

多元正态分布的极大似然估计

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多元正态分布

给定多元形式下的正态分布概率密度函数 ($\mathbf{x} \sim N_p(\mathbf{u}, \Sigma)$):

$$f(\mathbf{x}) = \frac{1}{(2\pi)^{\frac{p}{2}} |\Sigma|^{\frac{1}{2}}} \exp\left[-\frac{1}{2}(\mathbf{x} - \mathbf{u})^T \Sigma^{-1}(\mathbf{x} - \mathbf{u})\right]$$

其中 $\mathbf{x} \in \mathbb{R}^p, \mathbf{u} \in \mathbb{R}^p, \Sigma \in \mathcal{S}_+^p$ 。

接着对于样本集 $\mathcal{D} = \{(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_n, y_n)\}$, 其中 $\mathbf{x}_i \in \mathbb{R}^n, y_i \in \mathbb{R}, |\mathcal{D}| = n$ 。

参照定义, 给出极大似然函数:

$$\begin{aligned} L(\mathbf{u}, \Sigma) &= \prod_{i=1}^n f(\mathbf{x}_i) \\ &= \frac{1}{(2\pi)^{\frac{pn}{2}} |\Sigma|^{\frac{n}{2}}} \exp\left[-\frac{1}{2} \sum_{i=1}^n (\mathbf{x}_i - \mathbf{u})^T \Sigma^{-1}(\mathbf{x}_i - \mathbf{u})\right] \end{aligned}$$

对其取对数, 并进行化简:

$$\ln L(\mathbf{u}, \Sigma) = -\frac{pn}{2} \ln(2\pi) - \frac{n}{2} \ln |\Sigma| - \frac{1}{2} \sum_{i=1}^n (\mathbf{x}_i - \mathbf{u})^T \Sigma^{-1}(\mathbf{x}_i - \mathbf{u})$$

矩阵求导公式

给出可能会用到的求导公式:

$$\frac{\partial \ln |\mathbf{X}|}{\partial \mathbf{X}} = \mathbf{X}^{-T} \quad (1)$$

$$\frac{\partial \ln(\boldsymbol{\lambda}^T \mathbf{X}^{-1} \boldsymbol{\lambda})}{\partial \mathbf{X}} = -(\mathbf{X}^{-1} \boldsymbol{\lambda} \boldsymbol{\lambda}^T \mathbf{X}^{-1})^T \quad (2)$$

$$\frac{\partial (\boldsymbol{\lambda} - \mathbf{x})^T \Sigma^{-1}(\boldsymbol{\lambda} - \mathbf{x})}{\partial \mathbf{x}} = [(\boldsymbol{\lambda} - \mathbf{x})^T (\Sigma^{-T} + \Sigma^{-1})]^T \quad (3)$$

接着记 $\frac{\partial f(\mathbf{X})}{\partial \mathbf{X}} = \mathbf{A}$, 其中 $\mathbf{X} \in \mathbb{R}^{n \times n}$, $f: \mathbb{R}^{n \times n} \rightarrow \mathbb{R}$ 。那么对于对称矩阵 $\mathbf{X} \in S^n$, 我们有如下形式:

$$\frac{\partial f(\mathbf{X})}{\partial \mathbf{X}} = \mathbf{A}^T + \mathbf{A} - \mathbf{A} \circ \mathbf{E} \quad (4)$$

其中 \circ 为 Hadamard product, 此处不给出证明。

已知 $\mathbf{X} \in S^n$, 则式(1), (2) 可以变成如下的 (5), (6):

$$\frac{\partial \ln |\mathbf{X}|}{\partial \mathbf{X}} = \mathbf{X}^{-1} + \mathbf{X}^{-T} - \mathbf{X}^{-T} \circ \mathbf{E} \quad (5)$$

$$\begin{aligned} \frac{\partial \ln(\boldsymbol{\lambda}^T \mathbf{X}^{-1} \boldsymbol{\lambda})}{\partial \mathbf{X}} &= -(\mathbf{X}^{-1} \boldsymbol{\lambda} \boldsymbol{\lambda}^T \mathbf{X}^{-1}) \\ &\quad - (\mathbf{X}^{-1} \boldsymbol{\lambda} \boldsymbol{\lambda}^T \mathbf{X}^{-1})^T \\ &\quad + (\mathbf{X}^{-1} \boldsymbol{\lambda} \boldsymbol{\lambda}^T \mathbf{X}^{-1})^T \circ \mathbf{E} \end{aligned} \quad (6)$$

由式(3)可得:

$$\begin{aligned} \frac{\partial \ln L(\mathbf{u}, \boldsymbol{\Sigma})}{\partial \mathbf{u}} &= -\frac{1}{2} \sum_{i=0}^n \frac{\partial (\mathbf{x}_i - \mathbf{u})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x}_i - \mathbf{u})}{\partial \mathbf{u}} \\ &= \sum_{i=1}^n \boldsymbol{\Sigma}^{-1} (\mathbf{x}_i - \mathbf{u}) \\ &= \boldsymbol{\Sigma}^{-1} \sum_{i=1}^n (\mathbf{x}_i - \mathbf{u}) \end{aligned} \quad (7)$$

由式 (5),(6)可得:

$$\begin{aligned}
\frac{\partial \ln L(\mathbf{u}, \Sigma)}{\partial \Sigma} &= -\frac{n}{2} \frac{\partial \ln |\Sigma|}{\partial \Sigma} - \frac{1}{2} \sum_{i=0}^n \frac{\partial (\mathbf{x}_i - \mathbf{u})^T \Sigma^{-1} (\mathbf{x}_i - \mathbf{u})}{\partial \Sigma} \\
&= \frac{1}{2} \left[\sum_{i=0}^n (\Sigma^{-1} (\mathbf{x}_i - \mathbf{u}) (\mathbf{x}_i - \mathbf{u})^T \Sigma^{-1} - n \Sigma^{-1})^T \right. \\
&\quad + \sum_{i=0}^n (\Sigma^{-1} (\mathbf{x}_i - \mathbf{u}) (\mathbf{x}_i - \mathbf{u})^T \Sigma^{-1} - n \Sigma^{-1}) \\
&\quad \left. - \sum_{i=0}^n (\Sigma^{-1} (\mathbf{x}_i - \mathbf{u}) (\mathbf{x}_i - \mathbf{u})^T \Sigma^{-1} - n \Sigma^{-1}) \circ \mathbf{E} \right] \\
&= \sum_{i=0}^n (\Sigma^{-1} (\mathbf{x}_i - \mathbf{u}) (\mathbf{x}_i - \mathbf{u})^T \Sigma^{-1} - n \Sigma^{-1}) \\
&\quad - \frac{1}{2} \sum_{i=0}^n (\Sigma^{-1} (\mathbf{x}_i - \mathbf{u}) (\mathbf{x}_i - \mathbf{u})^T \Sigma^{-1} - n \Sigma^{-1}) \circ \mathbf{E}
\end{aligned} \tag{8}$$

令式(7)等于 $\mathbf{0}$, 左右两边乘以 Σ^{-1} , 可得 $\hat{\mathbf{u}} = \sum_{i=0}^n \mathbf{x}_i = \bar{\mathbf{x}}$

把结果带入式(8), 并令式(8)等于 $\mathbf{0}$, 可得 $\hat{\Sigma} = \frac{1}{n} \sum_{i=1}^n (\mathbf{x}_i - \bar{\mathbf{x}})(\mathbf{x}_i - \bar{\mathbf{x}})^T$
可以看出多元正态分布的极大似然估计值 $\hat{\mathbf{u}}$ 为样本的均值, 而 $\hat{\Sigma}$ 为样本的协方差矩阵