

Warm up

**Conditional expectation of  $g(Y)$  given  $X = x$ :** Integrate  $g$  against the conditional density:

$$E(g(Y) | X = x) = \int g(y) f_Y(y | X = x) dy$$

ex

Suppose

$$f_Y(y | X = x) = 2x + 2y - 4xy \quad \text{for } 0 \leq x \leq 1 \text{ and } 0 \leq y \leq 1$$

Find  $E(Y | X = x)$

$$\int_{y=0}^{y=1} y(2x + 2y - 4xy) dy = \int_0^1 (2xy + 2y^2 - 4xy^2) dy$$

$$= \left. \frac{2xy^2}{2} + \frac{2y^3}{3} - \frac{4xy^3}{3} \right|_{y=0}^{y=1} = x + \frac{2}{3} - \frac{4x}{3}$$

$$= \left[ \frac{2}{3} - \frac{x}{3}, 0 \leq x \leq 1 \right]$$

## Last time

Average conditional expectation: (rule of iterated expect.)  
 $E(Y) = E(E(Y|X))$

$$E(Y) = \int E(Y|X=x) f_X(x) dx$$

In above problem, given  $f_X(x) = 1$ ,

find  $E(Y)$ .

$$E(Y) = E\left(\frac{2}{3} - \frac{x}{3}\right) = \int_{x=0}^{x=1} \left(\frac{2}{3} - \frac{x}{3}\right) \cdot 1 dx$$

$$= \left. \frac{2}{3}x - \frac{x^2}{6} \right|_0^1 = \frac{2}{3} - \frac{1}{6} = \frac{3}{6} = \boxed{\frac{1}{2}}$$

**Multiplication rule:** The joint density is the product of the marginal and the conditional

$$f(x, y) = f_X(x) f_Y(y|X=x)$$

$$= f_Y(y) f_X(x|Y=y)$$

mixed joint  $X$  cont,  $Y$  discrete

$$\text{joint}(x, y) = f_X(x) P(Y=y|X=x) = P_Y(Y=y) f_X(x|Y=y)$$

$$\Rightarrow f_X(x|Y=y) = \frac{f_X(x) P(Y=y|X=x)}{P(Y=y)}$$

Bayesian Statistics treats unknown parameters of distributions (ex  $\text{Ber}(p)$ ) as a RV.

i.e  
 $X \sim \text{Unit}(0,1) \leftarrow \text{RV}$

$$I_1, I_2 \sim \text{Ber}(X)$$

$$I_1 | X=x, I_2 | X=x \stackrel{\text{iid}}{\sim} \text{Ber}(x)$$

$f_X(x | I_1=1) = \frac{\overset{\text{Likelihood}}{P(I_1=1 | X=x)} \cdot \overset{\text{Prior}}{f_X(x)}}{\underset{\text{constant}}{P(I_1=1)}}$

Posterior  $\propto$  Likelihood  $\cdot$  Prior

A likelihood which is binomial and a prior that is beta always has a beta posterior (i.e.  $\text{lik} = \text{binomial}$ ,  $\text{prior} = \text{beta}$  are conjugate pairs).

The old posterior can now become the new prior if you get more data (say you know  $I_1=1, I_2=1$ ).

Today

- ① Sec 6.4 covariance and the variance of sum

# ① Sec 6.4 Covariance and variance of a sum

$$X, Y, S = X + Y$$

$$\text{mean } \mu_X, \mu_Y, \mu_S = \mu_X + \mu_Y$$

$$\begin{aligned} S - \mu_S &= X + Y - (\mu_X + \mu_Y) \quad D_S \text{ deviation from mean} \\ &= (X - \mu_X) + (Y - \mu_Y) = D_X + D_Y \end{aligned}$$

$$\begin{aligned} \text{Var}(S) &= E((S - \mu_S)^2) = E((D_X + D_Y)^2) \\ &= E(D_X^2 + D_Y^2 + 2D_X D_Y) \\ &= E(D_X^2) + E(D_Y^2) + 2E(D_X D_Y) \\ &\quad \underbrace{\quad}_{\text{Var}(X)} \quad \underbrace{\quad}_{\text{Var}(Y)} \quad \underbrace{\quad}_{\text{Cov}(X, Y)} \end{aligned}$$

Def<sup>n</sup> The Covariance of  $X$  and  $Y$  is  
 $\text{Cov}(X, Y) = E((X - \mu_X)(Y - \mu_Y))$

# Bilinearity Properties

Proved end of lecture,

- (a)  $\text{Cov}(X+Y, Z) = \text{Cov}(X, Z) + \text{Cov}(Y, Z)$   
(b)  $\text{Cov}(aX, bY) = ab \text{Cov}(X, Y)$

More generally

$$\begin{aligned} \text{Cov}\left(\sum_{i=1}^n a_i X_i, \sum_{j=1}^m b_j Y_j\right) \\ = \sum_{i=1}^n \sum_{j=1}^m a_i b_j \text{Cov}(X_i, Y_j) \end{aligned}$$

Proved end of lecture,

Thm  $\text{Cov}(X, Y) = E(XY) - E(X)E(Y)$

$$\text{Var}(X) = E((X - \mu_X)^2)$$

$$\text{Var}(X) = E(X^2) - (E(X))^2$$

Easy facts

$$\text{Cov}(X, X) = E(X^2) - E(X)^2 = \text{Var}(X)$$

$$\text{Cov}(X, X) = \text{Var}(X)$$

$$\text{Cov}(X, Y) = \text{Cov}(Y, X)$$

$$\text{Cov}(X, c) = 0$$

Constant

|||

Simplify

$$\begin{aligned} & \text{Cov}(X - 5Y, 3X + Y - Z + 10) \\ &= 3\text{Var}(X) + \text{Cov}(X, Y) - \text{Cov}(X, Z) + 0 - 15\text{Cov}(X, Y) \\ & - 5\text{Var}(Y) + 5\text{Cov}(Y, Z) + 0 \end{aligned}$$

$$\text{Cov}(X, -Z) = \text{Cov}(X, \textcircled{-1}Z) = -\text{Cov}(X, Z)$$

$$\text{Cov}(-5Y, 3X) = -15\text{Cov}(Y, X) = -15\text{Cov}(X, Y)$$

Recall  $X, Y$  independent

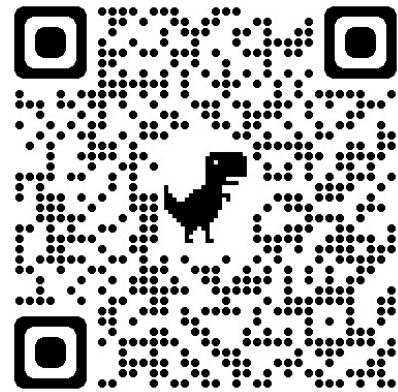
$$\Rightarrow E(XY) = E(X)E(Y)$$

$\text{Cov}(X, Y) = 0$  if  $X, Y$  independent,

Hence if  $X, Y$  indep,

$$\begin{aligned} \text{Var}(X+Y) &= \text{Var}(X) + \text{Var}(Y) + 2\overset{=0}{\text{Cov}(X, Y)} \\ &= \text{Var}(X) + \text{Var}(Y) \end{aligned}$$

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# Stat 134

1. Consider a  $\text{Poisson}(\lambda)$  process. Let  $T_r \sim \text{gamma}(r, \lambda)$  be the  $r$ th arrival time.  $\text{Cov}(T_1, T_3)$  equals:

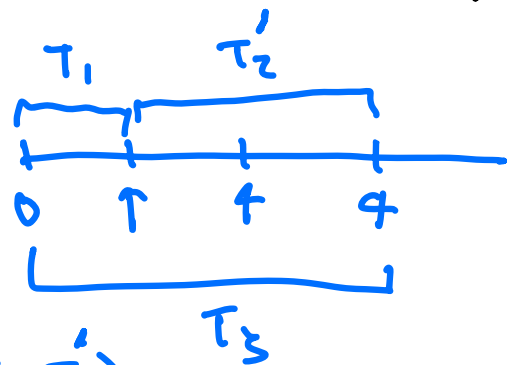
a  $\lambda$

b  $\lambda^2$

**c**  $1/\lambda^2$

d none of the above

Recall  $\text{Var}(T_r) = \frac{r}{\lambda^2}$



$$\text{Cov}(T_1, T_3) = \text{Cov}(T_1, T_1 + T_2')$$

$$= \text{Cov}(T_1, T_1) + \text{Cov}(T_1, T_2')$$

$$\stackrel{||}{=} \text{Var}(T_1)$$

$$\boxed{\frac{1}{\lambda^2}}$$

$\stackrel{||}{=} 0$  since  $T_1, T_2'$  indep

ex

Toss a fair coin 30 times.

Let  $X = \# \text{ heads in the first 20 tosses}$

$Y = \# \text{ heads in the last 20 tosses}$

Find  $\text{Cov}(X, Y)$  # heads in first 10 tosses



$$\begin{aligned} \text{Cov}(X, Y) &= \text{Cov}(A+V, V+B) \\ &= \underbrace{\text{Cov}(A, V)}_0 + \underbrace{\text{Cov}(A, B)}_0 + \underbrace{\text{Cov}(V, V)}_{\text{Var}(V)} + \underbrace{\text{Cov}(V, B)}_0 \end{aligned}$$

$$V \sim \text{Bin}(10, \frac{1}{2})$$

$$\text{Var}(V) = npq = 10 \left(\frac{1}{2}\right)\left(\frac{1}{2}\right) = \boxed{\frac{10}{4}}$$



Extra

Average conditional probability:

$$P(B) = \int P(B|X=x) f_X(x) dx$$

(6 pts) Suppose  $Y \sim \text{Pois}(X)$ , where  $X \sim \text{Exp}(\lambda)$ . That is, given  $X = x$ ,  $Y \sim \text{Pois}(x)$ .

(a) (3 pts) Show that the unconditional distribution of  $Y$  is Geometric  $(\frac{\lambda}{\lambda+1})$  on  $\{0, 1, 2, \dots\}$ .

$Y \sim \text{geom}(p)$  on  $0, 1, 2, \dots$

← # failures until first success

$$P(Y=n) = q^n p$$

$$p = \frac{\lambda}{\lambda+1} \quad q = 1 - \frac{\lambda}{\lambda+1} = \frac{1}{\lambda+1}$$

$$P(Y=n) = \left(\frac{1}{\lambda+1}\right)^n \left(\frac{\lambda}{\lambda+1}\right) \leftarrow \text{want this}$$

$$P(Y=n) = \int_{x=0}^{x=\infty} P(Y=n|X=x) f_X(x) dx$$

$$= \int_0^{\infty} \frac{e^{-x} x^n}{n!} \cdot \lambda e^{-\lambda x} dx$$

$$= \frac{\lambda}{n!} \int_0^{\infty} x^{(n+1)-1} e^{-(\lambda+1)x} dx$$

$$= \frac{\lambda}{n!} \cdot \frac{n!}{(\lambda+1)^{n+1}} = \left(\frac{1}{\lambda+1}\right)^n \left(\frac{\lambda}{\lambda+1}\right) \quad \text{as required}$$

variable part of Gamma  $(n+1, \lambda+1)$

## Appendix

### Bilinearity Properties

Thm

(a)  $\text{Cov}(X+Y, Z) = \text{Cov}(X, Z) + \text{Cov}(Y, Z)$

(b)  $\text{Cov}(aX, bY) = ab \text{Cov}(X, Y)$

Pf/

a)

$$\begin{aligned}\text{Cov}(X+Y, Z) &= E((X+Y - \mu_{X+Y})(Z - \mu_Z)) \\ &= E((X - \mu_X) + (Y - \mu_Y))(Z - \mu_Z)\end{aligned}$$

$$\begin{aligned}&= E((X - \mu_X)(Z - \mu_Z) + (Y - \mu_Y)(Z - \mu_Z)) \\ &= E((X - \mu_X)(Z - \mu_Z)) + E((Y - \mu_Y)(Z - \mu_Z)) \\ &= \text{Cov}(X, Z) + \text{Cov}(Y, Z). \quad \square\end{aligned}$$

$$\begin{aligned}\text{b) } \text{Cov}(aX, bY) &= E((aX - \mu_{aX})(bY - \mu_{bY})) \\ &= E((aX - a\mu_X)(bY - b\mu_Y)) \\ &= E(ab(X - \mu_X)(Y - \mu_Y)) \\ &= abE((X - \mu_X)(Y - \mu_Y)) \\ &= ab \text{Cov}(X, Y) \quad \square\end{aligned}$$

## Appendix

Thm  $\text{Cov}(X, Y) = E(XY) - E(X)E(Y)$

Pf

$$\begin{aligned}\text{Cov}(X, Y) &= E(D_X D_Y) = E((X - \mu_X)(Y - \mu_Y)) \\ &= E(XY - \mu_X Y - X \mu_Y + \mu_X \mu_Y) \\ &= E(XY) - \mu_X \mu_Y - \mu_X \mu_Y + \mu_X \mu_Y \\ &= E(XY) - E(X)E(Y)\end{aligned}$$

□