Finite State-Space Markov Chains

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Stochastic Processes: Handout 1

IIA Module 3M1: Mathematical Methods

Stochastic Processes: Finite State-Space Markov Chains

Overview

- This part of the course will focus of Markov Chains and related applications
- Course has three parts
 - finite-space Markov chains: 2 lectures
 - continuous state-space systems: 1.5 lectures
 - Monte Carlo Markov chains: 1.5 lectures
- Handouts available from 3M1 web-site
- Insufficient time to examine all topics in detail
 - useful starting point for 4th year modules

Applications of Markov Chains

- Markov Chains under-pin the work in many areas
 - Google Page ranker (very large Markov chain!)
 - Information theory (Entropy)
 - Speech and Language Processing (acoustic models/language models)
 - Physics (statistical mechanics/thermodynamics)
 - Economics and finance (asset pricing)
 - Queueing theory

Pushkin's "Eugene Onegin"





- Andrej Markov (left) analysed Alexander Pushkin's novel Eugene Onegin
 - probability of vowels following consonants (and vice-versa)
 - first analysis of chains (sequences) of (stochastic) events

Language Modelling

Use current word and previous words (state) to predict the next word

Unigram - No information

- Every enter now severally so, let
- Will rash been and by I the me loves gentle me not slavish page, the and hour; ill let

Bigram - Current word

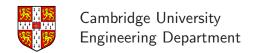
- What means, sir. I confess she? then all sorts, he is trim, captain.
- The world shall- my lord!

Trigram - Current and previous word

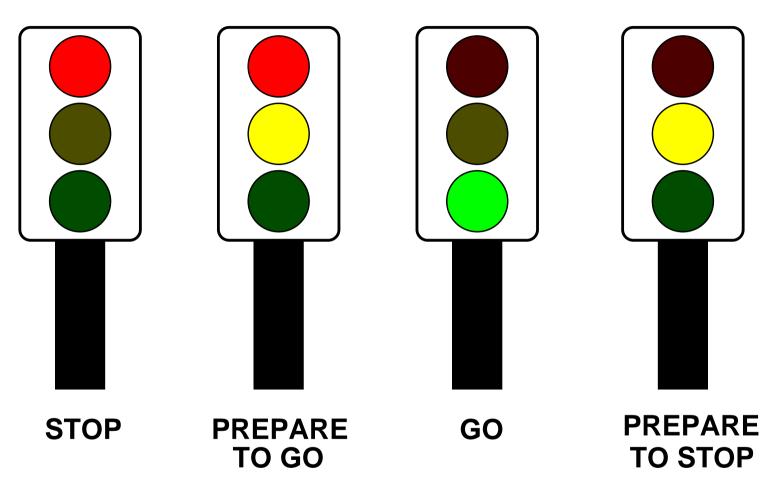
- Indeed the duke; and had a very good friend.
- Sweet prince, Fallstaff shall die. Harry of Monmouth's grave.

4-gram - - Current and previous two words

- It cannot be but so.
- Enter Leonato's brother Antonio, and the rest, but seek the weary beds of people sick.



Traffic Lights



- Simple deterministic process:
 - four distinct states (numbered 1 to 4) with associated meaning!

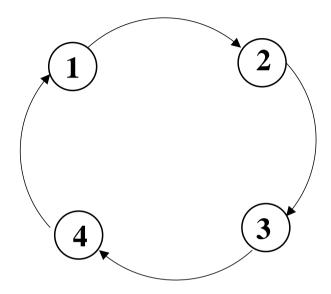
Traffic Lights

• Can describe the process as a state-diagram with transitions

a deterministic process

The current state completely determines the next state

Transition Matrix Description



• The transition matrix for this process is very simple:

$$\texttt{red} \to \texttt{red/amber} \to \texttt{green} \to \texttt{amber} \to \texttt{red}$$

$$\texttt{state} \ 1 \to \texttt{state} \ 2 \to \texttt{state} \ 3 \to \texttt{state} \ 4 \to \texttt{state} \ 1$$

ullet More generally each element of the transition matrix ${f P}$ indicates:

$$P_{j,k} = \operatorname{Prob}(j \to k), \quad \sum_{k} P_{j,k} = 1, \quad P_{j,k} \ge 0$$

Traffic Lights in a City

- Consider a (large) city with a large number of traffic lights
- At any instance in time traffic lights (drivers) can be split into
 - red (driving stopped)
 - red/amber (driver preparing to go)
 - green (driver going)
 - amber (driver preparing to stop)
- ullet The initial distribution of traffic lights, ${f x}^{(0)}$, will be written as

$$\mathbf{x}^{(0)} = \left[\begin{array}{cc} P(X_0 = \mathtt{red}) & P(X_0 = \mathtt{red/amber}) & P(X_0 = \mathtt{green}) & P(X_0 = \mathtt{amber}) \end{array}\right]$$

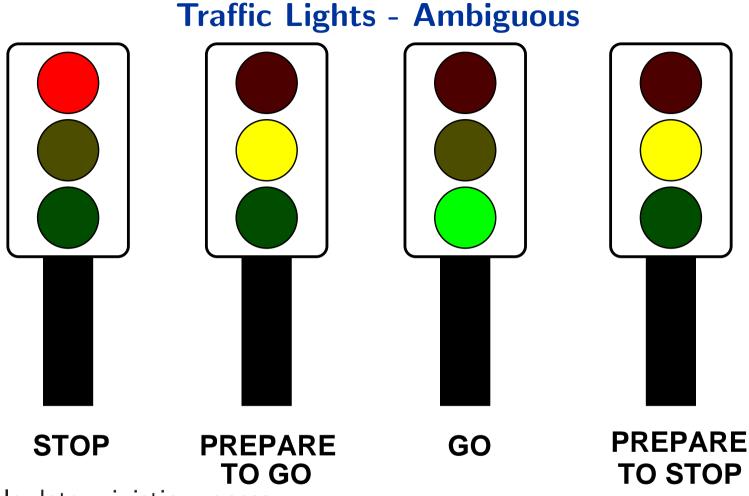
- at any instance in time a particular traffic light will be in one of the states
- the transition matrix will describe the state movement of each light

Markov Chain (Formal Definition)

A Stochastic Process X_0, X_1, X_2, \ldots is a Markov chain if and only if for all times $i \geq 1$

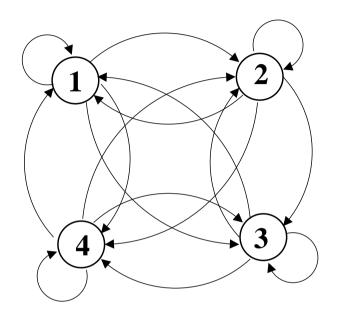
$$P(X_{i+1}|X_0=j_0,X_1=j_1,\ldots,X_i=j_i)=P(X_{i+1}|X_i=j_i)$$

- ullet There are a set of random variables X_0, X_1, X_2, \ldots indexed by time
 - each random variable takes a value from the state-space, \mathcal{S} .



- Simple deterministic process:
 - but not a Markov chain
 - amber state ambiguity

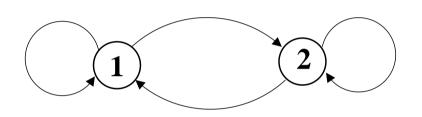
Faulty Traffic Lights



- How would we want to characterise these sorts of processes:
 - what happens after n steps?
 - what happens if the system is run for a long-time?
 - if light in red, how long until light expected to be in green?
 - will lights end up in a periodic cycle?

Population of California

- Simple model of population change in California: each year
 - -1/10 of the population outside California moves in
 - -2/10 of the population inside California moves out



$$\mathbf{P} = \begin{bmatrix} 0.9 & 0.1 \\ 0.2 & 0.8 \end{bmatrix}$$

- Two states:
 - state 1 outside California
 - state 2 inside California

What is the probability of an individual's location? What happens to the population of California in the log-run?

"Path" Probability

What is the probability of the following sequence for an individual

$$ext{inside} o ext{outside} o ext{inside} o ext{inside}$$

- initial distribution $\mathbf{x}^{(0)} = \begin{bmatrix} 0.5 & 0.5 \end{bmatrix}$
- Need to compute

$$\begin{split} P(X_0 = \texttt{inside}, X_1 = \texttt{outside}, X_2 = \texttt{inside}, X_3 = \texttt{inside}) \\ &= P(X_0 = \texttt{inside}) P(X_1 = \texttt{outside} | X_0 = \texttt{inside}) \\ &\times P(X_2 = \texttt{inside} | X_1 = \texttt{outside}) P(X_3 = \texttt{inside} | X_2 = \texttt{inside}) \\ &= x_2^{(0)} \times 0.2 \times 0.1 \times 0.8 = 0.5 \times 0.2 \times 0.1 \times 0.8 \\ &= 0.008 \end{split}$$

Population of California after n Years

Initial distribution of population at year zero (row vector):

$$\left[\begin{array}{cc} x_1^{(0)} & x_2^{(0)} \end{array}\right]$$

• After 1 year population:

$$\begin{bmatrix} 0.9x_1^{(0)} + 0.2x_2^{(0)} & 0.1x_1^{(0)} + 0.8x_2^{(0)} \end{bmatrix} = \begin{bmatrix} x_1^{(0)} & x_2^{(0)} \end{bmatrix} \begin{bmatrix} 0.9 & 0.1 \\ 0.2 & 0.8 \end{bmatrix}$$
$$= \begin{bmatrix} x_1^{(1)} & x_2^{(1)} \end{bmatrix}$$

• After *n* years population (by simple recursion):

$$\begin{bmatrix} x_1^{(n)} & x_2^{(n)} \end{bmatrix} = \begin{bmatrix} x_1^{(n-1)} & x_2^{(n-1)} \end{bmatrix} \begin{bmatrix} 0.9 & 0.1 \\ 0.2 & 0.8 \end{bmatrix} = \begin{bmatrix} x_1^{(0)} & x_2^{(0)} \end{bmatrix} \begin{bmatrix} 0.9 & 0.1 \\ 0.2 & 0.8 \end{bmatrix}^n$$

Theorem

For any time n

$$\mathbf{x}^{(n)} = \mathbf{x}^{(0)} \mathbf{P}^n$$

where $\mathbf{x}^{(n)}$ is the vector denoting the distribution of X_n .

- ullet Possible to get the transition matrix for n-steps:
 - raise the 1-step matrix to the power n

Chapman-Kolmogorov Equations

From the previous discussion we can say

$$P(X_{i+n} = k | X_i = j) = (\mathbf{P}^n)_{j,k} = P_{j,k}^n$$

Chapman-Kolmogorov equations: possible to write as

$$P(X_{i+n} = k | X_i = j) = P_{j,k}^n$$

$$= (\mathbf{P}^s \mathbf{P}^{n-s})_{j,k}$$

$$= \sum_{l} P_{j,l}^s P_{l,k}^{n-s}$$

$$= \sum_{l} P(X_{i+s} = l | X_i = j) P(X_{i+n} = k | X_{i+s} = l)$$

Limiting and Stationary Distributions

• For the faulty traffic lights, what happens after a very large number of steps

$$\mathbf{P} = \begin{bmatrix} 1/2 & 1/6 & 1/6 & 1/6 \\ 1/6 & 1/2 & 1/6 & 1/6 \\ 1/6 & 1/6 & 1/2 & 1/6 \\ 1/6 & 1/6 & 1/6 & 1/2 \end{bmatrix}$$

$$\mathbf{P} = \begin{bmatrix} 1/2 & 1/6 & 1/6 & 1/6 \\ 1/6 & 1/2 & 1/6 & 1/6 \\ 1/6 & 1/6 & 1/2 & 1/6 \\ 1/6 & 1/6 & 1/6 & 1/2 \end{bmatrix} \qquad \mathbf{P}^{\infty} = \begin{bmatrix} 1/4 & 1/4 & 1/4 & 1/4 \\ 1/4 & 1/4 & 1/4 & 1/4 \\ 1/4 & 1/4 & 1/4 & 1/4 \\ 1/4 & 1/4 & 1/4 & 1/4 \end{bmatrix}$$

• So for an initial distribution $\mathbf{x}^{(0)}$, the "final" distribution is

$$\mathbf{x}^{(\infty)} = \mathbf{x}^{(0)} \mathbf{P}^{\infty} = \begin{bmatrix} 1/4 & 1/4 & 1/4 \end{bmatrix}$$

 whatever initial distribution of start states, after "long-enough" equally likely to be in any state!

Limiting Distributions/Regular Ergodic (Definition)

For a Markov Process with transition matrix \mathbf{P} , a limiting distribution, $\mathbf{x}^{(\infty)}$, is one which for any initial distribution $\mathbf{x}^{(0)}$ satisfies $\mathbf{x}^{(0)}\mathbf{P}^n=\mathbf{x}^{(\infty)}$ as $n\to\infty$

• For any limiting distribution:

$$\mathbf{x}^{(\infty)}\mathbf{P} = \mathbf{x}^{(\infty)}$$

- any distribution that satisfies this expression is a stationary distribution

A Markov Chain is regular ergodic if there exists a unique stationary distribution which is the limit point for all initial distributions and which puts positive mass on every element of the state-space S.

- For an $n \times n$ transition matrix **P** then
 - 1. $\lambda = 1$ is an eigenvalue of **P**
 - 2. all eigenvalues satisfy $|\lambda| \leq 1$

Transition Matrix Eigenvalues (Proof)

• For one to be an eigenvalue: $det(\mathbf{P} - \mathbf{I}) = 0$

$$\mathbf{P} - \mathbf{I} = \begin{bmatrix} P_{1,1} - 1 & P_{1,2} & \cdots & P_{1,N} \\ P_{2,1} & P_{2,2} - 1 & \cdots & P_{2,N} \\ \vdots & \vdots & \ddots & \vdots \\ P_{N,1} & P_{N,2} & \cdots & P_{N,N} - 1 \end{bmatrix}$$

- summing first N-1 columns yields (negative) last column
- not a full-rank matrix, determinant is zero, one is an eigenvalue
- ullet Select an eigenvalue λ and (right) eigenvector $oldsymbol{v}$ of ${f P}$
 - select largest (magnitude) element of eigenvector v_k ($|v_k| \ge |v_i|, \forall i$)
 - consider equality $\mathbf{P} oldsymbol{v} = \lambda oldsymbol{v}$ and element k

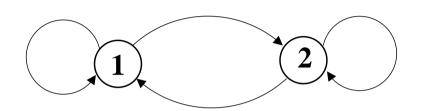
$$|\lambda||v_k| = |\lambda v_k| = |P_{k,1}v_1 + \dots + P_{k,N}v_N|$$

$$\leq P_{k,1}|v_1| + \dots + P_{k,N}|v_N|$$

$$\leq P_{k,1}|v_k| + \dots + P_{k,N}|v_k| = (P_{k,1} + \dots + P_{k,N})|v_k| = |v_k|$$

California Population

Interested in the stationary distribution for California population



$$\mathbf{P} = \begin{bmatrix} 0.9 & 0.1 \\ 0.2 & 0.8 \end{bmatrix}$$

- in the limit

$$\mathbf{P}^{\infty} = \begin{bmatrix} 0.9 & 0.1 \\ 0.2 & 0.8 \end{bmatrix}^{\infty} = \begin{bmatrix} 2/3 & 1/3 \\ 2/3 & 1/3 \end{bmatrix}$$

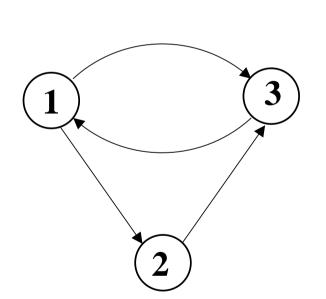
- stationary (limiting) distribution

$$\mathbf{x}^{(\infty)} = \begin{bmatrix} 2/3 & 1/3 \end{bmatrix}$$

- California has a third of the population of the US!
- does not depend on the initial population distribution.

Random Surfer

- Consider an individual who surfs the web in the following fashion
 - d fraction of the time randomly selects a link on the current page
 - -(1-d) fraction of the time randomly selects a web-page
- This is a Markov Process simple 3 web-page example



$$\mathbf{P} = \begin{bmatrix} 0 & 1/2 & 1/2 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \end{bmatrix}$$

When used with random surfer

$$\tilde{\mathbf{P}} = \frac{(1-d)}{3} \begin{bmatrix} 1 & 1 & 1 \\ 1 & 1 & 1 \\ 1 & 1 & 1 \end{bmatrix} + d\mathbf{P}$$

What is the stationary distribution?

PageRank

- Pagerank (Pr) important component for original Google search
 - there is a total of N webpages
 - \mathcal{L}_i set of pages that link to i, $\operatorname{Ln}(j)$ number of links from j

$$\Pr(i) = (1 - d)/N + d\left(\sum_{j \in \mathcal{L}_i} \Pr(j)/\operatorname{Ln}(j)\right)$$

• Consider stationary distribution and element i for random surfer

$$x_i^{(\infty)} = (\mathbf{x}^{(\infty)}\mathbf{P})_i = (1 - d)/N + d\left(\sum_{j \in \mathcal{L}_i} x_j^{(\infty)} P_{j,i}\right)$$

- $P_{j,i}=1/\mathrm{Ln}(j)$ probability of random move from j to i $x_i^{(\infty)}$ is the PageRank of page i

Stationary Distributions and (Left) EigenVectors

• To find a stationary distribution consider

$$\mathbf{P}^{T}\mathbf{x}^{(\infty)T} = \mathbf{x}^{(\infty)T}$$

- for the transition matrix $|\lambda| \leq 1$,
- need to find (right) eigenvector(s) of \mathbf{P}^{T} when $\lambda=1$
- For California distribution

$$\mathbf{P}^{\mathsf{T}} = \begin{bmatrix} 0.9 & 0.2 \\ 0.1 & 0.8 \end{bmatrix} = \begin{bmatrix} 0.894 & -0.707 \\ 0.447 & 0.707 \end{bmatrix} \begin{bmatrix} 1.0 & 0 \\ 0 & 0.7 \end{bmatrix} \begin{bmatrix} 0.894 & -0.707 \\ 0.447 & 0.707 \end{bmatrix}^{-1}$$

- eigenvector for $\lambda=1$ is $\begin{bmatrix} 0.894 & 0.447 \end{bmatrix}^{\mathrm{T}}$
- need to satisfy sum to one constraint:

$$\mathbf{x}^{(\infty)} = \begin{bmatrix} 2/3 & 1/3 \end{bmatrix}$$

EigenValues/EigenVectors (Reminder)

- \bullet Possible to express **P** in terms of eigenvectors/eigenvalues (note not symmetric)
 - set-up as (standard) right eigenvectors

$$\mathbf{P}^{\mathsf{T}} = \mathbf{V}\mathbf{D}\mathbf{V}^{-1}$$

- \mathbf{V} is the matrix of eigenvectors
- D is the diagonal matrix of eigenvalues
- ullet Interested in finding $\mathbf{P}^{\mathsf{T}\infty}$

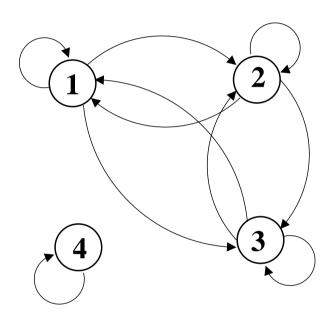
$$\mathbf{P}^{\mathsf{T}\infty} = \mathbf{V}\mathbf{D}^{\infty}\mathbf{V}^{-1}$$

- \mathbf{D}^{∞} will only be non-zero for $|\lambda| = 1$ (note $|\lambda| \leq 1$)
- Simple to find eigenvector for $\lambda = 1$ solving

$$\mathbf{P}^{\mathsf{T}}\mathbf{v} = \mathbf{v}$$

Modified Faulty Traffic Lights

• The traffic lights are now faulty with the following Markov Chain:



$$\begin{bmatrix} 1/2 & 1/4 & 1/4 & 0 \\ 1/4 & 1/2 & 1/4 & 0 \\ 1/4 & 1/4 & 1/2 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

• The stationary distributions for this Markov Chain are:

$$\mathbf{x}^{(\infty)} = \begin{bmatrix} 1/3 & 1/3 & 1/3 & 0 \end{bmatrix}, \quad \mathbf{x}^{(\infty)} = \begin{bmatrix} 0 & 0 & 0 & 1 \end{bmatrix}$$

Modified Faulty Traffic Lights (cont)

- Clearly the process is not regular ergodic:
 - two stationary distributions;
 - the stationary distribution depends on the initial distribution;
- Consider an initial distribution of the form

$$\mathbf{x}^{(0)} = \begin{bmatrix} a & 0.0 & 0.0 & 1 - a \end{bmatrix}$$

the final distribution of traffic lights is

$$\mathbf{x}^{(\infty)} = \begin{bmatrix} a/3 & a/3 & a/3 & 1-a \end{bmatrix}$$

Communicate (Definition)

Two states of a Markov chain j and k communicate if, and only if, there exists integers m and n such that

$$P_{j,k}^m > 0$$
 and $P_{k,j}^n > 0$

- Intuitively this means:
 - there must be a route from state j to k (in m steps)
 - and there must be a route from state k to j (in n steps)

Recurrent Set (Definition)

Let S denote the set of all possible states of a Markov chain. A subset $\tilde{S} \subseteq S$ is a recurrent set if

- 1. all pairs of states in $\tilde{\mathcal{S}}$ communicate
- 2. if $j \in \tilde{\mathcal{S}}$ and $k \notin \tilde{\mathcal{S}}$ then $P^i_{j,k} = 0$ for all $i \geq 0$
- Intuitively this means:
 - it is possible to move between all states in a recurrent set
 - cannot move from a member of a recurrent state to not a member

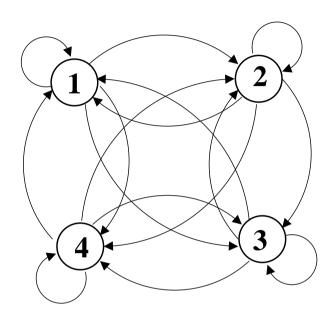
Recurrent and Transient / Irreducible (Definitions)

If a state belongs to a recurrent set then it is recurrent, otherwise it is transient

- Possible to show
 - 1. if state k is recurrent then $P(\text{return to state } k|X_o=k)=1$
 - 2. if we start in a transient state then eventually we will leave that state and never return to it

If all states in a Markov chain communicate with each other then the Markov chain is irreducible

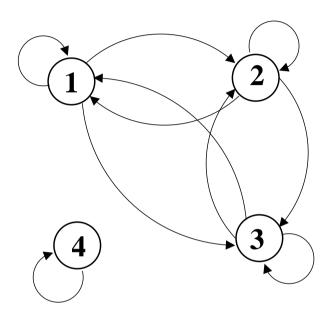
Faulty Traffic Lights



$$\begin{bmatrix} 1/2 & 1/6 & 1/6 & 1/6 \\ 1/6 & 1/2 & 1/6 & 1/6 \\ 1/6 & 1/6 & 1/2 & 1/6 \\ 1/6 & 1/6 & 1/6 & 1/2 \end{bmatrix}$$

- This Markov chain is irreducible
 - all states communicate with each other

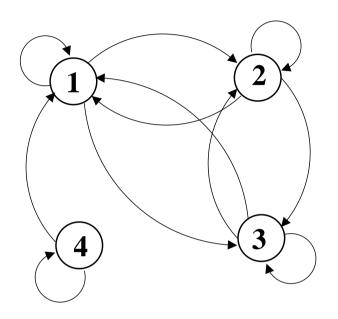
Modified Faulty Traffic Lights



$$\begin{bmatrix} 1/2 & 1/4 & 1/4 & 0 \\ 1/4 & 1/2 & 1/4 & 0 \\ 1/4 & 1/4 & 1/2 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

- States 1, 2 and 3 form a recurrent set
- State 4 forms a recurrent set

Modified Faulty Traffic Lights (2)



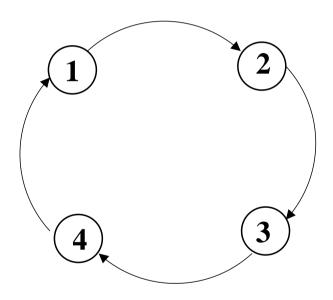
$$\begin{bmatrix} 1/2 & 1/4 & 1/4 & 0 \\ 1/4 & 1/2 & 1/4 & 0 \\ 1/4 & 1/4 & 1/2 & 0 \\ 1/2 & 0 & 0 & 1/2 \end{bmatrix}$$

- States 1, 2 and 3 form a recurrent set
- State 4 is a transient state
 - eventually state 4 is left and never returned to

Waiting Time

• I arrive when the traffic lights are at red: how long do I expect to wait

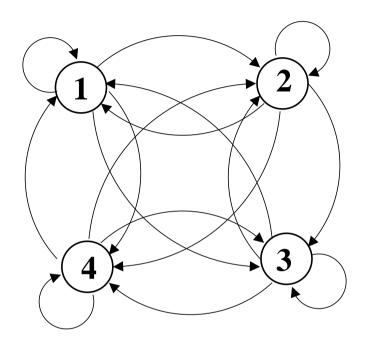
$$q_R = \mathcal{E}[\mathsf{time} \; \mathsf{to} \; \mathsf{wait} \; \mathsf{until} \; \mathsf{first} \; \mathsf{green} | X_0 = \mathsf{red}]$$

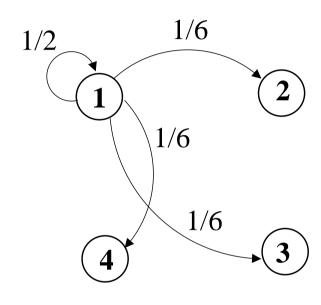


0.0	1.0	$\cap \cap$
	1.0	0.0
0.0	0.0	1.0
0.0	0.0	0.0
	0.0	0.0 0.0

- For standard traffic lights, simple 2!
 - state 1 is red, state 3 is green

Waiting Time - Faulty Traffic Lights





- Starting in red possible transitions on the right
 - if light in red, then still expect to wait q_R
 - if light goes to red/amber then expect to wait q_{RA}
 - if light goes to amber then expect to wait q_A
 - if light goes to green then wait finished!!

Waiting Time - Faulty Traffic Lights

Set-up equation to describe scenario

$$q_R = 1/2(1+q_R) + 1/6(1+q_{RA}) + 1/6(1+q_A) + 1/6$$

ullet Now need expressions for q_{RA} and q_A

$$q_{RA} = 1/2(1+q_{RA}) + 1/6(1+q_R) + 1/6(1+q_A) + 1/6$$

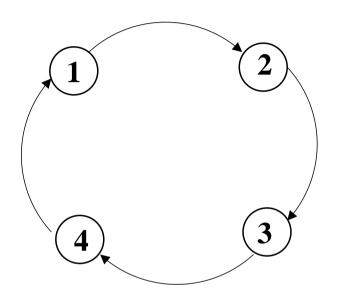
$$q_A = 1/2(1+q_A) + 1/6(1+q_R) + 1/6(1+q_{RA}) + 1/6$$

• Series of linear equations - simply solve to get:

$$q_R = 6$$
, $q_A = 6$, $q_{RA} = 6$, $q_G = 0$

Periodicity

• Consider the standard, working, traffic lights



• If the traffic lights start at red then the process is

$$\texttt{red} \to \texttt{red/amber} \to \texttt{green} \to \texttt{amber} \to \texttt{red}$$

$$\texttt{state} \ 1 \to \texttt{state} \ 2 \to \texttt{state} \ 3 \to \texttt{state} \ 4 \to \texttt{state} \ 1$$

- this has a period of 4

Periodicity (Definition)

The period of a state k of a Markov chain is the greatest common divisor of the set

$$\{i \ge 0 : P_{k,k}^i > 0\}$$

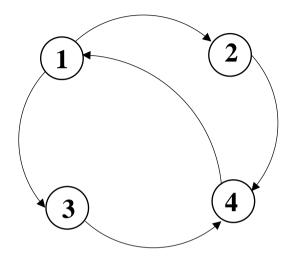
A state k is aperiodic if it has period 1.

- A Markov chain is aperiodic if all states are aperiodic
 - a state k is aperiodic if for some time i

$$P_{k,k}^{i} > 0$$
 and $P_{k,k}^{i+1} > 0$

 if a chain is irreducible and aperiodic then it is regular ergodic it will have a limiting distribution

Periodicity - Example 1 from UW



$$\begin{bmatrix} 0 & 1/2 & 1/2 & 0 \\ 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 1 \\ 1 & 0 & 0 & 0 \end{bmatrix}$$

• Possible transitions are

$$\begin{array}{c} 1 \rightarrow 2 \rightarrow 4 \rightarrow 1 \\ 1 \rightarrow 3 \rightarrow 4 \rightarrow 1 \end{array}$$

• Looking at the eigenvalues of P

$$\lambda^4 - \lambda = 0; \quad \lambda = 0, 1, e^{2\pi i/3}, e^{4\pi i/3}$$

– three eigenvalues satisfy $|\lambda|=1$

Periodicity - Example 1 from UW

• The transition matrices for 2,3 and 4 steps are:

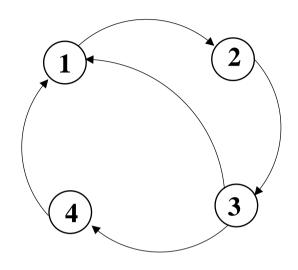
$$\mathbf{P}^{2} = \begin{bmatrix} 0 & 0 & 0 & 1 \\ 1 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 \\ 0 & 1/2 & 1/2 & 0 \end{bmatrix}, \mathbf{P}^{3} = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1/2 & 1/2 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix} \mathbf{P}^{4} = \begin{bmatrix} 0 & 1/2 & 1/2 & 0 \\ 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 1 \\ 1 & 0 & 0 & 0 \end{bmatrix}$$

- The process is not regular ergodic
 - it does not settle down to a unique stationary distribution

$$\mathbf{x}^{(0)}\mathbf{P} = \mathbf{x}^{(0)}\mathbf{P}^4 = \mathbf{x}^{(0)}\mathbf{P}^7 = \dots$$

- states are periodic with period 3

Periodicity - Example 2 from UW



$$\begin{bmatrix} 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 1/2 & 0 & 0 & 1/2 \\ 1 & 0 & 0 & 0 \end{bmatrix}$$

Possible transitions are

$$\begin{array}{c} 1 \rightarrow 2 \rightarrow 3 \rightarrow 1 \\ 1 \rightarrow 2 \rightarrow 3 \rightarrow 4 \rightarrow 1 \end{array}$$

- Possible to return to state 1 after 3/4/6/7/8 ...
 - the period of state 1 is 1 it is aperiodic

Periodicity - Example 2 from UW

Looking at the eigenvalues of P

$$\lambda^4 - \lambda/2 - 1/2 = 0; \quad \lambda = 1, -0.176 \pm 0.861i, -0.648$$

- one eigenvalues satisfy $|\lambda|=1$
- The transition matrices for 2,3,4,5,6 and 7 steps are:

$$\mathbf{P}^2 = \begin{bmatrix} 0 & 0 & 1 & 0 \\ 1/2 & 0 & 0 & 1/2 \\ 1/2 & 1/2 & 0 & 0 \\ 0 & 1 & 0 & 0 \end{bmatrix}, \mathbf{P}^3 = \begin{bmatrix} 1/2 & 0 & 0 & 1/2 \\ 1/2 & 1/2 & 0 & 0 \\ 0 & 1/2 & 1/2 & 0 \\ 0 & 0 & 1 & 0 \end{bmatrix} \mathbf{P}^4 = \begin{bmatrix} 1/2 & 1/2 & 0 & 0 \\ 0 & 1/2 & 1/2 & 0 \\ 1/4 & 0 & 1/2 & 1/4 \\ 1/2 & 0 & 0 & 1/2 \end{bmatrix}$$

$$\mathbf{P}^{5} = \begin{bmatrix} 0 & 1/2 & 1/2 & 0 \\ 1/4 & 0 & 1/2 & 1/4 \\ 1/2 & 1/4 & 0 & 1/4 \\ 1/2 & 1/2 & 0 & 0 \end{bmatrix}, \mathbf{P}^{6} = \begin{bmatrix} 1/4 & 0 & 1/2 & 1/4 \\ 1/2 & 1/4 & 0 & 1/4 \\ 1/4 & 1/2 & 1/4 & 0 \\ 0 & 1/2 & 1/2 & 0 \end{bmatrix} \mathbf{P}^{7} = \begin{bmatrix} 1/2 & 1/4 & 0 & 1/4 \\ 1/4 & 1/2 & 1/4 & 0 \\ 1/4 & 0 & 1/2 & 1/4 \end{bmatrix}$$

Periodicity - Example 2 from UW

- For this system looking at the leading diagonals $(P_{k,k}^i > 0)$
 - states 1,2,3 non-zero steps: 3,4
 - state 4 non-zero steps: 4,7
- Lowest common divisor is 1 for all states:
 - system is aperiodic
 - system also irreducible (all states communicate)
 - system is therefore regular ergodic
- Unique stationary (limiting) distribution

$$\mathbf{P}^{\infty} = \begin{bmatrix} 2/7 & 2/7 & 2/7 & 1/7 \\ 2/7 & 2/7 & 2/7 & 1/7 \\ 2/7 & 2/7 & 2/7 & 1/7 \\ 2/7 & 2/7 & 2/7 & 1/7 \end{bmatrix}$$

Periodicity and Eigenvalues (Reference)

- ullet Consider a Markov Chain with transition matrix ${f P}$ that has a period 3
 - there will be three eigenvalues of ${f P}$ with magnitude 1
 - the values of these eigenvalues will be: $1, e^{2\pi i/3}, e^{4\pi i/3}$

Example 1 from UW is a Markov Chain of this form:

$$\mathbf{P} = \begin{bmatrix} 0 & 1/2 & 1/2 & 0 \\ 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 1 \\ 1 & 0 & 0 & 0 \end{bmatrix}; \quad \text{eigenvalues: } 1, e^{2\pi i/3}, e^{4\pi i/3}, 0$$

- ullet More generally for a Markov Chain with period δ
 - there will be δ eigenvalues of ${f P}$ with magnitude 1
 - the values of these eigenvalues will be: $1, e^{2\pi i/\delta}, e^{4\pi i/\delta}, \dots$

Ergodic Theorem (Reference) / Detailed Balance

If X_o, X_1, \ldots is a regular ergodic Markov chain with stationary distribution π , then for any function f(x) as $N \to \infty$

$$\frac{1}{N} \sum_{i=1}^{N} f(X_i) \to \sum_{k \in \mathcal{S}} \pi_k f(k)$$

where S is the state-space of the Markov chain.

A transition matrix P and a distribution π are in detailed balance if

$$\pi_j P_{j,k} = \pi_k P_{k,j}$$

where S is the state-space of the Markov chain.

Detailed Balance & Stationary Distribution

A distribution π satisfies detailed balance with transition matrix P

$$\pi_j P_{j,k} = \pi_k P_{k,j}$$

What happens to this distribution after 1-step?

$$ilde{m{\pi}} = m{\pi} \mathbf{P}$$

Now consider element k of the distribution

$$\tilde{\pi}_k = \sum_j \pi_j P_{j,k} = \sum_j \pi_k P_{k,j} = \pi_k \left(\sum_j P_{k,j}\right) = \pi_k$$

ullet π is a stationary distribution of P

Summary

- Finite-state (discrete) Markov chains
 - current state completely determines probability of next state
- Range of attributes to characterise Markov Chains
 - limiting and stationary distributions
 - wait-times
 - periodicity and nature of states (communicate, transient etc)
- Next lecture examines continuous state-space systems