}) (\mathbf{x}_i - \mathbf{u})^T \Sigma^{-1} - n \Sigma^{-1}) \circ \mathbf{E} \right] \\
&= \sum_{i=0}^n (\Sigma^{-1} (\mathbf{x}_i - \mathbf{u}) (\mathbf{x}_i - \mathbf{u})^T \Sigma^{-1} - n \Sigma^{-1}) \\
&\quad - \frac{1}{2} \sum_{i=0}^n (\Sigma^{-1} (\mathbf{x}_i - \mathbf{u}) (\mathbf{x}_i - \mathbf{u})^T \Sigma^{-1} - n \Sigma^{-1}) \circ \mathbf{E}
\end{aligned} \tag{8}$$

let (7) equals $\mathbf{0}$, multiply the left and right sides by Σ^{-1} , then $\hat{\mathbf{u}} = \sum_{i=0}^n \mathbf{x}_i = \bar{\mathbf{x}}$

Take the result into equation (8), and let (8) equals $\mathbf{0}$, then $\hat{\Sigma} = \frac{1}{n} \sum_{i=1}^n (\mathbf{x}_i - \bar{\mathbf{x}})(\mathbf{x}_i - \bar{\mathbf{x}})^T$

It can be seen that the great likelihood estimate of the multivariate normal distribution $\hat{\mathbf{u}}$ is the sample mean, $\hat{\Sigma}$ is the covariance matrix of the sample.