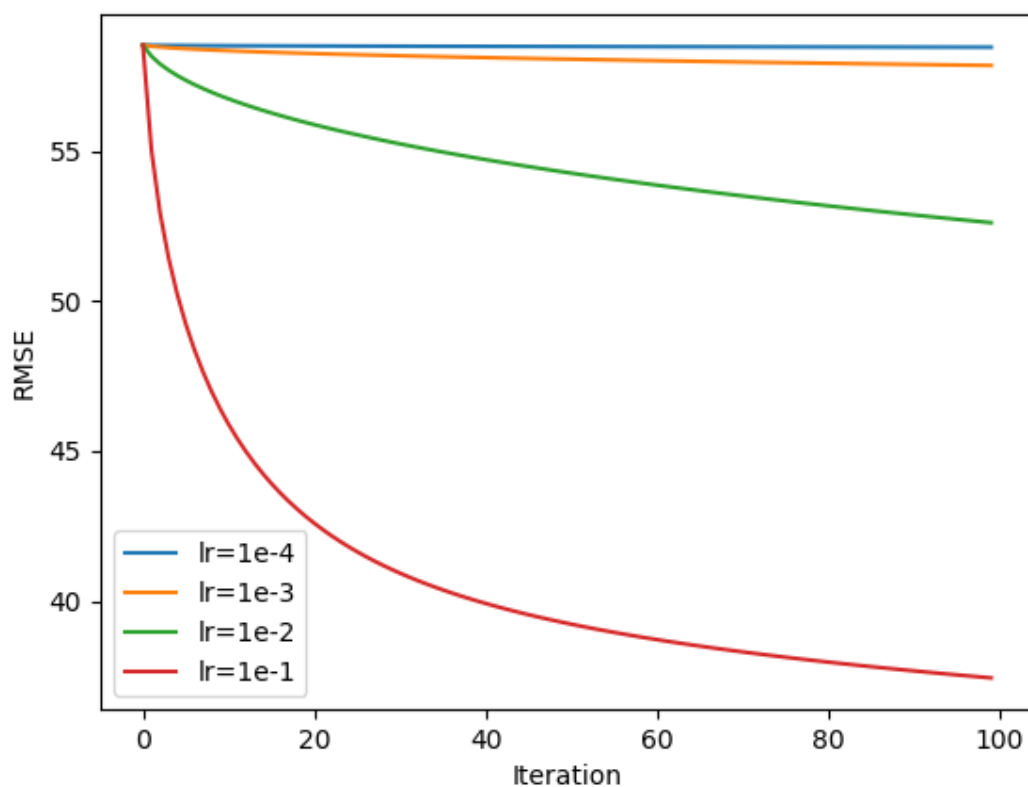
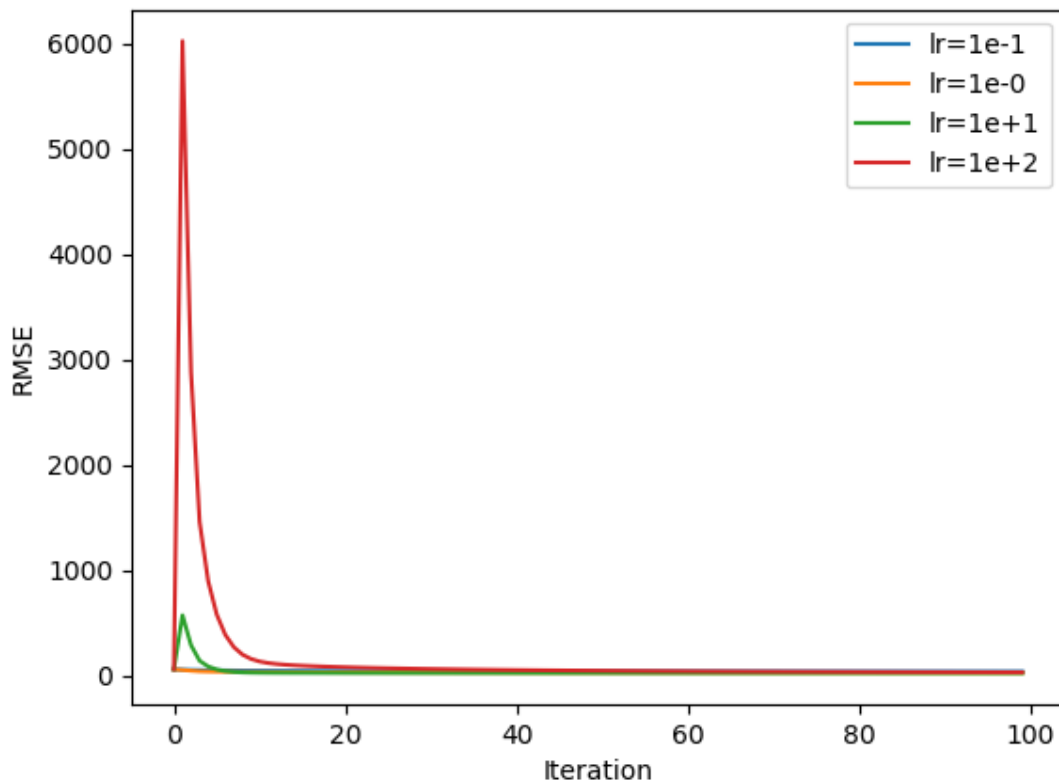


# Homework 1 Report - PM2.5 Prediction

學號：R06922132 系級：資工所碩二 姓名:何羿辰

1. (1%) 請分別使用至少 4 種不同數值的 **learning rate** 進行 **training**（其他參數需一致），對其作圖，並且討論其收斂過程差異。





本題採用七種 learning rate 0.0001/0.001/0.01/0.1/1/10/100，從圖中可以看出 learning rate 為 0.1 時收斂的最快，如果 learning rate 再繼續提高到 1.0 或 10.0 時 RMSE 會呈現一個暴增的山峰然後才慢慢收斂，原因是因為本題採用 gradient descent，雖然過程中會修正每一步的位移量，但太極端的 learning rate 則需要經過多個回合的修正才能開始收斂。

不過本圖也可看出在 gradient descent 的方法裡，learning rate 過高會比過低的收斂速度來的快。

**2. (1%) 請分別使用每筆 data9 小時內所有 feature 的一次項（含 bias 項）以及每筆 data9 小時內 PM2.5 的一次項（含 bias 項）進行 training，比較並討論這兩種模型的 root mean-square error（根據 kaggle 上的 public/private score）。**

| Score       | Private | Public  |
|-------------|---------|---------|
| All Feature | 8.73094 | 8.35278 |
| PM2.5       | 9.71723 | 9.61511 |

根據結果可以發現 feature 全取比只取 PM2.5 不管在 private 還是 public 的結果都還要好，因此可以判斷出除了 PM2.5 以外至少還有一個 feature 會影響未來 PM2.5 的數值。

3. (1%)請分別使用至少四種不同數值的 regularization parameter  $\lambda$  進行 training（其他參數需一至），討論及討論其 RMSE(training, testing)（testing 根據 kaggle 上的 public/private score）以及參數 weight 的 L2 norm。

| $\lambda$ | Train RMSE | L2 norm    | Private Score | Public Score |
|-----------|------------|------------|---------------|--------------|
| 10.0      | 25.008061  | 353.875457 | 9.42694       | 8.79099      |
| 0.1       | 23.115629  | 692.218842 | 8.66211       | 8.29411      |
| 0.01      | 23.113938  | 702.882058 | 8.66601       | 8.30637      |
| 0.00001   | 23.113909  | 704.103165 | 8.66655       | 8.30788      |

根據結果可以發現  $\lambda$  越大則 L2 norm 越小，原因是因為 minimize error 的過程中若  $\lambda$  越大則 weight 必須越小才能最小化，但 training RMSE 則會上升，原因是因為阻礙了 weight 的更新(變化較平滑)，而比較 Private 與 Public 的結果可以發現  $\lambda$  從大至小的 score 表現先降後升，原因與老師所講的結論一樣，function 不夠平滑與太平滑都不好。

## 4 (1%)

### (4-a)

Given  $t_n$  is the data point of the data set  $\mathcal{D} = \{t_1, \dots, t_N\}$ . Each data point  $t_n$  is associated with a weighting factor  $r_n > 0$ .

The sum-of-squares error function becomes:

$$E_D(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^N r_n (t_n - \mathbf{w}^T \mathbf{x}_n)^2$$

Find the solution  $\mathbf{w}^*$  that minimizes the error function.

$$\mathbf{w}^* = \arg \min_{\mathbf{w}} E_D(\mathbf{w}) \rightarrow \text{compute } \frac{dE_D}{d\mathbf{w}} \big|_{\mathbf{w}=\mathbf{w}^0} \rightarrow \text{find } \frac{\partial E_D}{\partial \mathbf{w}} = 0$$

對其微分展開得  $\frac{\partial E_D}{\partial \mathbf{w}} = \sum_{n=1}^N r_n (t_n - \mathbf{w}^T \mathbf{x}_n)(-\mathbf{x}_n) = \mathbf{0}$ ，故只需找到  $\mathbf{w}^*$  使其滿足左式即可

最小化 error function

$$\frac{\partial E_D}{\partial w_j} = \sum_{n=1}^N \frac{\partial}{\partial w_j} \left( \frac{1}{2} r_n (t_n - \mathbf{w}^T \mathbf{x}_n)^2 \right) = \sum_{n=1}^N r_n x_{n,j} (t_n - \mathbf{w}^T \mathbf{x}_n) = 0$$

$$\Rightarrow \sum_{n=1}^N r_n x_{n,j} t_n = \left( \sum_{n=1}^N r_n x_{n,j} x_n \right) \mathbf{w}$$

令  $\mathbf{A} = \left( \sum_{n=1}^N r_n x_{n,j} x_n \right)$  ,  $\mathbf{b} = \sum_{n=1}^N r_n x_{n,j} t_n$  , 則  $\mathbf{Aw} = \mathbf{b}$  解出  $\mathbf{w} = \mathbf{A}^{-1} \mathbf{b}$  即可 , 若  $\mathbf{A}$  不可逆則採用  $\sum_{n=1}^N r_n (t_n - \mathbf{w}^T \mathbf{x}_n)(-\mathbf{x}_n) = 0$  解聯立方程式一樣可求出  $\mathbf{w}^*$  , 底下 4-b 即是採用解聯立方方法求出

(4-b)

Following the previous problem(2-a), if

$$\mathbf{t} = [t_1 t_2 t_3] = [0 \quad 10 \quad 5], \mathbf{X} = [\mathbf{x}_1 \mathbf{x}_2 \mathbf{x}_3] = \begin{bmatrix} 2 & 5 & 5 \\ 3 & 1 & 6 \end{bmatrix}$$

$$r_1 = 2, r_2 = 1, r_3 = 3$$

Find the solution  $\mathbf{w}^*$  .

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令  $\mathbf{w}^* = [w_1 \ w_2]^T$  , 代入 4-a 所得公式可得兩個聯立方程式 , 展開解聯立方程式得

$$\mathbf{w}^* = \begin{bmatrix} 2.282752536391707 \\ -1.135862373180415 \end{bmatrix}$$

5 (1%)

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Given a linear model:

$$y(x, \mathbf{w}) = w_0 + \sum_{i=1}^D w_i x_i$$

with a sum-of-squares error function:

$$E(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^N (y(x_n, \mathbf{w}) - t_n)^2$$

where  $t_n$  is the data point of the data set  $\mathcal{D} = \{t_1, \dots, t_N\}$

Suppose that Gaussian noise  $\epsilon_i$  with zero mean and variance  $\sigma^2$  is added independently to each of the input variables  $x_i$ .

By making use of  $\mathbb{E}[\epsilon_i \epsilon_j] = \delta_{ij} \sigma^2$  and  $\mathbb{E}[\epsilon_i] = 0$ , show that minimizing  $E$  averaged over the noise distribution is equivalent to minimizing the sum-of-squares error for noise-free input variables with the addition of a weight -decay regularization term, in which the bias parameter  $w_0$  is omitted from the regularizer.

Hint

•

$$\delta_{ij} = \begin{cases} 1(i=j), \\ 0(i \neq j). \end{cases}$$

將 $x_n$ 加上 noise:  $\widetilde{x}_n = x_n + \epsilon_n$  · 則 minimizing E averaged over the noise distribution 為

$$\min\left\{\frac{1}{2}\sum_{n=1}^N\left(w_0 + \sum_{i=1}^D w_i(x_i + \epsilon_i) - t_n\right)^2\right\} = \min\left\{\frac{1}{2}\sum_{n=1}^N\left(w_0 + \sum_{i=1}^D w_i x_i + \sum_{i=1}^D w_i \epsilon_i - t_n\right)^2\right\}$$

令 $A_n = w_0 + \sum_{i=1}^D w_i x_i - t_n$  · 並將常數項 $\frac{1}{2}$ 省略不影響 min · 繼續展開

$$\text{原式} = \min\left\{\sum_{n=1}^N\left(A_n + \sum_{i=1}^D w_i \epsilon_i\right)^2\right\} = \min\left\{\sum_{n=1}^N\left(A_n^2 + 2 * A_n * \sum_{i=1}^D w_i \epsilon_i + \left(\sum_{i=1}^D w_i \epsilon_i\right)^2\right)\right\}$$

其中期望值 $\mathbb{E}[A_n^2] = A_n^2$  ·  $\mathbb{E}[2 * A_n * \sum_{i=1}^D w_i \epsilon_i] = 2A_n \sum_{i=1}^D w_i \mathbb{E}[\epsilon_i] = 2A_n \sum_{i=1}^D w_i * 0 = 0$  ·

$$\mathbb{E}\left[\left(\sum_{i=1}^D w_i \epsilon_i\right)^2\right] = \mathbb{E}\left[(w_1^2 \epsilon_1^2 + w_2^2 \epsilon_2^2 + \dots + w_D^2 \epsilon_D^2)\right] + 2\mathbb{E}[(w_1 \epsilon_1 w_2 \epsilon_2 + w_1 \epsilon_1 w_3 \epsilon_3 + \dots + w_{D-1} \epsilon_{D-1} w_D \epsilon_D)] \\ = (w_1^2 \delta_{11} \sigma^2 + w_2^2 \delta_{22} \sigma^2 + \dots + w_D^2 \delta_{DD} \sigma^2) + 2(w_1 w_2 \delta_{12} \sigma^2 + w_1 w_3 \delta_{13} \sigma^2 + \dots + w_{D-1} w_D \delta_{D-1D} \sigma^2) = \sum_{i=1}^D w_i^2 \sigma^2 + 0 = \sigma^2 \sum_{i=1}^D w_i^2 \text{ 帶回原式}$$

$$\text{原式} = \min\left\{\sum_{n=1}^N\left(A_n^2 + \sigma^2 \sum_{i=1}^D w_i^2\right)\right\} = \min\left\{\sum_{n=1}^N\left((w_0 + \sum_{i=1}^D w_i x_i - t_n)^2 + \sigma^2 \sum_{i=1}^D w_i^2\right)\right\}$$

前面項 $(w_0 + \sum_{i=1}^D w_i x_i - t_n)^2$ 是 minimizing the sum-of-squares error for noise-free input variables · 而後面項  $\sigma^2 \sum_{i=1}^D w_i^2$ 是 addition of a weight-decay regulation term · 故兩者結果相等得證#

## 6 (1%)

$\mathbf{A} \in \mathbb{R}^{n \times n}$ ,  $\alpha$  is one of the elements of  $\mathbf{A}$ , prove that

$$\frac{d}{d\alpha} \ln|\mathbf{A}| = \text{Tr}\left(\mathbf{A}^{-1} \frac{d}{d\alpha} \mathbf{A}\right)$$

where the matrix  $\mathbf{A}$  is a real, symmetric, non-singular matrix.

Hint:

- The determinant and trace of  $\mathbf{A}$  could be expressed in terms of its eigenvalues.

Derivation of Jacobi's formula by Laplace:

$$d|\mathbf{A}| = \sum_j A_{ij} \text{adj}^T(\mathbf{A})_{ij} \text{ 又 } |\mathbf{A}| = F(A_{11}, A_{12}, \dots, A_{21}, A_{22}, \dots, A_{nn})$$

$$\Rightarrow d|\mathbf{A}| = \sum_i \sum_j \frac{\partial F}{\partial A_{ij}} dA_{ij} \Rightarrow \frac{d|\mathbf{A}|}{dA_{ij}} = \frac{\partial \sum_k A_{ik} \text{adj}^T(\mathbf{A})_{ik}}{\partial A_{ij}} = \sum_k \frac{\partial A_{ik} \text{adj}^T(\mathbf{A})_{ik}}{\partial A_{ij}} = \sum_k \frac{\partial A_{ik}}{\partial A_{ij}} \text{adj}^T(\mathbf{A})_{ik} +$$

$$\sum_k \frac{\partial \text{adj}^T(\mathbf{A})_{ik}}{\partial A_{ij}} A_{ik} = \sum_k \frac{\partial A_{ik}}{\partial A_{ij}} \text{adj}^T(\mathbf{A})_{ik} + 0 = \sum_k \frac{\partial A_{ik}}{\partial A_{ij}} \text{adj}^T(\mathbf{A})_{ik} = \text{adj}^T(\mathbf{A})_{ij} \Rightarrow d|\mathbf{A}| =$$

$$\text{Tr}(\text{adj}(\mathbf{A})d\mathbf{A})$$

$$\text{因為 } \mathbf{A} \text{ 可逆} \Rightarrow \frac{d|\mathbf{A}|}{d\alpha} = |\mathbf{A}| \text{Tr}(\mathbf{A}^{-1} \frac{d\mathbf{A}}{d\alpha})$$

$$\begin{aligned}\frac{d}{d\alpha}\ln|A| &= \text{Tr}\left(\text{adj}(\ln A)\frac{d\ln A}{d\alpha}\right) = \text{Tr}\left(\ln|A| * \frac{1}{\ln A} * \frac{d\ln A}{d\alpha}\right) = \ln|A| \text{Tr}\left(\ln A^{-1} \frac{d\ln A}{d\alpha}\right) \\ &= \ln|A| \text{Tr}\left(\ln A^{-1} \frac{d\ln A}{dA} * \frac{dA}{d\alpha}\right) = \ln|A| \text{Tr}\left(\frac{1}{\ln A} * A^{-1} * \frac{dA}{d\alpha}\right)\end{aligned}$$

設 A 的 eigenvalues 為  $\lambda_1, \lambda_2, \lambda_3, \dots, \lambda_n$  . 則  $|A| = \lambda_1 \lambda_2 \lambda_3 \dots \lambda_n$  .  $\text{Tr}(A) = \lambda_1 + \lambda_2 + \lambda_3 + \dots + \lambda_n$  繼續

續展開

$$\frac{d}{d\alpha}\ln|A| = \ln(\lambda_1 \lambda_2 \lambda_3 \dots \lambda_n) * \frac{1}{\ln(\lambda_1) + \ln(\lambda_2) + \ln(\lambda_3) + \dots + \ln(\lambda_n)} * \text{Tr}\left(A^{-1} * \frac{dA}{d\alpha}\right)$$

$$\text{又 } \ln(\lambda_1 \lambda_2 \lambda_3 \dots \lambda_n) = \ln(\lambda_1) + \ln(\lambda_2) + \ln(\lambda_3) + \dots + \ln(\lambda_n)$$

$$\text{故 } \frac{d}{d\alpha}\ln|A| = \frac{1}{1} \text{Tr}\left(A^{-1} * \frac{d}{d\alpha}A\right) = \text{Tr}\left(A^{-1} * \frac{d}{d\alpha}A\right) \text{ 得證\#}$$