

Simulation Study of Automated Thickness Analyzing Machine (ATAM)

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Background and Motivation

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- 1 **Several diseases have been found to be related to white matter loss in brain, such as Autism (Vidal 2006) and ADHD (Luders 2009). Here, we measure white matter loss based the thickness of the mid-sagittal slice of the corpus callosum (CC), the largest white matter structure in brain.**

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- 2 **Our objective: To analyze the link between white matter thickness along the CC, and Expanded Disability Status Scale (EDSS) score in patients with multiple sclerosis (MS). Specifically, if there are regions of the CC that are predictive of EDSS score, we want to find out where they are.**

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- ③ **Today we'll present a simulation study on our proposed method.**

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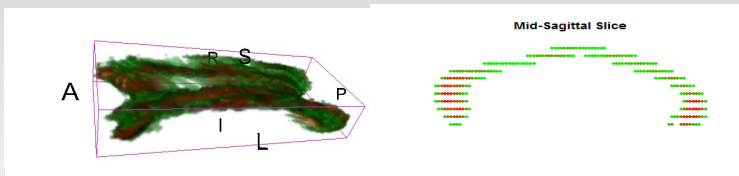
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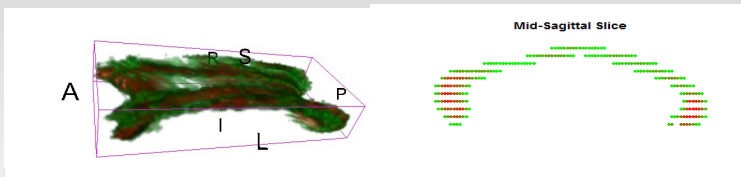
Figure: CC 3D rendering and mid-sagittal slice



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Color indicates corresponding Fractional Anisotropy (FA) value.

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- ④ **Hypothesis testing.**

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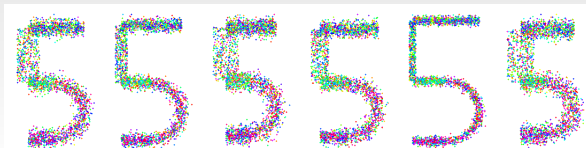
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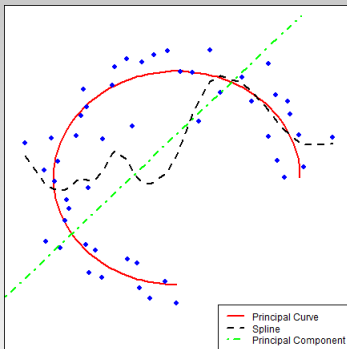
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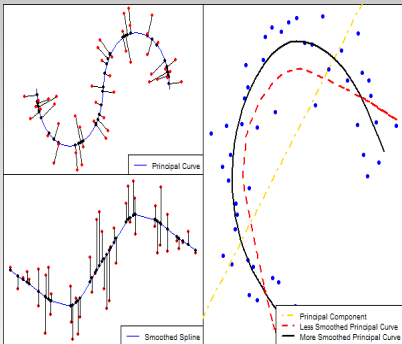
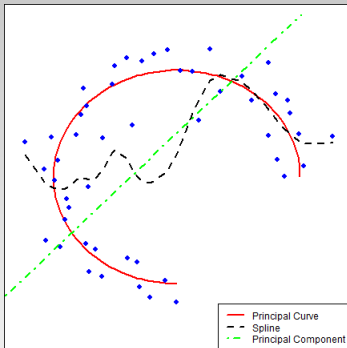
- **Case 2:** $I=300$, response $Y_i \sim \text{Poisson}(\lambda = 3)$. Thickness of image i in the vertical bar is related to Y_i .



Principal curve



Principal curve



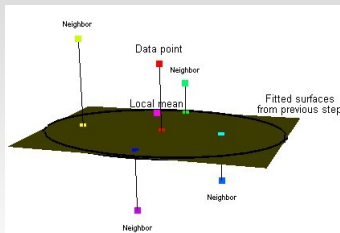
Principal Curve Algorithm

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- 1 The Projection Step. The points are projected onto the curve of the previous iteration.

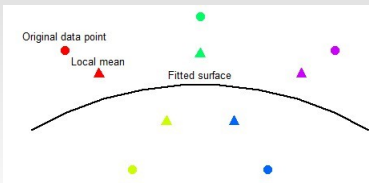
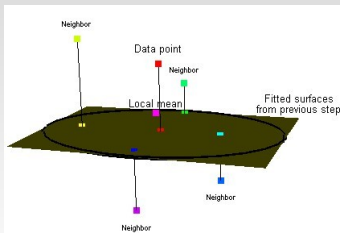
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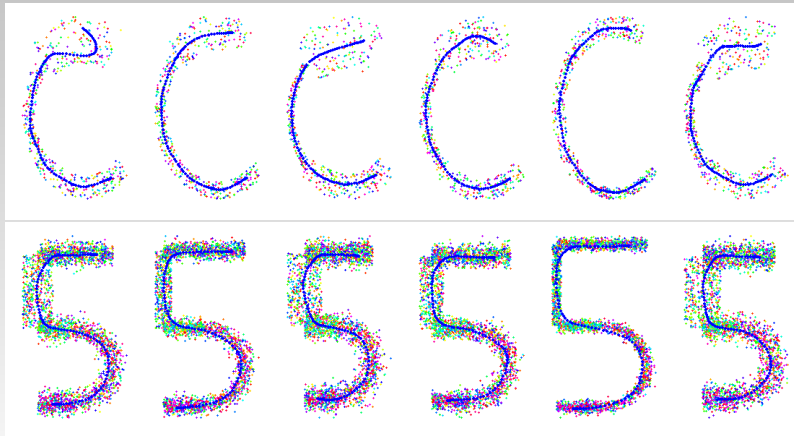


Principal Curve Algorithm

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- ② **The Conditional Expectation Step.** For each data point $\mathbf{x}^{(i)}$, we calculate a locally average $\bar{\mathbf{x}}^{(i)}$.
- ③ **The Smoothing Step.** Fitting a fast TPS using all the local average data point $\bar{\mathbf{x}}^{(i)}$, obtain f^{new} .



Principal curve fitting result



Obtaining Thickness along the principal curve

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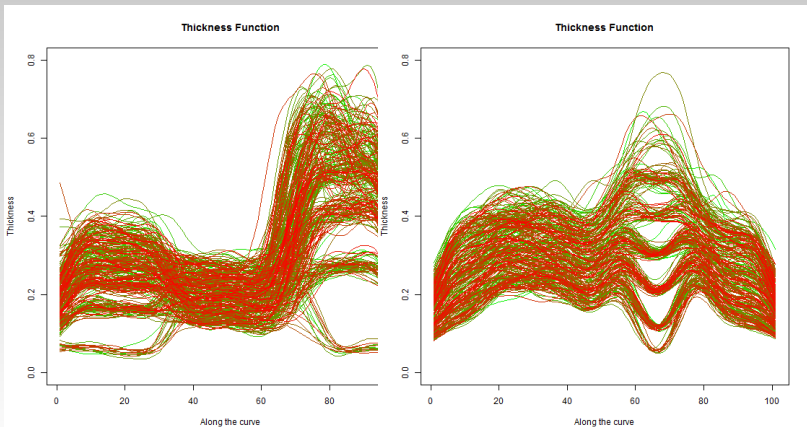
- $\mathbf{Thick}^*(t) = \mathbf{Quantile}(\{2 \times \mathbf{dist}_{t_j}, |t_j - t| \leq .02\}, 0.95).$

Obtaining Thickness along the principal curve

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Regression Notation

Y_i = Outcome (simulated EDSS score)

$X_i(t)$ = The thickness function along the track of the structure

$\beta(t)$ = Function representing effect of thickness on outcome, at different points.

$t \in [0, 1]$.

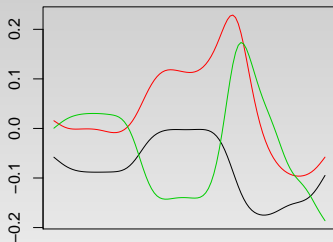
$$Y_i \sim \text{Poisson}(\theta_i)$$

$$\theta_i = \beta_0 + \int_0^1 X_i(t)\beta(t)dt$$

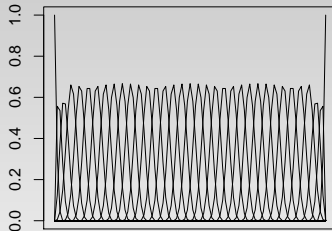
Regression Notation

Estimate the functions $X_i(t)$ and $\beta(t)$ using basis functions.
For X , we'll use it's principle components. For β , we'll use a b-spline basis.

Principle Components for X



b-spline basis for β



$$X_i(t) \approx \sum_{k=1}^{K_x} \xi_{ik} \Psi_k(t)$$

$$\beta(t) \approx \sum_{l=1}^{K_\beta} \beta_l \varphi_l(t)$$

Regression Notation

$$\begin{aligned}\int_0^1 X_i(t)\beta(t)dt &\approx \int_0^1 \left(\sum_{k \in \{A\}} \xi_{ik} \Psi_k(t) \right) \left(\sum_{j \in \{B\}} \beta_j \varphi_j(t) \right) dt \\ &= \int_0^1 \left(\sum_{(k,j) \in \{A\} \times \{B\}} \xi_{ik} \Psi_k(t) \varphi_j(t) \beta_j \right) dt \\ &= \sum_{(k,j) \in \{A\} \times \{B\}} \xi_{ik} \left(\int_0^1 \Psi_k(t) \varphi_j(t) dt \right) \beta_j \\ &=: \xi_i \mathbf{J} \beta^T\end{aligned}$$

- Where $\xi_i = [\xi_{i1}, \dots, \xi_{iK_X}]^T$, $\beta = [\beta_1, \dots, \beta_{K_\beta}]^T$, and \mathbf{J} is a $K_x \times K_\beta$ matrix with $(i, j)^{th}$ element equal to $\int_0^1 \Psi_i(t) \varphi_j(t) dt$.

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- For this project we'll assume $\xi_i \mathbf{J}$ is fixed, and will put a penalty prior on β .

Full Model

$$Y_i \sim \text{Poisson}(\theta_i)$$

$$\theta_i = \beta_0 + \xi_i \mathbf{J} \beta^T$$

$$\beta_0, \beta_1 \sim N(0, 100)$$

$$\beta_j \sim N(\beta_{j-1}, 1/\tau_\beta)$$

$$\tau_\beta \sim \Gamma(.001, .001)$$

¹(Crainiceanu and Goldsmith 2010; Brezger, Kneib, and Lang 2005; Lang and Brezger 2004; Goldsmith, Feder, Crainiceanu, Caffo, and Reich, 2010)

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This could be made to account for measurement error by letting:

$$W_i(t) = \sum_{k=1}^{K_x} \xi_{ik} \Psi_k(t) + \epsilon_i(t), \text{ and assigning priors for } \epsilon_i(t) \text{ and } \xi_{ik}.$$

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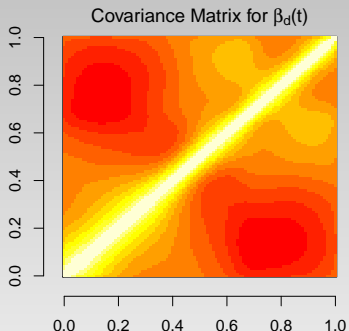
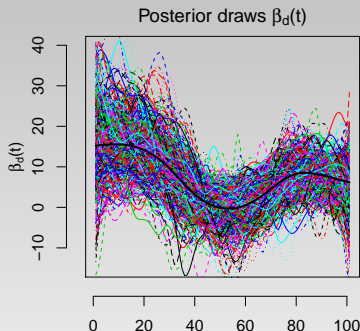
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Fit this in WinBUGS. All parameters mixed well except τ_β .¹

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Hypothesis Testing

Given draws ($\beta_d(t)$) from the posterior:

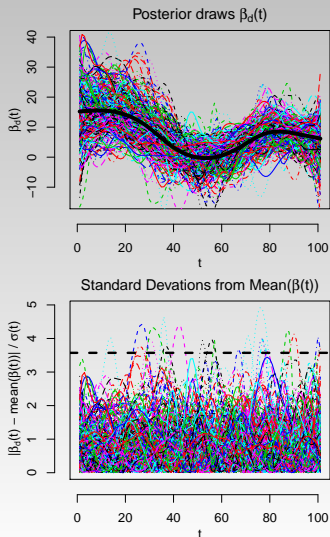


- Pointwise credible intervals give the range along the curve where 95% of the points will fall
 - To estimate this, we need the quantiles for $\beta_d(t)$.
- Joint credible intervals give the area within which 95% of the curves will be completely contained
 - Roughly speaking, we need the quantiles for each curve's largest deviation from the mean curve.

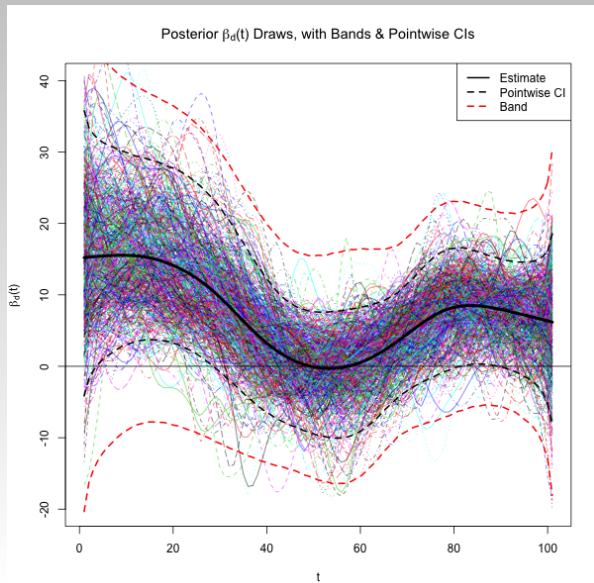
Hypothesis Testing

Let:

- $\hat{\beta}(t) = \text{mean}(\beta_d(t))$
- $\sigma(t)^2 = \text{var}(\beta_d(t))$
- $S_d(t) = \frac{|\beta_d(t) - \hat{\beta}(t)|}{\sigma(t)}$
 - Assume $P(S_d(t) > s)$ is constant over t , given s
- $m_d = \max_t \{S_d(t)\}$
- $q(m_d, .95)$ is the 95% quantile of m_d across d .
- Reconstruct joint CIs using $\hat{\beta}(t) \pm q(m_d, .95)\sigma(t)$.



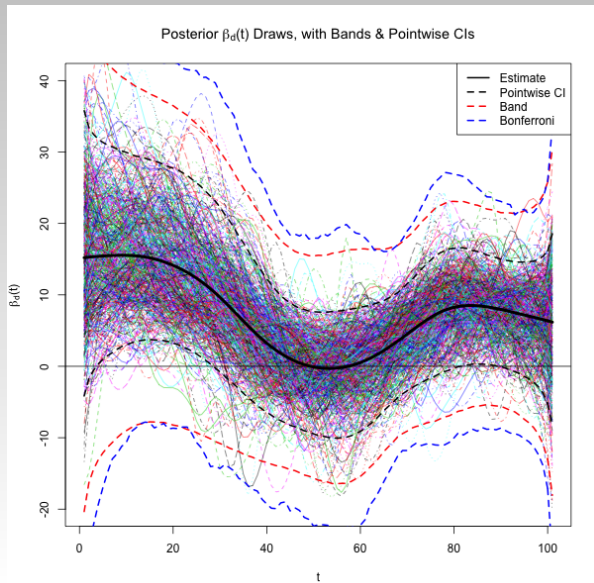
Pointwise & Joint Credible Intervals - C Shape



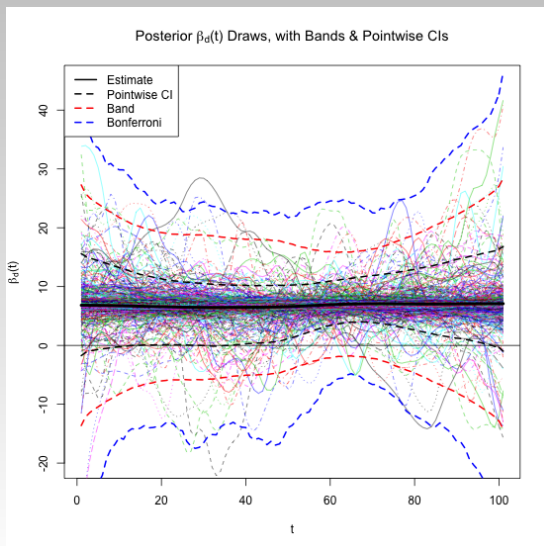
Comparison to a Bonferroni

- Pointwise intervals above are calculated above using the $(.5\alpha)$ & $(1-.5\alpha)$ quantiles at each point t (with $\alpha = .05$).
- We could also create a confidence band using a Bonferroni style adjustment, using the $(.5\alpha/100)$ and $(1 - \frac{.5\alpha}{100})$

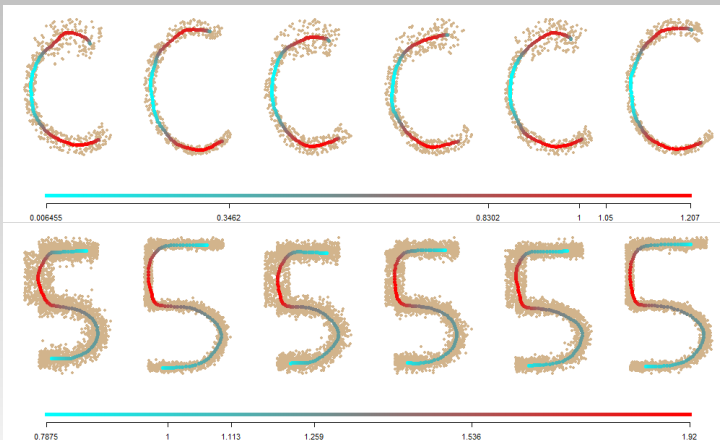
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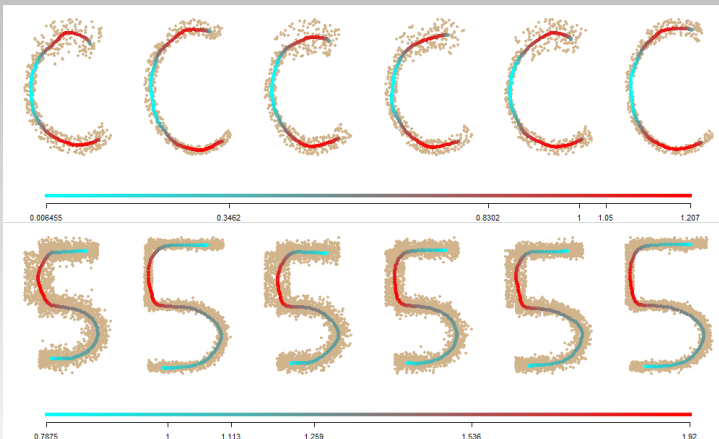
Pointwise & Joint Credible Intervals - 5 Shape



Back Mapping



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- Color indicates value $\frac{\hat{\beta}(t)}{\frac{1}{2} \text{Width}_{CI}(t)}$

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- Results are slightly sensitive to number of splines chosen.
- Pipeline is good,
- **but hard.**
 - **Make sure one step are not throwing garbage into the next one.**

Acknowledgement

*Thanks Dr. Crainiceanu, C., Gellar, J and Huang, L. for their help.
Thanks for everyone's patience!*