

# Variance, Moments, Function of a r.v.

MATH/STAT 394: Probability I Summer 2021 A Term

Introduction to Probability
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## Practice solution

#### **Practice**

Consider that the time to wait your bus is modeled by  $X \sim \text{Exp}(\lambda)$ .

- 1. What is the 90/100th quantile for  $\lambda = 1$ ?
- 2. Let's say that the 90/100th quantile is 20min ( $\lambda$  is unknown). Give an upper bound on the prob. that you wait more 30min

## **Solution**

- 1. Recall that the c.d.f. of  $X \sim \text{Exp}(\lambda)$  is  $F(t) = 1 e^{-\lambda t}$ 
  - To find the pth quantile, it suffices to inverse F.
  - Indeed, denote  $F^{-1}$  the inverse of F, then

$$\mathbb{P}(X \le F^{-1}(p)) = F(F^{-1}(p)) = p$$

So  $F^{-1}$  gives us the quantile. Here  $p=1-e^{-\lambda t}\iff t=-\frac{\log(1-p)}{\lambda}$ .

- Therefore  $x_{90/100} \approx 2.30$  for  $\lambda = 1$
- 2. If we know that the 90/100th quantile is 20min, then

$$\mathbb{P}(X \ge 30) \le \mathbb{P}(X \ge 20) = 1 - 90/100 = 10/100$$

# Recap

## Median Quantiles

The median of a r.v. X is a value m s.t.

$$\mathbb{P}(X \geq m) \geq 1/2$$
  $\mathbb{P}(X \leq m) \geq 1/2$ 

Function of a discrete r.v.

• The pth quantile of a r.v. X is a value  $x_p$  s.t.

$$\mathbb{P}(X \le x_p) \ge p$$
  $\mathbb{P}(X \ge x_p) \ge 1 - p$ 

### Variance

Variance

$$Var(X) = \mathbb{E}\left((X - \mathbb{E}[X])^2\right) = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

- Standard deviation:  $\sqrt{Var(X)}$
- for  $X \sim \text{Ber}(p)$ , Var(X) = p(1-p)

Variance

Variance

Moments

Function of a discrete r.v.

Variance computations

# Variance of a uniform dist.

### **Exercise**

Let  $X \sim \text{Unif}([a, b])$  with a < b, what is Var(X)?

#### Solution

- Use the formulation  $Var[X] = \mathbb{E}[X^2] \mathbb{E}[X]^2$  (often more easy to compute)
- ullet We have that  $\mathbb{E}[X]=rac{a+b}{2}$  and

$$\mathbb{E}[X^2] = \int_{-\infty}^{+\infty} x^2 f(x) dx = \frac{1}{b-a} \int_a^b x^2 dx = \frac{b^3 - a^3}{3(b-a)} = \frac{1}{3} (b^2 + ba + a^2)$$

Function of a discrete r.v.

Therefore

$$Var[X] = \mathbb{E}[X^2] - \mathbb{E}[X]^2 = \frac{1}{3}(b^2 + ba + a^2) - \frac{(a+b)^2}{4} = \frac{(b-a)^2}{12}$$

• Note that as b-a increases, the variance naturally increases

## Other variances

Variance

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### Lemma

If 
$$X \sim \text{Bin}(n, p)$$
,  $\text{Var}(X) = np(1 - p)$ 

Note that it is simply n times the variance of a Bernoulli. 1

#### Proof Idea:

- 1. Use the formula  $Var[X] = \mathbb{E}[X^2] \mathbb{E}[X]^2$ .  $\mathbb{E}[X]$  is known so it remains to compute  $\mathbb{E}[X^2]$
- 2. Formulate the computation of  $\mathbb{E}[X^2]$  from the p.m.f.
- 3. Make appropriate change of var.
- 4. Use binomial theorem to simplify some parts

Full proof given in supp. slides

Try by yourself before looking at the solution!

<sup>&</sup>lt;sup>1</sup>This can be properly justified (see MATH/STAT395)

## Other variances

Variance

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#### Lemma

If 
$$X \sim \text{Geom}(p)$$
,  $\text{Var}(X) = (1-p)/p^2$ 

#### Proof Idea:

- 1. Use the formula  $Var[X] = \mathbb{E}[X^2] \mathbb{E}[X]^2$ .  $\mathbb{E}[X]$  is known so it remains to compute  $\mathbb{E}[X^2]$
- 2. Try to compute  $\mathbb{E}[X(X-1)]$ , the variance is then given by using that  $\mathbb{E}[X^2] = \mathbb{E}[X(X-1)] + \mathbb{E}[X]$
- 3. Formulate  $\mathbb{E}[X(X-1)]$  as computing the second derivative of  $h: t \to \sum_{t=0}^{+\infty} t^k$ for |t| < 1
- 4. Use that h(t) = 1/1 t to deduce the value of the second derivative, and therefore the variance

Full proof given in supp. slides

Try by yourself before looking at the solution!

Variance

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### Lemma

If 
$$X \sim \text{Exp}(\lambda)$$
,  $\text{Var}(X) = 1/\lambda^2$ 

### Proof Idea:

- 1. Use the formula  $Var[X] = \mathbb{E}[X^2] \mathbb{E}[X]^2$ .  $\mathbb{E}[X]$  is known so it remains to compute  $\mathbb{E}[X^2]$
- 2. Compute  $\mathbb{E}[X^2]$  by integration by part

Full proof given in supp. slides

Try by yourself before looking at the solution!

# Other variances

### Classical variances

- for  $X \sim \text{Ber}(p)$ , Var(X) = p(1-p)
- for  $X \sim \text{Bin}(n, p)$ , Var(X) = np(1 p)
- for  $X \sim \text{Geom}(p)$ ,  $\text{Var}(X) = (1-p)/p^2$
- for  $X \sim \text{Unif}([a, b])$ ,  $\text{Var}(X) = (b a)^2/12$
- for  $X \sim \text{Exp}(\lambda)$ ,  $\text{Var}(X) = 1/\lambda^2$

### Note

 You should be able to easily compute variances of random variables from the definition of the variance and the p.m.f./p.d.f.

Function of a discrete r v

- These are mostly calculus exercises
- Beyond these exercises, the variance is a key property that will allow us to charc. r.v.

# Variance properties

### Motivation

• The variance is **not linear!** Instead we have the following prop.

### Lemma

For a r.v. X and  $a, b \in \mathbb{R}$ .

$$Var(aX + b) = a^2 Var(X)$$

### Proof

$$Var(aX + b) = \mathbb{E}[(aX + b - \mathbb{E}[aX + b])^{2}]$$

$$= \mathbb{E}[(aX + b - a\mathbb{E}[X] - b)^{2}] = \mathbb{E}[a^{2}(X - \mathbb{E}[X])] = a^{2} Var(X)$$

## Takeaway:

- Adding a constant to the r.v. does not change the variance
- Multiplying a r.v. by a constant a you get a standard deviation multiplied by a
- Remember the standard deviation is  $\sqrt{Var}(X)$
- Note that the standard deviation has the same 'unit' as the r.v. or the mean, while the variance has the squared of this unit

## **Null variance**

Variance

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### Motivation

• The following theorem formalizes the intuition that if a r.v. does not vary (i.e. Var(X) = 0) then it must be a constant

### Theorem

For a r.v. X,  $\mathsf{Var}(X) = 0$  if and only if  $\mathbb{P}(X = \mathsf{a}) = 1$  for some constant  $\mathsf{a} \in \mathbb{R}$ 

### Proof

- Clearly if  $\mathbb{P}(X = a) = 1$  then  $\mathbb{E}[X] = a$  and Var(X) = 0
- On the other hand, we consider simply the discrete case, then

$$0 = \operatorname{Var}(X) = \sum_{k} (k - \mu)^2 \mathbb{P}(X = k)$$

where  $\mu = \mathbb{E}[X]$ 

• This is only possible if all the terms are zero, that is  $(k-\mu)^2\mathbb{P}(X=k)=0$ , which is equivalent to

$$k = \mu$$
 or  $\mathbb{P}(X = k) = 0$ 

• Thus the only value k of X with  $\mathbb{P}(X=k)>0$  is  $k=\mu$  and hence  $\mathbb{P}(X=\mu)=1$ .

# Outline

Variance

Moments

Function of a discrete r.v.

Variance computations

## **Moments**

#### Motivation

- The expectation gives a first summary
- Taking the square of the r.v, we get the variance, and more interesting information
- Can we go on like that and define key properties of the r.v. as

$$\mathbb{E}[X^k]$$
 or  $\mathbb{E}[(X-\mu)^k]$ 

• Yes, and it is a great idea, explored in more details in MATH/STAT395

### **Moments**

### Definition

The **nth moment** of a r.v. X is

$$\mathbb{E}[X^n]$$

Function of a discrete r v

The **nth centered moment** of a r.v. X is

$$\mathbb{E}[(X - \mathbb{E}[X])^n]$$

## Notes

- The first moment is the mean
- The second moment is the squared mean
- The second centered moment is the variance
- The third centered moment is called the skewness It informs us about the asymmetry of the r.v. For example, if a r.v. is symmetric the skewness is 0
- The fourth centered moment is called the kurtosis It is a measure 'tailedness' or how 'heavy' (likely) the tails of the distribution are

Function of a discrete r.v.

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Variance

Variance

Moments

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Variance computations

## Distribution of a function of a r.v.

### Motivation

Variance

- As already observed, numerous r.v. of interest may be expressed as Y = g(X) for X a classical r.v.
- We have seen how to compute  $\mathbb{E}[Y]$ , i.e.,
  - $\mathbb{E}[g(X)] = \sum_{k \in \mathcal{X}} g(k) \mathbb{P}(X = k)$  for X discrete
  - $\mathbb{E}[g(X)] = \int_{-\infty}^{+\infty} g(x)f(x)dx$  for X continuous
- Now this may not be sufficient (just having the mean may not fully characterize your dist.)
- Here we are going to see how to compute the p.m.f./p.d.f. of a function of a r.v.

## Inverse of a function

## Motivation

• The main idea is to be able to map back values of Y onto values of X

Function of a discrete r v

Recall then the definition of an inverse

## Definition

Let E, F be two sets. A function  $g: E \to F$  is invertible if

for any 
$$y \in F$$
, there exists a unique  $x \in E$  s.t.  $y = g(x)$ 

We denote the inverse of g as  $g^{-1}$  which satisfies for any  $x \in E$ ,  $y \in F$ 

$$g^{-1}(g(x)) = x$$
  $g(g^{-1}(y)) = y$ 

## Example

- The function exp has inverse log
- The function  $x^2$  is not invertible on  $\mathbb{R}$
- Any strictly increasing/decreasing function  $g: \mathbb{R} \to \mathbb{R}$  is invertible

# Distribution of a function of a r.v.

# Simple case

- Consider g invertible and X a r.v. taking values in  $\mathcal{X}$
- The r.v. Y = g(X) takes values in  $\{g(k) : k \in \mathcal{X}\}$
- The p.m.f. of Y is then given by

$$\mathbb{P}(Y = g(k)) = \mathbb{P}(g^{-1}(Y) = k) = \mathbb{P}(X = k)$$

Function of a discrete r.v.

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Now what if g is not invertible?

## Practice next lecture

Variance

### **Practice**

Let X be a discrete r.v. s.t.

$$\mathbb{P}(X = k) = \frac{\lambda^k}{k!} e^{-\lambda} \quad \text{for } k \in \{0, 1, \ldots\}$$

- 1. Check that  $\sum_{k=0}^{+\infty} \mathbb{P}(X=k) = 1$
- 2. Compute  $\mathbb{E}[X]$ , Var[X] for  $\lambda = 1$

Reminder: The Taylor series expansion of exp is  $e^x = \sum_{n=0}^{+\infty} \frac{x^k}{k!}$  for any  $x \in \mathbb{R}$ .

Variance

Variance

Moments

Function of a discrete r.v.

Variance computations

# Variance of a geometric r.v.

### Exercise

Consider  $X \sim \text{Geom}(p)$ . What is Var(X)?

## Solution

• Idea: Use that  $\frac{d}{d^2t}\sum_{k=0}^{+\infty}t^k=\frac{d}{d^2t}\frac{1}{1-t}$ , which gives for any 0< p<1

$$\sum_{k=2}^{+\infty} k(k-1)t^{k-2} = \frac{2}{(1-t)^3}$$

· Here we have that

$$\mathbb{E}[X^2] = \mathbb{E}[X] + \mathbb{E}[X(X-1)]$$

and denoting q = 1 - p

$$\mathbb{E}[X(X-1)] = \sum_{k=2}^{+\infty} k(k-1)q^{k-1}p = pq\sum_{k=2}^{+\infty} k(k-1)q^{k-2} = \frac{2pq}{(1-q)^3} = \frac{2p(1-p)}{p^3}$$

• Therefore

$$Var[X] = \mathbb{E}[X^2] - \mathbb{E}[X]^2 = \frac{1}{p} + \frac{2p(1-p)}{p^3} - \frac{1}{p^2} = \frac{1-p}{p^2}$$

# Variance of a binomial r.v.

# Exercise

Let  $X \sim \text{Bin}(n, p)$ , compute Var(X).

# Solution

$$\mathbb{E}[X^2] = \sum_{k=1}^n k^2 \frac{n!}{k!(n-k)!} p^k (1-p)^{n-k}$$

$$= \sum_{k=1}^n k \frac{n!}{(k-1)!(n-k)!} p^k (1-p)^{n-k}$$

$$= \sum_{k=1}^n (k-1) \frac{n!}{(k-1)!(n-k)!} p^k (1-p)^{n-k} + \sum_{k=1}^n k \frac{n!}{(k-1)!(n-k)!} p^k (1-p)^{n-k}$$

$$= \sum_{k=1}^n \frac{n!}{(k-2)!(n-k)!} p^k (1-p)^{n-k} + \sum_{k=1}^n k \frac{n!}{(k-1)!(n-k)!} p^k (1-p)^{n-k}$$

$$= n(n-1)p^2 + np$$

For the last line, for the second term you recognize the same term as the one computed for the expectation. For the first term, make a change of variable k-2-j, then factorizes  $p^2$  and n(n-1) and use the binomial theorem

# Variance of an exponential r.v.

# **Exercise**

Let  $X \sim \mathsf{Exp}(\lambda)$ , compute  $\mathsf{Var}(X)$ 

## Solution

• We have (recall that  $\lambda > 0$  always)

$$\mathbb{E}[X^2] = \int_0^{+\infty} x^2 \lambda e^{-\lambda x} dx$$

$$= \left[ -x^2 e^{-\lambda x} \right]_0^{+\infty} + 2 \int_0^{+\infty} x e^{-\lambda x} dx$$

$$= 0 + 2 \left[ -\frac{x}{\lambda} e^{-\lambda x} \right]_0^{+\infty} + 2 \int_0^{+\infty} \frac{1}{\lambda} e^{-\lambda x} dx$$

$$= \frac{2}{\lambda^2}$$

So

$$Var(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2 = \frac{2}{\lambda^2} - \frac{1}{\lambda^2} = \frac{1}{\lambda^2}$$