

## 1 Concentration inequality for Poisson distribution

Let  $X$  be the sum of 20 i.i.d. Poisson random variables  $X_1, \dots, X_{20}$  with  $\mathbb{E}[X_1] = 1$ . Use the following techniques to upper bound  $\Pr(X \geq 26)$ .

1. Markov's Inequality
2. Chebyshev's Inequality
3. Chernoff Bound

(Hint: if  $X, Y$  are independent Poisson random variables with parameter  $\lambda_1, \lambda_2$ , then  $Z = X + Y$  is a Poisson random variable with parameter  $\lambda_1 + \lambda_2$ ).

**Solution:** From the hint, we know that  $X$  is a Poisson random variable with parameter  $\lambda = 20$ . Thus we know that

$$\begin{aligned}\mathbb{E}[X] &= 20 \\ \text{Var}[X] &= 20 \\ \forall t > 0, \mathbb{E}[\exp(sX)] &= \sum_{n=0}^{\infty} \frac{\lambda^n \exp(-\lambda)}{n!} \exp(sn) \\ &= \exp(-\lambda) \sum_{n=0}^{\infty} \frac{(\lambda \exp(s))^n}{n!} \\ &= \exp(-\lambda) \exp(\lambda \exp(s)) \\ &= \exp(\lambda(\exp(s) - 1)) \\ &= \exp(20(\exp(s) - 1))\end{aligned}$$

1. Using Markov's Inequality:

$$\Pr(X \geq a) \leq \frac{\mathbb{E}[X]}{a},$$

for all  $a > 0$ . So,

$$\Pr(X \geq 26) \leq \frac{20}{26} \approx 0.769.$$

2. Using Chebyshev's Inequality:

$$\Pr(|X - \mathbb{E}[X]| \geq c) \leq \frac{\text{Var}[X]}{c^2},$$

so we have

$$\Pr(|X - 20| \geq 6) \leq \frac{20}{36} \approx 0.5556.$$

3. Using the Chernoff Bound we have

$$\Pr(X \geq 26) \leq \min_{s \geq 0} e^{-26s} e^{20(e^s - 1)} = \min_{s \geq 0} e^{-26s + 20e^s} e^{-20}.$$

To minimize this function we only have to minimize  $-26s + 20e^s$ . Taking the derivative and setting it equal to zero, we have

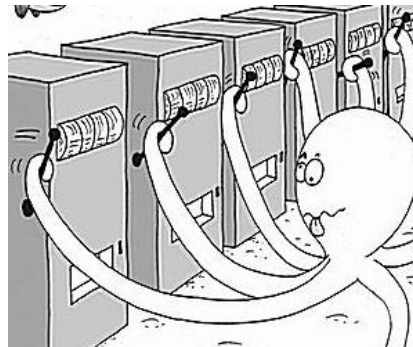
$$\begin{aligned} -26 + 20e^s &= 0, \\ e^s &= \frac{26}{20}, \\ s &= \ln \frac{26}{20} \approx 0.26236. \end{aligned}$$

Plugging in  $s$  to the above equation, we have

$$\Pr(X \geq 26) \leq e^{-26s + 20e^s - 20} \Big|_{s=\ln(13/10)} \approx 0.4398.$$

## 2 Multi-Armed Bandits: UCB and Hoeffding's inequality

In the Tuesday's lecture, we began talking about multi-armed bandits. In the multi-armed bandits setting, we consider a decision-maker who is given  $K$  options to choose from. We refer to these options as “arms”. Associated with each arm is a probability distribution over rewards. Initially, this distribution is unknown to the decision-maker. The decision-maker chooses an arm, usually referred to as pulling an arm, and receives a reward sampled from the corresponding reward distribution. This process is repeated over and over again.



The problem we want to solve is to decide which arm to pull at each time step. One possible algorithm for deciding which arm to pull is the Upper Confidence Bound (UCB) algorithm presented in lecture. If we assume that the reward of each arm is bounded (e.g. the slot machine returns between \$0 and \$100), then the UCB algorithm has bounded regret over time.

In this discussion, we study the derivation of the UCB algorithm using the Hoeffding bound. We will assume that the reward of each arm is bounded.

1. We first set up the framework of a multi-armed bandit problem. Suppose you have a set of  $K$  “arms”,  $\mathcal{A} = \{1, 2, \dots, K\}$ . Each arm  $a \in \mathcal{A}$  has its own reward distribution  $X_a \sim \mathbb{P}_a$  with mean  $\mu_a = \mathbb{E}[X_a]$ . Define the number of times arm  $a$  has been pulled up to and including time  $t$  as  $T_a(t)$ . In these problems we do not know  $\mu_a$  but we would like to efficiently find the arm with the maximum mean by creating an algorithm that balances *exploration* of the arms with *exploitation* of the best possible arm. The efficiency of the algorithm is measured by a theoretical quantity known as regret, which measures how well the algorithm performs in expectation against an ‘oracle’ that knows the means of all the arms and always pulls the arm with highest mean.

We will now derive the upper confidence bound that yields the UCB algorithm. The general formula for constructing an upper confidence bound for the true mean  $\mu_a$  of an “arm”  $a$ , given  $T_a(t)$  samples  $X_a^{(1)}, \dots, X_a^{(T_a(t))}$ , is to find a value of  $C_a(T_a(t), \delta)$  such that:

$$P(\mu_a < \hat{\mu}_{a,T_a(t)} + C_a(T_a(t), \delta)) > 1 - \delta \quad (1)$$

where  $\hat{\mu}_{a,T_a(t)}$  is the sample mean of the reward from arm  $a$ , given by  $\hat{\mu}_{a,T_a(t)} = \frac{1}{T_a(t)} \sum_{i=1}^{T_a(t)} X_a^{(i)}$ . In words, Equation (1) says that with probability at least  $1 - \delta$ , the true mean  $\mu_a$  is less than the estimated mean  $\hat{\mu}_{a,T_a(t)}$  plus an upper confidence bound  $C_a(T_a(t), \delta)$ .

- (a) Suppose that you know that the reward of any arm is between 0 and 1. That is:

$$X_a \in [0, 1]$$

Construct an upper confidence bound  $C_a(T_a(t), \delta)$  for the mean of arm  $a$ , after observing  $T_a(t)$  samples from arm  $a$ .

**Solution:** Hoeffding’s inequality bounds the difference between  $\hat{\mu}_{a,T_a(t)}$  and  $\mu_a$ . Hoeffding’s inequality works in two directions:

$$P(\hat{\mu}_{a,T_a(t)} - \mu_a \geq \epsilon) \leq e^{-2T_a(t)\epsilon^2} \quad (2)$$

$$P(\hat{\mu}_{a,T_a(t)} - \mu_a \leq -\epsilon) \leq e^{-2T_a(t)\epsilon^2} \quad (3)$$

The goal is to find the unknown quantity  $C_a(T_a(t), \delta)$  from Equation (1) in terms of  $\delta$  and  $T_a(t)$ . We now rearrange Equation (1) so that we can apply Hoeffding’s inequality.

$$\begin{aligned} P(\mu_a < \hat{\mu}_{a,T_a(t)} + C_a(T_a(t), \delta)) &> 1 - \delta \\ P(\hat{\mu}_{a,T_a(t)} - \mu_a > -C_a(T_a(t), \delta)) &> 1 - \delta \end{aligned}$$

Subtracting both sides from 1,

$$\begin{aligned} 1 - P(\hat{\mu}_{a,T_a(t)} - \mu_a > -C_a(T_a(t), \delta)) &\leq \delta \\ \implies P(\hat{\mu}_{a,T_a(t)} - \mu_a \leq -C_a(T_a(t), \delta)) &\leq \delta \end{aligned}$$

This looks a lot more like the form of Hoeffding's inequality in Equation (3). Remember that the goal here is to compute the unknown quantity  $C_a(t, \delta)$  in terms of  $\delta$  and  $t$ . If we replace  $\epsilon$  in Hoeffding's inequality (Equation (3)) with  $C_a(T_a(t), \delta)$ , we have

$$P(\hat{\mu}_{a, T_a(t)} - \mu_a \leq -C_a(T_a(t), \delta)) \leq e^{-2T_a(t)C_a(T_a(t), \delta)^2} \quad (4)$$

To compute the value of  $C_a(T_a(t), \delta)$  that achieves the desired probability bound of  $\delta$  from Equation (1), we can set the right hand side of Equation (4) to  $\delta$ :

$$\begin{aligned} e^{-2T_a(t)C_a(T_a(t), \delta)^2} &= \delta \\ \implies C_a(T_a(t), \delta) &= \sqrt{\frac{-\log \delta}{2T_a(t)}} \end{aligned}$$

Plugging this value of  $C_a(T_a(t), \delta)$  back into Equation (1), we have

$$P\left(\mu_a < \hat{\mu}_{a, T_a(t)} + \sqrt{\frac{-\log \delta}{2T_a(t)}}\right) > 1 - \delta.$$

- (b) Suppose we set  $\delta = \frac{1}{t^3}$ . This controls the probability that the true mean  $\mu_a$  is greater than our upper confidence bound  $C_a(T_a(t), \delta)$  on the estimated mean  $\hat{\mu}_{a, T_a(t)}$ . What rule does the UCB algorithm use to choose an arm  $A_t$  at each iteration  $t$ ?

**Solution:** The UCB algorithm optimistically chooses the arm with the highest estimated mean plus upper confidence bound. At time step  $t$ , we have access to mean and upper confidence bound estimates from up to time  $t - 1$ . Therefore, at time step  $t$ , the UCB algorithm chooses arm  $A_t$  as follows:

$$A_t = \operatorname{argmax}_{a \in \mathcal{A}} \hat{\mu}_{a, T_a(t-1)} + C_a(T_a(t-1), \delta)$$

Plugging in the upper confidence bound  $C_a(T_a(t-1), \delta)$  from part (a), we have

$$A_t = \operatorname{argmax}_{a \in \mathcal{A}} \hat{\mu}_{a, T_a(t-1)} + \sqrt{\frac{-\log \delta}{2T_a(t-1)}}$$

Substituting in our choice of small  $\delta = \frac{1}{t^3}$ ,

$$\begin{aligned} A_t &= \operatorname{argmax}_{a \in \mathcal{A}} \hat{\mu}_{a, T_a(t-1)} + \sqrt{\frac{-\log \frac{1}{t^3}}{2T_a(t-1)}} \\ \implies A_t &= \operatorname{argmax}_{a \in \mathcal{A}} \hat{\mu}_{a, T_a(t-1)} + \sqrt{\frac{3 \log t}{2T_a(t-1)}}. \end{aligned}$$