

(Question 1): (40 points) Suppose data  $X_i, i = 1, \dots, n$  satisfy exponential distribution with pdf  $f(x, \lambda) = \lambda \exp(-\lambda x)$ . We want to test the

$$H_0 : \lambda = \lambda_0, \text{ versus } H_1 : \lambda = \lambda_1$$

here  $\lambda_0, \lambda_1$  are known numbers. Find the Uniform most powerful test of this problem. You do not need to specify the exact rejection region in this problem.

$$L(X_1, X_2, \dots, X_n | \mu_0) = \lambda_0^n \cdot \exp(-\lambda_0 \cdot \sum_{i=1}^n X_i)$$

$$L(X_1, X_2, \dots, X_n | \mu_1) = \lambda_1^n \cdot \exp(-\lambda_1 \cdot \sum_{i=1}^n X_i)$$

$$\begin{aligned} \text{Define } \lambda(x) &= \frac{L(X_1, \dots, X_n | \mu_0)}{L(X_1, \dots, X_n | \mu_1)} \\ &= \left(\frac{\lambda_0}{\lambda_1}\right)^n \cdot \exp((\lambda_1 - \lambda_0) \cdot \sum_{i=1}^n X_i) \end{aligned}$$

By Neyman-Pearson Lemma, we apply the likelihood-ratio test.

For a given number  $k (k > 0)$ ,

$$\lambda(x) \leq k \Leftrightarrow n \cdot \log \lambda_0 - n \log \lambda_1 + (\lambda_1 - \lambda_0) \sum_{i=1}^n X_i \leq \log k$$

$$\Leftrightarrow n \cdot \log \lambda_0 - n \cdot \log \lambda_1 + n(\lambda_1 - \lambda_0) \bar{X} \leq \log k$$

Given that  $\lambda_0, \lambda_1$  are fixed, we have:

$$\text{we have: } \begin{cases} \bar{X} \leq c, & \text{if } \lambda_0 < \lambda_1 \\ \bar{X} \geq c, & \text{if } \lambda_0 > \lambda_1 \end{cases}$$

① Suppose  $\lambda_0 < \lambda_1$

$$P_{\lambda_0}(\lambda(x) \leq k) = \alpha$$

$$\Rightarrow P_{\lambda_0}(\bar{X} \leq c) = \alpha \Rightarrow P_{\lambda_0}\left(\frac{\bar{X} - \frac{1}{\lambda_0}}{\frac{1}{\lambda_0}/\sqrt{n}} \leq \frac{c - \frac{1}{\lambda_0}}{\frac{1}{\lambda_0}/\sqrt{n}}\right) = \alpha$$

$$\Rightarrow \frac{c - \frac{1}{\lambda_0}}{\frac{1}{\lambda_0}/\sqrt{n}} = -z_\alpha$$

$$c = \frac{1}{\lambda_0} - \frac{1}{\lambda_0 \sqrt{n}} z_\alpha. \quad R = \{\bar{X} : \bar{X} \leq \frac{1}{\lambda_0} - \frac{1}{\lambda_0 \sqrt{n}} z_\alpha\}$$

The uniform most powerful test is:  $\varphi(x)$  where

$$\varphi(x) = 1_{x \in R} = \begin{cases} 1, & \text{if } x \in R \\ 0, & \text{if } x \notin R. \end{cases}$$

② Suppose  $\lambda_0 > \lambda_1$

$$P_{\lambda_0}(\lambda(x) \leq k) = \alpha$$

$$\Rightarrow P_{\lambda_0}(\bar{X} \geq c) = \alpha \Rightarrow P_{\lambda_0}\left(\frac{\bar{X} - \frac{1}{\lambda_0}}{\frac{1}{\lambda_0}/\sqrt{n}} \geq \frac{c - \frac{1}{\lambda_0}}{\frac{1}{\lambda_0}/\sqrt{n}}\right) = \alpha$$

$$\Rightarrow \frac{c - \frac{1}{\lambda_0}}{\frac{1}{\lambda_0}/\sqrt{n}} = z_\alpha$$

$$c = \frac{1}{\lambda_0} + \frac{1}{\lambda_0 \sqrt{n}} z_\alpha. \quad R = \{ \bar{x} : \bar{x} \geq \frac{1}{\lambda_0} + \frac{1}{\lambda_0 \sqrt{n}} z_\alpha \}$$

The uniform most powerful test is:  $\varphi(x)$  where

$$\varphi(x) = 1_{x \in R} = \begin{cases} 1, & \text{if } x \in R \\ 0, & \text{if } x \notin R. \end{cases}$$

(Question2): (60 points) Suppose data  $X_1, \dots, X_n$  satisfy normal distribution  $N(\mu, \sigma^2)$  with unknown  $\mu$  and known  $\sigma^2$ .

(a) Suppose we want to test the two-point hypothesis

$$H_0 : \mu = \mu_0 \text{ versus } H_1 : \mu = \mu_1$$

here  $\mu_0, \mu_1$  are known numbers. Derive the form of the uniform most powerful test.

(Hint: you need to separate  $\mu_1 > \mu_0$  and  $\mu_0 > \mu_1$  situation.)

(b) Determine the rejection region  $R_\alpha$  so that the type-1 error equals  $\alpha$ .

(c) Calculate the power of test under  $H_1$ .

(d) Determine the  $p$ -value of the test.

$$(a) \quad \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim Z = N(0, 1).$$

$$L(X_1, \dots, X_n | \mu = \mu_0) = \frac{1}{\sqrt{2\pi\sigma^2}^n} \cdot \exp\left\{-\frac{n}{2\sigma^2} \sum_{i=1}^n X_i^2 + 2n\mu_0\bar{X} - n\mu_0^2\right\} / 2\sigma^2$$

$$L(X_1, \dots, X_n | \mu = \mu_1) = \frac{1}{\sqrt{2\pi\sigma^2}^n} \cdot \exp\left\{-\frac{n}{2\sigma^2} \sum_{i=1}^n X_i^2 + 2n\mu_1\bar{X} - n\mu_1^2\right\} / 2\sigma^2$$

$$\text{Define } \lambda(x) = \frac{L(X_1, \dots, X_n | \mu_0)}{L(X_1, \dots, X_n | \mu_1)}$$

$$= \exp\{[2n(\mu_0 - \mu_1)\bar{X} + n(\mu_1^2 - \mu_0^2)] / 2\sigma^2\}$$

By Neyman-Pearson Lemma, we apply the likelihood-ratio test.

For a given number  $k$  ( $k > 0$ ),

$$\lambda(x) \leq k \iff (2n(\mu_0 - \mu_1)\bar{x} + n(\mu_1^2 - \mu_0^2)) / 2\sigma^2 \leq \log k$$

Since  $\mu_0, \mu_1$  are fixed,

$$\begin{cases} \bar{x} \leq c & \text{if } \mu_0 > \mu_1 \\ \bar{x} \geq c & \text{if } \mu_0 < \mu_1 \end{cases}$$

① Suppose  $\mu_0 > \mu_1$ .

$$P_{\mu_0}(\lambda(x) \leq k) = \alpha$$

$$\Rightarrow P_{\mu_0}(\bar{X} \leq c) = P_{\mu_0}\left(\frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \leq \frac{c - \mu_0}{\sigma/\sqrt{n}}\right) = \alpha.$$

$$\Rightarrow \frac{c - \mu_0}{\sigma/\sqrt{n}} = -z_\alpha$$

$$c = \mu_0 - \frac{\sigma}{\sqrt{n}} z_\alpha$$

$$R = \{ \bar{X} : \bar{X} \leq \mu_0 - \frac{\sigma}{\sqrt{n}} z_\alpha \}.$$

The uniform most powerful test is:  $\varphi(x)$  where

$$\varphi(x) = 1_{x \in R} = \begin{cases} 1, & \text{if } x \in R \\ 0, & \text{if } x \notin R. \end{cases}$$

② Suppose  $\mu_0 < \mu_1$

$$P_{\mu_0}(\lambda(x) \geq k) = \alpha$$

$$\Rightarrow P_{\mu_0}(\bar{X} \geq c) = P_{\mu_0}\left(\frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \geq \frac{c - \mu_0}{\sigma/\sqrt{n}}\right) = \alpha$$

$$\frac{c - \mu_0}{\sigma/\sqrt{n}} = z_\alpha$$

$$c = \mu_0 + \frac{\sigma}{\sqrt{n}} z_\alpha$$

$$R = \{ \bar{X} : \bar{X} \geq \mu_0 + \frac{\sigma}{\sqrt{n}} z_\alpha \}$$

The uniform most powerful test is:  $\varphi(x)$  where

$$\varphi(x) = 1_{x \in R} = \begin{cases} 1, & \text{if } x \in R \\ 0, & \text{if } x \notin R. \end{cases}$$

(b) By (a), if  $\mu_0 > \mu_1$ ,  $R_\alpha = \{ \bar{X} : \bar{X} \leq \mu_0 - \frac{\sigma}{\sqrt{n}} z_\alpha \}$ ;  
if  $\mu_0 < \mu_1$ ,  $R_\alpha = \{ \bar{X} : \bar{X} \geq \mu_0 + \frac{\sigma}{\sqrt{n}} z_\alpha \}$

(c) Under  $H_1$ , the power function  $K(\mu_1) = 1 - \beta(\mu_1)$

$$\frac{\bar{X} - \mu_1}{\sigma/\sqrt{n}} = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} + \frac{\mu_0 - \mu_1}{\sigma/\sqrt{n}} \sim N(0,1)$$

$$\textcircled{1} \mu_0 > \mu_1, R_\alpha = \{ \bar{X} : \bar{X} \leq \mu_0 - \frac{\sigma}{\sqrt{n}} z_\alpha \}$$

$$\beta(\mu_1) = P(\bar{X} > \mu_0 - \frac{\sigma}{\sqrt{n}} z_\alpha)$$

$$= P\left(\frac{\bar{X} - \mu_1}{\sigma/\sqrt{n}} > \frac{\sqrt{n}(\mu_0 - \mu_1)}{\sigma} - z_\alpha\right)$$

$$= 1 - \Phi\left(\frac{\sqrt{n}(\mu_0 - \mu_1)}{\sigma} - z_\alpha\right)$$

$$K(\mu_1) = 1 - \beta(\mu_1)$$

$$= \Phi\left(\frac{\sqrt{n}(\mu_0 - \mu_1)}{\sigma} - z_\alpha\right)$$

$$\textcircled{2} \mu_0 < \mu_1, R_\alpha = \{ \bar{X} : \bar{X} \geq \mu_0 + \frac{\sigma}{\sqrt{n}} z_\alpha \}$$

$$\beta(\mu_1) = P(\bar{X} < \mu_0 + \frac{\sigma}{\sqrt{n}} z_\alpha)$$

$$= P\left(\frac{\bar{X} - \mu_1}{\sigma/\sqrt{n}} < \frac{\sqrt{n}(\mu_0 - \mu_1)}{\sigma} + z_\alpha\right)$$

$$= \Phi\left(\frac{\sqrt{n}(\mu_0 - \mu_1)}{\sigma} + z_\alpha\right)$$

$$K(\mu_1) = 1 - \beta(\mu_1) = 1 - \Phi\left(\frac{\sqrt{n}(\mu_0 - \mu_1)}{\sigma} + z_\alpha\right)$$

(d)  $\hat{p} := \inf\{\alpha: \bar{x} \in R_\alpha\}$ . Let  $z_0 = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}}$

①  $\mu_0 > \mu_1$ ,  $R_\alpha = \{\bar{x}: \bar{x} \leq \mu_0 - \frac{\sigma}{\sqrt{n}} z_\alpha\}$

$$\Rightarrow \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}} \leq -z_\alpha \quad \text{i.e., } z_0 \leq -z_\alpha$$

$$\Rightarrow \Phi(z_0) \leq \alpha$$

$$\Rightarrow \hat{p} = \Phi\left(\frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}}\right)$$

②  $\mu_0 < \mu_1$ ,  $R_\alpha = \{\bar{x}: \bar{x} \geq \mu_0 + \frac{\sigma}{\sqrt{n}} z_\alpha\}$

$$\Rightarrow \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}} \geq z_\alpha$$

$$\Phi(z_0) \geq 1 - \alpha$$

$$\Rightarrow \hat{p} = 1 - \Phi\left(\frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}}\right)$$