

Warmup

The joint distribution of X and Y is drawn below:

		$\frac{3}{8}$	$\frac{1}{2}$	$\frac{1}{8}$	$P(X)$
					$P(Y)$
1		$\frac{1}{4}$	$\frac{1}{3}$	$\frac{1}{12}$	$\frac{2}{3}$
0		$\frac{1}{8}$	$\frac{1}{6}$	$\frac{1}{24}$	$\frac{1}{3}$
Y	X	0	1	2	

a) X and Y are independent ✓

b) If we divide both rows by their marginal probability we get the same answer.

c) $P(X = x|Y = 0) = P(X = x|Y = 1)$ ✓

d) All of the above

check $P(X,Y) = P(X)P(Y)$
for all cells
ex $\frac{1}{4} = \frac{3}{8} \cdot \frac{2}{3}$

note that
 $P(X|Y=\frac{2}{3}) = \frac{P(X,Y=\frac{2}{3})}{P(Y=\frac{2}{3})}$
by Bayes' rule
so b) and c) are equivalent

Last time

Sec 3.1 Random Variables

The event $(X=x, Y=y)$ is the intersection of events $X=x$ and $Y=y$. ↖ sometimes written (x, y)

$P(X+Y=s)$ is the sum of $P(X=x, Y=y)$ for all x, y such that $x+y=s$

i.e
$$P(X+Y=s) = \sum_{(x,y): x+y=s} P(x,y) = \sum_{\text{all } x} P(x, s-x)$$

Independence of (X, Y, Z) means

$$P(X=x, Y=y, Z=z) = P(X=x)P(Y=y)P(Z=z) \quad \text{for all } x \in X, \\ y \in Y, z \in Z.$$

Today

- ① Sec 3.1 Sums of independent Poissons is Poisson
- ② Sec 3.2 Expectation of a RV.

① Sum of independent Poisson is Poisson

informal argument:

$$X_1 \sim \text{Bin}(1000, \frac{1}{1000}) \approx \text{Pois}(1)$$

$$X_2 \sim \text{Bin}(2000, \frac{1}{1000}) \approx \text{Pois}(2)$$

$$X_1 + X_2 \sim ? \quad \text{Bin}(3000, \frac{1}{1000}) \approx \text{Pois}(3)$$

✓ # heads in 3000 coin tosses
at a $p = \frac{1}{1000}$ coin

Proven in appendix to these notes

Claim If $X \sim \text{Pois}(\mu)$ and $Y \sim \text{Pois}(\lambda)$
are independent then

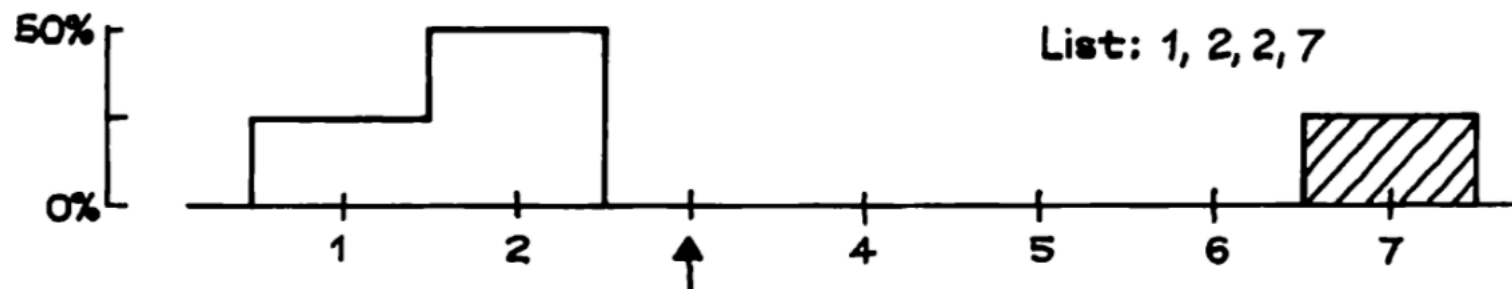
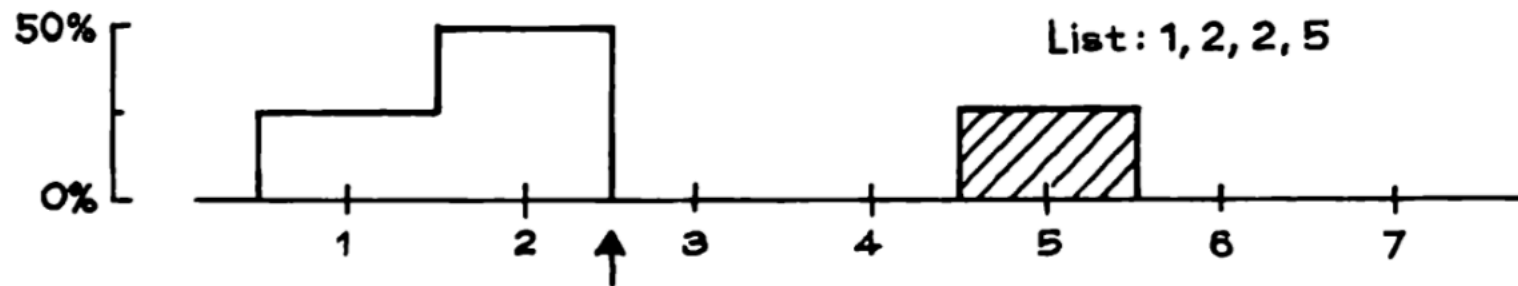
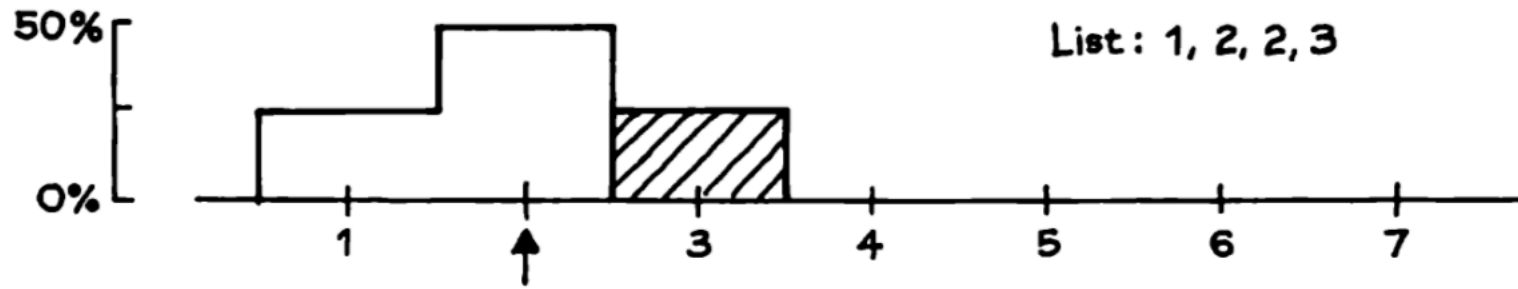
$$S = X + Y \sim \text{Pois}(\mu + \lambda).$$

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Sec 3.2 Expectation

$$E(X) = \sum_{x \in X} x \cdot P(X=x)$$

$$E(X) = 1 \cdot \frac{1}{4} + 2 \cdot \frac{1}{2} + 3 \cdot \frac{1}{4} = 2$$



$$E(X) = \sum_{x \in X} x \cdot P(X=x)$$

Properties of Expectation — P167 Pitman

$$(1) E(c) = c$$

$$(2) E(X+Y) = E(X) + E(Y) \quad (X, Y \text{ don't need to be independent})$$

$$(3) E(aX + b) = aE(X) + b$$

Indicators

An indicator is a RV that has only 2 values 1 (w/ prob p) and 0 (with prob $1-p$),

$$I = \begin{cases} 1 & \text{prob } p \\ 0 & \text{prob } 1-p \end{cases}$$

— Same as a Bernoulli p trial.

$$E(I) = 1 \cdot p + 0 \cdot (1-p) = p$$

RV that are counts can often be written as a sum of indicators.

$$\text{ex } X \sim \text{Bin}(n, p)$$

↖ # successes in n Bernoulli p trials,

$$\text{ex } X = \# \text{ heads in } n \text{ flips of } p \text{ coin}$$

$$X = I_1 + I_2 + \dots + I_n$$

$$\text{where } I_j = \begin{cases} 1 & \text{if } j^{\text{th}} \text{ trial success} \\ 0 & \text{else} \end{cases} \quad \text{— } p$$

$$E(X) = \underbrace{E(I_1)}_p + \dots + \underbrace{E(I_n)}_p = \boxed{np}$$

indicators are independent since trials are indep.

eg $X = \# \text{ aces in a poker hand}$ from a deck of cards
 $X \sim HG(n=5, N=52, b=4)$

a) what are the range of values of X ?

0, 1, 2, 3, 4

b) write X as a sum of indicators χ_i .

$$X = I_1 + I_2 + I_3 + I_4 + I_5$$

c) How is I_z defined? $P = \frac{4}{52} = \frac{1}{13}$

$$I_2 = \begin{cases} 1 & \text{if } 2^{\text{nd}} \text{ and } 1^{\text{st}} \text{ are} \\ 0 & \text{else} \end{cases}$$

d) Find $E(I_2) = 1/3$

e) Find $E(X) = E(I_1) + E(I_2) + E(I_3) + E(I_4) + E(I_5)$

$\parallel \quad \parallel \quad \parallel \quad \parallel$
 $\frac{1}{13} \quad \frac{1}{13} \quad \frac{1}{13} \quad \frac{1}{13}$

$= \boxed{\frac{5}{13}}$

Another more complicated solution?

Note

You may define $I_2 = \begin{cases} 1 & \text{if get 2 ones} \\ 0 & \text{else} \end{cases}$

so

$$X = I_1 + 2I_2 + 3I_3 + 4I_4$$

This is also correct but more complicated.

$$E(I_1) = \frac{\binom{4}{1} \binom{48}{4}}{\binom{52}{5}}$$

$$E(I_3) = \frac{\binom{4}{3} \binom{48}{2}}{\binom{52}{5}}$$

$$E(I_2) = \frac{\binom{4}{2} \binom{48}{3}}{\binom{52}{5}}$$

$$E(I_4) = \frac{\binom{4}{4} \binom{48}{1}}{\binom{52}{5}}$$

$$\begin{aligned} \text{so } E(X) &= \frac{1}{\binom{52}{5}} \left[\binom{4}{1} \binom{48}{4} + 2 \cdot \binom{4}{2} \binom{48}{3} + \right. \\ &\quad \left. 3 \cdot \binom{4}{3} \binom{48}{2} + 4 \cdot \binom{4}{4} \binom{48}{1} \right] \\ &= 5 \cdot \left(\frac{4}{52} \right) \leftarrow \text{I checked this in R} \end{aligned}$$

Note $E(X) = 1 \cdot P(X=1) + 2 \cdot P(X=2) + 3 \cdot P(X=3) + 4 \cdot P(X=4)$

ex

A drawer contains s black socks and s white socks ($s > 0$). I pull two socks out at random without replacement and call that my first pair. Then I pull out two socks at random without replacement and call that my second pair. I proceed in this way until I have s pairs and the drawer is empty. Find the expected number of pairs in which two socks are different colors.

$$X = \# \text{ of mismatched pairs } \rightarrow (\text{out of } s)$$
$$I_2 = \begin{cases} 1 & \text{if the second pair is mismatch} \\ 0 & \text{else} \end{cases}$$
$$X = I_1 + \dots + I_s$$
$$E(X) = s \cdot \frac{\binom{s}{1}\binom{s}{1}}{\binom{2s}{2}}$$
$$P = \frac{2s}{2s} \cdot \frac{s}{2s-1}$$

1st sock 2nd sock

$$\text{or } P = \frac{\binom{s}{1}\binom{s}{1}}{\binom{2s}{2}}$$

Appendix

Claim If $X \sim \text{Pois}(\mu)$ and $Y \sim \text{Pois}(\lambda)$ are independent then

$$S = X + Y \sim \text{Pois}(\mu + \lambda).$$

To prove this you need to know 2 facts:

Recall binomial theorem

$$\begin{aligned}(a+b)^3 &= \binom{3}{3} a^3 b^0 + \binom{3}{2} a^2 b^1 + \binom{3}{1} a^1 b^2 + \binom{3}{0} a^0 b^3 \\ &= a^3 + 3a^2b + 3ab^2 + b^3\end{aligned}$$

$$(a+b)^n = \sum_{k=0}^n \binom{n}{k} a^k b^{n-k}$$

Recall $X \sim \text{Pois}(\mu)$

$$P(X=k) = \frac{e^{-\mu} \mu^k}{k!}$$

Pf/ $P(S=s)$ addition rule

$$\begin{aligned}P(S=s) &= P(X=0, Y=s) + P(X=1, Y=s-1) + \dots + P(X=s, Y=0) \\ &\stackrel{\text{summation notation}}{=} \sum_{k=0}^s P(X=k, Y=s-k) \\ &\stackrel{\text{independence of } X \text{ and } Y}{=} \sum_{k=0}^s P(X=k) P(Y=s-k)\end{aligned}$$

Poisson
formula

$$= \sum_{k=0}^s \frac{e^{-\mu} \mu^k}{k!} \cdot \frac{e^{-\lambda} \lambda^{s-k}}{(s-k)!}$$

$$\frac{s!}{s!} = 1$$

$$= e^{-(\lambda+\mu)} \frac{1}{s!} \sum_{k=0}^s \frac{s!}{k!(s-k)!} \mu^k \lambda^{s-k}$$

binomial
thm

$$= e^{-(\lambda+\mu)} \frac{1}{s!} (\mu+\lambda)^s$$

$$\Rightarrow S \sim \text{Pois}(\mu+\lambda).$$

Poisson
formula.

