

Stat 134 lec 41 (final review pt 1)

Warmup

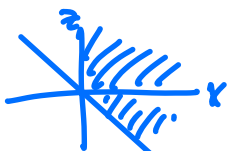
Let the joint (X, Y) be standard bivariate normal distribution with correlation $\rho = 1/2$. Which of the following statements is true? Mark all that apply.

- i. X and Y are independent.
- ii. $P(X < 0) = 1/2$.
- iii. $P(X > 0, Y > 0) = 1/4$.
- iv. X and Y are identically distributed.
- v. $2X + 3Y \sim N(0, 19)$
- vi. None of the above.

we saw last time

$$P(X > 0, Y > 0) = P(X > 0, \rho X + \sqrt{1-\rho^2} Z > 0)$$
$$= P(X > 0, Z > -\frac{\rho}{\sqrt{1-\rho^2}} X)$$

$\frac{1}{4}$



ii & iv & v are correct. Since the correlation is nonzero, X and Y cannot be independent. Since X is a standard normal it is symmetric. the probability is $1/4$ only if $\rho = 0$ and X and Y are both standard normal so have the same distribution. $2X + 3Y$ is normal with expectation 0 and variance $= 4\text{Var}(X) + 9\text{Var}(Y) + 2\text{Cov}(2X, 3Y) = 4 + 9 + 2 \cdot 2 \cdot 3 \cdot \frac{1}{2} = 19$

Announcement :

Tutor Wile will run a post MT2 review session Wed.
There will be another in-class review session Fri.

Defⁿ (Standard Bivariate Normal Distribution)

let X, Z iid $N(0,1)$, $-1 \leq \rho \leq 1$

$$Y = \rho X + \sqrt{1-\rho^2} Z \sim N(0,1)$$

$$\text{Corr}(X, Y) = \rho$$

We call the joint distribution (X, Y) the

Standard bivariate normal with $\text{Corr}(X, Y) = \rho$

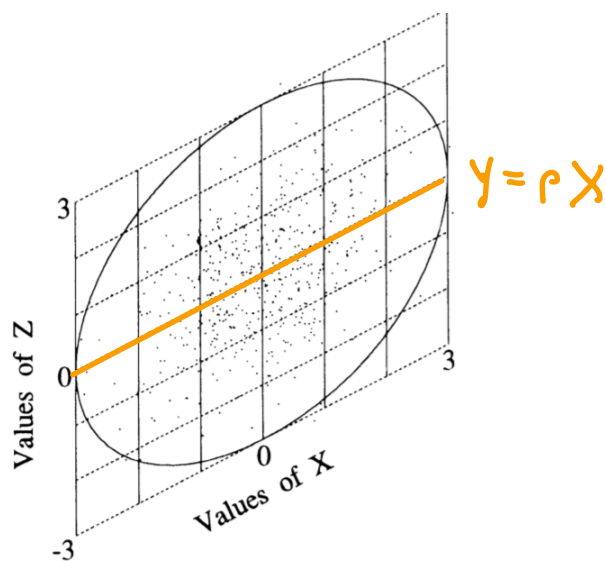
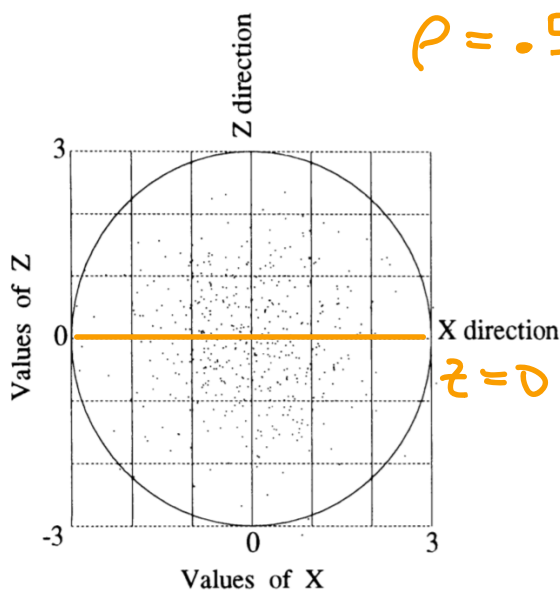
written $(X, Y) \sim \text{BVN}(0, 0, 1, 1, \rho)$



$$(1, 0) \rightarrow (1, \frac{1}{2})$$

$$(X, Z) \longrightarrow (X, Y) = (X, \rho X + \sqrt{1-\rho^2} Z)$$

$$\rho = .5$$



Main properties

$$\textcircled{1} (X, Y) \text{ std BVN iff } aX + bY \sim N(0, a^2 + b^2 + 2ab\rho)$$

② X, Y indep $\Rightarrow \text{Corr}(X, Y) = 0$ always

If (X, Y) BVN,

X, Y indep $\Leftrightarrow \text{Corr}(X, Y) = 0$

(10 pts) (BVN - Adam) Let A and B represent the cost of a television in states A and B , respectively, with $A \sim N(\mu_A, \sigma_A^2)$, $B \sim N(\mu_B, \sigma_B^2)$, $E[AB] = \mu_{AB}$. Furthermore, the joint distribution of A and B is bivariate normal.

(a) (4 pts) Find the correlation between A and B .

(b) (6 pts) Calculate $P(A > 2B)$. You may leave your answer in terms of Φ (the cdf of the standard normal distribution).

$$\begin{aligned} \text{(a) } \text{Corr}(A, B) &= \frac{\text{Cov}(A, B)}{\text{SD}(A)\text{SD}(B)} \\ &= \frac{E[AB] - E[A]E[B]}{\sqrt{\text{Var}(A)\text{Var}(B)}} \\ &= \frac{\mu_{AB} - \mu_A\mu_B}{\sigma_A\sigma_B} \end{aligned}$$

$$\text{(b) } P(A > 2B) = P(A - 2B > 0)$$

For convenience, define $C = A - 2B$.

By properties of bivariate normal, we know that C is normally distributed.

$$E[C] = E[A - 2B] = \mu_A - 2\mu_B$$

$$\text{Var}(C) = \text{Var}(A) + \text{Var}(-2B) + 2\text{Cov}(A, -2B) = \sigma_A^2 + 4\sigma_B^2 - 4(\mu_{AB} - \mu_A\mu_B)$$

$$\text{Therefore } P(A > 2B) = P(C > 0) = 1 - P(C < 0) = 1 - \Phi\left(\frac{0 - E[C]}{\sqrt{\text{Var}(C)}}\right)$$

$$= 1 - \Phi\left(\frac{-(\mu_A - 2\mu_B)}{\sqrt{\sigma_A^2 + 4\sigma_B^2 - 4(\mu_{AB} - \mu_A\mu_B)}}\right)$$

Conditional expectation

12. Let N have the Poisson distribution with mean μ . Let U_1, U_2, \dots be independent uniform $(0, 1)$ variables, independent of N .

Let $M|N = \min(U_1, U_2, \dots, U_N)$. If $N = 0$, define M to be 1.

a) Find $E(M|N)$.

b) Find $E(M)$.

c) Find the survival function of M .

d) Sketch the c.d.f. of M . and calculate $E(M)$ using the CDF.

a)

$$M|N = U_{(1)} \sim \text{Beta}(1, N)$$

$$E(U_{(1)}) = \frac{1}{1+N}$$

$$E(\text{Beta}(r, s)) = \frac{r}{r+s}$$

$$b) E(M) = E(E(M|N)) = E\left(\frac{1}{1+N}\right)$$

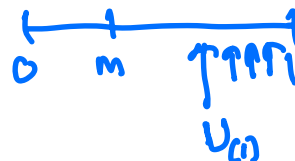
$$= \sum_{n=0}^{\infty} \frac{1}{1+n} \cdot \frac{e^{-\mu} \mu^n}{n!}$$

$$= \frac{1}{\mu} \sum_{n=0}^{\infty} \frac{e^{-\mu} \mu^{n+1}}{(n+1)!}$$

$$= \frac{1}{\mu} \left(\sum_{n=1}^{\infty} \frac{e^{-\mu} \mu^n}{n!} \right) = 1 - \frac{e^{-\mu} \mu^0}{0!}$$

$$= \frac{1}{\mu} (1 - e^{-\mu})$$

$$E(g(n)) = \sum_{n=0}^{\infty} g(n) P(N=n)$$



$$c) P(m > m) = \sum_{n=0}^{\infty} \underbrace{P(M > m | N=n)}_{\substack{\text{rule of} \\ \text{arg cond.} \\ \text{prob.}}} P(N=n)$$

$$\underbrace{P(U_1 > m) P(U_2 > m) \dots P(U_n > m)}_{\substack{\text{rule of} \\ \text{arg cond.} \\ \text{prob.}}}$$

$$= P(m > m | N=0) + \sum_{n=1}^{\infty} (1-m)^n \underbrace{e^{-m} \frac{m^n}{n!}}_{\substack{P(U_i > m)^n \\ (1-m)^n}} e^{-m}$$

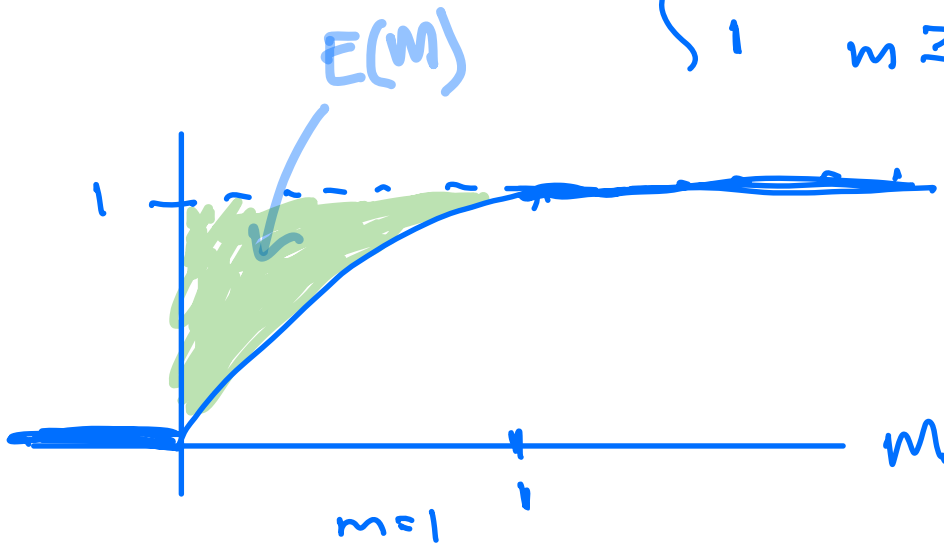
$\begin{array}{c} 1 \\ 0 \quad m \end{array}$
 $e^{-m+m} = e^{-m(1-m)}$

$$= e^{-m} + e^{-m} \sum_{n=1}^{\infty} \frac{e^{-m(1-m)} [m(1-m)]^n}{n!}$$

$$= e^{-m} + e^{-m} (1 - e^{-m(1-m)}) = \boxed{e^{-m}}$$

$$P(m > n) = \begin{cases} 1 & m < 0 \\ e^{-m} & 0 \leq m < 1 \\ 0 & m \geq 1 \end{cases}$$

$$d) F(m) = P(m \leq m) = \begin{cases} 0 & m < 0 \\ 1 - e^{-mm} & 0 \leq m < 1 \\ 1 & m \geq 1 \end{cases}$$



$$E(m) = \int_{m=0}^{\infty} (1 - F(m)) dm$$

$$= \int_0^1 e^{-mm} dm$$

$$= \left[\frac{e^{-mm}}{-m} \right]_{m=0}^{m=1}$$

$$= \frac{e^{-m}}{-m} - \frac{1}{-m} = \boxed{\frac{1 - e^{-m}}{m}}$$

Extra

Suppose that, conditionally on $N = n$ and $P = p$, B has a Binomial(n, p) distribution. Additionally, suppose P has a (continuous) uniform distribution on the interval $[0, \frac{1}{N}]$, and N is a random variable taking values in $\{1, 2, \dots\}$.

Calculate $E[B]$.

$$B | N=n, P=p \sim \text{Bin}(n, p)$$

$$B | N, P=p \sim \text{Bin}(N, p)$$

$$B | N=n, P \sim \text{Bin}(n, P)$$

$$\begin{aligned} E_P E_B(B | N=n, P) &= E_P(nP) \\ &= n E(P) \\ &= \frac{1}{2n} \end{aligned}$$

Don't use $E_N E_B(B | N, P=p) = E(Np)$ since we don't know $E(N)$.