

# Lecture 20. Inference on PGMs

COMP90051 Statistical Machine Learning

Semester 1, 2021  
Lecturer: Trevor Cohn



# This lecture

- Probabilistic inference: computing (conditional) marginals from joint distributions
  - \* Needed to learn (posterior update) in Bayesian ML
  - \* Exact inference: Elimination algorithm
  - \* Approximate inference: Sampling
- Statistical inference: Parameter estimation
  - \* Fully observed case: Factors decompose under MLE
  - \* Latent variables: Motivates the EM algorithm

# Probabilistic inference on PGMs

*Computing marginal and conditional distributions from the joint of a PGM using Bayes rule and marginalisation.*

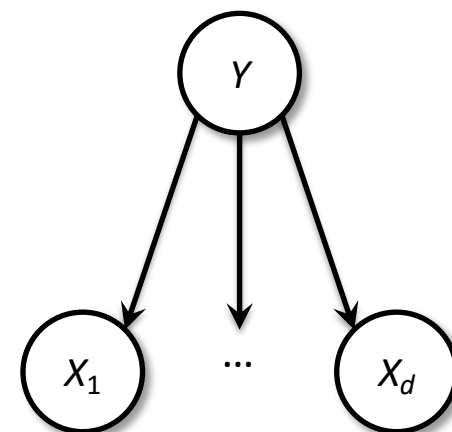
*This deck: how to do it efficiently.*

# Two familiar examples

- Naïve Bayes (frequentist/Bayesian)

- \* Chooses most likely class given data

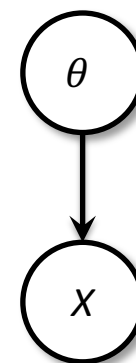
- \* 
$$\Pr(Y|X_1, \dots, X_d) = \frac{\Pr(Y, X_1, \dots, X_d)}{\Pr(X_1, \dots, X_d)} = \frac{\Pr(Y, X_1, \dots, X_d)}{\sum_y \Pr(Y=y, X_1, \dots, X_d)}$$



- Data  $X|\theta \sim N(\theta, 1)$  with prior  $\theta \sim N(0,1)$  (Bayesian)

- \* Given observation  $X = x$  update posterior

- \* 
$$\Pr(\theta|X) = \frac{\Pr(\theta, X)}{\Pr(X)} = \frac{\Pr(\theta, X)}{\sum_{\theta} \Pr(\theta, X)}$$



- Joint + Bayes rule + marginalisation  $\rightarrow$  anything

# Nuclear power plant

- **Alarm sounds**; meltdown?!

$$\Pr(HT|AS = t) = \frac{\Pr(HT, AS=t)}{\Pr(AS=t)}$$

$$= \frac{\sum_{FG, HG, FA} \Pr(AS=t, FA, HG, FG, HT)}{\sum_{FG, HG, FA, HT'} \Pr(AS=t, FA, HG, FG, HT')}$$

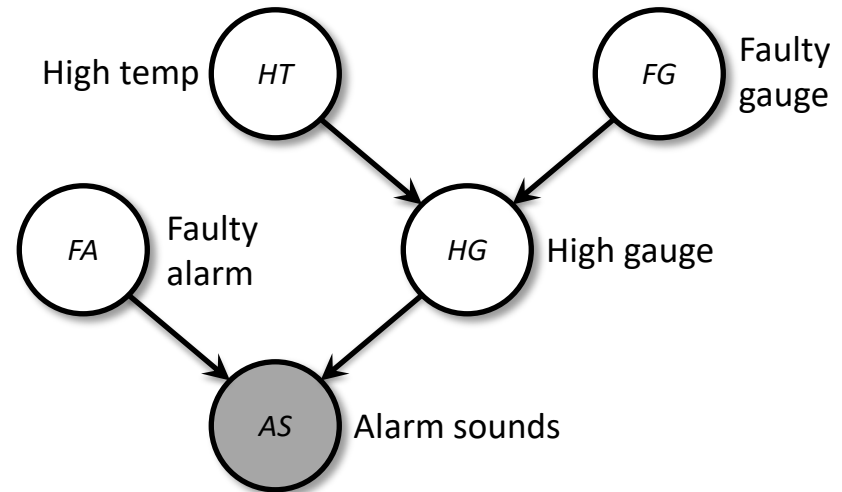
- Numerator (denominator similar)

expanding out sums, joint *summing once over  $2^5$  table*

$$= \sum_{FG} \sum_{HG} \sum_{FA} \Pr(HT) \Pr(HG|HT, FG) \Pr(FG) \Pr(AS = t|FA, HG) \Pr(FA)$$

distributing the sums as far down as possible *summing over several smaller tables*

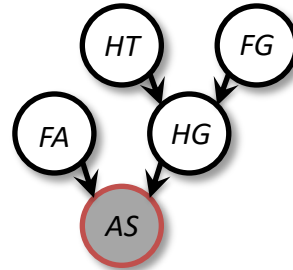
$$= \Pr(HT) \sum_{FG} \Pr(FG) \sum_{HG} \Pr(HG|HT, FG) \sum_{FA} \Pr(FA) \Pr(AS = t|FA, HG)$$



# Nuclear power plant (cont.)

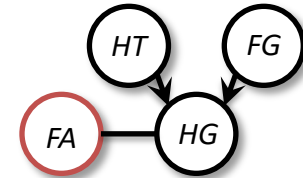
$$= \Pr(HT) \sum_{FG} \Pr(FG) \sum_{HG} \Pr(HG|HT, FG) \sum_{FA} \Pr(FA) \Pr(AS = t|FA, HG)$$

**eliminate AS:** since AS observed, really a no-op



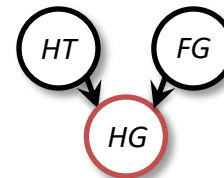
$$= \Pr(HT) \sum_{FG} \Pr(FG) \sum_{HG} \Pr(HG|HT, FG) \sum_{FA} \Pr(FA) m_{AS}(FA, HG)$$

**eliminate FA:** multiplying 1x2 by 2x2



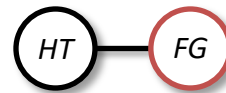
$$= \Pr(HT) \sum_{FG} \Pr(FG) \sum_{HG} \Pr(HG|HT, FG) m_{FA}(HG)$$

**eliminate HG:** multiplying 2x2x2 by 2x1



$$= \Pr(HT) \sum_{FG} \Pr(FG) m_{HG}(HT, FG)$$

**eliminate FG:** multiplying 1x2 by 2x2



$$= \Pr(HT) m_{FG}(HT)$$



Multiplication of tables, followed by summing, is actually matrix multiplication

$$m_{FA}(HG) =$$

FA	
f	t
0.6	0.4

	HG	
	f	t
f	1.0	0
t	0.8	0.2

X

# Elimination algorithm

**Eliminate** (Graph  $G$ , Evidence nodes  $E$ , Query nodes  $Q$ )

1. Choose node ordering  $I$  such that  $Q$  appears last
2. Initialise empty list **active**
3. For each node  $X_i$  in  $G$ 
  - a) Append  $\Pr(X_i | \text{parents}(X_i))$  to **active**
4. For each node  $X_i$  in  $E$ 
  - a) Append  $\delta(X_i, x_i)$  to **active**
5. For each  $i$  in  $I$ 
  - a) potentials = Remove tables referencing  $X_i$  from **active**
  - b)  $N_i$  = nodes other than  $X_i$  referenced by tables
  - c) Table  $\phi_i(X_i, X_{N_i})$  = product of tables
  - d) Table  $m_i(X_{N_i}) = \sum_{X_i} \phi_i(X_i, X_{N_i})$
  - e) Append  $m_i(X_{N_i})$  to **active**
6. Return  $\Pr(X_Q | X_E = x_E) = \phi_Q(X_Q) / \sum_{x_Q} \phi_Q(X_Q)$

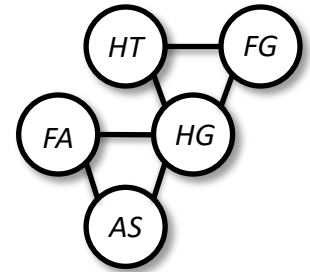
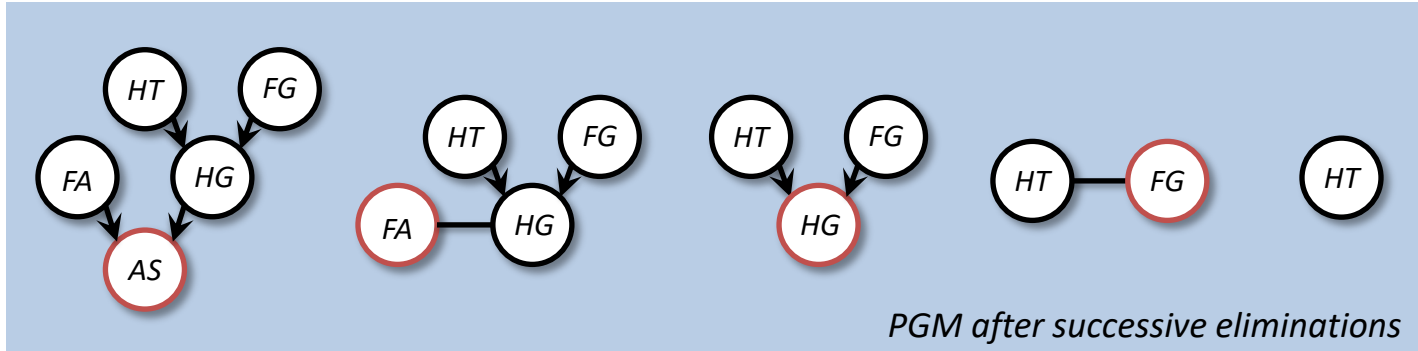
initialise

evidence

marginalise

normalise

# Runtime of elimination algorithm



"reconstructed" graph  
From process called  
**moralisation**

- Each step of elimination
  - \* Removes a node
  - \* Connects node's remaining neighbours  
→ **forms a clique** in the "reconstructed" graph  
(cliques are exactly r.v.'s involved in each sum)
- Time complexity **exponential in largest clique**
- Different elimination orderings produce different cliques
  - \* **Treewidth**: minimum over orderings of the largest clique
  - \* Best possible time complexity is exponential in the treewidth e.g.  $O(2^{\text{tw}})$



# Mini Summary

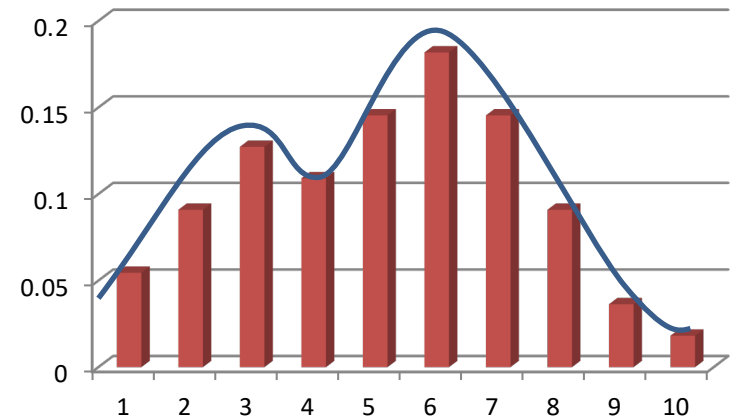
(Exact) probabilistic inference on PGMs

- What? Marginalise out variables, Condition
- Why? Example: Bayesian posterior updates!
- How? The elimination algorithm
- How long? Time exponential in treewidth

Next time: Approximate PGM probabilistic inference

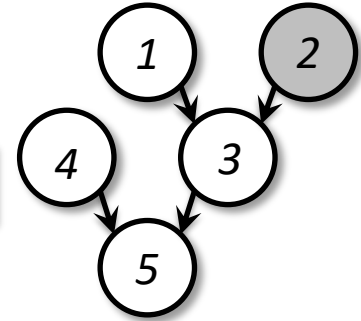
# Probabilistic inference by simulation

- Exact probabilistic inference can be expensive/impossible
  - \* Integration may not have analytical solution!
- Can we approximate numerically?
- Idea: **sampling methods**
  - \* Approximate **distribution** by **histogram of a sample**
  - \* We can't trivially sample: (1) only know desired distribution up to a (normalising) constant (2) naïve sampling approaches are inefficient in high dimensions.



# Gibbs sampling

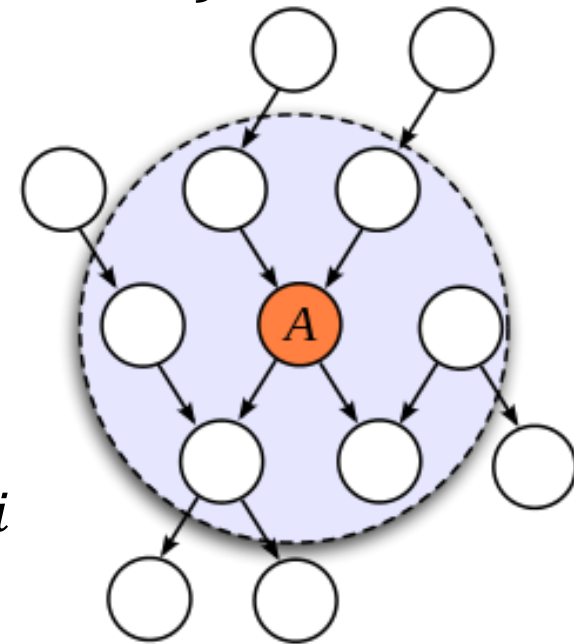
Divide and conquer: Sampling single variable at a time.



- Given: D-PGM on  $d$  random variables  
 Given: evidence values  $\mathbf{x}_E$  over variables  $E \subset \{1, \dots, d\}$   
 Goal: many approximately independent samples from joint conditioned on  $\mathbf{x}_E$
- 1. Initialise with a starting  $\mathbf{X}^{(0)} = (X_1^{(0)}, \dots, X_d^{(0)})$  with  $\mathbf{X}_E^{(0)} = \mathbf{x}_E$
- 2. Repeat many times
  - a) Pick non-evidence node  $X_j$  uniformly at random
  - b) Sample single node  $X'_j \sim p(X_j | X_1^{(i-1)}, \dots, X_{j-1}^{(i-1)}, X_{j+1}^{(i-1)}, \dots, X_d^{(i-1)})$
  - c) Save entire joint sample  $\mathbf{X}^{(i)} = (X_1^{(i-1)}, \dots, X_{j-1}^{(i-1)}, X'_j, X_{j+1}^{(i-1)}, \dots, X_d^{(i-1)})$
- Need not update nodes in random order, e.g. **parents first order**  
 But do need to be able to **sample from conditionals** (e.g. conjugacy)

# Markov blanket

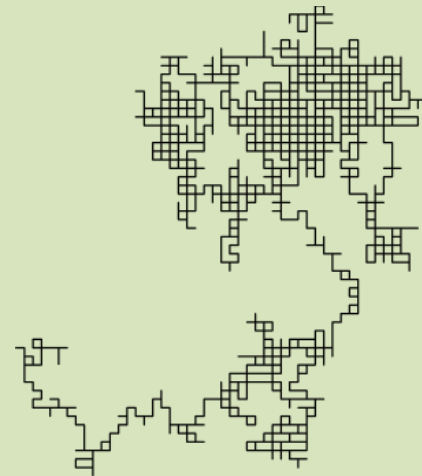
- Intuition: all the nodes that you directly depend on.  
*Not just your parents/children!*
- Consider node  $X_i$  in D-PGM on nodes  $N = \{1, \dots, d\}$
- Markov blanket  $\text{MB}(i)$  of  $X_i$ :
  - \* Nodes  $B \subseteq N \setminus \{i\}$  such that...
  - \*  $X_i$  independent of  $\mathbf{X}_{\bar{B} \setminus \{i\}}$  given  $\mathbf{X}_B$
  - \*  $p(X_i \mid X_1, \dots, X_{i-1}, X_{i+1}, \dots, X_d) = p(X_i \mid \text{MB}(X_i))$
- In D-PGM Markov blanket is:
  - \* Parents of  $i$ , children of  $i$ , parents of children of  $i$
  - \*  $p(X_i \mid \text{MB}(X_i)) \propto p(X_i \mid X_{\pi_i}) \prod_{k: i \in \pi_k} p(X_k \mid X_{\pi_k})$



public domain

# Markov Chain Monte Carlo (MCMC)

- Gibbs sampling produces a chain of samples  $\mathbf{X}^{(1)}, \mathbf{X}^{(2)}, \dots$  approximating draws from  $p(\mathbf{X}_{\bar{E}} | \mathbf{X}_E = \mathbf{x}_E)$
- How good an approximation? Independent draws possible?
- Samples form a Markov chain: Each  $\mathbf{X}^{(i)}$  depends only  $\mathbf{X}^{(i-1)}$ 
  - \* States are all possible values taken by joint samples
  - \* Initial distribution  $\mathbf{p}_0$  of state  $\mathbf{X}^{(0)}$  given by initialisation process
  - \* Transition probability matrix  $\mathbf{T}$  given by PGM conditional probabilities
  - \* Combines to: distribution  $\mathbf{p}_i = (\mathbf{T})^i \mathbf{p}_0$  of state  $\mathbf{X}^{(i)}$ .
- **Burn in:** Run Gibbs long enough and  $\mathbf{X}^{(i)} \sim p(\mathbf{X}_{\bar{E}} | \mathbf{X}_E = \mathbf{x}_E)$ 
  - \* “Limiting distribution”  $\lim_{i \rightarrow \infty} \mathbf{p}_i$  is  $p(\mathbf{X}_{\bar{E}} | \mathbf{X}_E = \mathbf{x}_E)$  under condition that no entry of  $\mathbf{T}$  is zero (“ergodicity” – may not always hold)
  - \* Solution: throw away first few thousand samples
- **Thinning:** Want saved full samples to be independent
  - \* Neighbouring  $\mathbf{X}^{(i)}, \mathbf{X}^{(i+1)}$  are highly correlated. **Intuition why?**
  - \* Solution: only keep every 100 or so samples



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# Initialising Gibbs: Forward Sampling

- Set all evidence nodes to observed values
- Remaining nodes, parent-first order
  - \* Node has no parents? Sample from its D-PGM marginal
  - \* Sample node given previously sampled parents
- However Markov chain theory tells us MCMC converges irrespective of initial sample's distribution
  - \* The limiting distribution – the “equilibrium distribution” – is a property of the transition matrix (the PGM's joint) not the initial distribution

# Now what??



- With our  $\mathbf{X}^{(1)}, \dots, \mathbf{X}^{(T)}$  in hand after running Gibbs for a while with burn-in and thinning...
- These form “i.i.d.” sample of  $p(\mathbf{X}_{\bar{E}} | \mathbf{X}_E = \mathbf{x}_E)$
- We can do heaps!
  - a) Can approximate the distribution via a histogram of these samples (make bins, form counts).
  - b) Marginalising out variables == Dropping components from samples
  - c) Expectations: Estimating by sample mean of samples

# Mini Summary

Approximate probabilistic inference on PGMs

- Why? Summation/integration may be costly
- Why? Integration may be impossible analytically
- Briefly: Gibbs sampling

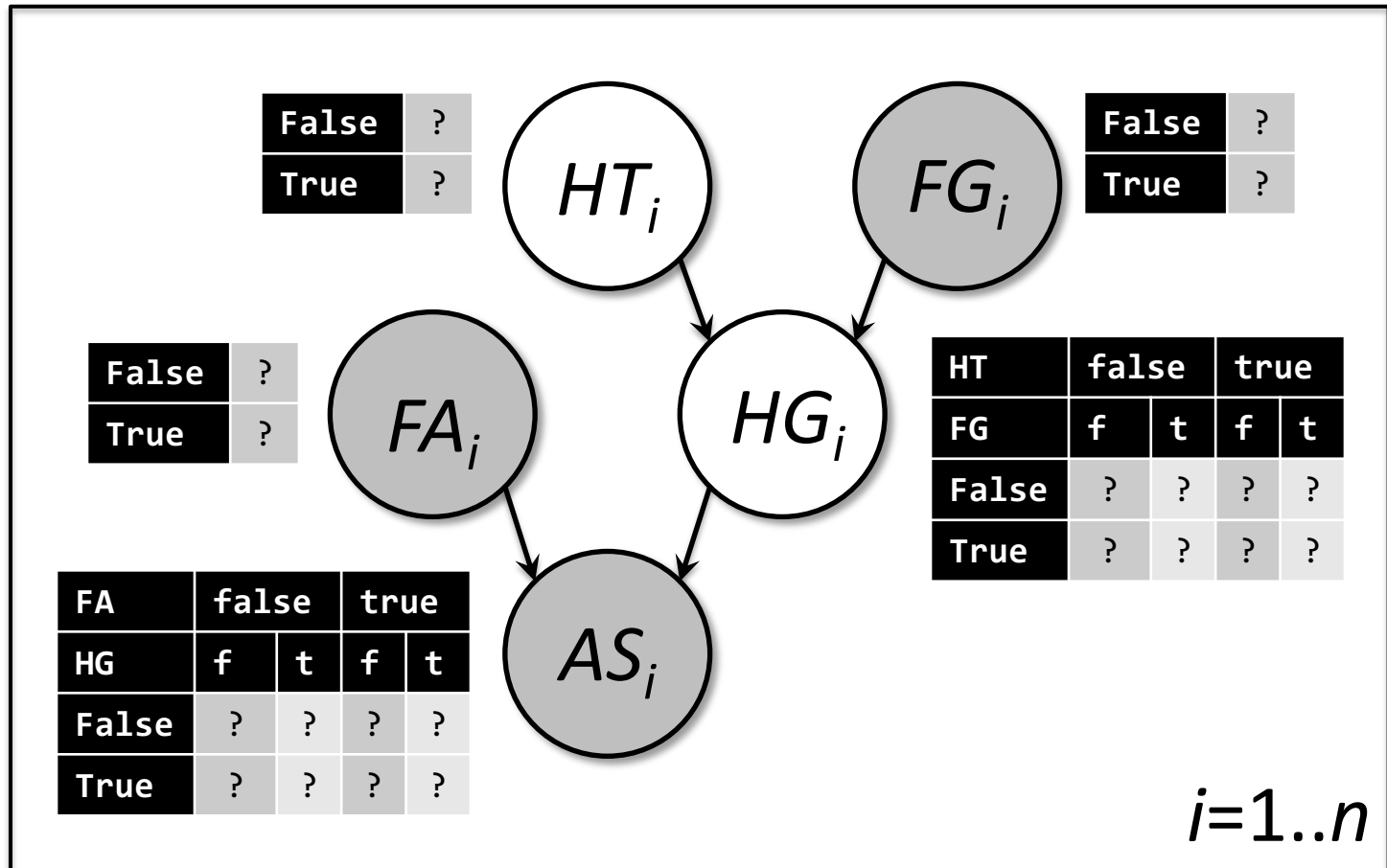
Next time: Statistical inference on PGMs



# Statistical inference on PGMs

*Learning from data – fitting probability tables to observations (eg as a frequentist; a **Bayesian would just use probabilistic inference** to update prior to posterior)*

# Have PGM, Some observations, No tables...



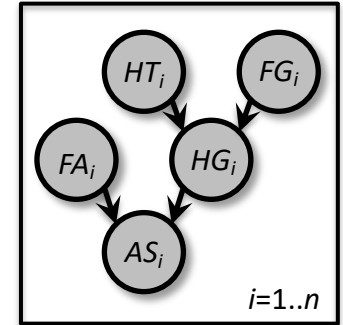
# Fully-observed case is “easy”

- Max-Likelihood Estimator (MLE) says

- \* If we observe *all* r.v.'s  $\mathbf{X}$  in a PGM independently  $n$  times  $\mathbf{x}_i$

- \* Then maximise the *full* joint

$$\arg \max_{\theta \in \Theta} \prod_{i=1}^n \prod_j p(X^j = x_i^j | X^{\text{parents}(j)} = x_i^{\text{parents}(j)})$$



- Decomposes easily, leads to counts-based estimates

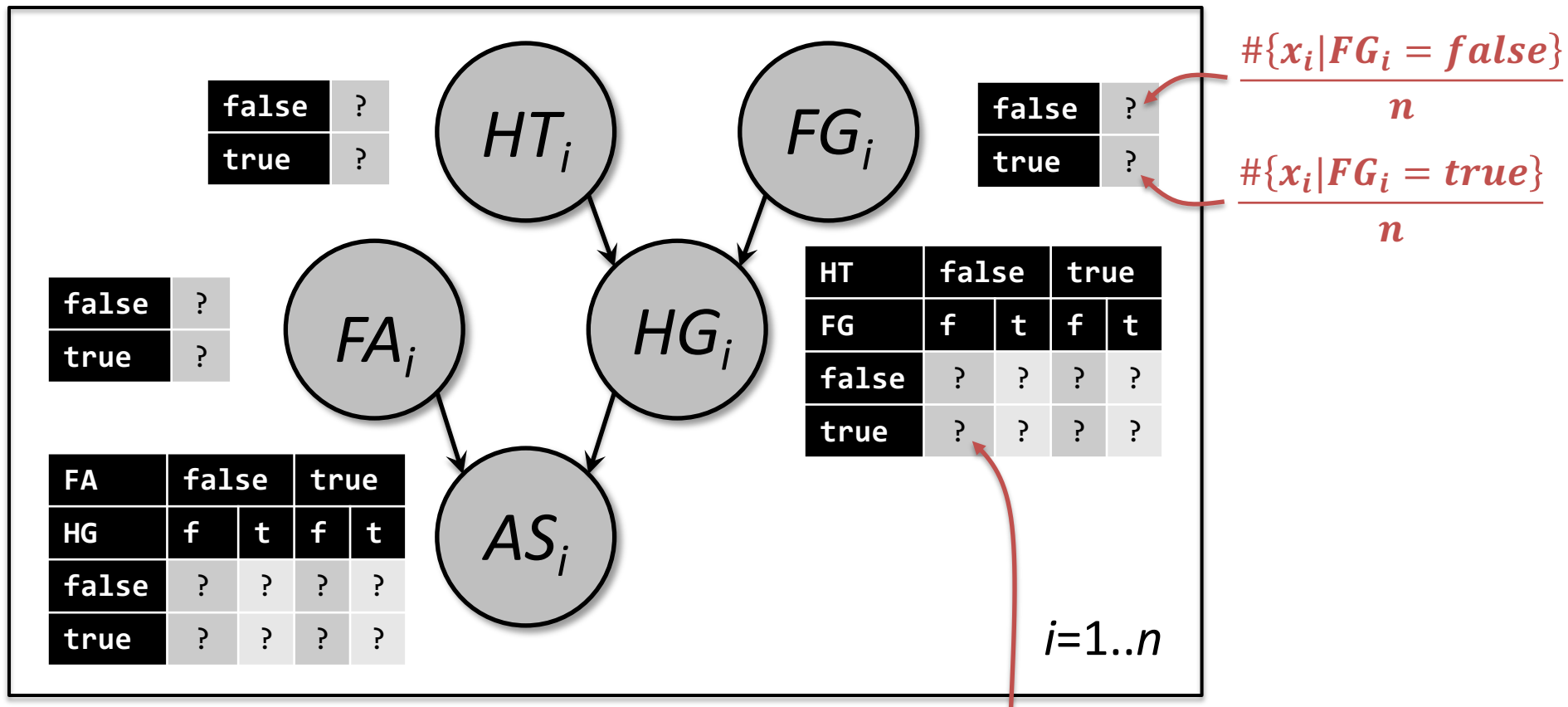
- \* Maximise log-likelihood instead; becomes sum of logs

$$\arg \max_{\theta \in \Theta} \sum_{i=1}^n \sum_j \log p(X^j = x_i^j | X^{\text{parents}(j)} = x_i^{\text{parents}(j)})$$

- \* Big maximisation of all parameters together, *decouples into small independent problems*

- Example is training a naïve Bayes classifier

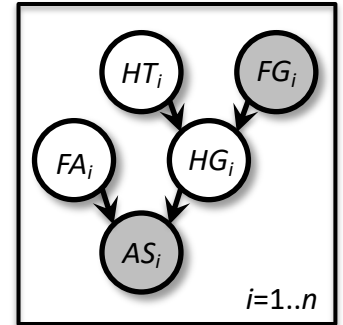
# Example: Fully-observed case



$$\frac{\#\{x_i | HG_i = true, HT_i = false, FG_i = false\}}{\#\{x_i | HT_i = false, FG_i = false\}}$$

# Presence of unobserved variables trickier

- But most PGMs you'll encounter will have latent, or unobserved, variables



- What happens to the MLE?
  - \* Maximise likelihood of observed data only
  - \* Marginalise full joint to get to desired “partial” joint
  - \*  $\arg \max_{\theta \in \Theta} \prod_{i=1}^n \sum_{\text{latent } j} \prod_j p(X^j = x_i^j | X^{\text{parents}(j)} = x_i^{\text{parents}(j)})$
  - \* This won't decouple – oh-no's!!

→ Use **EM algorithm**!

# Summary

- Probabilistic inference on PGMs
  - \* What is it and why do we care?
  - \* Elimination algorithm; complexity via cliques
  - \* Monte Carlo approaches as alternate to exact integration
- Statistical inference on PGMs
  - \* What is it and why do we care?
  - \* Straight MLE for fully-observed data
  - \* EM algorithm for mixed latent/observed data

Next time: deeper dive into HMMs and more