

## Continuous Random Variables

TF: Justin Zhu ([justinzhu@college.harvard.edu](mailto:justinzhu@college.harvard.edu)), Credits to Timothy Kang

### Practice Problems

#### Example 1. Uniform Power.

Let  $U \sim \text{Unif}(-1, 1)$

- (a) Compute  $E(U)$ ,  $\text{Var}(U)$ ,  $E(U^4)$ .
- (b) Find the PDF and CDF of  $U^2$ . Is it also a uniform distribution?

#### Solution

When trying to compute the PDF of a transformation, we usually work from the CDF first. Note that the PDF of  $U$  is  $f(x) = \frac{1}{2}$ .

- (a)  $E(U) = 0$  because the distribution is symmetric about 0. We need to calculate  $E(U^2)$  for the variance, so we have:

$$E(U^2) = \int_{-1}^1 u^2 \cdot \frac{1}{2} du = \left[ \frac{1}{6} u^3 \right]_{-1}^1 = \frac{1}{3}$$

$$\text{Therefore, } \text{Var}(U) = E(U^2) - E(U)^2 = \boxed{\frac{1}{3}}$$

Next, we do the same for  $E(U^4)$ :

$$E(U^4) = \int_{-1}^1 u^4 \cdot \frac{1}{2} du = \left[ \frac{1}{10} u^5 \right]_{-1}^1 = \boxed{\frac{1}{5}}$$

- (b) We first find the CDF.

$$P(U^2 < k) = P(-\sqrt{k} < U < \sqrt{k}) = \frac{2\sqrt{k}}{2} = \sqrt{k}$$

, which we can easily calculate graphically. If this wasn't possible to do graphically, then we would integrate the PDF of  $U$  between  $-k$  and  $k$ .

Therefore, the PDF is:

$$\frac{d}{dk} P(U^2 < k) = \boxed{\frac{1}{2\sqrt{k}}}$$

This is definitely **not** a uniform distribution. This also shows that  $U^k$  is not uniform anymore for any  $k > 1$ .

#### Example 2. Normal Squared.

Let  $Z \sim N(0, 1)$  with CDF  $\Phi$ . The PDF of  $Z^2$  is the function given by:

$$g(w) = \frac{1}{\sqrt{2\pi w}} e^{-w/2}$$

with a support of  $w \geq 0$ .

- (a) Find expressions for  $E(Z^4)$  as integrals in two different ways, one based on the PDF of  $Z$  and the other based on the PDF of  $Z^2$ .
- (b) Find  $E(Z^2 + Z + \Phi(Z))$ .
- (c) Find the CDF of  $Z^2$  in terms of  $\Phi$ ; do not find the PDF of  $g$ .

**Solution**

(a) Let  $W = Z^2$ , so  $W^2 = Z^4$ . By LOTUS,

$$E(Z^4) = \int_{-\infty}^{\infty} z^4 \varphi(z) dz = \int_0^{\infty} w^2 g(w) dw,$$

where  $\varphi(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2}$  is the PDF of  $Z$ , and  $g$  is as above. (Using techniques from Chapter 6, it turns out that this reduces to a very simple answer:  $E(Z^4) = 3$ .)

(b) By linearity, this is  $E(Z^2) + E(Z) + E(\Phi(Z))$ . The second term is 0 and the first term is 1 since  $E(Z) = 0$ ,  $\text{Var}(Z) = 1$ . The third term is  $1/2$  since by universality of the Uniform,  $\Phi(Z) \sim \text{Unif}(0, 1)$ . Thus, the value is  $3/2$ .

(c) For  $w \leq 0$ , the CDF of  $Z^2$  is 0. For  $w > 0$ , the CDF of  $Z^2$  is

$$P(Z^2 \leq w) = P(-\sqrt{w} \leq Z \leq \sqrt{w}) = \Phi(\sqrt{w}) - \Phi(-\sqrt{w}) = 2\Phi(\sqrt{w}) - 1.$$

**Example 3. Universality of the Uniform.**

Let  $U \sim \text{Unif}(0, 1)$ , and let  $X = -(\log(1 - U))^{1/3}$ . Find the CDF and PDF of  $X$ .

**Solution**

We calculate  $P(x \leq X)$ , but use the substitution given to try and solve a Uniform distribution's CDF instead:

$$\begin{aligned} P(X \leq x) &= P(-(\log(1 - U))^{1/3} \leq x) \\ &= P(\log(1 - U)^{1/3} \geq -x) \\ &= P(\log(1 - U) \geq -x^3) \\ &= P(1 - e^{-x^3} \geq U) \\ &= P(U \leq 1 - e^{-x^3}) \\ &= \boxed{1 - e^{-x^3}} \end{aligned}$$

which is the CDF of  $X$ .

The PDF of  $x$  is the derivative, so:

$$f(x) = \frac{\partial}{\partial x} (1 - e^{-x^3}) = 3x^2 e^{-x^3}$$

**Example 4. Continuous RV manipulation.**

Let  $X$  be a continuous r.v. with CDF  $F$  and PDF  $f$ .

- (a) Find the conditional CDF  $X$  given that  $X > a$  (where  $a$  is a constant with  $P(X > a) \neq 0$ ).
- (b) Find the conditional PDF of  $X$  given  $X > a$  (this is the derivative of the conditional CDF).

- (c) Check that the conditional PDF from (b) is a valid PDF, by showing directly that it is nonnegative and integrates to 1.

**Solution**

(a) We have  $P(X \leq x|X > a) = 0$  for  $x \leq a$ . For  $x > a$ ,

$$P(X \leq x|X > a) = \frac{P(a < X \leq x)}{P(X > a)} = \frac{F(x) - F(a)}{1 - F(a)}.$$

(b) The derivative of the conditional CDF is  $f(x)/(1 - F(a))$  for  $x > a$ , and 0 otherwise.

(c) We have  $f(x)/(1 - F(a)) \geq 0$  since  $f(x) \geq 0$ . And

$$\int_a^\infty \frac{f(x)}{1 - F(a)} dx = \frac{1}{1 - F(a)} \int_a^\infty f(x) dx = \frac{1 - F(a)}{1 - F(a)} = 1.$$