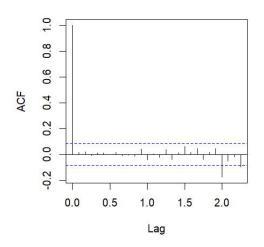
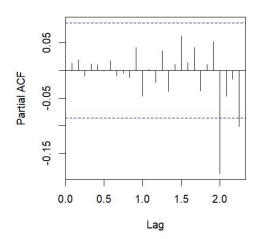
```
1.
```

```
> aic_diffADL <- numeric()</pre>
 for (i in 1:15) {
     adl_model <- dynlm(diff_Yt ~ lag(diff_Yt, i)
+ diff_OIL_P + lag(diff_OIL_P, i)
+ diff_OIL_PROD + lag(diff_OIL_PROD, i)
+ diff_OIL_STOCKS + lag(diff_OIL_STOCKS,
+
+
 i)
                                + diff_WORLD_IP + lag(diff_WORLD_IP, i)
+ diff_US_CPI+ lag(diff_US_CPI, i)
+
+
                                   data = data
     aic_diffADL[i] <- AIC(adl_model)</pre>
> best_AICorder <- which.min(aic_diffADL)</pre>
> print(best_AICorder)
[1] 9
> Bic_diffADL <- numeric()</pre>
> for (i in 1:15) {
     adl_model <- dynlm(diff_Yt ~ lag(diff_Yt, i)
+ diff_OIL_P + lag(diff_OIL_P, i)
+ diff_OIL_PROD + lag(diff_OIL_PROD, i)
+ diff_OIL_STOCKS + lag(diff_OIL_STOCKS,
+
+
+
+
 i)
                                + diff_WORLD_IP + lag(diff_WORLD_IP, i)
                                + diff_US_CPI+ lag(diff_US_CPI, i)
                                   data = data
     Bic_diffADL[i] <- BIC(adl_model)</pre>
  }
> best_BICorder <- which.min(Bic_diffADL)</pre>
> print(best_BICorder)
[1] 9
stargazer(ADL_9,type = "text")
                                         Dependent variable:
                                                 diff_Yt
                                                    518
Observations
                                                   0.532
R2
                                                   0.472
Adjusted R2
                                      0.526 (df = 458)
8.823*** (df = 59; 458)
Residual Std. Error
F Statistic
                                   *p<0.1; **p<0.05; ***p<0.01
Note:
```

#### Series ADL\_9\$residuals

#### Series ADL\_9\$residuals

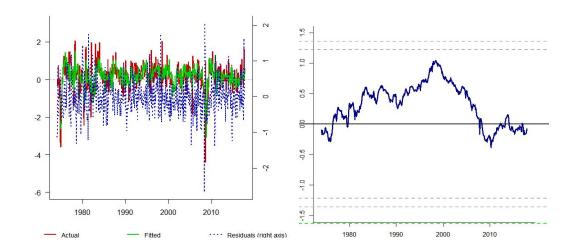




### > LjungBox(ADL\_9\$residuals,90,59)

1 2 3 4	1 2 3 4	0.0931 0.2966	NA NA	
2	2	0.2966 0.3427	NA NA	
1	4	0.4115	NA NA	
5	7	0 4730	NA NA	
6	6	0.4731	NA NA	
5 6 7	5 6 7	0.6307	NA NA	
8	8	0.4731 0.6307 0.6791 0.6891 0.7780	NA	
8 9	9	0.6891	NA	
10	10	0.7780	NA	
11	11	1.7067	NA	
12	10 11 12 13 14	1.7067 2.7902 2.7928 3.1070 3.8005	NA	
13	13	2.7928	NA	
14	14	3.1070	NA	
15	15	3.8005	NA	
16	16	4.5570	NA	
11 12 13 14 15 16 17 18 19 20 21 22 23 24 25	15 16 17 18 19 20 21 22 23 24 25 26	4.5570 4.6039 6.6305 6.7634 7.8589 8.6177	NA	
18	18	6.6305	NA	
19	19	6.7634	NA	
20	20	7.8589	NA	
21	21	8.6177	NA	
22	22	8.7558	NA	
23	23	8.7558 9.9298 26.8169 28.2201 28.4317 33.8216 35.7241 41.3947 42.8718 42.9850 43.8020 43.8723 43.9069	NA	
24	24	26.8169	NA	
25	25	28.2201	NA	
26	26	28.4317	NA	
27	27	33.8216	NA	
28	28	35.7241	NA	
29	29	41.3947	NA	
28 29 30 31 32	27 28 29 30 31 32 33 34	42.8/18	NA	
31	31	42.9850	NA	
32	32	43.8020	NA	
33	33	43.8/23	NA	
34	34	43.9069	NA	
35	35 36 37 38	43.9092	NA	
36 37	36	44.6362	NA	
3/	3/	44.7052	NA	
38	38	45.0020	NA	
39	39	45.3004	NA	
40	40	46.1999	NA	

41 42 43	41 42 43	46.2028 47.0674 47.2669	NA NA NA
44	44	47.2669 47.2743	NA
45	45	49.3025	NA
45 46	46	47.2669 47.2743 49.3025 50.1197 50.2510	NA
47	47	50.2510	NA
48	48	51.1866	NA
49	49	52.1228	NA
50	50	52.2342	NA
51	51	53.0309	NA
52	52	53.5732	NA
47 48 49 50 51 52 53 54 55 57 58	47 48 49 50 51 52 53 54 55 56 57 58	49.3025 50.1197 50.2510 51.1866 52.1228 52.2342 53.0309 53.5732 53.6836 58.8572 58.9393 60.3026 61.6636 63.1226	NA
54	54	58.85/2	NA NA
55	55	58.9393	NA
56	56	60.3026	NA
5 / 5 Ω	Ο/ 50	61.0030 62 1226	NA NA
50 50	50 50	64 0594	NA NA
60	60	64 3398	0.3272
60 61 62 63	60 61 62 63	63.1226 64.0594 64.3398 64.3539 64.9216 65.1769 66.1696 66.6154 66.6158 67.6358 68.5380 68.6327 68.8658	0 3600
62	62	64.9216	0.3752
63	63	65.1769	0.4009
64	64	66.1696	0.4019
65 66 67 68 69	65 66 67 68	66.6154	0 4212
66	66	66.9905	0.4212 0.4428 0.4629 0.4896 0.4931 0.5239 0.5239
67	67	67.4130	0.4629
68	68	67.6358	0.4896
69	69	68.5380	0.4931
70	70	68.6327	0.5239
71	71	68.5380 68.6327 68.8658 73.7267	0.5239 0.5497 0.4214 0.4424
72	72	73.7267	0.4214
73	73	74.0893	
74	74	74.1404	0.4735
75	/5	74.1825	0.5050
70 71 72 73 74 75 76 77 78 79 80	70 71 72 73 74 75 76 77 78 79	73.7267 74.0893 74.1404 74.1825 75.9113 75.9221 77.2344 78.5154	0.5050 0.4813 0.5133 0.5032
7/	7/	75.9221	0.5133
/ ŏ	/ ŏ	77.2344	0.5032 0.4942
/ Y	79 80	78.5154 78.5307	0.4942 0.5255
0U	٥0	/0.330/	0.3233



### Can you do standard inference based on the estimation results?

No I cannot do standard inference, because now the model is ADL in first difference and I did nt test if there are cointegration relationshps. The model would be misspecification if there is cointegration relationship which will make the estimation inconsistent and not asymtotically n ormal, thus standard inference is not possible.

# Compute the dynamic response of the growth rate of U.S. industrial production to a 10% increase in the oil price.

From the IRF\_calc, it can calculate two months dynamic impact of the 10% increase in the oil price. Pay attention, the 10% is permanent increase in oil price and transitory in oil price in fir st difference (and it is 10 unit shock).

## Outline possible disadvantages of the ADL model you have estimated.

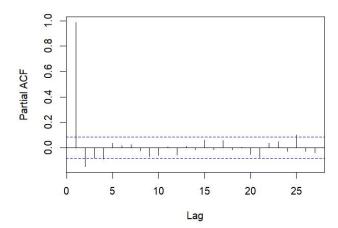
First, it used all the variables in first differences without testing for the cointegrations. if there is cointegrations, the model will be misspecified and inconsistent, and the estimations is not a symptotically normal, standard inference is not possible,

Second, it ignore the endogenity, which will cause the endogenous problem in the model, the error term and the explanatory variables will be correlated and the model is inconsistent. Also the endogenous problem will make the casual inference incorrectly identified.

In addition, all the variables choose the same lag, in fact this is not possible and there will be overfitting problem in the specification.

```
2.
 static_model_inlevel <- lm(Yt ~ OIL_P + OIL_PROD + OIL_STOCK</pre>
                                WORLD_IP + US_CPI)
 stargazer(static_model_inlevel, type = "text")
                          Dependent variable:
                               -0.069***
OIL_P
                                (0.005)
                                 0.094
OIL_PROD
                                (0.070)
                                -0.023
OIL_STOCKS
                                (0.032)
                               0.689***
WORLD_IP
                                (0.054)
                                 0.093
US_CPI
                                (0.057)
                                22.971
Constant
                               (70.752)
Observations
                                  528
                                 0.972
Adjusted R2
                                 0.972
Residual Std. Error
                          4.856 (df = 522)
F Statistic
                     3,669.979*** (df = 5; 522)
                     *p<0.1; **p<0.05; ***p<0.01
Note:
```

#### Series static\_model\_inlevel\_residuals



```
> AR2_residuals_static = arima(static_model_inlevel_residuals,
                             order=c(2,0,0), include.mean = TRUE)
> AR4_residuals_static = arima(static_model_inlevel_residuals,
                             order=c(4,0,0), include.mean = TRUE)
> LR_test_residuals_AR2_AR4 <- lrtest(AR2_residuals_static, AR</p>
4_residuals_static)> print(LR_test_residuals_AR2_AR4)
                        Likelihood ratio test
Model 1: arima(x = static_model_inlevel_residuals, order = c(2,
 0, 0),
    include.mean = TRUE)
Model 2: arima(x = static_model_inlevel_residuals, order = c(4,
    include.mean = TRUE)
  #Df
      LogLik Df Chisq Pr(>Chisq)
    4 -592.88
    6 - 589.93
               2 5.8936
                             0.05251 .
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
> LjungBox(AR2_residuals_static$residuals, 25, 2)
lag Ljung-Box p-value
           0.0622
                        NA
12345678
           2.9459
                       NA
          6.1464
     3
                   0.1047
     4
           6.6499
                   0.1556
     5
           6.7265
                   0.2418
     6
7
                   0.2983
           7.2505
           7.2570
                   0.4026
     8
                   0.2747
           9.8638
9
     9
          10.5997
                   0.3042
10
    10
         10.5997
                   0.3895
11
    11
         12.9352
                   0.2976
12
    12
         12.9416
                   0.3733
13
    13
         13.3746
                   0.4193
14
    14
         14.0105
                   0.4489
15
    15
         14.0460
                   0.5220
16
    16
         16.9694
                   0.3876
17
    17
         16.9738
                   0.4561
18
    18
         17.0937
                   0.5167
19
    19
         17.4898
                   0.5567
20
    20
         20.9262
                   0.4015
21
    21
         21.8423
                   0.4086
22
    22
                   0.3681
         23.6081
23
    23
         24.8282
                   0.3592
24
          32.5530
    24
                   0.1139
25
    25
         32.5537
                   0.1426
> ADF1_cointegration = ur.df(static_model_inlevel_residuals, l
ags = 1, type = "none"
> summary(ADF1_cointegration)
```

```
# Augmented Dickey-Fuller Test Unit Root Test #
Test regression none
call:
lm(formula = z.diff \sim z.lag.1 - 1 + z.diff.lag)
Residuals:
           1Q
               Median
              0.0336 0.4606
-4.0737 -0.4240
Coefficients:
          -0.012811
                            -1.900 0.058001
z.lag.1
                              3.578 0.000378 ***
z.diff.lag 0.154756
                    0.043249
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.7439 on 524 degrees of freedom
Multiple R-squared: 0.02841, Adjusted R-squared: F-statistic: 7.66 on 2 and 524 DF, p-value: 0.0005258
                             Adjusted R-squared: 0.0247
Value of test-statistic is: -1.8999
Critical values for test statistics:
     1pct 5pct 10pct
tau1 -2.58 -1.95 -1.62
```

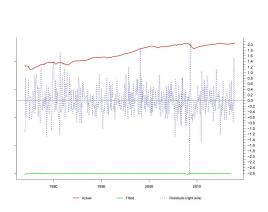
## Test for cointegration (taking the order of integration of the variables into account).

From the ADF test on residuals, the ADF test value is -1.8999 and the critical value by MacKi nnon is -3.7429 - 8.352/528 - 13.41/528 = -3.784116 (variables=3, no trend, T=528). so we accept the H0 of the error term is not I(0) and there is no cointegration. and this may because the 6 variables may have different orders of integration.

### Can you do standard inference using the estimation r esults?

No, there is no cointegration, there will be spurious regression in the model which bring incon sistent result, and it is not asymptotically normal, the standard inference is not possible.

```
> staticResid = ts(static_model_inlevel$residuals, start = 19
74, frequency = 12)
```





#### > LjungBox(ECMresid, 40,12)

1 2 3	1 2 3	0.0773 0.2580 0.3135	NA NA NA	21 22 23	21 22 23	8.9312 9.1004 10.3545	0.9897 0.9928 0.9889
4	4	0.3548 0.3941	NA	24 25	24	26.9931 28.3217	0.3048 0.2933
5 6	5 6	0.3941	NA	26	25 26	28.5413	0.2933
7	7	0.5002	NA NA	27	27	33.9127	0.3324
8	8	0.5969	NA NA	28	28	35.8945	0.1000
9	9	0.6613	NA NA	29	29	41.8512	0.1433
10	10	0.7327	NA NA	30	30	43.5425	0.0524
11	11	1.6658	NA	31	31	43.6407	0.0655
12	12	2.7088	NA NA	32	32	44.5164	0.0697
13	13	2.7169	0.9988	33	33	44.6104	0.0855
14	$\overline{14}$	2.9556	0.9991	34	34	44.6591	0.1044
15	15	3.7468	0.9985	35	35	44.6593	0.1270
16	16	4.4214	0.9980	36	36	45.3428	0.1367
17	17	4.4850	0.9989	37	37	45.4711	0.1600
18	18	6.8264	0.9915	38	38	45.8681	0.1782
19	19	7.0386	0.9940	39	39	46.2124	0.1989
20	20	8.2491	0.9901	40	40	47.0401	0.2065

Adjusted R2 0.478
Residual Std. Error 0.522 (df = 457)
F Statistic 8.898\*\*\* (df = 60; 457)

R2

Note:

# Test for cointegration in this ECM and explain in wh at respect this test differs from the one under question 2?

0.539

\*p<0.1; \*\*p<0.05; \*\*\*p<0.01

From the cointegration test in ECM, we accept of H0 of a=0, there is no cointegration. (we can not use standard critical value because this is not a normal distribution **under H0**), so we us e Mackinnon critical value as a proxy.

The differences is, by this way, we test the significance of the coefficient of the error corre ction term to see if it exists. And if it is significant, then there is cointegration relationship. A nd the Engle Granger two step approach is conducted by the ADF test on the residuals to see if it is I(0), if it is I(0), there is cointegration.

And by this way, we have more power than ADF test to reject there is no cointegration. And we can also avoid the common factor restrictions on the ECM.

```
4. 
> ADL_10 = dynlm(Yt ~ L(Yt,1:10)
+ OIL_P + L(OIL_P,1:10)
+ OIL_PROD + L(OIL_PROD,1:10)
+ OIL_STOCKS + L(OIL_STOCKS, 1:10)
+ WORLD_IP+ L(WORLD_IP,1:10)
+ US_CPI + L(US_CPI,1:10)+)
```

> LjungBox(ADL\_10\$residuals, 100,65)

```
1
       1
              0.0153
                             NA
65
      65
            67.4197
                             NA
66
      66
            67.4837
                        0.4262
             68.1374
67
      67
                        0.4383
            68.3513
68
      68
                        0.4652
69
            69.5975
      69
                        0.4572
70
      70
             69.6344
                        0.4898
            69.9234
                        0.5139
71
      71
72
      72
            73.8817
                        0.4165
      73
73
             73.9802
                        0.4459
74
      74
             73.9804
                        0.4788
75
      75
             74.1088
                        0.5074
            74.1088
75.6316
75.6873
76.7854
77.5475
77.5687
78.9699
79.4866
80.6818
      76
76
                        0.4903
77
      77
                        0.5210
78
      78
                        0.5177
      79
80
                        0.5252
0.5562
79
80
                        0.5431
81
      81
82
83
                        0.5581
      82
      83
                        0.5516
            82.4574
82.5150
82.5477
84
      84
                        0.5272
85
      85
                        0.5561
86
      86
                        0.5854
             82.5979
87
      87
                        0.6136
88
             89.3336
      88
                        0.4403
89
      89
             91.2048
                        0.4153
90
      90
             91.2052
                        0.4447
91
      91
             92.6792
                        0.4313
92
            92.7865
      92
                        0.4574
            94.2175
93
      93
                        0.4452
            94.9859
94
      94
                        0.4521
            94.9893
95
      95
                        0.4810
            95.9079
96
      96
                        0.4835
             95.9093
97
      97
                        0.5122
            98.6600
98
      98
                        0.4623
99
      99
            98.6628
                        0.4907
            98.7161
100 100
                        0.5175
```

### Is this a valid approach?

Yes, it is valid. In principle, we should use the ADL model in first difference if the variables is not stationary and there is no cointegration. However, when conducting cointegration test. use the residuals from static model will ignore the dynamics which will lead a biased and not even asymptotically normal distribution result, also there may be less power to reject the H0 of no cointegration. so, there may be type II error. And if incorrectly accept the H0 of no cointegration and use the model in first difference, the model will be misspecified and inconsistent.

and by estimating the ADL model in level, the result will be biased but consistent, because the real DGP will be nested in our model, so this is also a valid approach.

## Can you do standard inference based on the estimation results?

No, because there may be type II error in our cointegration test, we are not sure if there is error correction term in **real data generating process**, so in general, the standard inference is not possible, unless you know about whether there is cointegration.

# What is the link between this ADL model in levels, the ADL in first differences (question 1) and the error correction model (question 3)?

The ADL in levels can be reparameterization to the error correction model which equals the A DL in first differences plus error correction term.

5.

```
> ECMcoint = nls(formula, data = ECMdata, start = list(theta1
= thetalinit,
+ delta0 = delta0init, delta1 = delta1init, kappa0 = kappa0ini
t, kappa1 = kappa1init,
+ Oilstock0 = Oilstock0init, Oilstock1 = Oilstock1init,
+ worldip0 = worldip0init, worldip1 = worldip1init, uscpi0 = u
scpi0init, uscpi1 = uscpi1init,
+ alpha = alphainit, + beta0 = beta0init, beta1 = beta1init, b
eta2 = beta2init, beta3 = beta3init, + beta4 = beta4init, beta5
= beta5init))

Error in nls(formula, data = ECMdata, start = list(theta1 = theta1init, :
Singular gradient
```

```
> formula = diff_Yt ~ theta1*diff_Yt_1 + delta0*diff_OIL_P +
delta1*diff_OIL_P_1 ++ kappa0*diff_OIL_PROD + kappa1*diff_OI
L_PROD_1 + + alpha*( Yt_1 - beta0 - beta1*OIL_P_1 - beta2*OI
L_PROD_1 + +
L_PROD_1 )>
> ECMcoint = nls(formula, data = ECMdata, start = list(theta1
= theta1init, + delta0 = delta0init, delta1 = delta1init, kapp
a0 = kappa0init, kappa1 = kappa1init, + alpha = alphainit, + be
ta0 = beta0init, beta1 = beta1init, beta2 = beta2init))
> summary(ECMcoint)
Formula: diff_Yt ~ theta1 * diff_Yt_1 + delta0 * diff_OIL_P +
delta1 *
    diff_OIL_P_1 + kappa0 * diff_OIL_PROD + kappa1 * diff_OIL_
PROD_1 +
    alpha * (Yt_1 - beta0 - beta1 * OIL_P_1 - beta2 * OIL_PROD
Parameters:
          Estimate Std. Error t value Pr(>|t|)
.287e-01 4.126e-02 7.966 1.05e-14
                                    7.966 1.05e-14 ***
         3.287e-01
theta1
                      3.963e-03
                                             0.6742
delta0
         1.667e-03
                                    0.421
delta1
         6.731e-03
                      3.966e-03
                                    1.697
                                             0.0903
                      1.922e-02
         8.395e-02
                                    4.368 1.51e-05
kappa0
         3.759e-03
kappa1
                      1.959e-02
                                    0.192
                                             0.8479
        -3.184e-03
                      2.750e-03
alpha
                                   -1.158
                                             0.2474
beta0
        -1.317e+03
                      9.494e+02
                                   -1.387
                                             0.1660
        -5.194e-01
                      4.197e-01
                                   -1.238
beta1
                                             0.2165
beta2
         1.750e+00
                      8.028e-01
                                   2.179
                                             0.0298 *
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.6622 on 517 degrees of freedom
Number of iterations to convergence: 3
Achieved convergence tolerance: 4.749e-08
  (Three observations have been deleted)
```

### Test for cointegration in this ECM.

The t-statistic value of alpha is -1.158, compared to the MacKinnon critical value as a proxy of the real distribution. it accept H0 of there is no cointegration.

Of course there still can be type II error, because when firstly using the nonlinear least squa re method to set the result of the ECM from the static OLS regression as the initial value to estimate this error correction model of order(9) with full variables, there is singular gr adient in the coefficient matrix. And this is very likely because of the multicolinearity in such a complex model. So I have to reduce the lag and even the variables to avoid multicolinearity. But by doing that, I omitted some dynamics and the error term will be autocorrelated.

### What are the advantages of this approach?

By this one step approach, first, it have more powers than the ADF test to reject the H0.

Second, it doesn't ignore the dynamics so because used the cointegrating vector within the E CM, so if there is cointegration, it can do standard inference on the coefficients of the lev el variables (beta) because now it is unbiased and normal distributed.

Third, by testing within the ECM, it does not impose common factor restrictions on the ECM.

# Can you do standard inference based on the estimation results?

Yes, because by ECM it will give me biased but consistent result, standard inference on the coefficient of the differenced variables is appropriate for it is stationary and asymptotically normal. And now the alpha is not significant compared to its critical value, and there is no cointegration, so the error correction term can be dropped out, standard inference on the level variables within the error correction term is not appropriate but is relevant now. So the standard inference is appropriate.