

Problem 1

$$(a) \ell(\theta) = x_1 \log(2+\theta) + (x_2+x_3) \log(1-\theta) + x_4 \log \theta + c$$

$$\frac{\partial \ell(\theta)}{\partial \theta} = \frac{x_1}{2+\theta} + \frac{x_2+x_3}{1-\theta} (-1) + \frac{x_4}{\theta}$$

$$= \frac{125}{2+\theta} + \frac{41}{\theta-1} + \frac{33}{\theta} = 0$$

$$\Rightarrow 125\theta(\theta-1) + 41\theta(2+\theta) + 33(\theta-1)(2+\theta) = 0$$

$$\Rightarrow 199\theta^2 - 10\theta - 66 = 0 \quad (\hat{\theta} > 0)$$

$$\hat{\theta} = \frac{10 + \sqrt{100 + 4 \times 66 \times 199}}{2 \times 199} \approx 0.60157$$

(b) E-step:

$$Q(\theta, \theta^{(t)}) = E[\ell(\theta) | x, \theta^{(t)}]$$

$$= E[(x_2+x_4) \log(\theta) + (x_2+x_3) \log(1-\theta) | x, \theta^{(t)}]$$

$$= E(x_1 | x, \theta^{(t)}) \log(\theta) + x_4 \log(\theta) + (x_2+x_3) \log(1-\theta)$$

$$= x_1 \cdot \left(\frac{\frac{1}{2}\theta^{(t)}}{\frac{1}{2} + \frac{1}{4}\theta^{(t)}} \right) \log(\theta) + x_4 \log(\theta) + (x_2+x_3) \log(1-\theta)$$

M-step:

$$\theta^{(t+1)} = \arg \max_{\theta} Q(\theta, \theta^{(t)})$$

$$\frac{\partial Q}{\partial \theta} = \frac{x_1 \left(\frac{\frac{1}{2}\theta^{(t)}}{\frac{1}{2} + \frac{1}{4}\theta^{(t)}} \right)}{\theta} + \frac{x_4}{\theta} + \frac{x_2+x_3}{\theta-1} = 0.$$

$$\Rightarrow (\theta-1) \left(x_1 \left(\frac{\frac{1}{2}\theta^{(t)}}{\frac{1}{2} + \frac{1}{4}\theta^{(t)}} \right) + x_4 \right) + \theta(x_2+x_3) = 0.$$

$$\Rightarrow \theta^{(t+1)} = \frac{x_1 \left(\frac{\frac{1}{2}\theta^{(t)}}{\frac{1}{2} + \frac{1}{4}\theta^{(t)}} \right) + x_4}{x_1 \left(\frac{\frac{1}{2}\theta^{(t)}}{\frac{1}{2} + \frac{1}{4}\theta^{(t)}} \right) + x_2 + x_3 + x_4}$$

$$(c) \text{ in (a), } \hat{\theta} \approx 0.60157$$

$$\text{in (b), for } \hat{\theta}_0 = -50, -1, 0, 1, 100,$$

$$\text{all } \hat{\theta} \approx 0.60157$$

So the answer in (a) and (b) are the same.

Problem 2:

$$\log L_c = -\log(|\Sigma|) - \sum E[(x_i - \mu)^T \Sigma^{-1} (x_i - \mu) | y_{obs}, \Sigma]$$

$$\begin{pmatrix} x_1 \\ x_v \end{pmatrix} = \begin{pmatrix} \text{---}, x_{1,p+1}, \dots, x_{1,p+q}, x_{1,p+q+1}, \dots, x_{1,p+q+r} \\ x_2, \dots, x_p, \text{---}, x_{2,p+q+1}, \dots, x_{2,p+q+r} \end{pmatrix}$$

Let $n = p+q+r$

$$E[x_{i1} | x_{i2}] = \mu_1^{(k)} + \sigma_{12}^{(k)} \sigma_2^{(k)-2} (x_{i2} - \mu_2^{(k)})$$

$$E[x_{i2} | x_{i1}] = \mu_2^{(k)} + \sigma_{21}^{(k)} \sigma_1^{(k)-2} (x_{i1} - \mu_1^{(k)})$$

$$E[x_{i1}^2 | x_{i2}] = (\mu_1^{(k)} + \sigma_{12}^{(k)} \sigma_2^{(k)-2} (x_{i2} - \mu_2^{(k)}))^2 + \sigma_1^{(k)2} - \sigma_{12}^{(k)2} \sigma_2^{(k)-2}$$

$$E[x_{i2}^2 | x_{i1}] = (\mu_2^{(k)} + \sigma_{21}^{(k)} \sigma_1^{(k)-2} (x_{i1} - \mu_1^{(k)}))^2 + \sigma_2^{(k)2} - \sigma_{21}^{(k)2} \sigma_1^{(k)-2}$$

for i from 1 to p , $E[(x_i - \mu)^T \Sigma^{-1} (x_i - \mu) | x_{obs}, \Sigma]$

$$= \underbrace{\left[\frac{\sigma_{12}^{(k)} (x_{i2} - \mu_2^{(k)})}{\sigma_2^{(k)2}} + \mu_1^{(k)} - \mu_1 \right]}_A \Sigma^{-1} A^T$$

for i from $p+1$ to $p+q$, $E[(x_i - \mu)^T \Sigma^{-1} (x_i - \mu) | x_{obs}, \Sigma]$

$$= \underbrace{\left[x_{i1} - \mu_1, \frac{\sigma_{21}^{(k)} (x_{i1} - \mu_1^{(k)})}{\sigma_1^{(k)2}} + \mu_2^{(k)} - \mu_2 \right]}_B \Sigma^{-1} B^T$$

for i from $p+q+1$ to $p+q+r$, $E[(x_i - \mu)^T \Sigma^{-1} (x_i - \mu) | x_{obs}, \Sigma]$

$$= (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)$$

$$\text{so } Q = -\frac{n}{2} \log(|\Sigma|) - \frac{1}{2} \sum_{i=1}^p A \Sigma^{-1} A^T - \frac{1}{2} \sum_{i=p+1}^{p+q} B \Sigma^{-1} B^T - \frac{1}{2} \sum_{i=p+q+1}^{p+q+r} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)$$

M-step:

$$\nabla \frac{\partial Q}{\partial \mu_1} = \left(\sum_{i=1}^p \left[\frac{\sigma_{12}^{(k)} (x_{i2} - \mu_2^{(k)})}{\sigma_2^{(k)2}} + \mu_1^{(k)} \right] + \sum_{i=p+1}^{p+q} [x_{i1}] + \sum_{i=p+q+1}^{p+q+r} x_{i1} - \sum_{i=1}^{p+q+r} \mu_1 \right) = 0$$

$$\therefore \mu_1^{(k+1)} = \frac{1}{n} \left[\sum_{i=1}^p \left[\frac{\sigma_{12}^{(k)} (x_{i2} - \mu_2^{(k)})}{\sigma_2^{(k)2}} + \mu_1^{(k)} \right] + \sum_{i=p+1}^{p+q} x_{i1} \right]$$

$$\nabla \frac{\partial Q}{\partial \mu_2} = \left(\sum_{i=1}^p x_{i2} + \sum_{i=p+1}^{p+q} \left[\frac{\sigma_{21}^{(k)} (x_{i1} - \mu_1^{(k)})}{\sigma_1^{(k)2}} + \mu_2^{(k)} \right] + \sum_{i=p+q+1}^{p+q+r} x_{i2} - \sum_{i=1}^{p+q+r} \mu_2 \right) = 0$$

$$\therefore \mu_2^{(k+1)} = \frac{1}{n} \left[\sum_{i=1}^p x_{i2} + \sum_{i=p+1}^{p+q} \left[\frac{\sigma_{21}^{(k)} (x_{i1} - \mu_1^{(k)})}{\sigma_1^{(k)2}} + \mu_2^{(k)} \right] + \sum_{i=p+q+1}^{p+q+r} x_{i2} \right]$$

$$\nabla \frac{\partial Q}{\partial \Sigma^{-1}} = \frac{n}{2} \Sigma - \frac{1}{2} \left[\sum_{i=1}^p A A^T + \sum_{i=p+1}^{p+q} B B^T + \sum_{i=p+q+1}^{p+q+r} (x_i - \mu)(x_i - \mu)^T \right] = 0$$

$$\therefore \Sigma^{(k+1)} = \frac{1}{n} \left\{ \sum_{i=1}^p \left[\frac{\sigma_{12}^{(k)} (x_{i2} - \mu_2^{(k)})}{\sigma_2^{(k)2}} + \mu_1^{(k)} - \mu_1^{(k+1)} \right] \left[\frac{\sigma_{12}^{(k)} (x_{i2} - \mu_2^{(k)})}{\sigma_2^{(k)2}} + \mu_1^{(k)} - \mu_1^{(k+1)}, x_{i2} - \mu_2^{(k+1)} \right]^T \right. \\ \left. + \sum_{i=p+1}^{p+q} \left[x_{i1} - \mu_1^{(k+1)}, \frac{\sigma_{21}^{(k)} (x_{i1} - \mu_1^{(k)})}{\sigma_1^{(k)2}} + \mu_2^{(k)} - \mu_2^{(k+1)} \right]^T \left[x_{i1} - \mu_1^{(k+1)}, \frac{\sigma_{21}^{(k)} (x_{i1} - \mu_1^{(k)})}{\sigma_1^{(k)2}} + \mu_2^{(k)} - \mu_2^{(k+1)} \right] \right. \\ \left. + \sum_{i=p+q+1}^{p+q+r} (x_i - \mu^{(k+1)}) (x_i - \mu^{(k+1)})^T \right\}, \quad \mu^{(k+1)} = \begin{pmatrix} \mu_1^{(k+1)} \\ \mu_2^{(k+1)} \end{pmatrix}$$

Thus we can use this update formular to estimate μ and Σ .

Problem 3

Since $f(x) \propto e^{-x}$, $0 < x < 2$

let $f(x) = ce^{-x}$

$$\int_{\mathbb{R}} f(x) dx = 1 \Rightarrow \int_0^2 ce^{-x} dx = 1$$

$$\Rightarrow -ce^{-x} \Big|_0^2 = 1$$

$$-c(e^{-2} - 1) = 1$$

$$c = \frac{1}{1 - e^{-2}} = \frac{e^2}{e^2 - 1}$$

$$\therefore f(x) = \frac{e^{2-x}}{e^2 - 1}, \quad 0 < x < 2$$

$$F(x) = \int_0^x f(x) dx = \frac{1}{e^2 - 1} \int_0^x e^{2-x} dx = \frac{1}{e^2 - 1} (-1) e^{2-x} \Big|_0^x$$

$$= \frac{e^{2-x} - e^2}{1 - e^2}$$

$$= \frac{1 - e^{-x}}{1 - e^{-2}}$$

$$\text{So } F(x) = \begin{cases} 0 & x \leq 0 \\ \frac{1 - e^{-x}}{1 - e^{-2}} & 0 < x < 2 \\ 1 & x \geq 2 \end{cases}$$

$$0 < x < 2$$

Inverse transformation: $(1 - e^{-2})F(x) = 1 - e^{-x}$

$$\Rightarrow 1 - (1 - e^{-2})F(x) = e^{-x}$$

$$\Rightarrow \log\{1 - (1 - e^{-2})F(x)\} = -x$$

$$\Rightarrow x = -\log\{1 - (1 - e^{-2})F(x)\}$$

$$\Leftrightarrow F^{-1}(x) = -\log\{1 - (1 - e^{-2})x\}$$

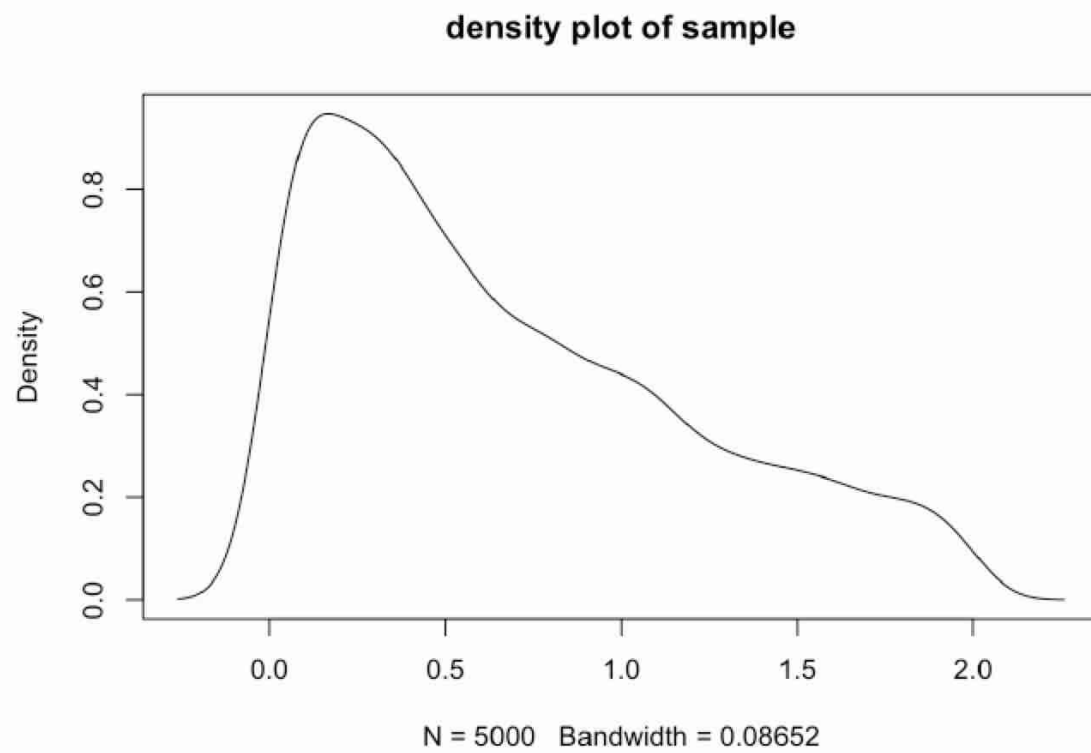
So sample $Y \sim \text{Unif}(0, 1)$ with 5000 observations

$$\Rightarrow X = -\log\{1 - (1 - e^{-2})Y\}$$

the histogram of sample X and estimated density plot are in R code.

Problem 1: The θ is: 0.6015713

Problem 3:



Problem 4

1.

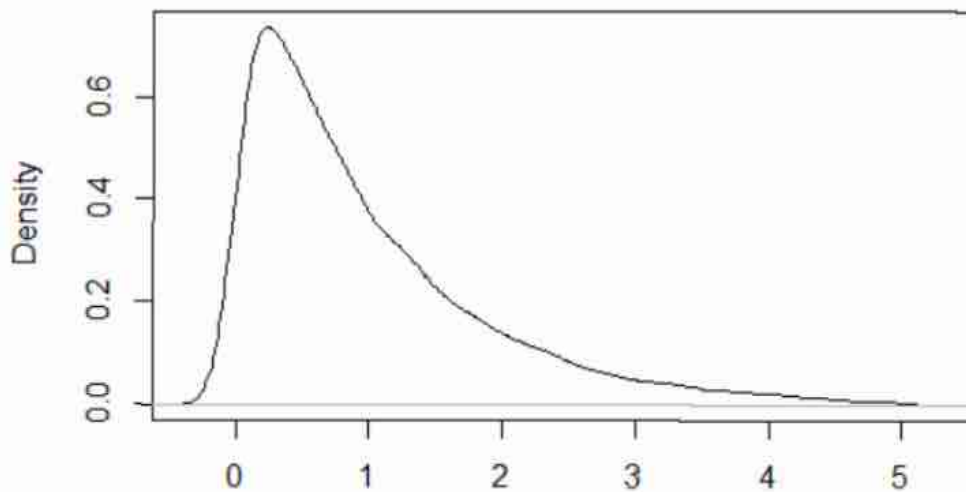
(1) $g_1(x) = e^{-x}$

We can get $X = -\log(1-U)$, $U \sim U(0,1)$;

Then, getting a sampling of 5000 random observations under the condition of range 0 to 5.

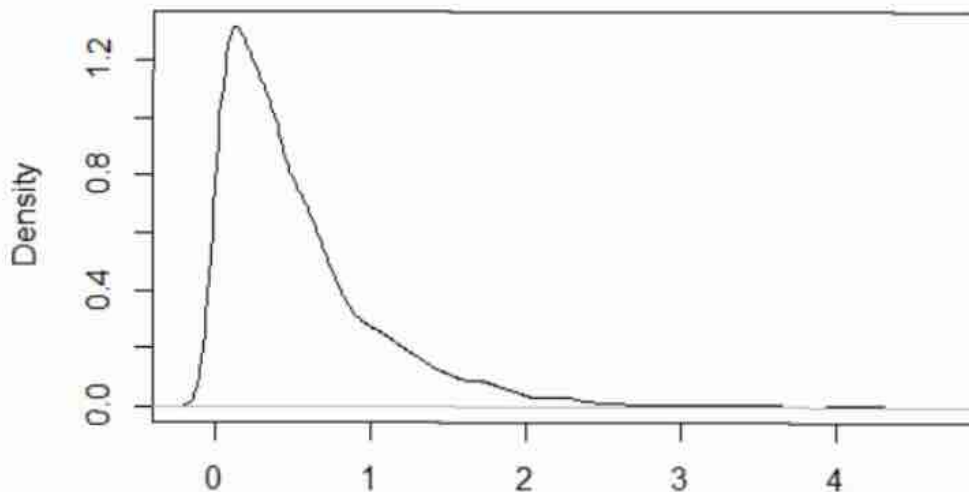
If $U \leq \frac{q(x)}{g_1(x)}$, $\partial = \sup \frac{q(x)}{g_1(x)} = 1$, we return X as the random variable from $f(x)$.

Density of g1



N = 5000 Bandwidth = 0.1317

Density of f_g1



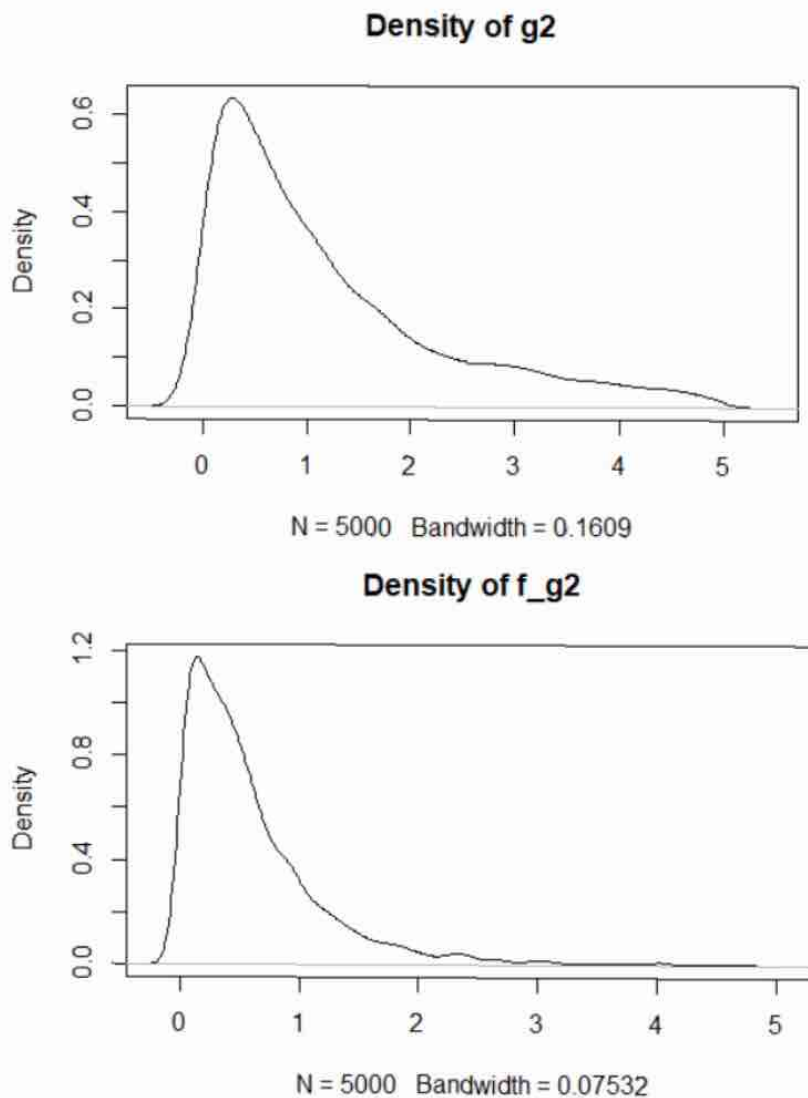
N = 5000 Bandwidth = 0.06872

$$(2) \quad g_2(x) = \frac{2}{\pi(1+x^2)}$$

We can get $X = \tan\left(\left(\frac{\pi}{2}\right) * U\right)$, $U \sim U(0,1)$;

Then, getting a sampling of 5000 random observations under the condition of range 0 to 5.

If, $U \leq \frac{q(x)}{g_2(x)}$, $\partial = \sup \frac{q(x)}{g_2(x)}$ (we just choose $\partial = \frac{\pi}{2}$ to have p_a close to 1), we return X as the random variable from f(x).



2.

The speed of sampling: to sample 5000 observations, g1 is faster than g2. And g2 needs more iterations than g1.

The results: the ratio of accepted value of g1 is:
0.6214496

the ratio of accepted value of g2 is:
0.3956271

Problem 5

(a) $g(x) \propto (2x^{\theta-1} + x^{\theta-\frac{1}{2}})e^{-x} \quad x > 0$

Let $g(x) = c(2x^{\theta-1} + x^{\theta-\frac{1}{2}})e^{-x}$

$\int_{\mathbb{R}} g(x) dx = 1 \Rightarrow \int_0^{+\infty} c(2x^{\theta-1} + x^{\theta-\frac{1}{2}})e^{-x} dx = 1$

$\Rightarrow 2c \int_0^{+\infty} x^{\theta-1} e^{-x} dx + c \int_0^{+\infty} x^{\theta-\frac{1}{2}} e^{-x} dx = 1$

$\Rightarrow 2c \Gamma(\theta) \underbrace{\int_0^{+\infty} \frac{1}{\Gamma(\theta) \cdot 1^\theta} x^{\theta-1} e^{-\frac{x}{1}} dx}_{=1} + c \Gamma(\theta+\frac{1}{2}) \underbrace{\int_0^{+\infty} \frac{1}{\Gamma(\theta+\frac{1}{2}) \cdot 1^{\theta+\frac{1}{2}}} x^{\theta-\frac{1}{2}} e^{-\frac{x}{1}} dx}_{=1} = 1$

So $2c\Gamma(\theta) + c\Gamma(\theta+\frac{1}{2}) = 1$

$c = \frac{1}{2\Gamma(\theta) + \Gamma(\theta+\frac{1}{2})}$

(b) $\frac{1}{\Gamma(\theta) \cdot 1^\theta} x^{\theta-1} e^{-\frac{x}{1}}$ is pdf of $\text{Gamma}(\theta, 1) \rightarrow A$

$\frac{1}{\Gamma(\theta+\frac{1}{2}) \cdot 1^{\theta+\frac{1}{2}}} x^{\theta-\frac{1}{2}} e^{-\frac{x}{1}}$ is pdf of $\text{Gamma}(\theta+\frac{1}{2}, 1) \rightarrow B$

$g(x) = \frac{1}{2\Gamma(\theta) + \Gamma(\theta+\frac{1}{2})} 2x^{\theta-1} e^{-x} + \frac{1}{2\Gamma(\theta) + \Gamma(\theta+\frac{1}{2})} x^{\theta-\frac{1}{2}} e^{-x}$

$= mA + nB \quad (m, n \text{ are constant})$

$m = \frac{2\Gamma(\theta)}{2\Gamma(\theta) + \Gamma(\theta+\frac{1}{2})}, \quad n = \frac{\Gamma(\theta+\frac{1}{2})}{2\Gamma(\theta) + \Gamma(\theta+\frac{1}{2})}$

So $g(x)$ is a mixture of Gamma distributions

(c) 1. sample $U \sim \text{uniform}(0, 1)$

2. if $U \leq m$, sample $X \sim \text{Gamma}(\theta, 1)$, return X

otherwise, sample $X \sim \text{Gamma}(\theta+\frac{1}{2}, 1)$.

(d) $g(x) = \sqrt{4+x} x^{\theta-1} e^{-x}$

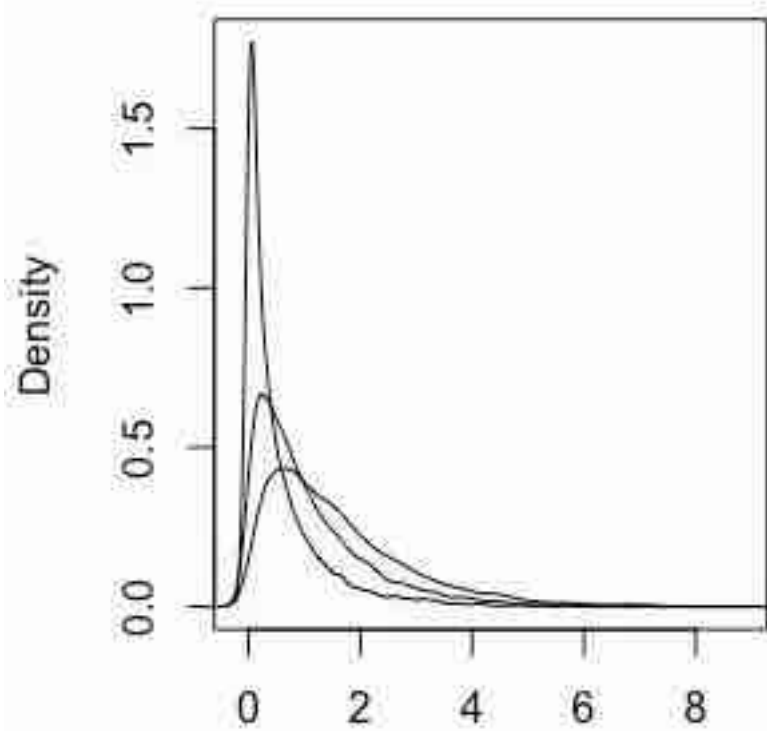
$\frac{g(x)}{g(x)} = \frac{\sqrt{4+x}}{\frac{2}{2\Gamma(\theta) + \Gamma(\theta+\frac{1}{2})} + \frac{1}{2\Gamma(\theta) + \Gamma(\theta+\frac{1}{2})} x^{\frac{1}{2}}} \leq \alpha$

$\lim_{\substack{x \rightarrow \infty \\ x > 0}} \frac{g(x)}{g(x)} = \left(\frac{1}{2\Gamma(\theta) + \Gamma(\theta+\frac{1}{2})} \right)^{-1} \therefore \alpha = 2\Gamma(\theta) + \Gamma(\theta+\frac{1}{2})$

1. sample $X \sim g(x), U \sim \text{uniform}(0, 1)$

2. if $U > \frac{g(x)}{\alpha g(x)}$, then go to step 1, otherwise return X .

sampling $f(x)$



$N = 5000$ Bandwidth = 0.08162

Problem 6

According to the density function, it will not be valid if $a \leq -1$, because it is not meaningful for $\int_0^1 x^a dx$.

If $a > -1$, we have $\int_0^1 \int_0^{\sqrt{1-x^2}} x^a y dy dx = \frac{1}{(a+1)(a+3)}$, a valid density.

And also, according to calculation, we can find that when x is from $\text{beta}(\alpha+1, 1)$ and y is from $\text{beta}(2, 1)$, and also they are independent from each other, we can get the joint distribution $f(x, y) \propto x^a y$. The calculation is as following:

$$\frac{\Gamma(\alpha+2)}{\Gamma(\alpha+1)\Gamma(1)} x^\alpha (1-x)^{1-1} * \frac{\Gamma(3)}{\Gamma(2)\Gamma(1)} y^{2-1} (1-y)^{1-1} = (\alpha+1)x^\alpha y \propto x^a y$$

$$\text{Let } g(x, y) = \frac{\Gamma(\alpha+2)}{\Gamma(\alpha+1)\Gamma(1)} x^\alpha (1-x)^0 \cdot \frac{\Gamma(3)}{\Gamma(2)\Gamma(1)} y^1 (1-y)^0$$

$$\alpha = \sup \frac{q(x, y)}{g(x, y)} = \sup \frac{x^2 y}{f(x, y)/k} = \sup \frac{f(x, y)}{\frac{(\alpha+1)(\alpha+3)}{f(x, y)/k}} = \frac{k}{(\alpha+1)(\alpha+3)}$$

$$\Rightarrow \frac{q(x, y)}{\alpha \cdot g(x, y)} = \frac{x^2 y}{\frac{k}{(\alpha+1)(\alpha+3)} \cdot \frac{f(x, y)}{k}} = 1 \quad \therefore \forall u \leq \frac{q(x, y)}{\alpha \cdot g(x, y)}$$

Thus, we can get the rejection algorithm:

1. sample $X \sim \text{beta}(\alpha+1, 1)$, $Y \sim \text{beta}(2, 1)$, independently.
2. if $X^2 + Y^2 > 1$, then go to step 1, otherwise return (X, Y) .