Sequential Learning in DGLMs

0. Set m_0, C_0, G_t (see example 4.2), W_t (see supplementary material).

Bernoulli logistic DGLM

- 1. At any time t-1, current information is summarized as $(heta_{t-1}|D_{t-1},I_{t-1})\sim (m_{t-1},C_{t-1})$
- 2. Evolutionary equation induces 1-step ahead prior: $(\theta_t|D_{t-1},I_{t-1})\sim (a_t,R_t), a_t=G_tm_{t-1},R_t=G_tC_{t-1}G_t^T+W_t$
- 3. Conjugate prior: $\pi_t \sim Beta(\alpha_t, \beta_t)$ (no action)
- 4. Solve (α_t, β_t) from $F_t'a_t =: f_t = \psi(\alpha_t) \psi(\beta_t)$, $F_t'R_tF_t =: q_t = \psi'(\alpha_t) + \psi'(\beta_t)$, where ψ and ψ' are digamma and trigamma function
- 5. 1-step ahead forecast: $\Pr[z_t = 1 | \mathcal{D}_{t-1}, \mathcal{I}_{t-1}] = \alpha_t / (\alpha_t + \beta_t)$
- 6. On observing z_t , we have the posterior. (no action)
- 7. Calculate $g_t = \psi\left(\alpha_t + z_t\right) \psi\left(\beta_t + 1 z_t\right)$, $p_t = \psi'\left(\alpha_t + z_t\right) + \psi'\left(\beta_t + 1 z_t\right)$ (z_t is observed!)
- 8. Posterior update $(m{ heta}_t | \mathcal{D}_t) \sim (\mathbf{m}_t, \mathbf{C}_t)$: $\mathbf{m}_t = \mathbf{a}_t + \mathbf{R}_t \mathbf{F}_t \left(g_t f_t \right) / g_t$ and

$$\mathbf{C}_t = \mathbf{R}_t - \mathbf{R}_t \mathbf{F}_t \mathbf{F}_t' \mathbf{R}_t' \left(1 - p_t/q_t\right)/q_t$$

(This is Kalman filtering)

Filtering and Forecasting

In DGLM, The positive count model component will be updated only when $z_t = 1$. When $z_t = 0$, the positive count value is treated as missing. I.e., update both Poisson and Bernoulli when $z_t = 1$, and update only Bernoulli when $z_t = 0$.

Forward filtering: Bernoulli and Poisson are updated separately.

Forecasting:

t+k: Implied mixture of Bernoulli and shifted Poisson (step 5 above, Appendix 3)

At time t, the k-step ahead forecast distribution has a pdf of the compositional form

$$p\left(y_{t+k}|\mathcal{D}_{t},\mathcal{I}_{t},\pi_{t+k}
ight)=\left(1-\pi_{t+k}
ight)\delta_{0}\left(y_{t+k}
ight)+\pi_{t+k}h_{t,t+k}\left(y_{t+k}
ight)$$

where

- (π_{t+k} | D_t, I_t) ~ Be(α_t⁰(k), β_t⁰(k)) and δ₀(y) is the Dirac delta function at zero.
- h_{t,t+k}(y_{t+k}) is the density of y_{t+k} = 1+x_{t+k} where x_{t+k} is negative binomial.

$$(x_{t+k} \mid \mathcal{D}_t, \mathcal{I}_t) \sim \text{Nb}\left(\alpha_t^+(k), \frac{\beta_0^+(k)}{1 + \beta_t^+(k)}\right).$$

 The defining parameters α_t⁰(k), β_t⁰(k), α_t⁺(k), β_t⁺(k) are computed from the binary and positive count DGLMs, respectively.

That is, the mixture places probability $1 - \pi_{t+k}$ on $y_{t+k} = 0$, and probability π_{t+k} on the implied shifted negative binomial distribution.

Full joint forecast of $y_{t+1:t+k}$: Monte Carlo samples from $p(y_{t+1:t+k}|D_t,I_t)$ (Appendix 3)

Try: (Maybe a "future possibilities" part in the presentation on Nov 19)

- 1. Normal DLM for simulated continuous data
- 2. performance of this model when it is not the true model (via simulation). e.g., when the true data generating process is simply poisson
- 3. performance of this model when the counts are not sparse and are large (so maybe it can be treated as continuous): e.g., 中国进出口总额数据, and compare it to normal DGLM or ARMA
- 4. W_t : no need to specify it carefully according to supplementary material. Try different specifications and look at their influence on the outcome (maybe in the univariate case)
- 5. different loss functions