

ECSE 2500
Lec 22
April 20

Topic: Conditional expectation
— Law of iterated expectation

Sum/Average of several RVs

— Law of large number

□ Conditional expectation

Recall the definition in the last class

Q: What is the expected value of the conditional expectation of Y given X (e.g. a function of RV X)?

$$E[g(X)] = \int_{-\infty}^{+\infty} g(x) f_X(x) dx$$

$$= \int_{-\infty}^{+\infty} E[Y|X=x] f_X(x) dx$$

$$= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} y f_{Y|X}(y|x) dy f_X(x) dx$$

$$= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} y f_{XY}(x,y) dx dy$$

$$= \int_{-\infty}^{\infty} y \cdot f_Y(y) dy$$

$$= E[Y]$$

To summarize, we have

$$E[E[Y|X]] = E[Y]$$

Reduce to the simplified case if $h(Y) = Y$.

The Law of iterated expectation !!

Composition/Nest of two Expectations

Slight generalization:

If we have a function $h(\cdot)$ of a RV Y ,

$$E[E[h(Y)|X]] = E[h(Y)]$$

That is, we can think of $E[h(Y)]$ in two steps:

First, compute $E[h(Y)|x]$ for each possible value x of RV X . Secondly, we average over the distributions of X .

Example

Consider the coin toss example in the
(Discrete RVs) last week. Flip 3 coins, and let

$X = \# \text{ of heads}$

$Y = \text{position of first head}$

In the last week, we have computed $P_{Y|X}^A$ s

		Y			
		0	1	2	3
X	0	1	0	0	0
	1	0	$\frac{1}{3}$	$\frac{1}{3}$	$\frac{1}{3}$
	2	0	$\frac{2}{3}$	$\frac{1}{3}$	0
	3	0	1	0	0

Given the conditional PMF in the table, we calculate

$$E[Y | X=0] = \sum_{y_j} y_j P(y_j | X=0)$$

$$= 1 \cdot 0 + 1 \cdot 0 + 2 \cdot 0 + 3 \cdot 0 = 0$$

$$E[Y | X=1] = 0 \cdot 0 + (1+2+3) \cdot \frac{1}{3} = 2$$

$$E[Y | X=2] = 0 \cdot 0 + \frac{2}{3} + 2 \cdot \frac{1}{3} + 3 \cdot 0 = \frac{4}{3}$$

$$E[Y|X=3] = 1 \cdot 1 = 1$$

Then we use the law of iterated expectation, we have

$$E[Y] = E[E[Y|X]]$$

$$= E[Y|X=0]P(X=0) + E[Y|X=1]P(X=1)$$

$$+ E[Y|X=2]P(X=2) + E[Y|X=3]P(X=3)$$

$$= 0 \cdot \frac{1}{8} + 1 \cdot \frac{3}{8} + 2 \cdot \frac{4}{3} \cdot \frac{3}{8} + 1 \cdot \frac{1}{8} = \frac{11}{8}$$

To do a sanity check,

$$E[Y] = \sum_{y_i} y_i P(Y=y_i)$$

$$= 0 \cdot \frac{1}{8} + 1 \cdot \frac{1}{2} + 2 \cdot \frac{1}{4} + 3 \cdot \frac{1}{8} = \frac{11}{8}$$

→ The law of iterated expectation is verified!

Example

Consider X is uniformly drawn from $[0, 1]$,
(Continuous RVs) and Y is uniformly drawn from $[0, x]$.
Q: What is the expected value of Y ?

The problem structure facilitates us to use
the law of iterated expectation!

① What is $E[Y|X]$?

$$\begin{aligned} E[Y|X] &= \int_{-\infty}^{+\infty} y \underbrace{f_{Y|X}(y|x)}_{\downarrow x} dy \\ &= \int_0^x \frac{y}{x} dy \\ &= \frac{1}{x} \frac{y^2}{2} \Big|_{y=0}^{y=x} = \frac{x}{2} \end{aligned}$$

② What is $E[Y]$?

$$E[Y] = E[E[Y|X]]$$

$$= \int_0^1 \frac{x}{2} \cdot f_X(x) dx$$

$$= \int_0^1 \frac{x}{2} dx = \frac{x^2}{4} \Big|_{x=0}^{x=1} = \frac{1}{4}$$

To do a sanity check, from previous notes,
we have calculated $f_Y(y) = -\log y$, and then

$$E[Y] = \int_0^1 y (-\log y) dy = \frac{y^2}{4} - \frac{y^2}{2} \log y \Big|_{y=0}^{y=1} = \frac{1}{4}$$

□ Sum / Mean of several random variables

In many situations, we observe a sequence of RVs (e.g., a series of coin flips) and want to process these to extract information from those RVs.

A natural way to do so is via the sum/mean of RVs.

$$\text{Sum : } S_n = X_1 + X_2 + \dots + X_n = \sum_{i=1}^n X_i$$

↑
number of RVs

$$\text{Mean : } M_n = \frac{1}{n} S_n = \frac{1}{n} \sum_{i=1}^n X_i$$

First we compute the expected value of these S_n and M_n

$$\begin{aligned} E[S_n] &= E[X_1 + X_2 + \dots + X_n] \\ &= \sum_{i=1}^n E[X_i] \end{aligned}$$

$$\begin{aligned} E[M_n] &= \frac{1}{n} E[X_1 + X_2 + \dots + X_n] \\ &= \frac{1}{n} \sum_{i=1}^n E[X_i] \end{aligned}$$

That is, the expected value of Sum of RVs is sum of Expected RVs.

Second, what about their variance?

$$\begin{aligned}\text{Var}(S_n) &= E[(S_n - E[S_n])^2] \\&= E[S_n^2] - (E[S_n])^2 \\&= E[(X_1 + X_2 + \dots + X_n)(X_1 + X_2 + \dots + X_n)] \\&\quad - (E[X_1 + X_2 + \dots + X_n])^2 \\&= E\left[\sum_{i=1}^n \sum_{j=1}^n X_i X_j\right] - \sum_{i=1}^n \sum_{j=1}^n E[X_i] E[X_j] \\&= \sum_{j=1}^n [E[X_j^2] - (E[X_j])^2] \\&\quad + \sum_{i=1}^n \sum_{j \neq i} [E[X_i X_j] - \underbrace{E[X_i] E[X_j]}_{=(X_i - E[X_i])(X_j - E[X_i])}] \\&= \sum_{j=1}^n \text{Var}(X_j) + \sum_{i=1}^n \sum_{j \neq i} \text{Cov}(X_i, X_j)\end{aligned}$$

That is, the variance of sum of RVs is the sum of all the possible covariances.

Special cases 1) $\text{Cov}(X_j, X_i) = 0$, for all $i \neq j$
all uncorrelated

$$\text{Var}(S_n) = \sum_{j=1}^n \text{Var}(X_j)$$

2) All RVs X_i are independent and identically distributed (iid). That is, each X_i has the same PDF $f_{X_i}(x)$ and their covariance are zero,

$$\text{(iid)} \xrightarrow{\text{implication}} \left\{ \begin{array}{l} E[X_i] = \mu, \text{ for all } i \\ \text{Var}(X_i) = \sigma^2, \text{ for all } i \\ \text{Cor}(X_i, X_j) = 0, \text{ for } i \neq j \end{array} \right.$$

Sum S_n

$$\left[\begin{array}{l} E[S_n] = \sum_{i=1}^n E[X_i] = n\mu \\ \text{Var}(S_n) = n\sigma^2 \end{array} \right]$$

Mean M_n

$$\left[\begin{array}{l} E[M_n] = \frac{1}{n} \sum_{i=1}^n E[X_i] = \mu \\ \text{Var}(M_n) = \frac{1}{n^2} \text{Var}(S_n) = \frac{\sigma^2}{n} \quad (*) \\ \text{Var}(aX) = a^2 \text{Var}(X) \end{array} \right]$$

The equation $(*)$ is important since it says that the variance of M_n gets smaller as the number of RVs increases!!

In the limit, for n iid. RVs,

$$\lim_{n \rightarrow \infty} \text{Var}(M_n) = 0,$$

$$\lim_{n \rightarrow \infty} E[M_n] = \mu.$$

So we get closer and closer to the right mean of X : when $n \rightarrow \infty$.