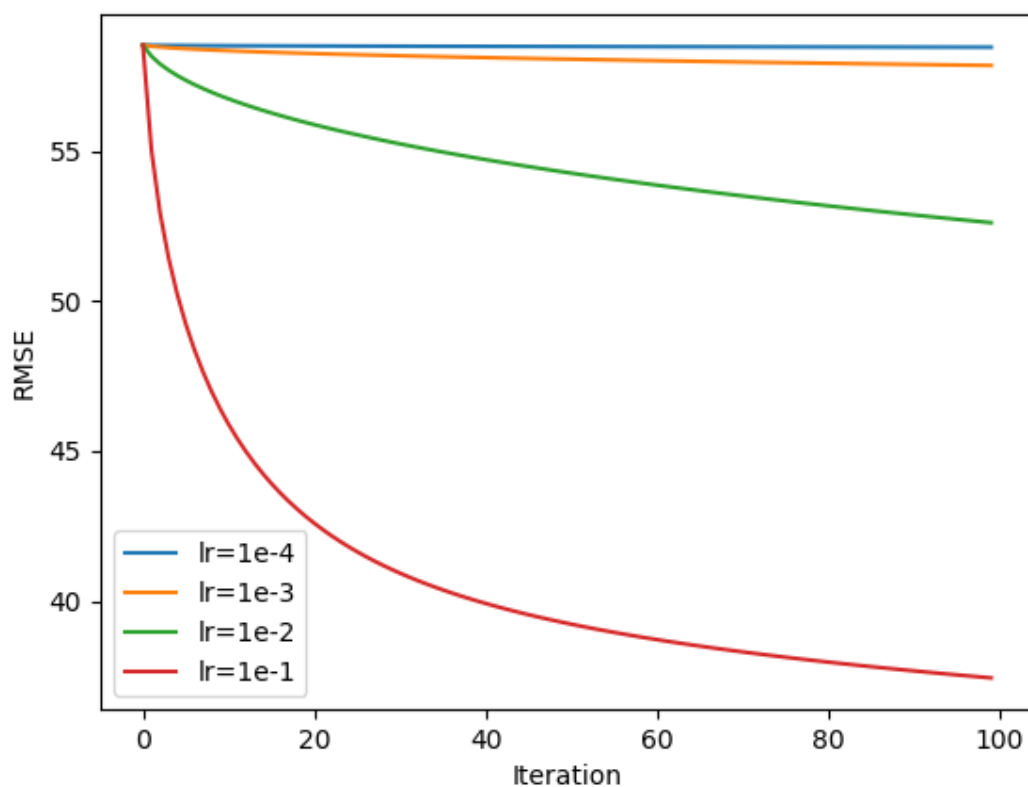
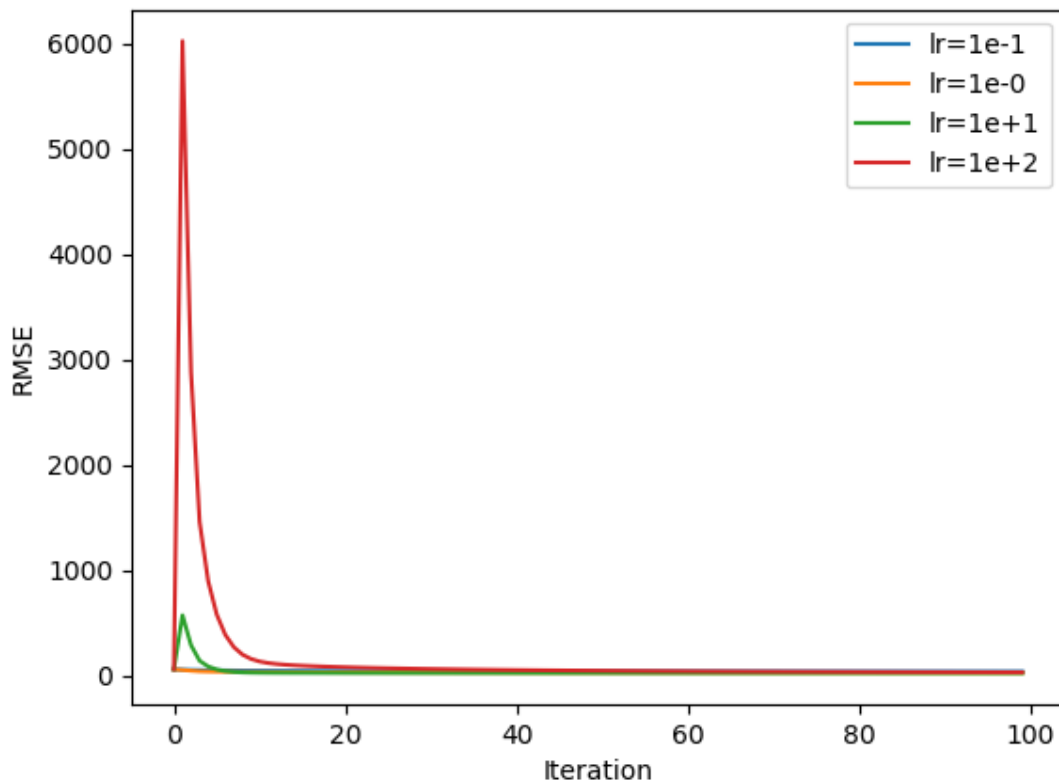


Homework 1 Report - PM2.5 Prediction

學號：R06922132 系級：資工所碩二 姓名:何羿辰

1. (1%) 請分別使用至少 4 種不同數值的 **learning rate** 進行 **training**（其他參數需一致），對其作圖，並且討論其收斂過程差異。





本題採用七種 learning rate 0.0001/0.001/0.01/0.1/1/10/100，從圖中可以看出 learning rate 為 0.1 時收斂的最快，如果 learning rate 再繼續提高到 1.0 或 10.0 時 RMSE 會呈現一個暴增的山峰然後才慢慢收斂，原因是因為本題採用 gradient descent，雖然過程中會修正每一步的位移量，但太極端的 learning rate 則需要經過多個回合的修正才能開始收斂。

不過本圖也可看出在 gradient descent 的方法裡，learning rate 過高會比過低的收斂速度來的快。

2. (1%) 請分別使用每筆 data9 小時內所有 feature 的一次項（含 bias 項）以及每筆 data9 小時內 PM2.5 的一次項（含 bias 項）進行 training，比較並討論這兩種模型的 root mean-square error（根據 kaggle 上的 public/private score）。

Score	Private	Public
All Feature	8.73094	8.35278
PM2.5	9.71723	9.61511

根據結果可以發現 feature 全取比只取 PM2.5 不管在 private 還是 public 的結果都還要好，因此可以判斷出除了 PM2.5 以外至少還有一個 feature 會影響未來 PM2.5 的數值。

3. (1%)請分別使用至少四種不同數值的 regularization parameter λ 進行 training（其他參數需一至），討論及討論其 RMSE(training, testing)（testing 根據 kaggle 上的 public/private score）以及參數 weight 的 L2 norm。

λ	Train RMSE	L2 norm	Private Score	Public Score
10.0	25.008061	353.875457	9.42694	8.79099
0.1	23.115629	692.218842	8.66211	8.29411
0.01	23.113938	702.882058	8.66601	8.30637
0.00001	23.113909	704.103165	8.66655	8.30788

根據結果可以發現 λ 越大則 L2 norm 越小，原因是因為 minimize error 的過程中若 λ 越大則 weight 必須越小才能最小化，但 training RMSE 則會上升，原因是因為阻礙了 weight 的更新(變化較平滑)，而比較 Private 與 Public 的結果可以發現 λ 從大至小的 score 表現先降後升，原因與老師所講的結論一樣，function 不夠平滑與太平滑都不好。

4 (1%)

(4-a)

Given t_n is the data point of the data set $\mathcal{D} = \{t_1, \dots, t_N\}$. Each data point t_n is associated with a weighting factor $r_n > 0$.

The sum-of-squares error function becomes:

$$E_D(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^N r_n (t_n - \mathbf{w}^T \mathbf{x}_n)^2$$

Find the solution \mathbf{w}^* that minimizes the error function.

collaborator:f06b22037 郭尚哲/r07944030 黃聖智

$\mathbf{w}^* = \arg \min_{\mathbf{w}} E_D(\mathbf{w}) \rightarrow \text{compute } \frac{dE_D}{d\mathbf{w}}|_{\mathbf{w}=\mathbf{w}^0} \rightarrow \text{find } \frac{\partial E_D}{\partial \mathbf{w}} = 0$

對其微分展開得 $\frac{\partial E_D}{\partial \mathbf{w}} = \sum_{n=1}^N r_n (t_n - \mathbf{w}^T \mathbf{x}_n)(-\mathbf{x}_n) = \mathbf{0}$ ，故只需找到 \mathbf{w}^* 使其滿足左式即可最小化 error function

$$\frac{\partial E_D}{\partial w_j} = \sum_{n=1}^N \frac{\partial}{\partial w_j} \left(\frac{1}{2} r_n (t_n - \mathbf{w}^T \mathbf{x}_n)^2 \right) = \sum_{n=1}^N r_n x_{n,j} (t_n - \mathbf{w}^T \mathbf{x}_n) = 0$$

$$\Rightarrow \sum_{n=1}^N r_n x_{n,j} t_n = \left(\sum_{n=1}^N r_n x_{n,j} x_n \right) \mathbf{w}$$

令 $\mathbf{A} = \left(\sum_{n=1}^N r_n x_{n,j} x_n \right)$, $\mathbf{b} = \sum_{n=1}^N r_n x_{n,j} t_n$, 則 $\mathbf{Aw} = \mathbf{b}$ 解出 $\mathbf{w} = \mathbf{A}^{-1} \mathbf{b}$ 即可 , 若 \mathbf{A} 不可逆則採用 $\sum_{n=1}^N r_n (t_n - \mathbf{w}^T \mathbf{x}_n)(-\mathbf{x}_n) = 0$ 解聯立方程式一樣可求出 \mathbf{w}^* , 底下 4-b 即是採用解聯立方法求出

(4-b)

Following the previous problem(2-a), if

$$\mathbf{t} = [t_1 t_2 t_3] = [0 \quad 10 \quad 5], \mathbf{X} = [\mathbf{x}_1 \mathbf{x}_2 \mathbf{x}_3] = \begin{bmatrix} 2 & 5 & 5 \\ 3 & 1 & 6 \end{bmatrix}$$

$$r_1 = 2, r_2 = 1, r_3 = 3$$

Find the solution \mathbf{w}^* .

令 $\mathbf{w}^* = [w_1 \ w_2]^T$, 代入 4-a 所得公式展開可得兩個聯立方程式 , 展開解聯立方程式得

$$\mathbf{w}^* = \begin{bmatrix} 2.282752536391707 \\ -1.135862373180415 \end{bmatrix}$$

5 (1%)

Given a linear model:

$$y(x, \mathbf{w}) = w_0 + \sum_{i=1}^D w_i x_i$$

with a sum-of-squares error function:

$$E(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^N (y(x_n, \mathbf{w}) - t_n)^2$$

where t_n is the data point of the data set $\mathcal{D} = \{t_1, \dots, t_N\}$

Suppose that Gaussian noise ϵ_i with zero mean and variance σ^2 is added independently to each of the input variables x_i .

By making use of $\mathbb{E}[\epsilon_i \epsilon_j] = \delta_{ij} \sigma^2$ and $\mathbb{E}[\epsilon_i] = 0$, show that minimizing E averaged over the noise distribution is equivalent to minimizing the sum-of-squares error for noise-free input variables with the addition of a weight-decay regularization term, in which the bias parameter w_0 is omitted from the regularizer.

Hint

$$\bullet \quad \delta_{ij} = \begin{cases} 1(i=j), \\ 0(i \neq j). \end{cases}$$

collaborator:f06b22037 郭尚哲/r07944030 黃聖智

將 x_n 加上 noise: $\tilde{x}_n = x_n + \epsilon_n$ ，則 minimizing E averaged over the noise distribution 為

$$\min \left\{ \frac{1}{2} \sum_{n=1}^N \left(w_0 + \sum_{i=1}^D w_i (x_i + \epsilon_i) - t_n \right)^2 \right\} = \min \left\{ \frac{1}{2} \sum_{n=1}^N \left(w_0 + \sum_{i=1}^D w_i x_i + \sum_{i=1}^D w_i \epsilon_i - t_n \right)^2 \right\}$$

令 $A_n = w_0 + \sum_{i=1}^D w_i x_i - t_n$ ，並將常數項 $\frac{1}{2}$ 省略不影響 min，繼續展開

$$\text{原式} = \min \left\{ \sum_{n=1}^N \left(A_n + \sum_{i=1}^D w_i \epsilon_i \right)^2 \right\} = \min \left\{ \sum_{n=1}^N \left(A_n^2 + 2 * A_n * \sum_{i=1}^D w_i \epsilon_i + \left(\sum_{i=1}^D w_i \epsilon_i \right)^2 \right) \right\}$$

其中期望值 $\mathbb{E}[A_n^2] = A_n^2$ ， $\mathbb{E}[2 * A_n * \sum_{i=1}^D w_i \epsilon_i] = 2 A_n \sum_{i=1}^D w_i \mathbb{E}[\epsilon_i] = 2 A_n \sum_{i=1}^D w_i * 0 = 0$ ，

$$\begin{aligned} \mathbb{E} \left[\left(\sum_{i=1}^D w_i \epsilon_i \right)^2 \right] &= \mathbb{E} \left[(w_1^2 \epsilon_1^2 + w_2^2 \epsilon_2^2 + \dots + w_D^2 \epsilon_D^2) \right] + 2 \mathbb{E} \left[(w_1 \epsilon_1 w_2 \epsilon_2 + w_1 \epsilon_1 w_3 \epsilon_3 + \dots + \right. \\ &\quad \left. + w_{D-1} \epsilon_{D-1} w_D \epsilon_D) \right] = (w_1^2 \delta_{11} \sigma^2 + w_2^2 \delta_{22} \sigma^2 + \dots + w_D^2 \delta_{DD} \sigma^2) + 2(w_1 w_2 \delta_{12} \sigma^2 + w_1 w_3 \delta_{13} \sigma^2 + \\ &\quad \dots + w_{D-1} w_D \delta_{D-1D} \sigma^2) = \sum_{i=1}^D w_i^2 \sigma^2 + 0 = \sigma^2 \sum_{i=1}^D w_i^2 \end{aligned}$$

帶回原式

$$\text{原式} = \min \left\{ \sum_{n=1}^N \left(A_n^2 + \sigma^2 \sum_{i=1}^D w_i^2 \right) \right\} = \min \left\{ \sum_{n=1}^N \left((w_0 + \sum_{i=1}^D w_i x_i - t_n)^2 + \sigma^2 \sum_{i=1}^D w_i^2 \right) \right\}$$

前面項 $(w_0 + \sum_{i=1}^D w_i x_i - t_n)^2$ 是 minimizing the sum-of-squares error for noise-free

input variables，而後面項 $\sigma^2 \sum_{i=1}^D w_i^2$ 是 addition of a weight-decay regulation term，

故兩者結果相等得證#

6 (1%)

$\mathbf{A} \in \mathbb{R}^{n \times n}$, α is one of the elements of \mathbf{A} , prove that

$$\frac{d}{d\alpha} \ln|\mathbf{A}| = \text{Tr}\left(\mathbf{A}^{-1} \frac{d}{d\alpha} \mathbf{A}\right)$$

where the matrix \mathbf{A} is a real, symmetric, non-singular matrix.

Hint:

- The determinant and trace of \mathbf{A} could be expressed in terms of its eigenvalues.

Derivation of Jacobi's formula by Laplace:

$$d|\mathbf{A}| = \sum_j A_{ij} \text{adj}^T(\mathbf{A})_{ij} \text{ 又 } |\mathbf{A}| = F(A_{11}, A_{12}, \dots, A_{21}, A_{22}, \dots, A_{nn})$$

$$\rightarrow d|\mathbf{A}| = \sum_i \sum_j \frac{\partial F}{\partial A_{ij}} dA_{ij} \rightarrow \frac{\partial |\mathbf{A}|}{\partial A_{ij}} = \frac{\partial \sum_k A_{ik} \text{adj}^T(\mathbf{A})_{ik}}{\partial A_{ij}} = \sum_k \frac{\partial A_{ik} \text{adj}^T(\mathbf{A})_{ik}}{\partial A_{ij}} = \sum_k \frac{\partial A_{ik}}{\partial A_{ij}} \text{adj}^T(\mathbf{A})_{ik} +$$

$$\sum_k \frac{\partial \text{adj}^T(\mathbf{A})_{ik}}{\partial A_{ij}} A_{ik} = \sum_k \frac{\partial A_{ik}}{\partial A_{ij}} \text{adj}^T(\mathbf{A})_{ik} + 0 = \sum_k \frac{\partial A_{ik}}{\partial A_{ij}} \text{adj}^T(\mathbf{A})_{ik} = \text{adj}^T(\mathbf{A})_{ij} \rightarrow d|\mathbf{A}| =$$

$$\text{Tr}(\text{adj}(\mathbf{A})d\mathbf{A})$$

$$\text{因為 } \mathbf{A} \text{ 可逆} \rightarrow \frac{d|\mathbf{A}|}{d\alpha} = |\mathbf{A}| \text{Tr}\left(\mathbf{A}^{-1} \frac{d\mathbf{A}}{d\alpha}\right)$$

$$\frac{d}{d\alpha} \ln|\mathbf{A}| = \text{Tr}\left(\text{adj}(\ln \mathbf{A}) \frac{d \ln \mathbf{A}}{d\alpha}\right) = \text{Tr}\left(\ln|\mathbf{A}| * \frac{1}{\ln \mathbf{A}} * \frac{d \ln \mathbf{A}}{d\alpha}\right) = \ln|\mathbf{A}| \text{Tr}\left(\ln \mathbf{A}^{-1} \frac{d \ln \mathbf{A}}{d\alpha}\right)$$

$$= \ln|\mathbf{A}| \text{Tr}\left(\ln \mathbf{A}^{-1} \frac{d \ln \mathbf{A}}{d\alpha} * \frac{d\mathbf{A}}{d\alpha}\right) = \ln|\mathbf{A}| \text{Tr}\left(\frac{1}{\ln \mathbf{A}} * \mathbf{A}^{-1} * \frac{d\mathbf{A}}{d\alpha}\right)$$

設 \mathbf{A} 的 eigenvalues 為 $\lambda_1, \lambda_2, \lambda_3, \dots, \lambda_n$. 則 $|\mathbf{A}| = \lambda_1 \lambda_2 \lambda_3 \dots \lambda_n$. $\text{Tr}(\mathbf{A}) = \lambda_1 + \lambda_2 + \lambda_3 + \dots + \lambda_n$ 繼

續展開

$$\frac{d}{d\alpha} \ln|\mathbf{A}| = \ln(\lambda_1 \lambda_2 \lambda_3 \dots \lambda_n) * \frac{1}{\ln(\lambda_1) + \ln(\lambda_2) + \ln(\lambda_3) + \dots + \ln(\lambda_n)} * \text{Tr}\left(\mathbf{A}^{-1} * \frac{d\mathbf{A}}{d\alpha}\right)$$

$$\text{又 } \ln(\lambda_1 \lambda_2 \lambda_3 \dots \lambda_n) = \ln(\lambda_1) + \ln(\lambda_2) + \ln(\lambda_3) + \dots + \ln(\lambda_n)$$

$$\text{故 } \frac{d}{d\alpha} \ln|\mathbf{A}| = \frac{1}{1} \text{Tr}\left(\mathbf{A}^{-1} * \frac{d\mathbf{A}}{d\alpha}\right) = \text{Tr}\left(\mathbf{A}^{-1} * \frac{d\mathbf{A}}{d\alpha}\right) \text{ 得證\#}$$