(wischern & Johnson)

- Multivariate Normal Distribution Functions,
- Conditional Distribution and its relation to regression model,
- Estimation of parameters.

1-> Multivariade normal Dist?

2-> Multiple Lineal Regression

3-> Multivariate Regression

# NORMAL DISTRIBUTION (Univariate)

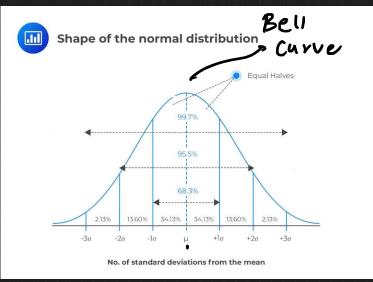
$$f(n; \mu, r) = \frac{1}{\sqrt{2\pi}} \exp \left\{ \frac{-1}{2\sigma r} (x - \mu)^{2} \right\} ; -\infty < \mu < \infty$$

M= Mean wherl

6 = Standard Deviation

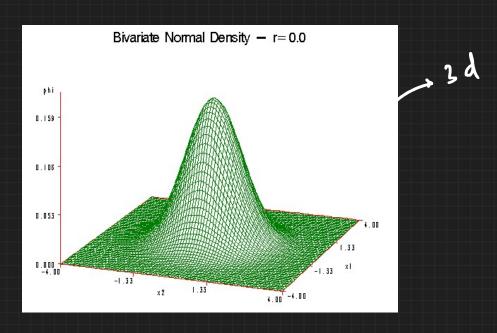
G= variance

- m < m < m



Mean = modl = median

The "regular" normal distribution has one random variable; A bivariate normal distribution is made up of two independent random variables. The two variables in a bivariate normal are both are normally distributed, and they have a normal distribution when both are added together. Visually, the bivariate normal distribution is a three-dimensional bell curve.



pat of Birariate Normal dista is

$$f(x_1, x_2) = \frac{1}{2\pi\epsilon_1\epsilon_2} \exp\left\{-\frac{Z}{2(1-\varsigma^2)}\right\}$$

where 
$$2 = (\frac{1}{4 - \frac{1}{4}} - \frac{2}{3}(\frac{1}{2} - \frac{1}{4})(\frac{1}{4} - \frac{1}{4}) + (\frac{1}{4} - \frac{1}{4})^{2}$$

$$g = Correlation$$
 $coefficients$ 
 $cov(24, 32)$ 
 $cov(24, 32)$ 

Matrix approach to Bivariate Dist's

$$x = \begin{bmatrix} x_1 \\ y_2 \end{bmatrix}$$
  $\mathcal{M} = \mathcal{E}(x) = \begin{bmatrix} \mathcal{M}_1 \\ \mathcal{M}_2 \end{bmatrix}$ 

$$\Sigma = \begin{bmatrix} V(M) & COV(NIM) \\ COV(NIM) & V(M) \end{bmatrix} = Var-Covariance matrix = \begin{bmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{21} & \sigma_{22} \end{bmatrix}$$

P = No . of independent variables

$$(x-u)^{2} = (x-u)(\sigma L)^{-1}(x-u)$$

$$= (x-u)^{2} \Sigma^{-1}(x-u)$$

pdf of Bivariate normal distribution (Matrix Form)
$$f(n) = \frac{1}{(2\pi)^{1/2}} \frac{\exp(\frac{1}{2}(n-u)^{1/2} + (n-u)^{1/2})}{(2\pi)^{1/2}}$$

det 
$$V = \det \begin{pmatrix} \sigma_1^2 & S\sigma_1\sigma_2 \\ S\sigma_1\sigma_1 & G_2^2 \end{pmatrix} = \sigma_1^* \sigma_1^* (1-S^2)$$

$$v'' = \frac{1}{det(v)} \begin{pmatrix} 6^{\frac{1}{2}} & -96^{\frac{1}{2}} 6^{\frac{1}{2}} \\ -96^{\frac{1}{2}} & 6^{\frac{1}{2}} \end{pmatrix} = \frac{1}{6^{\frac{1}{2}} 6^{\frac{1}{2}} (1-9^{\frac{1}{2}})} \begin{pmatrix} 6^{\frac{1}{2}} & -96^{\frac{1}{2}} 6^{\frac{1}{2}} \\ -96^{\frac{1}{2}} & 6^{\frac{1}{2}} \end{pmatrix}$$

$$=\frac{1}{(1-5^{2})}\begin{pmatrix} \sigma_{1}^{-2} & -36.5^{-1} \\ -36.5^{-1} & \sigma_{2}^{-2} \end{pmatrix}$$

If X and Y are bivariate normal, what is the necessary and sufficient condition for X and Y to be

independent?  

$$\rightarrow \frac{1}{2\pi r_1 r_2} = \frac{1}{2\pi r_1 r_$$

when 
$$8=0$$
)
$$f(x,y) = \frac{1}{2\pi 6_{1}6_{2}} \exp \left[-\frac{1}{2}\left(\frac{2-\mu_{1}}{6_{1}}\right)^{2} + \left(\frac{y-\mu_{2}}{6_{2}}\right)^{2}\right]$$

$$f(x,y) = \frac{1}{2\pi 6_{1}6_{2}} \exp \left\{-\frac{1}{2}\left(\frac{x-\mu_{1}}{6_{1}}\right)^{2}\right\} - \frac{1}{5\pi 6_{2}} \exp \left\{-\frac{1}{2}\left(\frac{y-\mu_{2}}{6_{2}}\right)^{2}\right\}$$

$$f(xy) = f_{K}(x)$$
.  $f_{Y}(J)$   
Hence  $x & y$  are independent.

Result:

If X is distributed as Np(M,Z), then any linear combination of variables a'X =  $a_1x_1 + a_2X_2 + ... + a_pX_p$  is distributed as N(M). Also, if a'X is distributed as N(M) for every a, then X must be Np(M)

du, a'za

$$P(n; \mu, \Sigma) = \frac{1}{(2\pi)^{3/2}|\Sigma|^{2}} \exp\{-\frac{1}{2}(n-\mu)^{2} \sum_{i=1}^{3} (n-\mu)^{3} \sum_{i=1}$$

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Consider the linear combination a'X of a multivariate normal random vector determined by the choice a' = [1, 0, ..., 0].

$$a'x = \begin{bmatrix} 1 & 0 & 0 & \cdots & 0 \end{bmatrix} \begin{bmatrix} x_1 \\ x_p \end{bmatrix} = x_1$$

$$a'x = \begin{bmatrix} 1 & 0 & \cdots & 0 \end{bmatrix} \begin{bmatrix} x_1 \\ x_p \end{bmatrix} = M_1$$

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$$a'x = \begin{bmatrix} 1 & 0 &$$

xi is distributed as 12(M1,611)
xi is distributed us 12(M1,611)

Example: (The distribution of two linear combinations of the components of a normal random vector) For X distributed as  $N3(u, \geq)$ . Find the distributed as

$$\begin{bmatrix} x_{1} - x_{2} \\ x_{2} - x_{3} \end{bmatrix} = \begin{bmatrix} 0 & 1 & 0 \\ 1 & 1 \\ x_{3} \end{bmatrix} \begin{bmatrix} x_{1} \\ x_{2} \\ x_{3} \end{bmatrix} = \begin{bmatrix} 0 & 1 & 1 \\ 1 & 1 \\ x_{2} \end{bmatrix} \begin{bmatrix} x_{1} \\ x_{2} \\ x_{3} \end{bmatrix} \begin{bmatrix} x_{1} \\ x_{3} \\ x_{3} \end{bmatrix} \begin{bmatrix} x_{1} \\ x_{2} \\ x_{3} \end{bmatrix} \begin{bmatrix} x_{1} \\ x_{1} \\ x_{2} \end{bmatrix} \begin{bmatrix} x_{1} \\ x_{3} \\ x_{3} \end{bmatrix} \begin{bmatrix} x_{1} \\ x_{1} \\ x_{2} \end{bmatrix} \begin{bmatrix} x_{1} \\ x_{3} \\ x_{3} \end{bmatrix} \begin{bmatrix} x_{1} \\ x_{1} \\ x_{2} \end{bmatrix} \begin{bmatrix} x_{1} \\ x_{1} \end{bmatrix} \begin{bmatrix} x_{1} \\ x_{1} \\ x_{2} \end{bmatrix} \begin{bmatrix} x_$$

# Result:

a. If  $X_1$  and  $X_2$  are independent, then  $Cov(x_1, x_2) = 0$ , a  $q_1 \times q_2$  matrix of zeros.

$$61 - 2612 + 621 - 622 - 622 - 613$$

$$612 + 623 - 622 - 2623 + 633$$

$$612 + 623 - 623 - 623 - 633$$

# Example:

(The equivalence of zero covariance and independence for normal variables)

Let X be N3(1,2) with

$$\begin{bmatrix}
 2 - \begin{bmatrix}
 4 & 1 & 0 \\
 1 & 3 & 0 \\
 0 & 0 & 1
 \end{bmatrix}
 =
 \begin{bmatrix}
 6_{11} & \delta_{11} & \delta_{13} \\
 \delta_{21} & \delta_{21} & \delta_{23} \\
 \delta_{31} & \delta_{31} & \delta_{33}
 \end{bmatrix}$$

Are X1 and X2 independent? What about (X1, X2) and X3?

$$\Rightarrow$$
 x1 x x2 ave not ind. Since  $6_{11} = 6_{21} \neq 0$ 
 $x = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$ ,  $z = \begin{bmatrix} 4 & 1 & 0 \\ 1 & 3 & 0 \\ 0 & 0 & 2 \end{bmatrix}$ 
 $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} 4 & 1 & 0 \\ 1 & 3 & 0 \\ 0 & 0 & 2 \end{bmatrix}$ 
 $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$  & x  $z = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$ 

The conditional density of x,, given that x,= 1/2 for any bivariate dist is defined by

f(n/n) = (conditioned probability of x1 = f(n, n)

given mat x2:m)

f(n)

$$f(x_1|x_0) \sim N\left(u_1 + \delta_{12} \left(x_2 - u_1\right), \delta_{11} - \delta_{12}^2\right)$$

Consider a bivariate normal pope with 14=0, 4=2

611-2,622=1 X 312=0.5

1) write down the bivariate normal density

り (ハール) で (ハール)

$$f(x) = \frac{1}{(21)^{\frac{1}{2}}} \sum_{i=1}^{2} \frac{e^{ix_i}}{2} \int_{-\frac{1}{2}}^{\frac{1}{2}} \frac{1}{3} \left( \frac{1}{-\frac{5i}{2}} - \frac{5ix_1}{2} \right) \left( \frac{x_1}{x_2 - 2} \right)^{\frac{1}{2}}$$

= 
$$\frac{1}{(m)[\frac{7}{3}]} \exp \left\{ \frac{7}{3} \left( \frac{2}{24} - \sqrt{24} \left( \frac{m-2}{2} \right) + 2 \left( \frac{m-2}{2} \right)^2 \right\} \right\}$$

Result: Let 24, 1/2. 24 be mutually independent with Xj distributed at Np(11, Z) Then

y 11 = c 1 × 1 + c 2 × L + · · + C n × n

×

$$= V_2 = b_1 \times_1 + b_2 \times_3 + \cdots + b_n \times_n - \mu Np(\frac{2}{3n}b_j \mu_j, (\frac{2}{3n}b_j^{\perp})^{\Sigma})$$

are jointly multivariate normal with covariance matrix

$$\begin{bmatrix}
\begin{pmatrix}
\hat{\Sigma} & c\hat{i} \\
y & c\hat{i}
\end{pmatrix} & (b' c) \Sigma \\
(b' c) & (\hat{\Sigma} & \hat{\Sigma} & \hat{\Sigma} \\
(b' c) & (\hat{\Sigma} & \hat{\Sigma} & \hat{\Sigma} & \hat{\Sigma}
\end{bmatrix}$$

VI & V2 are independent if (b'c) 2 = 0

Let X1, X2, X3 and X4 be independent and identically distributed (3x1) random vectors with

$$M = \begin{bmatrix} 3 \\ -1 \\ 1 \end{bmatrix} \qquad \Sigma = \begin{bmatrix} 3 \\ -1 \\ 1 \end{bmatrix} \qquad 1$$

Consider the linear combinations of random vectors

Find the mean vector and covariance matrix for each linear combination of vectors and also the covariance between them.

 $\sum_{i=1}^{n} c_{i} u_{i} = (c_{1} + c_{2} + c_{3} + c_{4}) \begin{pmatrix} 3 \\ 1 \end{pmatrix} = (\frac{1}{2} + \frac{1}{2} + \frac{1}{2} + \frac{1}{2}) \begin{pmatrix} 3 \\ 1 \end{pmatrix}$  $= 2 \begin{pmatrix} 3 \\ -1 \\ 1 \end{pmatrix} = \begin{pmatrix} 2 \\ -2 \\ 2 \end{pmatrix}$  $\left(\frac{2}{5\pi}c_{j}^{2}\right)^{2} = \left(\frac{1}{2}\right)^{2} + \left(\frac{1}{2}\right)^{2} + \left(\frac{1}{2}\right)^{2} + \left(\frac{1}{2}\right)^{2} + \left(\frac{1}{2}\right)^{2} = 1. \quad 2 = \begin{bmatrix} 3 & -1 & 1 \\ -1 & 1 & 0 \\ 1 & 0 & 2 \end{bmatrix}$ 

MOLITYARIATE NORMAL DISTRIBUTION

$$x_1 + x_2 + x_3 - 3x_4$$
 Comparing it with  $b_1 + b_2 + b_3 + b_4 + b_4 + b_4 + b_4 + b_5 + b_6 +$ 

(bicitbicith 3 cs + 64(4) 2 = 0

Villy are independent.

Ex. 1) Let  $x_1, x_2, x_3, x_4$  be independent  $Np(\mu, \Sigma)$  random vectors  $\lambda_1 = L \times 1 - L \times 2 + L \times 3 - L \times 4$ 

V2 = 1 x, +1 x2-1 x3-1 x4

whether v, & V, are independent.

[]. Let 
$$X$$
 be  $N_3(\mu, \Sigma)$  with  $\mathcal{H}' = (2 - 3 \ 1)$  and  $\Sigma = \begin{bmatrix} 1 & 1 & 1 \\ 1 & 2 & 2 \\ 1 & 2 & 2 \end{bmatrix}$ 

a) Find the distribution of 3x1-2x2+X3

b) find a 2x1 vector a such that x2 and x2-a'(x1) are independent.

a) Let 
$$a = (3 -2 i)^i$$
, then  $a^i x = 3x_1 - 2x_2 + x_3$   
 $a^i x \sim N_{\rho}(a^i + a^i = a)$ 

$$a'M = (3 -2 1) \binom{2}{-3} = 13$$

$$a' \times N_{\rho}(a' + a' \times a)$$

$$a' \times = (3 - 2 + 1) {2 \choose -3 \choose 1} = 13$$

$$a' \times = (3 - 2 + 1) {1 \choose 1} = 13$$

$$a' \times = (3 - 2 + 1) {1 \choose 1} = 13$$

The dist of 
$$3x_1 - 2x_2 + x_3$$
 is  $N_3(13, 9)$ 

b) Let  $a' = (a_1 \ a_2)$  hen  $Y = x_2 - a' \binom{x_1}{x_3} = -a_1 x_1 + x_2 - a_3 x_3$ 

$$A = \begin{bmatrix} 0 & 1 & 0 \\ -a_1 & 1 & -a_3 \end{bmatrix}$$
 then  $Ax = \begin{bmatrix} x_2 \\ y \end{bmatrix} NN(AM, A' IA)$ 

$$A IA = \begin{pmatrix} 0 & 1 & 0 \\ -a_1 & 1 & -a_3 \end{pmatrix} \begin{pmatrix} 1 & 1 & 1 \\ 1 & 2 & 2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 1 & 1 \\ 0 & -a_3 \end{pmatrix}$$

$$A IA = \begin{pmatrix} 0 & 1 & 0 \\ -a_1 & 1 & -a_3 \end{pmatrix} \begin{pmatrix} 1 & 1 & 1 \\ 1 & 2 & 2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 1 & 1 \\ 0 & -a_3 \end{pmatrix}$$

$$A IA = \begin{pmatrix} 0 & 1 & 0 \\ -a_1 & 1 & -a_3 \end{pmatrix} \begin{pmatrix} 1 & 1 & 1 \\ 1 & 2 & 2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_3 \end{pmatrix}$$

$$A IA = \begin{pmatrix} 0 & 1 & 0 \\ -a_1 & 1 & -a_3 \end{pmatrix} \begin{pmatrix} 1 & 1 & 1 \\ 1 & 2 & 2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_3 \end{pmatrix}$$

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$$A IA = \begin{pmatrix} 0 & 1 & 0 \\ -a_1 & 1 & -a_2 \end{pmatrix} \begin{pmatrix} 1 & 1 & 1 \\ 0 & -a_1 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix}$$

$$A IA = \begin{pmatrix} 0 & 1 & 0 \\ -a_1 & 1 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix}$$

$$A IA = \begin{pmatrix} 0 & 1 & 0 \\ 0 & -a_1 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 & -a_2 \end{pmatrix} \begin{pmatrix} 0 & -a_1 \\ 0 &$$

since, we want to have x, & 4 ind., this implies that -a,-2a,+3=0

$$\alpha = \begin{pmatrix} 1 \\ 1 \end{pmatrix} + C \begin{pmatrix} 3 \\ 0 \end{pmatrix}$$
 for  $C \in \mathbb{R}$ .

Conditional Distribution of Bivariate Normal Distribution

Reput;If f(m,m) is the birariate normal density, then

$$f(3u(3u)) N N \left( u_1 + \sigma_{12} (3_2 - u_2), \sigma_{11} - \sigma_{12}^{\perp} \right)$$

Relation to Conditional Distribution to Regression Model

If the joint dist of  $x \in y$  is a normal dist , then  $x^{2} - 2Sxy+y^{2}-5x^{2}+y^{2}-5x^{2}+y^{2}-2Sxy+y^{2}-5x^{2}-x^{2}-2Sxy+y^{2}-5x^{2}-x^{2}-2Sxy+y^{2}-5x^{2}-x^{2}-2Sxy+y^{2}-5x^{2}-x^{2}-2Sxy+y^{2}-5x^{2}-x^{2}-2Sxy+y^{2}-5x^{2}-x^{2}-2Sxy+y^{2}-5x^{2}-x^{2}-2Sxy+y^{2}-5x^{2}-x^{2}-2Sxy+y^{2}-5x^{2}-x^{2}-2Sxy+y^{2}-5x^{2}-x^{2}-2Sxy+y^{2}-5x^{2}-x^{2}-2Sxy+y^{2}-5x^{2}-x^{2}-2Sxy+y^{2}-5x^{2}-x^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y^{2}-2Sxy+y$ 

22-23xy+42

consider the bivariage dist with PDF as

consider the bivariate distribith PDF as
$$\frac{1}{2\pi \sigma_{N} \sigma_{N} \sqrt{1-3^{2}}} = \exp\left\{-\frac{1}{2(1-3^{2})} \left[ \left(\frac{x-\mu_{N}}{\sigma_{N}}\right)^{2} - \frac{1}{2} \left(\frac{x-\mu_{N}}{\sigma_{N}}\right) \left(\frac{y-\mu_{N}}{\sigma_{N}}\right)^{2} + \frac{1}{2\pi \sigma_{N}} \left(\frac{y-\mu_{N}}{\sigma_{N}}\right)^{2} \right] \right\}$$

$$f(n|y) = \frac{1}{2n(x6y)^{1-s2}} \exp \left(\frac{1}{2(1-s^2)}\right) \left(\frac{y-ny}{6y} - \frac{3}{2n}\right)^{2} - (1-s^2)\left(\frac{x-4n^2}{6n}\right)^{2}$$

suppose that weight (165) & height (inches) of undergraduate college men have MVN with

$$M = \begin{pmatrix} 175 \\ 71 \end{pmatrix}$$
  $\chi$   $\bar{Z} = \begin{pmatrix} 550 & 40 \\ 40 & 8 \end{pmatrix}$ 

Find mean & variance of conditional dist

-> Mean = 
$$M_1 + \sigma_{12} (\pi_2 - \mu_2) = 175 + \frac{40}{8} (\pi_2 - 71) = 5\pi_2 - 180$$

Variance = 
$$6_{11} - 6_{12}^2 = 550 - 40^2 = 350$$

$$m=70$$
,  $E(Y|X)=-180+5$   $Z=7$   $B0+B,X$   $Z=170$   $Z=170$   $Z=170$   $Z=170$   $Z=170$   $Z=170$   $Z=170$   $Z=170$   $Z=170$ 

candy company makes 3 size candy barx.

X1: Regular X3: Big Size

X2: Fun size

What is the prob. that

1) x124(5,4)

The prob. that the regular bar is more than 802 is P(X1>8) = P(Z78-5)

= P(271.5) normal table

= P(271.5) veing colonletor

P(x,78) = 0.0668

111) 
$$4\pi - 3\pi + 5\pi = is normany distributed
 $a' = (4 - 3 5)$   $a' \times 2 4\pi - 3\pi \times 5\pi = \pi$   
 $a' \times 2 (4 - 3 5) (5) = 46$$$

$$a'za = (4-35)(4-10)(4)$$

$$= (4-35)(19)$$

$$= (4-35)(19)$$

$$= 289$$

$$= p(2<1)$$

$$= p(1)$$

$$= 0.8413447$$