Homework 4

1.

To show that $||x||_A = \sqrt{x^T Ax}$ is a norm:

- (i) by property of A that A is positive definite $x^T A x \ge 0$, then $||x||_A$ is nonnegative.
- (ii) also with property of A, $x^T A x \ge 0$ for all $x \in \mathbb{R}^d$ and with equality with if and only if x = 0, $||x||_A = 0$ is if and only if x = 0. (iii) homogeneous:

For any $t \in R$,

$$\left| |tx| \right|_A = \sqrt{(tx)^T A(tx)} = \sqrt{t^2 x^T A x} = |t| \sqrt{x^T A x} = |t| \left| |x| \right|_A$$

(iv) the triangle inequality:

Firstly, we can show that since $|x||_A = \sqrt{x^T A x}$, then

$$||x||_A = \sqrt{x^T A x} = \sqrt{x^T U U^T x} = ||U^T x||_2$$

Hence, the prof of triangle inequality can be transferred to

$$\begin{aligned} \left| |x+y| \right|_{A} & \leq \left| |x| \right|_{A} + \left| |y| \right|_{A} \\ \Leftrightarrow \left| \left| U^{T}x + U^{T}y \right| \right|_{2} & \leq \left| \left| U^{T}x \right| \right|_{2} + \left| \left| U^{T}y \right| \right|_{2} \\ \Leftrightarrow \left| \left| \alpha + \beta \right| \right|_{2} & \leq \underbrace{\left| \left| \alpha \right| \right|_{2}}_{\alpha = U^{T}x} + \underbrace{\left| \left| \beta \right| \right|_{2}}_{\beta = U^{T}y} \end{aligned}$$

And here is to show ℓ_2 norm triangle inequality:

$$||\alpha + \beta||_{2} = \sqrt{\sum_{i=1}^{d} (\alpha_{i} + \beta_{i})^{2}}$$

$$= \sqrt{\sum_{i=1}^{d} (\alpha_{i}^{2} + \beta_{i}^{2} + 2\alpha_{i}\beta_{i})}$$

$$\leq \sqrt{\sum_{i=1}^{d} (\alpha_i^2 + \beta_i^2) + 2\sqrt{\sum_{i=1}^{d} \alpha_i^2 \sum_{i=1}^{d} \beta_i^2}}$$

$$= \sqrt{\left(\sqrt{\sum_{i=1}^{d} \alpha_i^2 + \sqrt{\sum_{i=1}^{d} \beta_i^2}}\right)^2}$$

$$= \sqrt{\sum_{i=1}^{d} \alpha_i^2 + \sqrt{\sum_{i=1}^{d} \beta_i^2}}$$

$$= \left||\alpha||_2 + ||\beta||_2$$

As for Cauchy Inequality part, here is detail

$$2\sum\nolimits_{i = 1}^d ({{\alpha _i}{\beta _i}}) \le 2\sqrt {\sum\nolimits_{i = 1}^d {{\alpha _i}^2} \sum\nolimits_{i = 1}^d {{\beta _i}^2} }$$

So as shown above it satisfies triangle inequality.

Based on 4 properties, the $||x||_A = \sqrt{x^T A x}$ is a norm.

And based on the HW1, we can know how it can yield to distance.

2.

(a)

With

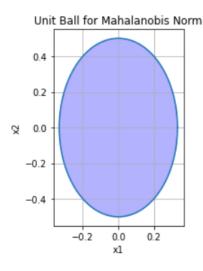
$$A = \begin{pmatrix} 9 & 0 \\ 0 & 4 \end{pmatrix}$$

Then

$$||x||_A = \sqrt{x^T A x} = \sqrt{9x_1^2 + 4x_2^2} \le 1$$

$$\Rightarrow$$
 $9x_1^2 + 4x_2^2 \le 1$

That is an ellipse and the figure drawn by Python is shown



And the blue area is for unit ball associated with Mahalanobis norm

$$K = \{x \in R^d \colon \big| |x| \big| \le 1\}$$

(b)

Hence,

Similar to what we've done in HW1, using properties of norm, we can obtain

$$||\theta x + (1 - \theta)y|| \le ||\theta x|| + ||(1 - \theta)y|| \text{ (triangle inequality)}$$
$$= |\theta|||x|| + |(1 - \theta)|||y|| \text{ (homogeneous)}$$

And since $x, y \in K$, we have $||x|| \le 1$, $||y|| \le 1$

$$|\theta| ||x|| + |(1 - \theta)| ||y|| = \theta ||x|| + (1 - \theta)||y||$$

$$\leq \theta + (1 - \theta) = 1$$

$$\Rightarrow ||\theta x + (1 - \theta)y|| \leq 1$$

$$\Rightarrow ||\theta x + (1 - \theta)y|| \in K, \text{ for any } 0 < \theta < 1$$

In other word, the unit ball with associated norm is a convex set

(c)

The conclusion is $K' \subseteq K$, prof is as follows For any $x \in K'$, then we have $||x||' \le 1$ Then,

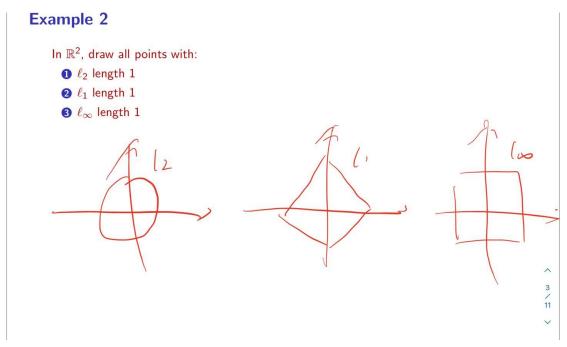
$$||x|| \le ||x||' \le 1$$
$$\Rightarrow x \in K$$

Therefore $K' \subseteq K$

(d)

It should be the unit ball with ℓ_1 norm.

Similar to the lecture,



this is a sketch when dimension is only 2 for simplicity. Obviously, for any dimension, it's still working, that ℓ_1 norm unit ball contains the least area.

By definition, we can easily know that

$$\sum\nolimits_{i=1}^{m}p_{i}=\sum\nolimits_{i=1}^{m}q_{i}=1, for\ every\ p\epsilon S$$

We use Lagrange formula to calculate,

$$\mathcal{L} = \sum_{p \in S} \sum\nolimits_{i=1}^{m} p_i ln \frac{p_i}{q_i} + \lambda \left(\sum\nolimits_{i=1}^{m} q_i - 1 \right)$$

Therefore, we calculate derivatives and let them be 0

$$\begin{cases} \frac{\partial \mathcal{L}}{\partial q_i} = \sum_{p \in S} \frac{-p_i}{q_i} + \lambda = 0, & \text{for every } 1 \leq i \leq m \\ \frac{\partial \mathcal{L}}{\partial \lambda} = \sum_{i=1}^m q_i - 1 = 0 \end{cases}$$

$$\Rightarrow \begin{cases} \frac{\sum_{p \in S} p_i}{q_i} = \lambda, & \text{for every } 1 \leq i \leq m \\ \sum_{i=1}^m q_i = 1 \end{cases}$$

$$\Rightarrow \begin{cases} \frac{\sum_{p \in S} p_i}{\lambda} = q_i, & \text{for every } 1 \leq i \leq m \\ \sum_{i=1}^m q_i = 1 \end{cases}$$

Then we combine these two and using the property mentioned at the beginning,

$$\sum_{i=1}^{m} q_i = \sum_{i=1}^{m} \frac{\sum_{p \in S} p_i}{\lambda} = \frac{\sum_{p \in S} \sum_{i=1}^{m} p_i}{\lambda}$$

$$\Rightarrow 1 = \frac{\sum_{p \in S} 1}{\lambda} = \frac{|S|}{\lambda}$$

$$\Rightarrow \lambda = |S|$$

Then

$$q_i = \frac{\sum_{p \in S} p_i}{\lambda} = \frac{\sum_{p \in S} p_i}{|S|}, for \ every \ 1 \le i \le m$$

Thus, we can obtain conclusion the desired q is

$$q = (q_1, q_2, \dots, q_m) = \left(\frac{\sum_{p \in S} p_1}{|S|}, \frac{\sum_{p \in S} p_2}{|S|}, \dots, \frac{\sum_{p \in S} p_m}{|S|}\right)$$

4.

given the hint using Jensen's inequality, and using f(z) = -lnz, we have

$$K(p,q)$$

$$= E_{X \sim p} \left(\frac{\ln p(x)}{\ln q(x)} \right)$$

$$= E_{X \sim p} - \left(\frac{\ln q(x)}{\ln p(x)} \right)$$

$$\geq -\ln \left(E_{X \sim p} \frac{q(x)}{p(x)} \right)$$

$$= -\ln \left(\sum_{x \in X} p(x) \frac{q(x)}{p(x)} \right)$$

$$= -\ln \left(\sum_{x \in X} q(x) \right) = 0$$

Therefore, we showed the nonnegativity.

$$K(p,q) \geq 0$$

5.

(a)

Using data in the table and the maximum likelihood estimation

$$\lambda = \frac{\sum_{k} N_{k} k}{\sum_{k} N_{k}} = \frac{0 \cdot 22 + 1 \cdot 66 + \dots + 8 \cdot 10}{500} = 3.154$$

$$\begin{bmatrix}
a = [22, 66, 106, 115, 85, 55, 28, 13, 10] \\
b = [i*a[i] & \text{for i in range}(1en(a))] \\
\text{sum}(b)
\end{bmatrix}$$
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$$[\text{sum}(b)/500]$$

3.154

(b)

Using the formula, $Pr(X = k) = \frac{e^{-\lambda}\lambda^k}{k!}$, and Python

```
import numpy as np
import math
k=[round(500*np.exp(-3.154)*(3.154**i)/math.factorial(i), 2) for i in range(len(a))]
k
```

[21.34, 67.31, 106.14, 111.59, 87.99, 55.51, 29.18, 13.15, 5.18]

We can have

k	0	1	2	3	4	5	6	7	8	≥9
N_k	21.34	67.31	106.14	111.59	87.99	55.51	29.18	13.15	5.18	2.61

And they are pretty close to original hence prove we' done right.

6.

(a)

Since $X \sim U_{\lambda}(\lambda)$, then pdf is

$$f_X(x) = \begin{cases} \frac{1}{\lambda}, 0 \le x \le \lambda \\ 0, others \end{cases}$$

And the cdf is

$$F_X(x) = \begin{cases} 0, x < 0 \\ \frac{x}{\lambda}, 0 \le x \le \lambda \\ 1, x > 1 \end{cases}$$

(b)

Using maximum likelihood estimation,

$$L(\lambda) = \prod_{i=1}^{n} f_{X_i}(x) = \begin{cases} \frac{1}{\lambda^n}, 0 \le x_1, x_2, \dots, x_n \le \lambda \\ 0, others \end{cases}$$

For this case, unlike usual way to compute derivative, we found that this function increased as λ decreased with $0 \le x_1, x_2, ..., x_n \le \lambda$. That is to say, the function is maximized when $\lambda = \max(x_1, x_2, ..., x_n)$. Therefore, using MLE, we can obtain $\lambda = \max(x_1, x_2, ..., x_n)$.

7.

(a)

$$EX = \int_{-\infty}^{\infty} x p(x) dx$$

$$= \int_{0}^{\infty} x \frac{\beta^{\alpha}}{\Gamma(\alpha)} x^{\alpha - 1} e^{-\beta x} dx$$

$$= \frac{1}{\Gamma(\alpha)} \int_{0}^{\infty} (x \beta)^{\alpha} e^{-\beta x} dx$$

$$= \frac{1}{\beta \Gamma(\alpha)} \int_{0}^{\infty} (\beta x)^{\alpha} e^{-\beta x} d(\beta x)$$

$$= \frac{1}{\beta \Gamma(\alpha)} \underbrace{\int_{0}^{\infty} t^{\alpha} e^{-t} dt}_{t = \beta x}$$

$$= \frac{1}{\beta \Gamma(\alpha)} \Gamma(\alpha + 1)$$
$$= \frac{1}{\beta \Gamma(\alpha)} \alpha \Gamma(\alpha)$$
$$= \frac{\alpha}{\beta}$$

(b)

Based on (a), we have $(EX)^2 = \frac{\alpha^2}{\beta^2}$, next we compute $E(X^2)$ Similar to (a),

$$E(X^{2}) = \int_{-\infty}^{\infty} x^{2} p(x) dx$$

$$= \int_{0}^{\infty} x^{2} \frac{\beta^{\alpha}}{\Gamma(\alpha)} x^{\alpha - 1} e^{-\beta x} dx$$

$$= \frac{\beta^{\alpha}}{\Gamma(\alpha)} \int_{0}^{\infty} x^{\alpha + 1} e^{-\beta x} dx$$

$$= \frac{\beta^{\alpha}}{\Gamma(\alpha)} \frac{1}{\beta^{\alpha + 2}} \int_{0}^{\infty} (\beta x)^{\alpha + 1} e^{-\beta x} d(\beta x)$$

$$= \frac{1}{\Gamma(\alpha)\beta^{2}} \underbrace{\int_{0}^{\infty} t^{\alpha + 1} e^{-t} dt}_{t = \beta x}$$

$$= \frac{1}{\Gamma(\alpha)\beta^{2}} \Gamma(\alpha + 2)$$

$$= \frac{1}{\Gamma(\alpha)\beta^{2}} (\alpha + 1) \alpha \Gamma(\alpha)$$

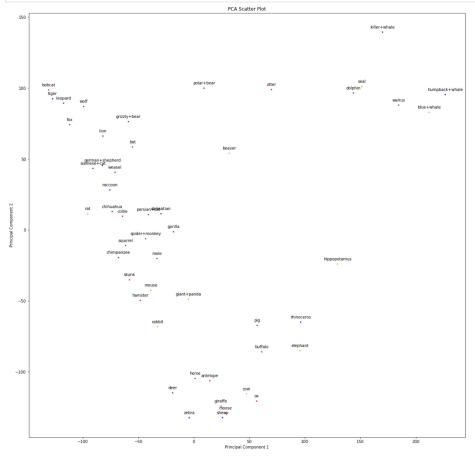
$$= \frac{(\alpha + 1)\alpha}{\beta^{2}}$$

Hence we can obtain variance

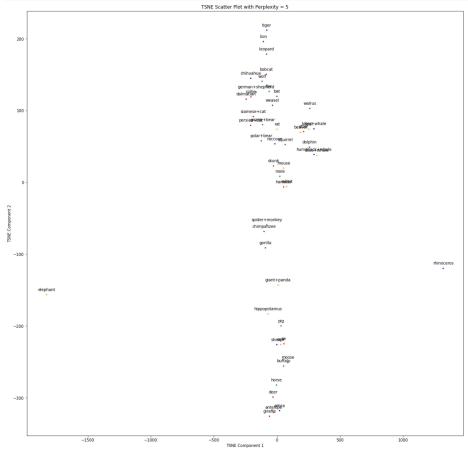
$$Var(X) = E(X^2) - (EX)^2$$

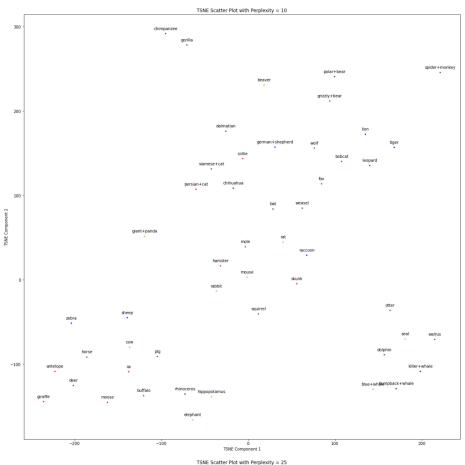
$$= \frac{(\alpha+1)\alpha}{\beta^2} - \frac{\alpha^2}{\beta^2}$$
$$= \frac{\alpha}{\beta^2}$$

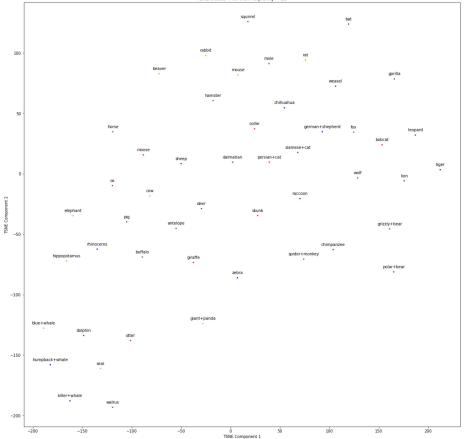
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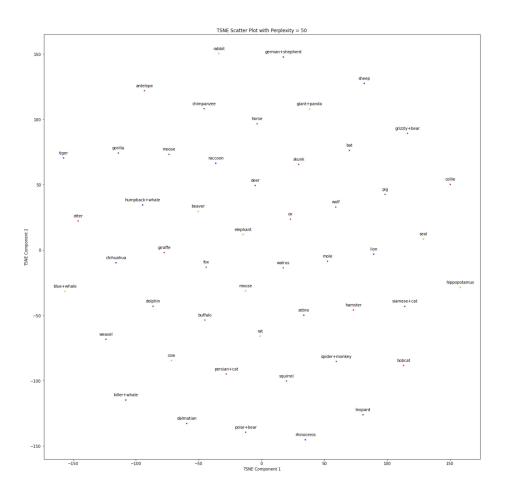












```
In [7]:

def getDistanceList(aList):
    theDistanceList(aList):
    for j in range(in(aList)-1):
        for j in range(in(aList)-1):
            theDistanceList append(op, linalg.norm(aList(j]))

return theDistanceList

def getDistortionFortNNE(0,Z_list):
    for perplexity, Z in zip([5, 10, 25, 50], Z_list):
        D_blat_TNNE = getDistanceList(Color (Linal (L
```

from what I've seen, the best is TSNE with perplexity 10 which acheives the lowest distortion

In []: