

Bayesian Statistical Methods: Homework 1

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Problem 1

I have decided to use the PyMC package to create a model and sample from the posterior distribution. The code is as follows:

```
# Model
from pymc import *
from numpy import array, empty
from numpy.random import randint

theta = pymc.Beta('theta', alpha=10, beta=10)
d = pymc.Binomial('d', n=np.array([20]), p=theta, value=np.array([3]), observed=True)

# Sampler
import hw1_model
from pymc import MCMC
from pylab import hist, show
from pymc.Matplot import plot

M = MCMC(hw1_model)
M.sample(iter=10000, burn=1000, thin=10)
plot(M)
```

The resulting posterior distributions are as follows:

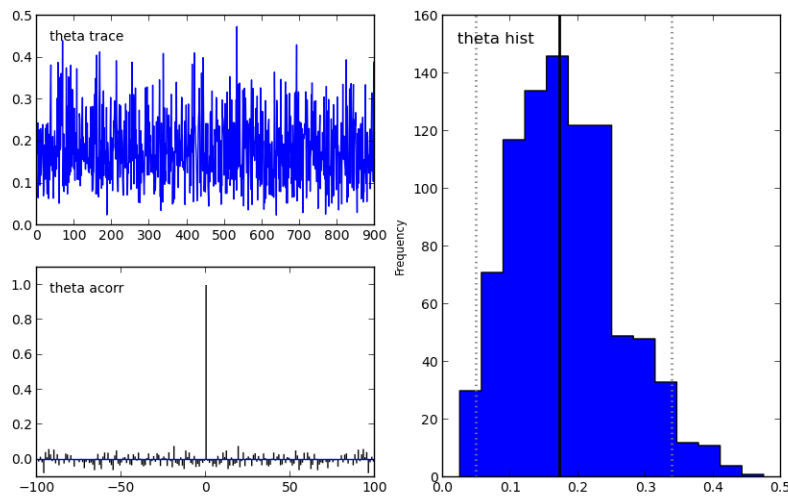
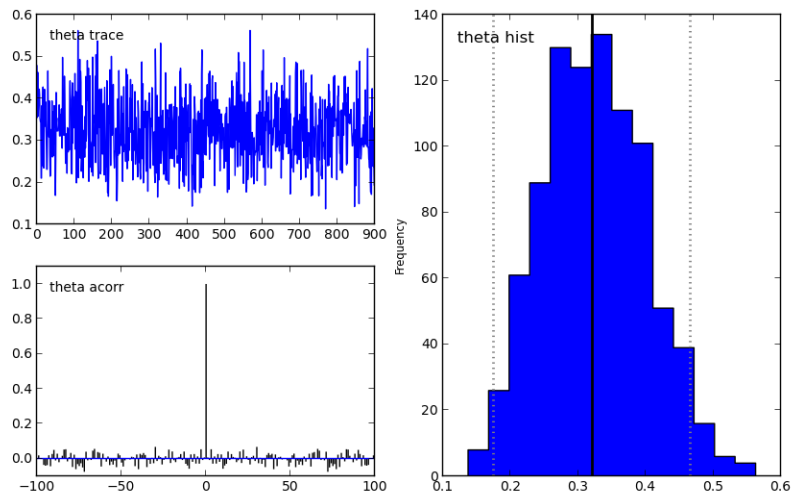
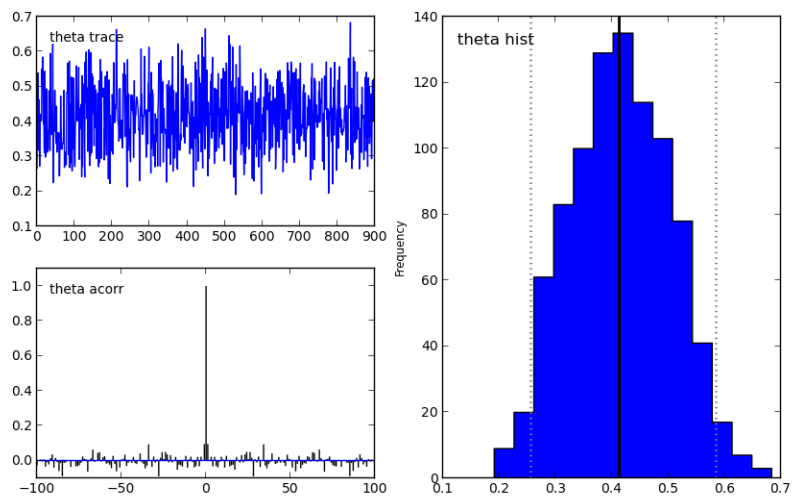


Figure 1: $\alpha = 1, \beta = 1$

It is notable that the mode of the posterior given an uninformed prior (Fig. 1), is closest to 0.15 which corresponds to $Y = 3$ after 20 trials. The influence of adding information to the prior in the form of the parameters α and β , can be seen in Fig. 2 and 3. In these instances the priors did not have a uniform shape.

Figure 2: $\alpha = 10, \beta = 10$ Figure 3: $\alpha = 10, \beta = 1$

Problem 2

For the case where a 20% response rate has been observed, α and β such that the mode of the prior distribution is 0.2.

$$0.2 = \frac{\alpha - 1}{\alpha + \beta - 2} \quad (1)$$

Then, we could scale α and β based on how much weight we wish to place on this previous observation.

Problem 3

Derive Jeffreys' prior for the parameter in the Binomial model.

$$Y \sim \text{Bin}(n, \theta) \quad (2)$$

Where Jeffreys' prior is given as

$$J(\theta) = -E\left[\frac{\partial^2 \log P(y|\theta)}{\partial \theta^2}\right] \quad (3)$$

and

$$P(y|\theta) = \binom{n}{y} \theta^y (1 - \theta)^{n-y} \quad (4)$$

Then, we derive Jeffreys' prior taking the following steps.

$$\log P(y|\theta) = \log \binom{n}{y} + y \log \theta + (n - y) \log(1 - \theta) \quad (5)$$

$$\frac{\partial^2 \log P(y|\theta)}{\partial \theta^2} = -\frac{y}{\theta^2} - \frac{n - y}{(1 - \theta)^2} \quad (6)$$

$$J(\theta) = -E\left[-\frac{y}{\theta^2} - \frac{n - y}{(1 - \theta)^2}\right] \quad (7)$$

$$J(\theta) = \frac{n\theta}{\theta^2} + \frac{n - n\theta}{(1 - \theta)^2} \quad (8)$$

Simplifying, we get:

$J(\theta) = \frac{n}{\theta(1 - \theta)} \quad (9)$
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Problem 4

$$P(\theta|y) \propto P(\theta)L(\theta|y) \quad (10)$$

For a Gamma distribution the probability density function is

$$P(\theta) = \frac{1}{\Gamma(a)b^a} \theta^{a-1} \exp\left(\frac{-\theta}{b}\right) \quad (11)$$

and the likelihood function n independent Poisson responses is

$$L(\theta|y) = \frac{\exp(-n\theta)\theta^{\sum y_i}}{\prod y_i!} \quad (12)$$

Combining Eqn. 10, 11 and 12 while leaving out terms that do not depend on θ .

$$P(\theta|y) \propto \theta^{a-1} \exp\left(\frac{-\theta}{b}\right) \exp(-n\theta) \theta^{\sum y_i} \quad (13)$$

Thus, the posterior distribution is a Gamma

$$P(\theta|y) \propto \theta^{a-1+\sum y_i} \exp\left(-\theta\left(\frac{1}{b} + n\right)\right) \quad (14)$$

with parameters $a_o = a + \sum y_i$ and $b_o = \frac{1}{n+(1/b)}$.

Problem 5

$$P\left(\frac{1}{\sigma^2}|y\right) \propto P\left(\frac{1}{\sigma^2}\right) L\left(\frac{1}{\sigma^2}|y\right) \quad (15)$$

The likelihood of a Normal with a known mean is

$$L\left(\frac{1}{\sigma^2}|y\right) = \left(\frac{1}{2\pi\sigma^2}\right)^{\frac{n}{2}} \exp\left[-\frac{1}{2\sigma^2} \sum (y_i - \mu)^2\right] \quad (16)$$

Using Eqn. 11 where θ is replaced by $\frac{1}{\sigma^2}$ and Eqn. 16, we can substitute into Eqn. ???. Leaving out terms that do not depend on $\frac{1}{\sigma^2}$, we get that

$$P\left(\frac{1}{\sigma^2}|y\right) \propto \left(\frac{1}{\sigma^2}\right)^{a-1} \exp\left(\frac{-1}{\sigma^2 b}\right) \left(\frac{1}{2\pi\sigma^2}\right)^{\frac{n}{2}} \exp\left[-\frac{1}{2\sigma^2} \sum (y_i - \mu)^2\right] \quad (17)$$

Simplifying, the posterior becomes

$$P\left(\frac{1}{\sigma^2}|y\right) \propto \left(\frac{1}{\sigma^2}\right)^{a-1-\frac{n}{2}} \exp\left[\frac{-1}{\sigma^2 b} - \frac{1}{2\sigma^2} \sum (y_i - \mu)^2\right] \quad (18)$$

To bring in prior information on $\frac{1}{\sigma^2}$ to the Gamma distribution, we could take the mean, variance, and number of samples from a previous set of observations of $\frac{1}{\sigma^2}$ and calculate new values for a and b based on the form of Eqn. 18.