第2章: 贝叶斯决策理论

刘成林(liucl@nlpr.ia.ac.cn) 2021年9月15日

助教: 赵梦彪(zhaomengbiao2017@ia.ac.cn)

郭宏宇(guohongyu2019@nlpr.ia.ac.cn)

朱 飞(zhufei2018@ia.ac.cn)



统计模式识别方法

生成模型

(Density-based, Bayes decision) 近似 $p(\mathbf{x}|\omega_i)$

Parametric

- √ Gaussian
- ✓ Dirichlet
- ✓ Bayesian network
- ✓ Hidden Markov model

Non-Parametric

- Histogram density
- ✓ Parzen window
- √ K-nearest neighbor

function) 近似p(ω_i |**x**)

判别模型

(discriminant/decision

- ✓ Neural network
- ✓ Logistic regression

Linear methods

- ✓ Decision tree
- ✓ Kernel (SVM)
- Boosting

Semi-Parametric

√ Gaussian mixture

注意discriminative和 discriminant的区别!

Discriminative model/learning Discriminant function/analysis



提纲

- 分类问题表示
- 最小错误率决策: 2类的例子
- 最小风险决策
 - 扩展: 开放集识别的贝叶斯决策
- 判别函数和决策面
- 高斯概率密度
 - 协方差矩阵的性质
 - 相关知识: 特征提取/降维
- 高斯密度下的判别函数
 - 扩展: 广义线性判别函数
- 分类错误率



导论:分类问题表示

- 类别: ω_i , i = 1,...,c
- 特征矢量 $\mathbf{x} = [x_1, ..., x_d] \in \mathbb{R}^d$
- 先验概率 $P(\omega_i)$ $\sum_{i=1}^{c} P(\omega_i) = 1$
- 概率密度函数(条件概率) $p(\mathbf{x} \mid \omega_i)$
- 后验概率

$$P(\omega_i \mid \mathbf{x}) = \frac{p(\mathbf{x} \mid \omega_i)P(\omega_i)}{p(\mathbf{x})} = \frac{p(\mathbf{x} \mid \omega_i)P(\omega_i)}{\sum_{j=1}^{c} p(\mathbf{x} \mid \omega_j)P(\omega_j)}$$
$$\sum_{j=1}^{c} P(\omega_i \mid \mathbf{x}) = 1$$

最小错误率决策: 2类的例子

- Salmon (ω_1) and sea bass (ω_2)
- If we have only prior probability
 - 例如,教室门口判断进来的是男生还是女生,没有任何传感器
 - Decide ω_1 if $P(\omega_1) > P(\omega_2)$, otherwise ω_2
 - Minimum error decision

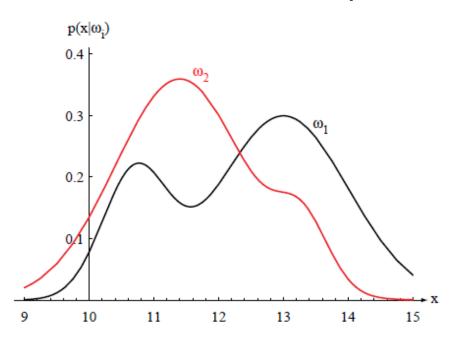
$$P(error) = \begin{cases} P(\omega_2) & \text{if we decide } \omega_1 \\ P(\omega_1) & \text{if we decide } \omega_2 \end{cases}$$

- 教室门口判断性别的例子: 错误率?



2类的例子

有传感器(特征)的情况
 Decision based on posterior probabilities



 $P(\omega_1|\mathbf{x})$ $P(\omega_1) = 2/3$ $P(\omega_2) = 1/3$ 0.8 $P(\omega_2) = 1/3$ 0.2 0.2 0.2 0.2 0.2 0.2 0.3 0.3 0.4 0.2 0.3 0.3 0.4 0.4

x轴:一维特征空间

$$P(\omega_i \mid \mathbf{x}) = \frac{p(\mathbf{x} \mid \omega_i)P(\omega_i)}{p(\mathbf{x})} = \frac{p(\mathbf{x} \mid \omega_i)P(\omega_i)}{\sum_{j=1}^{c} p(\mathbf{x} \mid \omega_j)P(\omega_j)}$$



Decision based on posterior probabilities

$$P(error|x) = \begin{cases} P(\omega_1|x) & \text{if we decide } \omega_2 \\ P(\omega_2|x) & \text{if we decide } \omega_1. \end{cases}$$

Decide ω_1 if $P(\omega_1|x) > P(\omega_2|x)$; otherwise decide ω_2

$$P(error|x) = \min [P(\omega_1|x), P(\omega_2|x)].$$

Evidence (a.k.a. likelihood)

Decide ω_1 if $p(x|\omega_1)P(\omega_1) > p(x|\omega_2)P(\omega_2)$; otherwise decide ω_2

- see
$$P(\omega_i \mid \mathbf{x}) = \frac{p(\mathbf{x} \mid \omega_i)P(\omega_i)}{p(\mathbf{x})}$$

教室门口判断性别的例子:用什么传感器(x)?



最小风险决策: 贝叶斯决策的一般形式

- 决策代价(loss, cost)
 - True class ω_j , decided as α_i $\lambda_{ij} = \lambda(\alpha_i \mid \omega_j)$
 - 有时 λ_{ij} 和 λ_{ji} 相差很大,比如医疗诊断的场合、工业检测、自动商店判断性别
- Condition risk

$$R(\alpha_i|\mathbf{x}) = \sum_{j=1}^{c} \lambda(\alpha_i|\omega_j) P(\omega_j|\mathbf{x})$$

Overall (expected) risk

$$R = \int R(\alpha(\mathbf{x})|\mathbf{x})p(\mathbf{x}) \ d\mathbf{x} \qquad \alpha(\mathbf{x}) \in \{\alpha_1, \dots, \alpha_c\}$$

Minimum risk decision (Bayes decision)

$$\arg\min_{i} R(\alpha_{i} \mid x)$$



- Minimum risk decision: 2-class case
 - Condition risk

$$R(\alpha_1|\mathbf{x}) = \lambda_{11}P(\omega_1|\mathbf{x}) + \lambda_{12}P(\omega_2|\mathbf{x})$$

$$R(\alpha_2|\mathbf{x}) = \lambda_{21}P(\omega_1|\mathbf{x}) + \lambda_{22}P(\omega_2|\mathbf{x})$$

Decision rule

$$R(\alpha_1 \mid x) < R(\alpha_2 \mid x) \leftrightarrow (\lambda_{21} - \lambda_{11})P(\omega_1 \mid x) > (\lambda_{12} - \lambda_{22})P(\omega_2 \mid x)$$

• Equivalently, decide ω_1 if

$$(\lambda_{21} - \lambda_{11}) \underline{p(\mathbf{x}|\omega_1)P(\omega_1)} > (\lambda_{12} - \lambda_{22})p(\mathbf{x}|\omega_2)P(\omega_2)$$
$$\frac{p(\mathbf{x}|\omega_1)}{p(\mathbf{x}|\omega_2)} > \frac{\lambda_{12} - \lambda_{22}}{\lambda_{21} - \lambda_{11}} \frac{P(\omega_2)}{P(\omega_1)}$$

(Likelihood ratio)



最小错误率分类

Zero-one loss

$$\lambda(\alpha_i | \omega_j) = \begin{cases} 0 & i = j \\ 1 & i \neq j \end{cases} i, j = 1, ..., c$$

$$R(\alpha_i | \mathbf{x}) = \sum_{j=1}^{c} \lambda(\alpha_i | \omega_j) P(\omega_j | \mathbf{x})$$

$$= \sum_{j \neq i} P(\omega_j | \mathbf{x})$$

$$= 1 - P(\omega_i | \mathbf{x})$$

Minimum error decision: Maximum a posteriori (MAP)

Decide
$$\omega_i$$
 if $P(\omega_i|\mathbf{x}) > P(\omega_i|\mathbf{x})$ for all $j \neq i$

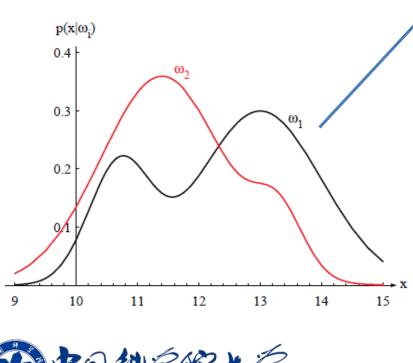


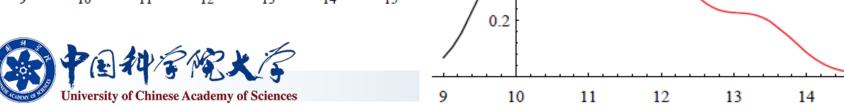
• 2-class case

– decide ω_1 if

$$\frac{p(\mathbf{x}|\omega_1)}{p(\mathbf{x}|\omega_2)} > \frac{\lambda_{12} - \lambda_{22}}{\lambda_{21} - \lambda_{11}} \frac{P(\omega_2)}{P(\omega_1)}$$

0-1 loss
$$\frac{p(\mathbf{x}|\omega_1)}{p(\mathbf{x}|\omega_2)} > \frac{P(\omega_2)}{P(\omega_1)}$$





Likelihood ratio

 $\overline{\mathcal{R}_{_1}}$

 $P(\omega_i|x)$

0.8

0.6

Decision

region

15

 $\overline{\mathcal{R}_{_1}}$

 ω_1

带拒识的决策

- 为什么要拒识?错误识别可能带来严重后果
 - 比如医疗诊断,金额识别

Ambiguity Rejection

Distance Rejection

In fact, the Tories made it worse now for the sick and needy than Labour had to make it in 1350. And as a percentage of social service expenditure, health had fallen from 28.5 to 23.1 percent.

English is outlier for a digit recognizer

経営不振に陥っているソニーのパソコン事業は国内 投資ファンド が買い取り、開発と製造をてがける新会社バイオをつくった。ソニーの国内向け通販サイトと同社の直営店を通じ、個人から注文をとってきた。

These are outlier for an English recognizer



带拒识的决策

- Formulation (Problem 13, Chapter 2)
 - C+1 classes

$$\lambda(\alpha_i \mid \omega_j) = \begin{cases} 0, & i = j \\ \lambda_s, & i \neq j \\ \lambda_r, & \text{reject} \end{cases}$$
 $\lambda_r < \lambda_s$

$$R(\alpha_i|\mathbf{x}) = \sum_{j=1}^{c} \lambda(\alpha_i|\omega_j) P(\omega_j|\mathbf{x})$$

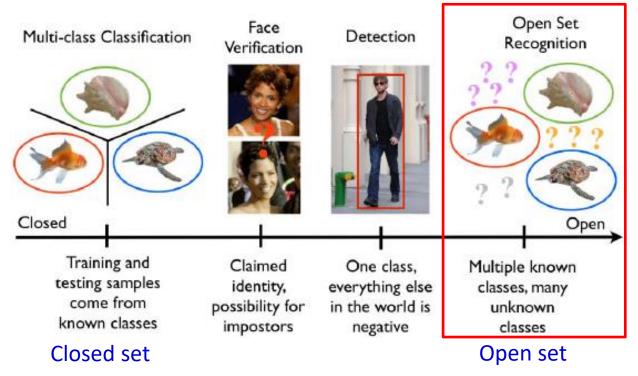
$$R_i(\mathbf{x}) = \begin{cases} \lambda_s [1 - P(\omega_i \mid \mathbf{x})], & i = 1, ..., c \\ \lambda_r, & \text{reject} \end{cases}$$

$$\arg\min_{i} R_{i}(\mathbf{x}) = \begin{cases} \arg\max_{i} P(\omega_{i} \mid \mathbf{x}), & \text{if } \max_{i} P(\omega_{i} \mid \mathbf{x}) > 1 - \lambda_{r} / \lambda_{s} \\ & \text{reject,} & \text{otherwise} \end{cases}$$



扩展: 开放集识别的贝叶斯决策

- 传统的分类器假设训练样本和测试样本都来自预设的C个类别(闭合集, Closed set)。
- 实际环境中测试样本可能不属于预设的C个类别(异常样本, outlier),这种情况称为开放集(Open set)。
- 开放集的难点是异常样本没有训练集,只能训练已知C类的分类器。



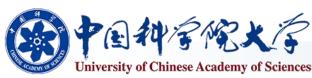
开放集分类贝叶斯决策

• 问题表示

- 已知类别: ω_i , i=1,...,c
- 先验概率 $\sum_{i=1}^{c} P(\omega_i) \leq 1$
- 后验概率 $\sum_{i=1}^{c} P(\omega_i \mid \mathbf{x}) \le 1 \qquad \sum_{i=1}^{c+1} P(\omega_i \mid \mathbf{x}) = 1$
- 条件概率密度 $p(\mathbf{x} \mid \omega_i), i = 1,...,c, p(\mathbf{x} \mid \omega_{c+1}) = ?$

• 分类决策

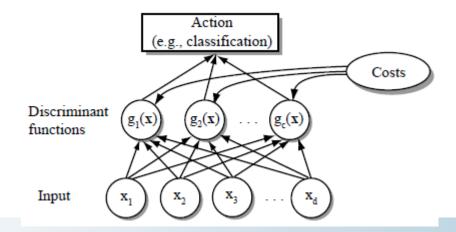
- 假设 $p(\mathbf{x} \mid \omega_{c+1}) = \rho$ ρ为很小的常数



判别函数、决策面

- 判别函数(Discriminant Function)
 - 表征模式属于每一类的广义似然度 $g_i(\mathbf{x})$, i=1,...,c
 - 分类决策 $\underset{i}{\operatorname{arg max}} g_{i}(\mathbf{x})$
 - E.g., conditional risk $g_i(\mathbf{x}) = -R(\alpha_i \mid \mathbf{x})$
 - Posterior probability $g_i(\mathbf{x}) = P(\omega_i \mid \mathbf{x})$
 - Likelihood $g_i(\mathbf{x}) = p(\mathbf{x} \mid \omega_i) P(\omega_i)$

$$g_i(\mathbf{x}) = \log p(\mathbf{x} \mid \omega_i) + \log P(\omega_i)$$





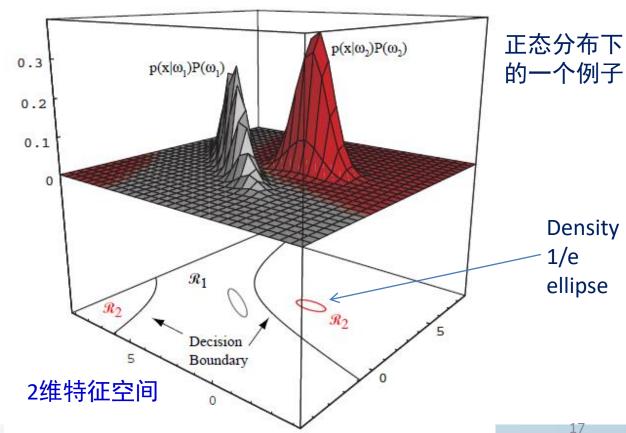
• 决策面(Decision surface)

- 特征空间中二类判别函数相等的点的集合

$$g(\mathbf{x}) \equiv g_1(\mathbf{x}) - g_2(\mathbf{x}) \qquad g(\mathbf{x}) = P(\omega_1|\mathbf{x}) - P(\omega_2|\mathbf{x}) = \mathbf{0}$$
$$g(\mathbf{x}) = \ln \frac{p(\mathbf{x}|\omega_1)}{p(\mathbf{x}|\omega_2)} + \ln \frac{P(\omega_1)}{P(\omega_2)}$$

有判别函数就可以分 类了,为什么还来求 决策面?

答:加深对特征空间和分类器性质的理解





贝叶斯决策用于模式分类

- Bayes决策的关键
 - 类条件概率密度估计
 - 先验概率: 从训练样本估计或假设等概率
 - 决策代价[λ_{ii}],一般为0-1代价
- 分类器设计
 - 收集训练样本
 - 用每一类的样本估计类条件概率密度 $p(\mathbf{x} \mid \omega_i)$
 - 估计类先验概率
 - 模型参数集: $\{p(\mathbf{x} \mid \omega_i, \theta_i), P(\omega_i)\}, i = 1,...,c$
- 分类过程
 - 计算测试样本x属于每一类的后验概率
 - 最大后验概率/最小风险决策



概率密度估计方法

• 参数法: 假定概率密度函数形式

$$p(\mathbf{x} \mid \omega_i) = p(\mathbf{x} \mid \theta_i)$$

- Distribution: Gaussian, Gamma, Bernouli
- Parameter estimation: maximum-likelihood (ML), Bayesian estimation
- 非参数法: 可以表示任意概率分布, 无函数形式
 - Parzen window, k-NN
 - 需要保存所有或大部分样本
- Semi-parametric, 近似任意概率分布, 有函数形式
 - Distribution: Gaussian mixture (GM)
 - Estimation: expectation-maximization (EM)



高斯密度函数

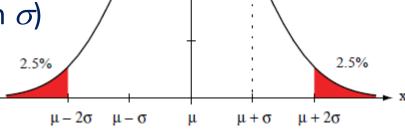
Gaussian density (normal distribution)

$$p(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right]$$

- Mean μ
- Variance σ^2 (standard deviation σ)

$$\mu \equiv \mathcal{E}[x] = \int_{-\infty}^{\infty} x p(x) \ dx$$

$$\sigma^2 \equiv \mathcal{E}[(x-\mu)^2] = \int_{-\infty}^{\infty} (x-\mu)^2 p(x) \ dx$$



p(x)

$$H(p(x)) = -\int p(x) \ln p(x) dx$$

- 在给定均值和方差的所有分布中,正态分布的熵最大(Problem 20, Chapter 2)
- 根据Central Limit Theorem, 大量独立随机变量之和趋近正态分布
- 实际环境中,很多类别的特征分布趋近正态分布



• Multivariate normal density $p(\mathbf{x}) \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$

$$- 公式要牢记 \qquad p(\mathbf{x}) = \frac{1}{(2\pi)^{d/2}|\Sigma|^{1/2}} \mathrm{exp}\left[-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^t \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\right]$$

- Mean
$$\mu \equiv \mathcal{E}[\mathbf{x}] = \int \mathbf{x} p(\mathbf{x}) d\mathbf{x}$$
 $\mu_i = \mathcal{E}[x_i]$

Covariance matrix

$$oldsymbol{\Sigma} \equiv \mathcal{E}[(\mathbf{x} - oldsymbol{\mu})(\mathbf{x} - oldsymbol{\mu})^t] = \int (\mathbf{x} - oldsymbol{\mu})(\mathbf{x} - oldsymbol{\mu})^t p(\mathbf{x}) \ d\mathbf{x}$$
 $\sigma_{11} \quad \sigma_{12} \quad \dots \quad \sigma_{1d} \quad \sigma_{21} \quad \sigma_{22} \quad \dots \quad \sigma_{2d} \quad$

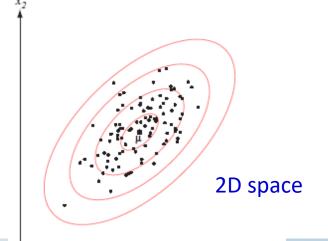
If x_i and x_j are statistically independent, $\sigma_{ij} = 0$

$$egin{bmatrix} \sigma_{11} & \sigma_{12} & ... & \sigma_{1d} \ \sigma_{21} & \sigma_{22} & ... & \sigma_{2d} \ dots & dots & \ddots & dots \ \sigma_{d1} & \sigma_{d2} & ... & \sigma_{dd} \end{bmatrix}$$

- 等密度点轨迹: hyperellipsoid
 - 特殊情况下为圆形或超球面
- Mahalanobis distance

$$r^2 = (\mathbf{x} - \boldsymbol{\mu})^t \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})$$





Break



协方差矩阵的性质

- 实对称矩阵
- Eigenvalues & eigenvecters (本征值,本征向量)

$$\Sigma \phi_{i} = \lambda_{i} \phi_{i} \qquad \Phi = [\phi_{1} \phi_{2} \cdots \phi_{d}] \qquad \Lambda = diag[\lambda_{1}, \lambda_{2}, \cdots, \lambda_{d}]$$
Orthonormal
$$\Phi^{T} \Phi = I$$

$$\Phi^{T} = \Phi^{-1} \qquad \phi_{i}^{t} \phi_{j} = \begin{cases} 1, & i = j \\ 0, & i \neq j \end{cases}$$

$$\Sigma \Phi = \Phi \Lambda \iff \Sigma = \Phi \Lambda \Phi^{T} \iff \Sigma = \sum_{i=1}^{d} \lambda_{i} \phi_{i} \phi_{i}^{T}$$

• 矩阵对角化

$$\Phi^{T} \Sigma \Phi = \Lambda = \begin{bmatrix} \lambda_{1} & 0 & \dots & 0 \\ 0 & \lambda_{2} & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \lambda_{d} \end{bmatrix} \longleftrightarrow \phi_{i}^{T} \Sigma \phi_{i} = \lambda_{i}$$



- 应用: Principal component analysis (PCA)
 - 一种降维(特征提取)方法
 - 将随机矢量投影到低维子空间, 使子空间投影的重建误差最小
 - 选择本征值最大的m (m<d)个本征向量作为子空间的基(basis)

线性空间中正交变换($\Phi^T\Phi = I$) 不影响欧氏距离

$$\sum_{j=1}^{d} [(\mathbf{x} - \boldsymbol{\mu})^{T} \boldsymbol{\phi}_{j}]^{2} = ||\mathbf{x} - \boldsymbol{\mu}||^{2}$$

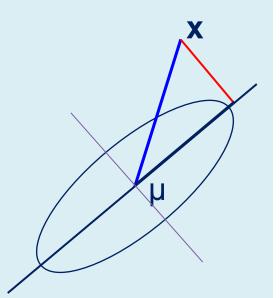
子空间投影 $\mu + \sum_{j=1}^{m} [(\mathbf{x} - \mu)^T \phi_j] \phi_j$

投影重建误差

$$r_E = ||\mathbf{x} - \mu||^2 - \sum_{j=1}^m [(\mathbf{x} - \mu)^T \phi_j]^2 = \sum_{j=m+1}^d [(\mathbf{x} - \mu)^T \phi_j]^2$$

期望

$$\mathcal{E}(r_E) = \mathcal{E}\left\{\sum_{j=m+1}^{d} [(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\phi}_j]^2\right\} = \mathcal{E}\left\{\sum_{j=m+1}^{d} \boldsymbol{\phi}_j^T (\mathbf{x} - \boldsymbol{\mu}) (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\phi}_j\right\}$$
$$= \sum_{j=m+1}^{d} \boldsymbol{\phi}_j^T \mathcal{E}[(\mathbf{x} - \boldsymbol{\mu}) (\mathbf{x} - \boldsymbol{\mu})^T] \boldsymbol{\phi}_j = \sum_{j=m+1}^{d} \boldsymbol{\phi}_j^T \boldsymbol{\Sigma} \boldsymbol{\phi}_j = \sum_{j=m+1}^{d} \boldsymbol{\lambda}_j$$



 $\min \sum_{j=m+1}^{u} \lambda_{j} \text{ or } \max \sum_{j=1}^{m} \lambda_{j}$ 意味着取 λ_{i} 最大的m个本征向量作为子空间基(basis)



线性变换的高斯分布

- 线性变换 $y = A^t x$
 - A^tA=1: 正交变换(坐标轴旋转)
 - 变换后的分布仍为正态分布

$$p(\mathbf{y}) \sim N(\mathbf{A}^t \boldsymbol{\mu}, \mathbf{A}^t \boldsymbol{\Sigma} \mathbf{A})$$

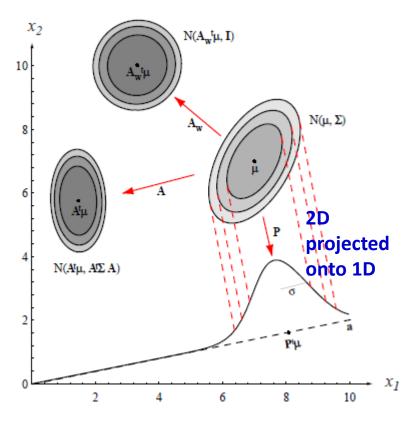
Diagonalization

$$A = \Phi$$

$$A^{t}\Sigma A = \Lambda$$

Whitening transform

$$\mathbf{A}_{w} = \mathbf{\Phi} \mathbf{\Lambda}^{-1/2}$$
 $A_{w}^{t} \Sigma A_{w} = \mathbf{\Lambda}^{-1/2} \mathbf{\Phi}^{t} \Sigma \mathbf{\Phi} \mathbf{\Lambda}^{-1/2}$
 $= \mathbf{\Lambda}^{-1/2} \mathbf{\Lambda} \mathbf{\Lambda}^{-1/2} = I$





相关知识:特征提取/降维

- Dimensionality reduction
- Feature extraction
 - Feature generation: original data d→x
 - Linear feature extraction $\mathbf{x} = \mathbf{A}^T \mathbf{d}$
- Feature selection (for reduction and performance)
 - Feature subset selection: a learning/optimization problem
- Feature transform (for extraction or reduction)
 - Linear transform $\mathbf{y} = \mathbf{A}^T \mathbf{x}$
 - Nonlinear transform: may increase dimensionality, e.g. kernel PCA, kernel LDA
- Handcrafted feature
- Feature learning
 - Automatic feature generation, e.g. convolutional neural network (CNN)



高斯密度下的判别函数

• 判别函数 $g_i(\mathbf{x}) = \ln p(\mathbf{x}|\omega_i) + \ln P(\omega_i)$

$$\begin{split} p(\mathbf{x} \mid \omega_i) &= \frac{1}{(2\pi)^{d/2} \mid \boldsymbol{\Sigma}_i \mid^{1/2}} \exp \left[-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_i) \boldsymbol{\Sigma}_i^{-1} (\mathbf{x} - \boldsymbol{\mu}_i) \right] \\ g_i(\mathbf{x}) &= -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_i)^t \boldsymbol{\Sigma}_i^{-1} (\mathbf{x} - \boldsymbol{\mu}_i) - \frac{d}{2} \ln 2\pi - \frac{1}{2} \ln |\boldsymbol{\Sigma}_i| + \ln P(\omega_i) \end{split}$$

- Quadratic discriminant function (QDF)
- 在不同covariance假设条件下得到一些特殊形式

• Case 1: $\Sigma_i = \sigma^2 I$ (去掉与类别无关项)

$$g_i(\mathbf{x}) = -\frac{\|\mathbf{x} - \boldsymbol{\mu}_i\|^2}{2\sigma^2} + \ln P(\omega_i)$$

- Euclidean distance $\|\mathbf{x} \boldsymbol{\mu}_i\|^2 = (\mathbf{x} \boldsymbol{\mu}_i)^t (\mathbf{x} \boldsymbol{\mu}_i)$
 - Nearest mean/distance classifier
- 展开二次式 $(x \mu_i)^t (x \mu_i)$

$$g_i(\mathbf{x}) = -\frac{1}{2\sigma^2} [\mathbf{x}^t \mathbf{x} - 2\boldsymbol{\mu}_i^t \mathbf{x} + \boldsymbol{\mu}_i^t \boldsymbol{\mu}_i] + \ln P(\omega_i)$$

- 忽略与类别无关项,得到线性判别函数

$$g_i(\mathbf{x}) = \mathbf{w}_i^t \mathbf{x} + w_{i0}$$

$$\mathbf{w}_i = \frac{1}{\sigma^2} \mu_i \qquad w_{i0} = \frac{-1}{2\sigma^2} \mu_i^t \mu_i + \ln P(\omega_i)$$

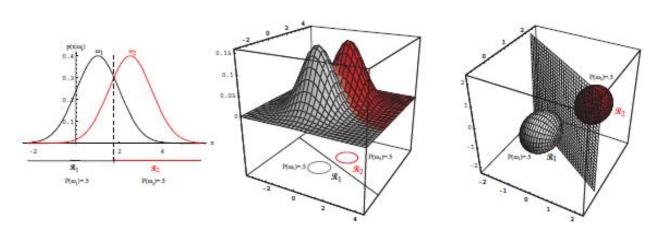


- Case 1: $\Sigma_i = \sigma^2 I$ (continued)
 - 二类决策面(判别函数相等的点构成)

$$\mathbf{g}_i(\mathbf{x}) = g_j(\mathbf{x})$$

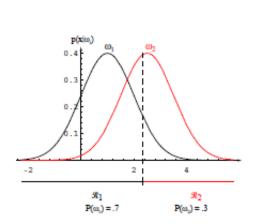
$$\longrightarrow \mathbf{w}^t(\mathbf{x} - \mathbf{x}_0) = 0 \qquad \mathbf{w} = \boldsymbol{\mu}_i - \boldsymbol{\mu}_j \qquad \qquad \mathbf{W}: 法向量 \\ \mathbf{v}_0 = \frac{1}{2}(\boldsymbol{\mu}_i + \boldsymbol{\mu}_j) - \frac{\sigma^2}{\|\boldsymbol{\mu}_i - \boldsymbol{\mu}_j\|^2} \ln \frac{P(\omega_i)}{P(\omega_j)} (\boldsymbol{\mu}_i - \boldsymbol{\mu}_j) \qquad \text{位置移向先验} \\ \mathbf{w}_0 = \frac{1}{2}(\boldsymbol{\mu}_i + \boldsymbol{\mu}_j) - \frac{\sigma^2}{\|\boldsymbol{\mu}_i - \boldsymbol{\mu}_j\|^2} \ln \frac{P(\omega_i)}{P(\omega_j)} (\boldsymbol{\mu}_i - \boldsymbol{\mu}_j) \qquad \text{使置移向先验} \\ \mathbf{w}_0 = \frac{1}{2}(\boldsymbol{\mu}_i + \boldsymbol{\mu}_j) - \frac{\sigma^2}{\|\boldsymbol{\mu}_i - \boldsymbol{\mu}_j\|^2} \ln \frac{P(\omega_i)}{P(\omega_j)} (\boldsymbol{\mu}_i - \boldsymbol{\mu}_j) \qquad \text{使置移向先验}$$

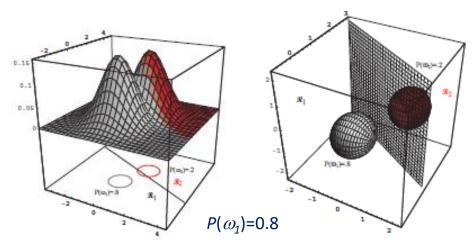
- 1D, 2D, 3D的情况
 - 当 $P(\omega_1)=P(\omega_2)$, 决策面为二类均值的等分面



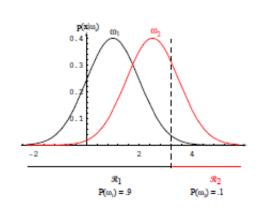


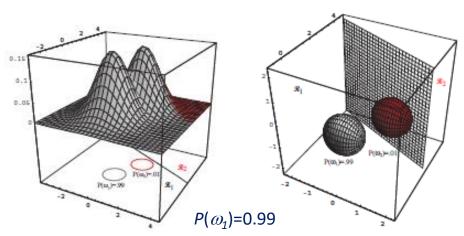
- 当先验概率变化,决策面发生平移





移向概率 小的类别





• Case 2: $\Sigma_i = \Sigma$ (所有类别共享协方差矩阵)

$$\begin{split} g_i(\mathbf{x}) &= -\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_i)^t \boldsymbol{\Sigma}_i^{-1}(\mathbf{x} - \boldsymbol{\mu}_i) - \frac{d}{2} \ln 2\pi - \frac{1}{2} \ln |\boldsymbol{\Sigma}_i| + \ln P(\omega_i) \\ & \Longrightarrow \quad g_i(\mathbf{x}) = -\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_i)^t \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu}_i) + \ln P(\omega_i) \end{split}$$

- 展开二次式 $(\mathbf{x} - \boldsymbol{\mu}_i)^t \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}_i)$

线性判别函数! $g_i(\mathbf{x}) = \mathbf{w}_i^t \mathbf{x} + w_{i0}$

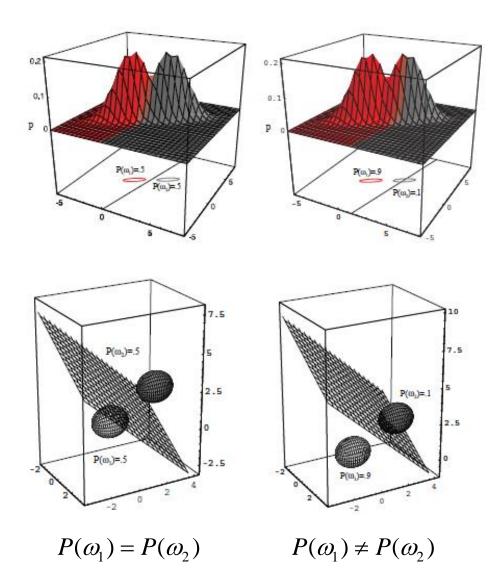
$$\mathbf{w}_i = \mathbf{\Sigma}^{-1} \boldsymbol{\mu}_i \qquad w_{i0} = -\frac{1}{2} \boldsymbol{\mu}_i^t \mathbf{\Sigma}^{-1} \boldsymbol{\mu}_i + \ln P(\omega_i)$$

- 二类决策面 $g_i(\mathbf{x}) = g_j(\mathbf{x})$

$$\mathbf{w}^{t}(\mathbf{x} - \mathbf{x}_{0}) = 0 \qquad \mathbf{w} = \mathbf{\Sigma}^{-1}(\boldsymbol{\mu}_{i} - \boldsymbol{\mu}_{j})$$
$$\mathbf{x}_{0} = \frac{1}{2}(\boldsymbol{\mu}_{i} + \boldsymbol{\mu}_{j}) - \frac{\ln\left[P(\boldsymbol{\omega}_{i})/P(\boldsymbol{\omega}_{j})\right]}{(\boldsymbol{\mu}_{i} - \boldsymbol{\mu}_{j})^{t}\mathbf{\Sigma}^{-1}(\boldsymbol{\mu}_{i} - \boldsymbol{\mu}_{j})}(\boldsymbol{\mu}_{i} - \boldsymbol{\mu}_{j})$$

- 注意跟μ₁-μ₂的关系,决策面不一定与之垂直
- 当 $P(\omega_1) = P(\omega_2)$, 决策面经过 $(\mu_1 + \mu_2)/2$



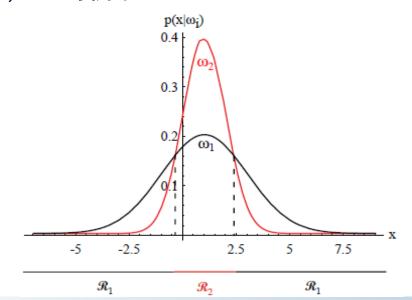




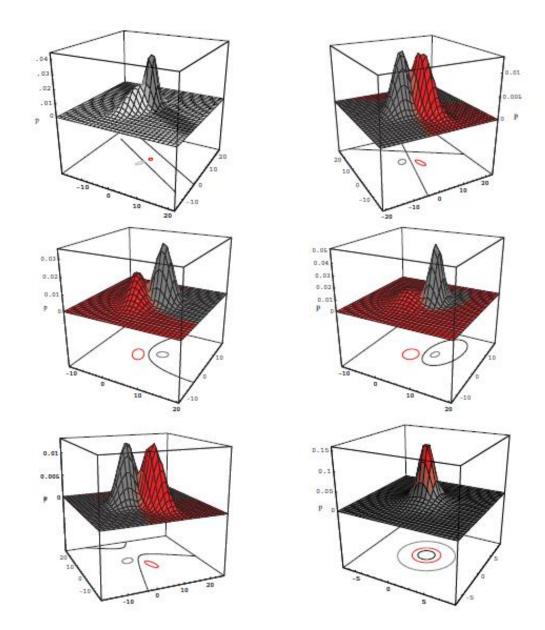
• Case 3: Σ_i = arbitrary

$$\begin{split} g_i(\mathbf{x}) &= -\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_i)^t \boldsymbol{\Sigma}_i^{-1}(\mathbf{x} - \boldsymbol{\mu}_i) - \frac{d}{2} \ln 2\pi - \frac{1}{2} \ln |\boldsymbol{\Sigma}_i| + \ln P(\omega_i) \\ g_i(\mathbf{x}) &= \mathbf{x}^t \mathbf{W}_i \mathbf{x} + \mathbf{w}_i^t \mathbf{x} + w_{i0} \\ \mathbf{W}_i &= -\frac{1}{2} \boldsymbol{\Sigma}_i^{-1} \qquad \mathbf{w}_i = \boldsymbol{\Sigma}_i^{-1} \boldsymbol{\mu}_i \\ w_{i0} &= -\frac{1}{2} \boldsymbol{\mu}_i^t \boldsymbol{\Sigma}_i^{-1} \boldsymbol{\mu}_i - \frac{1}{2} \ln |\boldsymbol{\Sigma}_i| + \ln P(\omega_i) \end{split}$$

- 二类决策面: g₁(x)=g₂(x), hyperquadratics
 - 等均值的情况下, 1D的例子

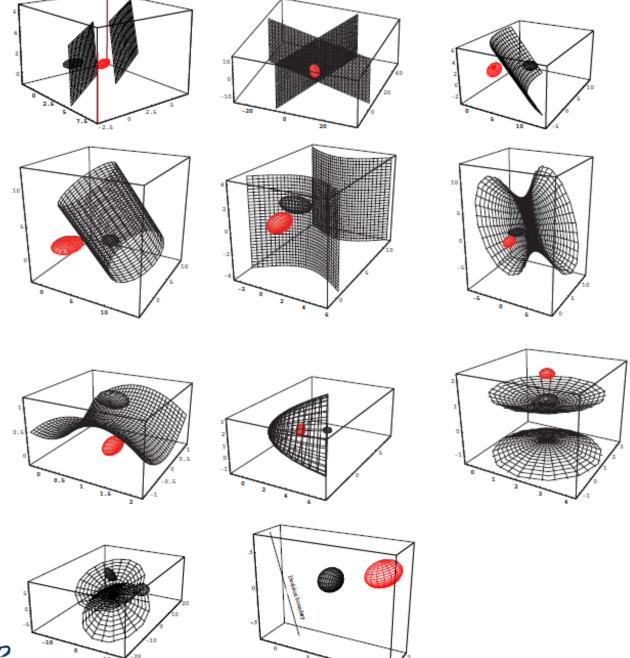


2D的例子 (z轴是概率密度)



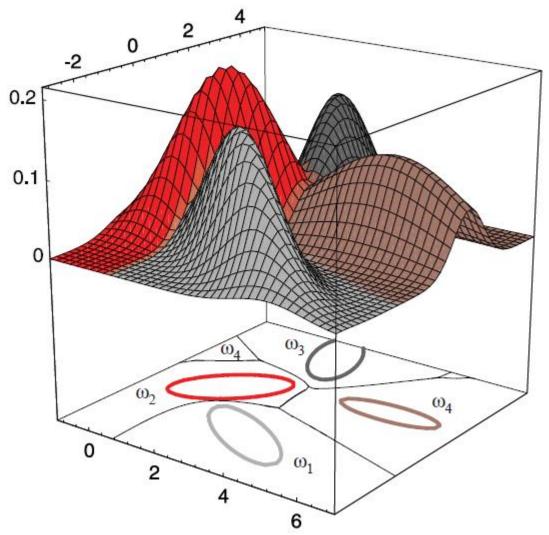


3D的例子





2D, 4类的例子





一个具体例子

$$-2$$
类,2D $P(\omega_1) = P(\omega_2) = 0.5$

$$\mu_1 = \begin{bmatrix} 3 \\ 6 \end{bmatrix}; \quad \Sigma_1 = \begin{pmatrix} 1/2 & 0 \\ 0 & 2 \end{pmatrix} \qquad \Sigma_1^{-1} = \begin{pmatrix} 2 & 0 \\ 0 & 1/2 \end{pmatrix}$$

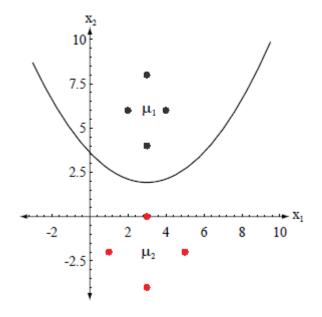
$$\mu_2 = \begin{bmatrix} 3 \\ -2 \end{bmatrix}; \quad \Sigma_2 = \begin{pmatrix} 2 & 0 \\ 0 & 2 \end{pmatrix} \qquad \Sigma_2^{-1} = \begin{pmatrix} 1/2 & 0 \\ 0 & 1/2 \end{pmatrix}$$

$$\Sigma_1^{-1} = \left(\begin{array}{cc} 2 & 0 \\ 0 & 1/2 \end{array}\right)$$

$$\Sigma_2^{-1} = \left(\begin{array}{cc} 1/2 & 0\\ 0 & 1/2 \end{array}\right)$$

- 决策面 g₁(x)=g₂(x)

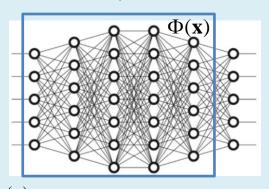
$$x_2 = 3.514 - 1.125x_1 + 0.1875x_1^2$$



线性判别函数、广义线性判别函数、神经网络

- 线性判别函数(LDF) $g_i(\mathbf{x}) = \mathbf{w}_i^T \mathbf{x} + w_{i0}$
 - 参数估计方法: Gaussian density, logistic regression,
 single-layer neural network, linear SVM, etc
- 广义线性判别函数 $g_i(\mathbf{x}) = \mathbf{w}_i^T \Phi(\mathbf{x}) + w_{i0}$
 - 扩充特征向量 $\Phi(\mathbf{x})$,如多项式特征,特征函数等
 - − Nonlinear SVM → kernel function $\Phi(\mathbf{x}_i) \cdot \Phi(\mathbf{x}_i) = k(\mathbf{x}_i, \mathbf{x}_j)$
 - 多层神经网络的输出层可看作是 广义线性判别函数
- LDF跟soft-max的关系
 - 由Gaussian density based LDF可知,

$$g_i(\mathbf{x}) \propto \log p(\mathbf{x} \mid \omega_i) P(\omega_i) \Rightarrow p(\omega_i \mid \mathbf{x}) = \frac{e^{g_i(\mathbf{x})}}{\sum_{j=1}^{c} e^{g_j(\mathbf{x})}}$$



分类错误率

• 2类的情况

$$P(error) = P(\mathbf{x} \in \mathcal{R}_2, \omega_1) + P(\mathbf{x} \in \mathcal{R}_1, \omega_2)$$

$$= P(\mathbf{x} \in \mathcal{R}_2 | \omega_1) P(\omega_1) + P(\mathbf{x} \in \mathcal{R}_1 | \omega_2) P(\omega_2)$$

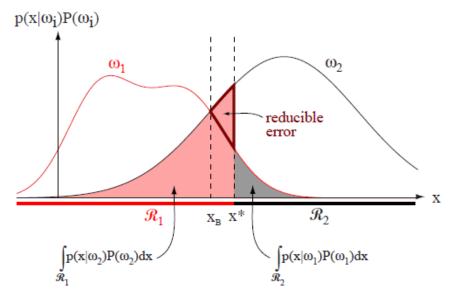
$$= \int_{\mathcal{R}_2} p(\mathbf{x} | \omega_1) P(\omega_1) d\mathbf{x} + \int_{\mathcal{R}_1} p(\mathbf{x} | \omega_2) P(\omega_2) d\mathbf{x}.$$

• 一般情况

$$P(correct) = \sum_{i=1}^{c} P(\mathbf{x} \in \mathcal{R}_{i}, \omega_{i})$$

$$= \sum_{i=1}^{c} P(\mathbf{x} \in \mathcal{R}_{i} | \omega_{i}) P(\omega_{i})$$

$$= \sum_{i=1}^{c} \int_{\mathcal{R}_{i}} p(\mathbf{x} | \omega_{i}) P(\omega_{i}) d\mathbf{x}$$



决策面为x_B时为最小错误率分类

• 最大后验概率决策(0-1 loss)的情况

University of Chinese Academy of Sciences

$$P(correct) = \int_{\mathbf{x}} \max_{i} p(\mathbf{x} \mid \omega_{i}) P(\omega_{i}) d\mathbf{x}$$

$$= \int_{\mathbf{x}} \max_{i} P(\omega_{i} \mid \mathbf{x}) p(\mathbf{x}) d\mathbf{x}$$

$$P(error) = \int_{\mathbf{x}} \left[1 - \max_{i} P(\omega_{i} \mid \mathbf{x}) \right] p(\mathbf{x}) d\mathbf{x}$$

$$P(\omega_{1} \mid x)$$

$$P(\omega_{2} \mid x)$$

$$1 - \max_{i} P(\omega_{i} \mid \mathbf{x})$$

40

讨论

- 贝叶斯分类器(基于贝叶斯决策的分类器)是最优的吗?
 - 最小风险、最大后验概率决策
 - 最优的条件: 概率密度、风险能准确估计
 - 具体的参数法、非参数法是贝叶斯分类器的近似,实际中难以达到最优
 - 判别模型:回避了概率密度估计,以较小复杂度估计 后验概率或判别函数
 - 一什么方法能胜过贝叶斯分类器: 在不同的特征空间才有可能!



下次课内容

- 第2章
 - 离散变量的贝叶斯决策
 - 复合模式分类
- 第3章
 - 最大似然参数估计
 - 贝叶斯估计

