

Warmup

A tub of blueberry muffin batter has
 $\lambda = 2 \text{ bb/in}^3$ intensity of bb,

bb muffin 1 $\hookrightarrow 3 \text{ in}^3$

bb muffin 2 $\hookrightarrow 4 \text{ in}^3$

a) on average how many bb is in
muffin 1? $\mu_1 = \lambda \cdot 3 = \boxed{6 \text{ bb}}$

b) Find $P(5 \text{ bb in each muffin})$

$$X_1 \sim \text{Poi}(6) \quad P(X_1=5, X_2=5)$$

$$X_2 \sim \text{Poi}(8)$$

$$= \boxed{\frac{e^{-6} 6^5}{5!} \cdot \frac{e^{-8} 8^5}{5!}}$$

c) Find $P(10 \text{ bb total in both muffins
together})$,

$$X_1 + X_2 \sim \text{Poi}(14)$$

$$P(X_1 + X_2 = 10) = \boxed{\frac{e^{-14} 14^{10}}{10!}}$$

Announcements

For Wednesday's review, write down questions in discussion board on b-course by Tuesday 8pm.

Last time

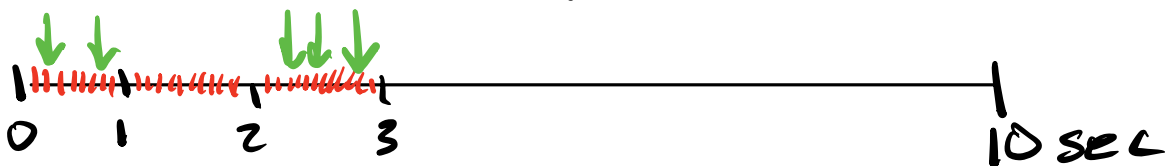
Sec 3.5 Poisson distribution

$$X \sim \text{Pois}(\mu)$$

$$P(X=k) = \frac{e^{-\mu} \mu^k}{k!}$$

$$E(X) = \text{Var}(X) = \mu.$$

Poisson Process or Poisson Random Scatter (PRS);
e.g. radioactive decay of Americium 241 in 10 seconds



Assumptions

- ① no two particles arrive at the same time,
(this allows us to divide 10 sec into n small
time intervals each with at most one arrival.
- ② X is a sum of n ^{large} iid
Bernoulli(p) trials,
 _{n small}

$\mu = n p$ is avg # of arrivals in 10 sec.
 $\lambda = \mu / 10$ is the arrival rate per second,

$X = \# \text{ arrivals in 10 seconds.}$

Suppose $\lambda = 4 \text{ arrivals/sec}$

then $\mu = \lambda \cdot 10 = 40 \Rightarrow X \sim \text{Pois}(40)$

Americium has a long half life.

$Y = \# \text{ arrivals in } 12070 \text{ sec.}$

$Y \sim \text{Pois}(\lambda \cdot 12070)$
 $\lambda = 4$

Today

② mid-term review

② midterm review

Which distributions are ^{exactly or} (approximately) a sum of a fixed number of independent Bernoulli trials?

Discrete

| name and range | $P(k) = P(X = k)$ for $k \in \text{range}$ | mean | variance |
|---|--|------------------------------------|--|
| uniform on $\{a, a+1, \dots, b\}$ $\{1, 2, \dots, n\}$ | $\frac{1}{b-a+1}$ | $\frac{a+b}{2}$ $\frac{n+1}{2}$ | $\frac{(b-a+1)^2 - 1}{12}$ $\frac{n^2 - 1}{12}$ |
| Bernoulli (p) on $\{0, 1\}$ ✓ | $P(1) = p; P(0) = 1 - p$ | p | $p(1 - p)$ |
| binomial (n, p) on $\{0, 1, \dots, n\}$ ✓ | $\binom{n}{k} p^k (1 - p)^{n-k}$ | np | $np(1 - p)$ |
| Poisson (μ) on $\{0, 1, 2, \dots\}$ ✓ | $\frac{e^{-\mu} \mu^k}{k!}$ | μ | μ |
| hypergeometric (n, N, G) on $\{0, \dots, n\}$ X = $I_1 + I_2 + \dots + I_n$ Sum of dependent indicators $I_2 = \begin{cases} 1 & \text{if 2nd card good} \\ 0 & \text{else} \end{cases}$ $\frac{6}{N}$ | $\frac{\binom{G}{k} \binom{N-G}{n-k}}{\binom{N}{n}}$ | $\frac{nG}{N}$ | $n \left(\frac{G}{N} \right) \left(\frac{N-G}{N} \right) \left(\frac{N-n}{N-1} \right)$ |
| geometric (p) on $\{1, 2, 3, \dots\}$ X | $(1 - p)^{k-1} p$ | $\frac{1}{p}$ | $\frac{1 - p}{p^2}$ |
| geometric (p) on $\{0, 1, 2, \dots\}$ X | $(1 - p)^k p$ | $\frac{1 - p}{p}$ | $\frac{1 - p}{p^2}$ |
| negative binomial (r, p) on $\{0, 1, 2, \dots\}$ X | $\binom{k+r-1}{r-1} p^r (1 - p)^k$ | $\frac{r(1 - p)}{p}$ | $\frac{r(1 - p)}{p^2}$ |

normal $\phi(x)$ n σ^2
 not necessarily, but the sum of iid Bernoulli trials is approx normal.

De Morgan's rule: $(A \cap B)^c = A^c \cup B^c$
 $\Rightarrow A \cap B = (A^c \cup B^c)^c$

$$\text{So } \boxed{P(A \cap B) = 1 - P(A^c \cup B^c)}$$

Inclusion exclusion formula:

Let A_1, A_2, A_3 be dependent RVs with

$$P(A_i) = .9 \text{ for } i=1, 2, 3.$$

Find a lower bound for $P(\bigcap_{i=1}^3 A_i)$

$$P(\bigcap_{i=1}^3 A_i) = 1 - P(\bigcup_{i=1}^3 A_i^c) \quad \text{De Morgan's rule,}$$

$$\begin{aligned} P(\bigcup_{i=1}^3 A_i^c) &= P(A_1^c) + P(A_2^c) + P(A_3^c) \\ &\quad - P(A_1^c A_2^c) - \quad - \\ &\quad + P(A_1^c A_2^c A_3^c) \end{aligned}$$

$$\leq \boxed{P(A_1^c) + P(A_2^c) + P(A_3^c)}$$

$$\Rightarrow P(\bigcap_{i=1}^3 A_i) \geq 1 - .3 = \boxed{.7}$$

$3(.1)$
 $= .3$

expectation question

An urn contains 90 marbles, of which there are 20 greens, 20 blacks and 50 red marbles. Tom draws marble without replacement until the 6th green marble. Let $X = \#$ of marbles drawn. Example: GGGBRBGGBRG with $x = 11$. Find $\mathbb{E}[X]$.

Hint First find the expected number of marbles until the 1st green marble.

What is the min and max of X ?

$X = \#$ marbles until first green. — $1 - 71$

$$X = I_1 + \dots + I_{70} + 1$$

\nwarrow green marble at the end.

$$\mathbb{E}(X) = 70\left(\frac{1}{21}\right) + 1 \quad I_i = \begin{cases} 1 & \text{if } i^{\text{th}} \text{ not before 1st green} \\ 0 & \text{else.} \end{cases}$$

$1 - G_2 - G_3 - G_6 - G_{20} -$

\nwarrow not here.

$Y = \#$ marbles until 6th green

$$\mathbb{E}(Y) = 6 \left(70 \left(\frac{1}{21} \right) + 1 \right)$$

\nwarrow by symmetry,

ex Conditional distribution, Poisson

8. Let X_1 and X_2 be independent random variables such that for $i = 1, 2$, the distribution of X_i is Poisson (μ_i). Let m be a fixed positive integer. Find the distribution of X_1 given that $X_1 + X_2 = m$. Recognize this distribution as one of the famous ones, and provide its name and parameters.

$$\left. \begin{array}{l} X_1 \sim \text{Poi}(\mu_1) \\ X_2 \sim \text{Poi}(\mu_2) \end{array} \right\} \text{ indep}$$

$$P(X_i = k) = \frac{e^{-\mu_i} \mu_i^k}{k!}$$

$X_1 | X_1 + X_2 = m$ takes value $0, 1, 2, \dots, m$

$$\begin{aligned} P(X_1 = k | X_1 + X_2 = m) &= \frac{P(X_1 = k, X_2 = m - k)}{P(X_1 + X_2 = m)} \\ &= \frac{P(X_1 = k) P(X_2 = m - k)}{P(X_1 + X_2 = m)} \\ &= \frac{e^{-\mu_1} \mu_1^k}{k!} \cdot \frac{e^{-\mu_2} \mu_2^{m-k}}{(m-k)!} \\ &= \frac{e^{-(\mu_1 + \mu_2)} (\mu_1 + \mu_2)^m}{m!} \\ &= \binom{m}{k} \left(\frac{\mu_1}{\mu_1 + \mu_2} \right)^k \left(\frac{\mu_2}{\mu_1 + \mu_2} \right)^{m-k} \end{aligned}$$

$$\Rightarrow X_1 | X_1 + X_2 = m \sim \text{Bin}\left(m, \frac{\mu_1}{\mu_1 + \mu_2}\right)$$

Problem 4 (conditional probability)

Two jars each contains r red marbles and b blue marbles. A marble is chosen at random from the first jar and placed in the second jar. A marble is then randomly chosen from the second jar. Find the probability this marble is red.

X = the first marble color

Y = the second marble color.

R = red

B = blue

$$\frac{r+1}{r+b+1} \cdot \frac{r}{r+b}$$

$$P(Y=R) = P(Y=R|X=R)P(X=R) \\ + P(Y=R|X=B)P(X=B)$$

$$\frac{r}{r+b+1} \cdot \frac{b}{r+b}$$