

随机过程第 1 周作业

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1. 设随机变量 X 服从参数为 1 的指数分布, 随机变量 $Y \sim N(0, 1)$, 且 X 与 Y 独立。试求随机变量 $Z = \sqrt{2X}|Y|$ 的分布密度函数。

解: 设 $U = |Y|$, 易知 X 和 U 的概率密度函数分别为

$$f_X(x) = \begin{cases} e^{-x} & x \geq 0 \\ 0 & \text{otherwise} \end{cases}$$
$$f_U(u) = \begin{cases} \frac{\sqrt{2}}{\sqrt{\pi}} e^{-\frac{u^2}{2}} & u \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

设 $V = U$, 则有 $\begin{cases} Z = \sqrt{2X}U \\ V = U \end{cases}$, 其逆映射为 $\begin{cases} X = \frac{Z^2}{2V^2} \\ U = V \end{cases}$ 。 $J = \begin{vmatrix} \frac{\partial X}{\partial Z} & \frac{\partial X}{\partial V} \\ \frac{\partial U}{\partial Z} & \frac{\partial U}{\partial V} \end{vmatrix} = \begin{vmatrix} \frac{Z}{V^2} & -\frac{Z^2}{V^3} \\ 0 & 1 \end{vmatrix} = \frac{Z}{V^2}$

则随机变量 Z, V 的联合概率密度函数为

$$f_{Z,V}(z, v) = f_{X,U}(x, u) \cdot |J| = f_X\left(\frac{z^2}{2v^2}\right) f_U(v) \frac{z}{v^2} = \frac{\sqrt{2}z}{\sqrt{\pi}v^2} e^{-\left(\frac{z^2}{2v^2} + \frac{v^2}{2}\right)}$$

则随机变量 Z 的边缘概率密度函数为

$$f_Z(z) = \int_0^{+\infty} f_{Z,V}(z, v) dv = \frac{\sqrt{2}z}{\sqrt{\pi}} \int_0^{+\infty} \frac{1}{v^2} e^{-\left(\frac{z^2}{2v^2} + \frac{v^2}{2}\right)} dv$$

设 $t = \frac{1}{v}$, 则

$$f_Z(z) = \frac{\sqrt{2}z}{\sqrt{\pi}} \int_0^{+\infty} \frac{1}{v^2} e^{-\left(\frac{z^2}{2v^2} + \frac{v^2}{2}\right)} dv = \frac{\sqrt{2}z}{\sqrt{\pi}} \int_0^{+\infty} e^{-\left(\frac{1}{2t^2} + \frac{z^2 t^2}{2}\right)} dt$$

设 $m = \sqrt{z}t$, 则

$$f_Z(z) = \frac{\sqrt{2}z}{\sqrt{\pi}} \int_0^{+\infty} e^{-\frac{z}{2}\left(\frac{1}{m^2} + m^2\right)} dm = \frac{\sqrt{2}z}{\sqrt{\pi}} e^{-z} \int_0^{+\infty} e^{-\frac{z}{2}\left(\frac{1}{m} - m\right)^2} dm$$

设 $I = \int_0^{+\infty} e^{-\frac{z}{2}\left(\frac{1}{m} - m\right)^2} dm$, 则

$$\begin{aligned} I &= \int_0^{+\infty} e^{-\frac{z}{2}\left(\frac{1}{m} - m\right)^2} dm = \int_0^{+\infty} \frac{m^2}{m^2 + 1} e^{-\frac{z}{2}\left(\frac{1}{m} - m\right)^2} d\left(m - \frac{1}{m}\right) \\ &= \int_0^{+\infty} e^{-\frac{z}{2}\left(\frac{1}{m} - m\right)^2} d\left(m - \frac{1}{m}\right) - \int_0^{+\infty} \frac{1}{m^2 + 1} e^{-\frac{z}{2}\left(\frac{1}{m} - m\right)^2} d\left(m - \frac{1}{m}\right) \\ &= \frac{\sqrt{2\pi}}{\sqrt{z}} - \int_0^{+\infty} \frac{1}{m^2} e^{-\frac{z}{2}\left(\frac{1}{m} - m\right)^2} dm = \frac{\sqrt{2\pi}}{\sqrt{z}} - \int_0^{+\infty} e^{-\frac{z}{2}\left(\frac{1}{r} - r\right)^2} dr \quad \left(r = \frac{1}{m}\right) \\ &= \frac{\sqrt{2\pi}}{\sqrt{z}} - I \end{aligned}$$

则

$$I = \frac{\sqrt{\pi}}{\sqrt{2z}}$$

则

$$f_Z(z) = \begin{cases} \frac{\sqrt{2z}}{\sqrt{\pi}} e^{-z} I = e^{-z}, & z \geq 0 \\ 0 & \end{cases}$$

2. 设随机变量 X_1, X_2 独立同分布, 服从参数为 $\lambda > 0$ 的指数分布。试证明随机变量

$$\frac{X_1}{X_1+X_2} \sim U[0, 1]。$$

解: 设 $\begin{cases} Y_1 = \frac{X_1}{X_1+X_2} \\ Y_2 = X_2 \end{cases}$, 其逆映射为 $\begin{cases} X_1 = \frac{Y_1 Y_2}{1-Y_1} \\ X_2 = Y_2 \end{cases}$ 。 $J = \begin{vmatrix} \frac{\partial X_1}{\partial Y_1} & \frac{\partial X_1}{\partial Y_2} \\ \frac{\partial X_2}{\partial Y_1} & \frac{\partial X_2}{\partial Y_2} \end{vmatrix} = \begin{vmatrix} \frac{Y_2}{(1-Y_1)^2} & \frac{Y_1}{1-Y_1} \\ 0 & 1 \end{vmatrix} = \frac{Y_2}{(1-Y_1)^2}$

当 $y_1, y_2 \in [0, 1]$ 时, 有

$$f_{Y_1, Y_2}(y_1, y_2) = |J| f_{X_1, X_2}(x_1, x_2) = \frac{y_2}{(1-y_1)^2} f_{X_1}\left(\frac{y_1 y_2}{1-y_1}\right) f_{X_2}(y_2) = \frac{\lambda^2 y_2}{(1-y_1)^2} e^{-\frac{\lambda y_2}{1-y_1}}$$

则当 $y_1 \in [0, 1]$ 随机变量 Y_1 的边缘概率密度函数为

$$f_{Y_1}(y_1) = \int_0^\infty f_{Y_1, Y_2}(y_1, y_2) dy_2 = \frac{\lambda^2}{(1-y_1)^2} \int_0^\infty y_2 e^{-\frac{\lambda y_2}{1-y_1}} dy_2$$

设 $I = \int_0^\infty y_2 e^{-\frac{\lambda y_2}{1-y_1}} dy_2$, 则

$$\begin{aligned} I &= \int_0^\infty y_2 e^{-\frac{\lambda y_2}{1-y_1}} dy_2 = \frac{y_1 - 1}{\lambda} \int_0^\infty y_2 d e^{-\frac{\lambda y_2}{1-y_1}} = \frac{y_1 - 1}{\lambda} \left(y_2 e^{-\frac{\lambda y_2}{1-y_1}} \Big|_{y_2=0}^{y_2=\infty} - \int_0^\infty e^{-\frac{\lambda y_2}{1-y_1}} dy_2 \right) \\ &= \frac{1-y_1}{\lambda} \int_0^\infty e^{-\frac{\lambda y_2}{1-y_1}} dy_2 = - \left(\frac{1-y_1}{\lambda} \right)^2 e^{-\frac{\lambda y_2}{1-y_1}} \Big|_{y_2=0}^{y_2=\infty} = \left(\frac{1-y_1}{\lambda} \right)^2 \end{aligned}$$

$$f_{Y_1}(y_1) = \frac{\lambda^2}{(1-y_1)^2} I = 1$$

易知当 $y_1 \notin [0, 1]$ 时, $f_{Y_1}(y_1) = 0$ 。

综上 $f_{Y_1}(y_1) = \begin{cases} 1 & y_1 \in [0, 1] \\ 0 & \text{otherwise} \end{cases}$, 即随机变量 $Y_1 = \frac{X_1}{X_1+X_2} \sim U[0, 1]$ 。

3. 设随机向量 (X, Y) 的两个分量相互独立, 且均服从标准正态分布 $N(0, 1)$ 。

a) 分别写出随机变量 $X+Y$ 和 $X-Y$ 的分布密度。

解: 设 $U = X+Y$, 由卷积公式有

$$f_U(u) = \int_{-\infty}^{+\infty} f_X(x) f_Y(u-x) dx = \frac{1}{2\pi} \int_{-\infty}^{+\infty} e^{-\frac{x^2}{2}} e^{-\frac{(u-x)^2}{2}} dx = \frac{1}{2\sqrt{\pi}} e^{-\frac{u^2}{4}}$$

即 $U \sim N(0, 2)$ 。

同理可证, $V = X-Y \sim U(0, 2)$, 即 $f_V(v) = \frac{1}{2\sqrt{\pi}} e^{-\frac{v^2}{4}}$

b) 试问: $X+Y$ 与 $X-Y$ 是否独立? 说明理由。

解: 设 $\begin{cases} U = X+Y \\ V = X-Y \end{cases}$, 其逆映射为 $\begin{cases} X = \frac{1}{2}(U+V) \\ Y = \frac{1}{2}(U-V) \end{cases}$ 。 $J = \begin{vmatrix} \frac{\partial X}{\partial U} & \frac{\partial X}{\partial V} \\ \frac{\partial Y}{\partial U} & \frac{\partial Y}{\partial V} \end{vmatrix} = \begin{vmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & -\frac{1}{2} \end{vmatrix} = -\frac{1}{2}$

随机变量 (U, V) 联合概率密度函数为

$$f_{U, V}(u, v) = f_{X, Y}\left(\frac{1}{2}(u+v), \frac{1}{2}(u-v)\right) = \frac{1}{4\pi} e^{-\frac{(u^2+v^2)}{4}} = f_U(u) f_V(v)$$

故随机变量 $X+Y, X-Y$ 独立。