

warmup

Suppose customers are arriving at a ticket booth at rate of five per minute, according to a Poisson arrival process. Find the probability that:

At least one customer arrives within 40 seconds after the arrival of the 13th customer

$$P(\underbrace{T_{14} - T_{13}}_{W \sim \text{Exp}(\lambda=5)} < \frac{2}{3}) = 1 - P(\underbrace{T_{14} - T_{13}}_W \geq \frac{2}{3}) = \boxed{1 - e^{-5 \cdot \frac{2}{3}}}$$

alternatively

$$= 1 - P(\text{no arrivals in time } t)$$

$$\approx 1 - e^{-5 \cdot \frac{2}{3}} \quad \text{by Poisson formula.}$$

Announcement:

Wednesday (lec 21) is a special lecture on moment generating functions (not in textbook),

Last time

Sec 4.2 Exponential Distribution

$$f(t) = \lambda e^{-\lambda t} \text{ where } \lambda = 1/E(T)$$

Sec 4.2 Gamma Distribution



$$T_1 \sim \text{Exp}(\lambda), \lambda > 0$$

$$f(t) = \begin{cases} \lambda e^{-\lambda t} & t \geq 0 \\ 0 & \text{else} \end{cases} \quad \leftarrow \text{Variable part}$$

$$T_r \sim \text{Gamma}(r, \lambda), \quad \lambda > 0, r > 0$$

$$f(t) = \begin{cases} \frac{1}{\Gamma(r)} \lambda^r t^{r-1} e^{-\lambda t} & t \geq 0 \\ 0 & \text{else} \end{cases} \quad \text{where } \Gamma(r) = \int_0^\infty t^{r-1} e^{-t} dt$$

$$r \in \{1, 2, 3, \dots\} \\ \text{then } \Gamma(r) = (r-1)!$$

$$T_r = w_1 + w_2 + \dots + w_r, \quad w_i \stackrel{\text{i.i.d.}}{\sim} \text{Exp}(\lambda)$$

$$E(w_1) = \frac{1}{\lambda} \Rightarrow E(T_r) = \frac{r}{\lambda}$$

$$\text{Var}(w_1) = \frac{1}{\lambda^2} \Rightarrow \text{Var}(T_r) = \frac{r}{\lambda^2}$$

Ex

A random variable X has non negative values and density $c x^4 e^{-3x}$ for $0 \leq x < \infty$, and some constant c .

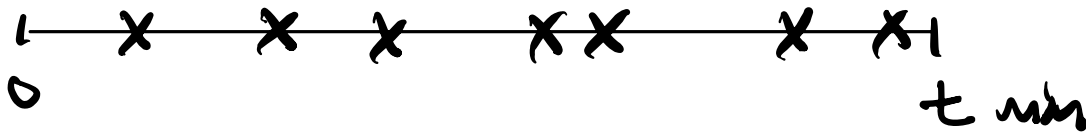
What distribution is X ? $X \sim \text{Gamma}(5, 3) \quad c = \frac{3^5}{4!}$

$$\text{Find } \text{Var}(X) = \frac{5}{9}$$

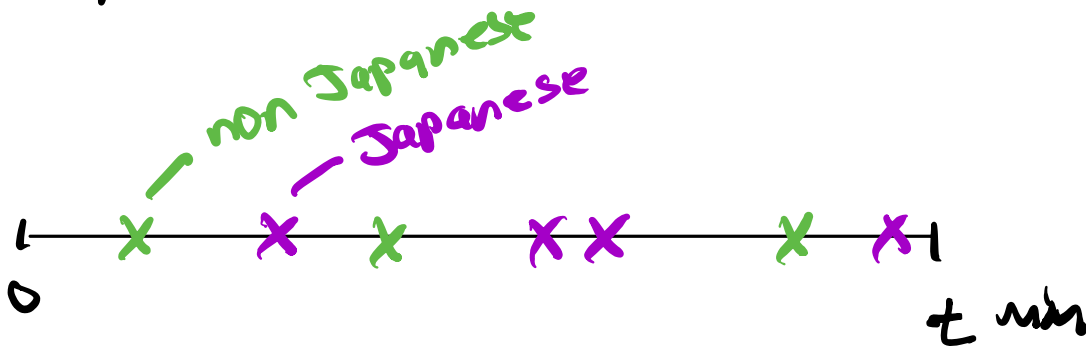
Sec 3.5 Poisson thinning

Cars arrive at a toll booth according to a Poisson process at a rate λ arrivals/min

$X = \# \text{ cars arriving at a toll booth in } t \text{ min. } X \sim \text{Pois}(\lambda t)$



The chance of a car arriving being a Japanese import is p .



$$\# \text{ cars} \sim \text{Pois}(\lambda t)$$

$$\# \text{ Japanese imports} \sim \text{Pois}(p\lambda t)$$

$$\# \text{ non Japanese} \sim \text{Pois}(2\lambda t)$$

} indep

Japanese car is a thinned poisson process $\text{Pois}(p\lambda t)$

Today

① more practice

② Sec 4.2 Competing Exponentials

③ Sec 4.2 Memoryless property

① More practice

A family is getting ready for their trip to Yosemite. Each person is in their room, packing their bags. For each person, the time it takes them to pack their bag is exponentially distributed and independent of the time it takes any other person. On average, it takes each parent 1 hour and each child 2 hours to get ready. In a family with 2 parents and 4 children, what is the probability that it takes the family more than 2 hours to get ready?

let $P_1, P_2, C_3, C_4, C_5, C_6$ be the packing time of the parents and children.

$$P_1, P_2 \stackrel{iid}{\sim} \text{Exp}(1)$$

$$C_1, C_2, C_3, C_4 \stackrel{iid}{\sim} \text{Exp}(\frac{1}{2}) \text{ since } \lambda = \frac{1}{E(C_i)}$$

$$M = \max(P_1, P_2, C_3, C_4, C_5, C_6)$$

$P(M > 2)$ is the probability at least one person isn't packed in 2 hrs.

$$P(M > 2) = 1 - P(M \leq 2) \quad \leftarrow \text{probability everyone is packed in 2 hrs.}$$

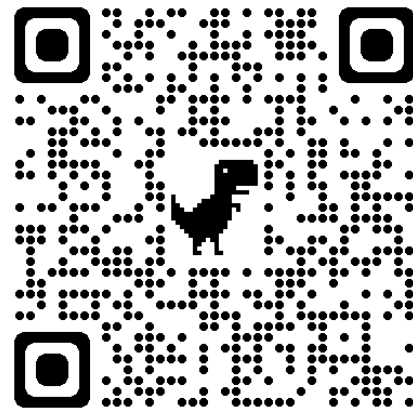
$$= 1 - P(P_1 \leq 2) P(P_2 \leq 2) P(C_3 \leq 2) \dots P(C_6 \leq 2)$$
$$\quad \quad \quad \overset{1 - e^{-1 \cdot 2}}{\quad} \quad \overset{1 - e^{-1 \cdot 2}}{\quad} \quad \overset{1 - e^{-\frac{1}{2} \cdot 2}}{\quad} \quad \overset{1 - e^{-\frac{1}{2} \cdot 2}}{\quad}$$

$$= \boxed{1 - (1 - e^{-2})^2 (1 - e^{-1})^4}$$

② Competition → Exponentiality

tinyurl.com/mav6-2023

Stat 134



1. GSI Brian and Yiming are each helping a student. Brian and Yiming see students at a rate of λ_B and λ_Y students per hour respectively.

Let

$B = \text{wait time for Brian} \sim \text{Exp}(\lambda_B)$

$Y = \text{wait time for Yiming} \sim \text{Exp}(\lambda_Y)$

} indep

What distribution is $T = \min(B, Y)$?

Hint: compute $P(T > t)$

a $\text{Exp}(\max(\lambda_B, \lambda_Y))$

b $\text{Exp}(\lambda_B - \lambda_Y)$

☒ c $\text{Exp}(\lambda_B + \lambda_Y)$

d none of the above

$$P(T > t) = P(B > t)P(Y > t) = e^{-\lambda_B t} e^{-\lambda_Y t} = e^{-(\lambda_B + \lambda_Y)t}$$

$$\Rightarrow T \sim \text{Exp}(\lambda_B + \lambda_Y)$$

SO

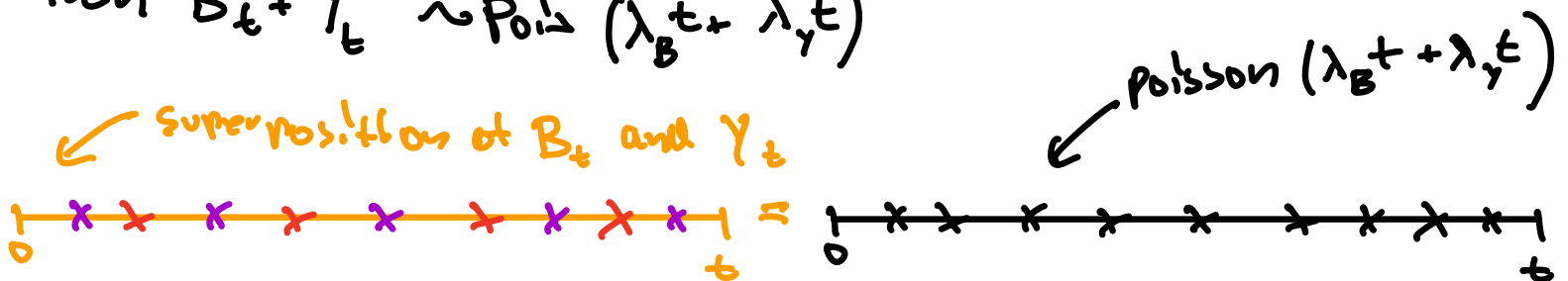
If X_1, \dots, X_n are indep exponentials
with rates $\lambda_1, \dots, \lambda_n$
then $\min(X_1, \dots, X_n) \sim \text{Exp}(\lambda_1 + \dots + \lambda_n)$

What is $P(B < Y)$?

superposition of Poisson random scatter:

let $B_t \sim \text{Pois}(\lambda_B t)$ and $Y_t \sim \text{Pois}(\lambda_Y t)$
be independent PRS corresponding to the number of
arrivals of Brian and Yiminy in time t .

Then $B_t + Y_t \sim \text{Pois}(\lambda_B t + \lambda_Y t)$



competing exponentials:

Let X = time until the first Brian arrival
 Y = time until the first Yiminy arrival

What is the chance, p , the first arrival is Brian?

$$P \approx \text{Prob}(B < Y)$$

$$= \text{Prob}(\text{an arrival is Brian})$$

$$B_y \text{ defn}, B_t \sim \text{Pois}(\lambda_B t)$$

By Poisson thinning

$$B_t \sim \text{Pois}(\rho(\lambda_B t + \lambda_Y t))$$

↑ Prob an arrival is Brian

$$\Rightarrow \lambda_B t = \rho(\lambda_B t + \lambda_Y t)$$

$$\Rightarrow \rho = \frac{\lambda_B t}{\lambda_B t + \lambda_Y t} = \boxed{\frac{\lambda_B}{\lambda_B + \lambda_Y}}$$

$$P(B < Y)$$

③ The memoryless property

Recall $P(T > k) = e^{-\lambda k}$

Let $T \sim \text{Exp}(\lambda)$

a) Find $P(T > 5) = e^{-5\lambda}$

b) Find $P(T > 13 | T > 8) = \frac{P(T > 13, T > 8)}{P(T > 8)} = \frac{P(T > 13)}{P(T > 8)}$
 $= \frac{e^{-13\lambda}}{e^{-8\lambda}} = e^{-5\lambda}$

The Exponential distribution has the memoryless property

Version 1

$$P(T > j+k | T > j) = P(T > k)$$

Version 2

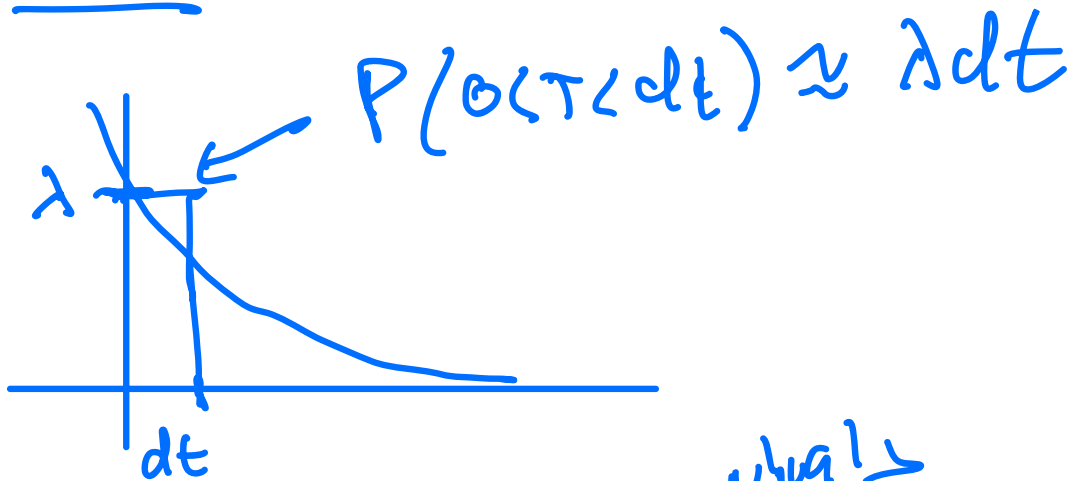
Small interval after time t .

$$P(T \in dt | T > t) = P(0 < T < dt)$$

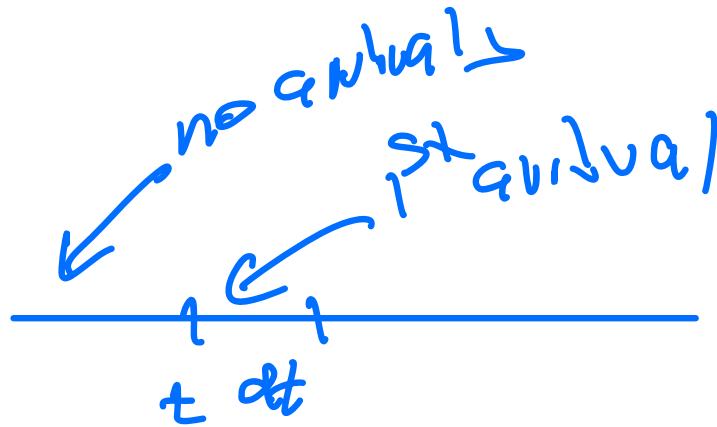
Small interval after time 0

Prove that $T \sim \text{Exp}(\lambda)$ satisfies
version 2:

RHS



LHS



$$P(T \in dt \mid T > t)$$

$$= P(1 \text{ arrival in } dt \mid \text{no arrival in } t)$$

$$= P(1 \text{ arrival in } dt) \text{ since arrivals are indep.}$$

$$= e^{-\lambda dt} \approx \lambda dt \text{ for small } \lambda,$$

Fact

— see appendix of lecture 12
for a proof.

Only two kinds of distributions are
memoryless:

geometric distribution of
nonnegative integers and
the exponential distribution
of nonnegative real numbers.