Computational Finance



Monte Carlo Methods

Brownian Motion

• We saw last week that the binomial tree implies for $X_t \equiv \log S_t$ that

$$X_{i\delta t} = X_{(i-1)\delta t} + R_i \iff \Delta X_i = R_i, \tag{\dagger}$$

where $R_i = \log u$ or $R_i = \log d$, with probabilities $\mathbb{Q}[u]$ and $\mathbb{Q}[d]$.

- Equation (†) is a stochastic difference equation.
- Its solution

$$X_T = \log S_0 + \sum_{i=1}^{N} R_i = \log S_0 + \sigma \sqrt{\delta t} (2k - N)$$

is called a binomial process, or in the special case with $\mathbb{E}[R_i] = 0$, a random walk.

• We also saw that if we let $N \to \infty$ (so that $\delta t \to 0$),

$$X_T - X_0 \stackrel{d}{\to} N(\mu T, \sigma^2 T), \quad \mu \equiv r - \frac{1}{2}\sigma^2.$$

• The argument can be repeated for every X_t , $t \leq T$, showing that

$$X_t - X_0 \stackrel{d}{\to} N(\mu t, \sigma^2 t),$$

and that for any $0 < t < T, X_t - X_0$ and $X_T - X_t$ are independent.

- As $\delta t \to 0$, $\{X_t\}_{t\geq 0}$ becomes a continuous time process: the indexing set is now given by the entire positive real line.
- This continuous time limit (with $\mu=0$ and $\sigma^2=1$) is called *Brownian motion*, or *Wiener process*.
- From now on, rather than modelling in discrete time and then letting $\delta t \to 0$, we will directly model in continuous time, using Brownian motion as a building block.

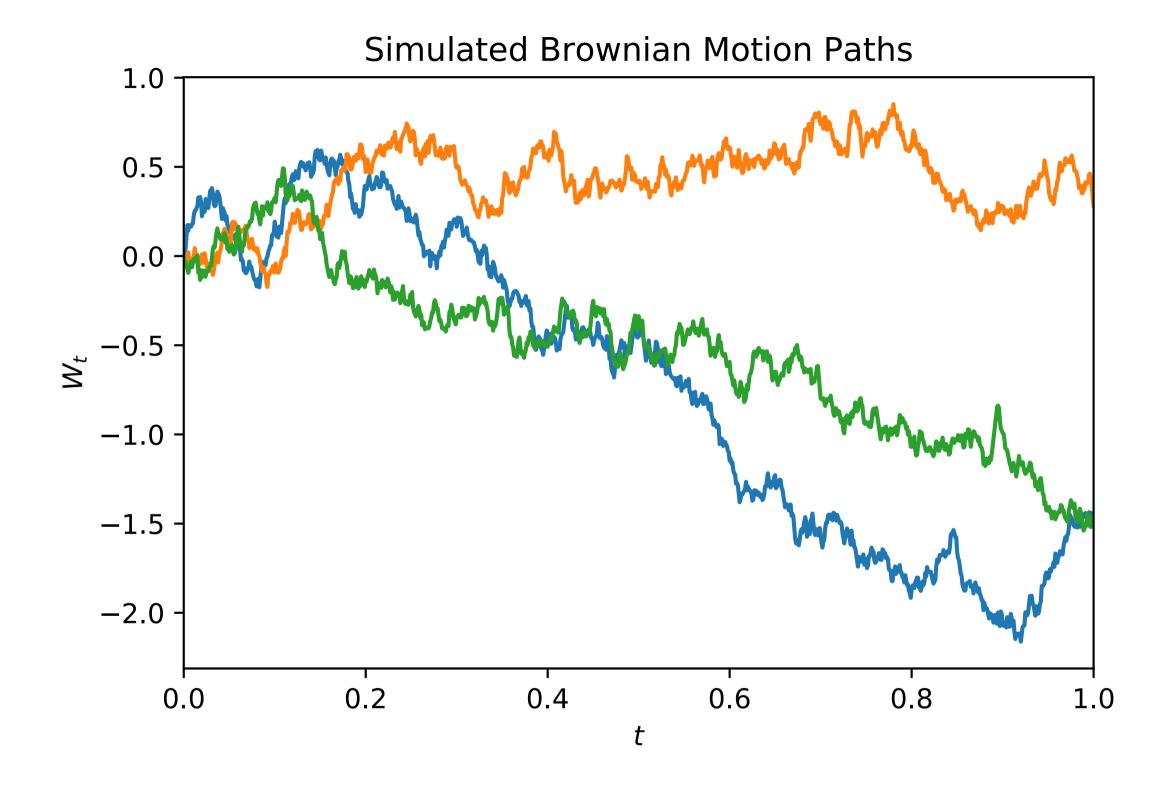
- ullet Definition of (standard) Brownian Motion: Stochastic process $\{W_t\}_{t\geq 0}$ satisfying
 - $W_0 = 0$;
 - The increments $W_t W_s$ are independent for all $0 \le s < t$;
 - $W_t W_s \sim N(0, t s)$ for all $0 \le s \le t$;
 - Continuous sample paths.
- This is standard Brownian motion, whereas $X_t = \sigma W_t$ is Brownian motion with variance σ^2 .
- Restriction that process start at zero may be loosened by considering $X_t = X_0 + \sigma W_t$.
- Brownian motion with drift: $X_t = X_0 + \mu t + \sigma W_t$, so that $\mathbb{E}[X_t] = X_0 + \mu t$, $\mathrm{Var}[X_t] = \sigma^2 t$.

- Properties of Brownian Sample Paths:
 - Continuity: by assumption, and also $W_{t+\delta t} W_t \sim N(0, \delta t) \rightarrow 0$ as $\delta t \downarrow 0$;
 - Nowhere differentiability: intuitively, this is seen from

$$\frac{W_t - W_{t-\delta t}}{\delta t} \sim N\left(0, \frac{1}{\delta t}\right), \quad \frac{W_{t+\delta t} - W_t}{\delta t} \sim N\left(0, \frac{1}{\delta t}\right);$$

left and right difference quotients do not have (common) limit as $\delta t \downarrow 0$.

• Self-similarity: Zooming in on a Brownian motion yields another Brownian motion: for any $c>0, X_t=\sqrt{c}W_{t/c}$ is a Brownian motion.

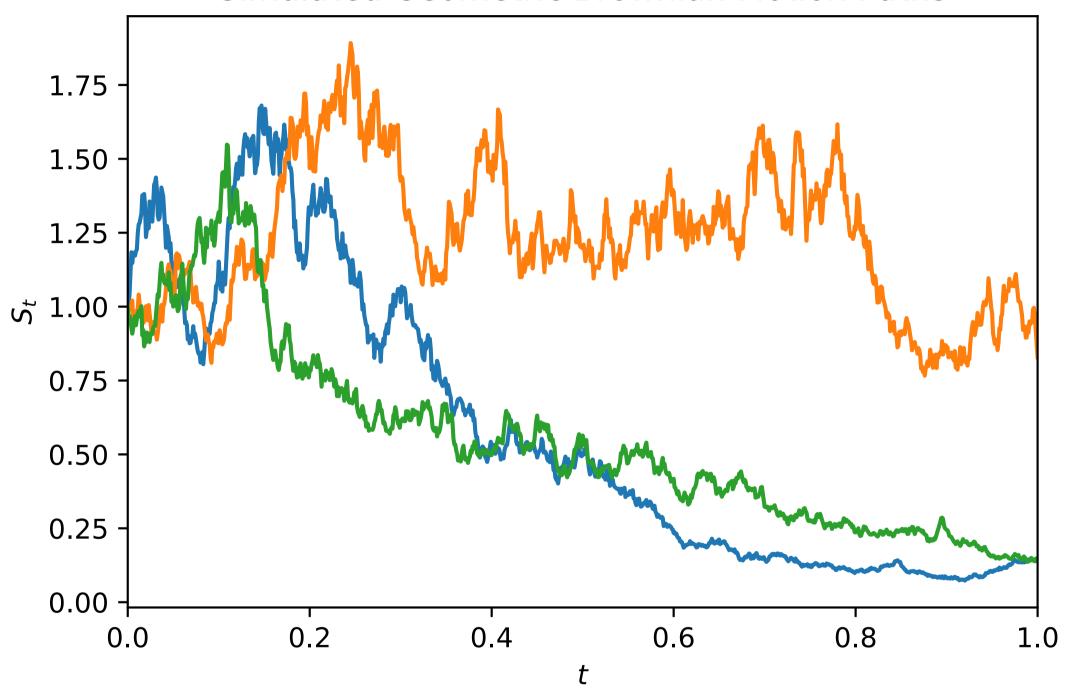


• Brownian motion itself is not a very useful model for stock prices, because it can become negative. Instead we model $X_t \equiv \log S_t$ as a Brownian motion with drift:

$$X_t = X_0 + \mu t + \sigma W_t$$
, so that
 $S_t = \exp(X_t)$
 $= S_0 \exp(\mu t + \sigma W_t)$.

- The resulting process for S_t is called Geometric Brownian motion (GBM).
- This implies that the log return $\log S_t \log S_s = X_t X_s$, s < t, is independent of X_s , with constant variance for fixed (t s).

Simulated Geometric Brownian Motion Paths



Continuous Time Martingales

- In continuous time, a process $\{X_t\}_{t\geq 0}$ is a martingale if
 - $\mathbb{E}[|X_t|] < \infty$, for all $t \ge 0$;
 - $\mathbb{E}[X_t | \mathcal{F}_s] = X_s$, for all $t > s \ge 0$, where \mathcal{F}_t denotes the information on X_t up to time t.
- E.g., for Brownian motion
 - $\mathbb{E}[|W_t|] < \infty$ because $W_t \sim N(0, t)$;
 - $\mathbb{E}[W_t | \mathcal{F}_s] = \mathbb{E}[W_s + (W_t W_s) | \mathcal{F}_s] = W_s + 0$ because of independent increments.

• For Geometric Brownian motion, $X_t = X_0 \exp(\mu t + \sigma W_t)$, so that $X_t = X_S \exp(\mu (t-s) + \sigma (W_t - W_s))$. Thus

$$\mathbb{E}[X_t | \mathcal{F}_s] = \mathbb{E}\left[X_s \exp(\mu(t-s) + \sigma(W_t - W_s)) | \mathcal{F}_s\right]$$

$$= X_s \exp(\mu(t-s)) \mathbb{E}\left[\exp(\sigma(W_t - W_s))\right]$$

$$= X_s \exp(\mu(t-s)) \exp\left(\frac{1}{2}\sigma^2(t-s)\right).$$

- The last line above follows because $\mathbb{E}[\exp(z)] = \mu + \frac{1}{2}\sigma^2$ if $z \sim N(\mu, \sigma^2)$. The distribution of $\exp(z)$ is called the *lognormal*.
- Hence GBM is a martingale if and only if $\mu = -\frac{1}{2}\sigma^2$.

Ito Processes

- Ito processes generalize Brownian motion with drift by allowing the drift and volatility to be time-varying and potentially stochastic.
- The trick is to describe the dynamics of a process with a *stochastic differential equation* (SDE), the continuous time equivalent of a stochastic difference equation.
- Take, for example, Brownian motion with drift, $X_t = X_0 + \mu t + \sigma W_t$.
- We know from calculus that

$$\int_{\tau}^{t} \mu ds = \mu \int_{\tau}^{t} ds = \mu(t - \tau).$$

• If we define $\int_{\tau}^{t}dW_{\scriptscriptstyle S}=W_{t}-W_{\tau}$, then we see that

$$X_t = X_0 + \int_0^t \mu ds + \int_0^t \sigma dW_s.$$

This is often written in differential form as

$$dX_t = \mu dt + \sigma dW_t.$$

Note that this is just short hand notation for the integral form.

• An Ito process generalizes this by allowing μ and σ to be time-varying and stochastic:

$$dX_t = \mu_t dt + \sigma_t dW_t. \tag{\dagger}$$

Again, this is just short-hand for

$$X_t = X_0 + \int_0^t \mu_s ds + \int_0^t \sigma_s dW_s,$$

where we define

$$\int_0^T \mu_s ds \equiv \lim_{n \to \infty} \sum_{i=0}^{N-1} \mu(t_i) \Delta t_{i+1}, \qquad \int_0^T \sigma(t) dW_t \equiv \lim_{n \to \infty} \sum_{i=0}^{N-1} \sigma(t_i) \Delta W_{t_{i+1}},$$

$$t_i \equiv iT/N, \Delta t_{i+1} \equiv t_{i+1} - t_i, \text{ and } \Delta W_{t_{i+1}} \equiv [W_{t_{i+1}} - W_{t_i}].$$

• Remarks:

- X_t is the sum of two integrals. The first is called a Riemann integral, the second is an Ito integral.
- **Do not** think of the integrals as an *area under the curve* like in high school. Your intuition for the Ito integral should be that we are summing infinitesimally small Brownian increments dW_t , each scaled by the instantaneous volatility σ_t .
- If μ_t and σ_t depend only on the current W_t , then (†) is called a stochastic differential equation. Example: $\mu(t,x) = \mu x$ and $\sigma(t,x) = \sigma x$, so that $dX_t = \mu X_t dt + \sigma X_t dW_t$.
- The *solution* to an SDE is an equation that describes X_t in terms of just W_t (i.e., X_t does not appear on the RHS). Often, Ito's lemma is helpful in finding it.

Ito's Lemma

- Ito's lemma answers the question: if X_t is an Ito process with given dynamics, then what are the dynamics of a function $f(t, X_t)$?
- It can be stated as follows: Let $\{X_t\}_{t\geq 0}$ be an Ito process satisfying $dX_t = \mu_t dt + \sigma_t dW_t$, and consider a function $f: \mathbb{R}^+ \times \mathbb{R} \to \mathbb{R}$ with continuous partial derivatives

$$\dot{f}(t,x) = \frac{\partial f(t,x)}{\partial t}, \qquad f'(t,x) = \frac{\partial f(t,x)}{\partial x}, \qquad f''(t,x) = \frac{\partial^2 f(t,x)}{\partial x^2}.$$

Then

$$df(t, X_t) = \dot{f}(t, X_t)dt + f'(t, X_t)dX_t + \frac{1}{2}f''(t, X_t)\sigma_t^2 dt.$$

• Example: Geometric Brownian Motion. Let

$$dS_t = S_t \mu dt + S_t \sigma dW_t, \qquad (\ddagger)$$
 and $X_t = f(S_t) = \log S_t$. Then $\dot{f}(S_t) = 0$, $f'(S_t) = 1/S_t$, $f''(S_t) = -1/S_t^2$, and
$$dX_t = df(S_t) = \dot{f}(S_t)dt + f'(S_t)dS_t + \frac{1}{2}f''(S_t)(S_t\sigma)^2 dt$$

$$= \frac{1}{S_t}dS_t - \frac{1}{2S_t^2}(S_t\sigma)^2 dt$$

$$= \frac{1}{S_t}(S_t\mu dt + S_t\sigma dW_t) - \frac{1}{2}\sigma^2 dt$$

$$= \nu dt + \sigma dW_t, \qquad \nu = \mu - \frac{1}{2}\sigma^2$$

• I.e., (‡) is the SDE for GBM: $S_t = \exp(X_t) = S_0 \exp(\nu t + \sigma W_t)$.

- Intuition (see Hull, 2012, Appendix to Ch. 13): In standard calculus, the total differential $df = \dot{f}(t,g(t))dt + f'(t,g(t))dg(t)$
 - is the linear part of a Taylor expansion; the remaining terms are of smaller order as $dt, dg(t) \rightarrow 0$, so the total differential is a local linear approximation to f.
- If $g(t) = X_t$, an Ito process, take a 2nd order Taylor approximation:

$$\delta f \approx f(t, X_t) \delta t + f'(t, X_t) \delta X_t + \frac{1}{2} \left[\frac{\partial^2 f}{\partial t^2} (\delta t)^2 + 2 \frac{\partial^2 f}{\partial t \partial X_t} (\delta t) (\delta X_t) + \frac{\partial^2 f}{\partial X_t^2} (\delta X_t)^2 \right].$$

- We have that $\delta X_t = (X_{t+\delta t} X_t) \approx \mu_t \delta t + \sigma_t \delta W_t \sim N(\mu_t \delta t, \sigma_t^2 \delta t)$. Thus, $\mathbb{E}[(\delta X_t)^2] \approx (\mu_t \delta t)^2 + \sigma_t^2 \delta t \approx \sigma_t^2 \delta t$; i.e., the 2nd order term is of the same order of magnitude as the 1st order term δt .
- It can be shown that as $\delta t \to 0$, $(\delta X_t)^2$ can be treated as non-stochastic: $(dX_t)^2 = \sigma_t^2 dt$. Together with $(dt)^2 = 0$ and $(dt)(dX_t) = 0$ this gives the result.

Simulating Ito Processes

Suppose we want to simulate sample paths of an Ito process described by the SDE

$$dX_t = \mu(t, X_t)dt + \sigma(t, X_t)dW_t.$$

- For pricing European claims, we only need draws for X_T , but for path-dependent options, we need the entire path $\{X_t\}_{t\in[0,T]}$.
- ullet A simple way is to discretize the model, for a small time step δt , as

$$\delta X_t = X_{t+\delta_t} - X_t \approx \mu(t, X_t) \delta t + \sigma(t, X_t) \delta W_t$$

where $\delta W_t \sim N(0, \delta t)$. This is known as the Euler scheme.

• With $\delta t = T/N$, we can sample the path at N discrete times $t_i = i\delta t$ as

$$X_{i+1} = X_i + \mu(t_i, X_i)\delta t + \sigma(t_i, X_i)\sqrt{\delta t}Z_i,$$

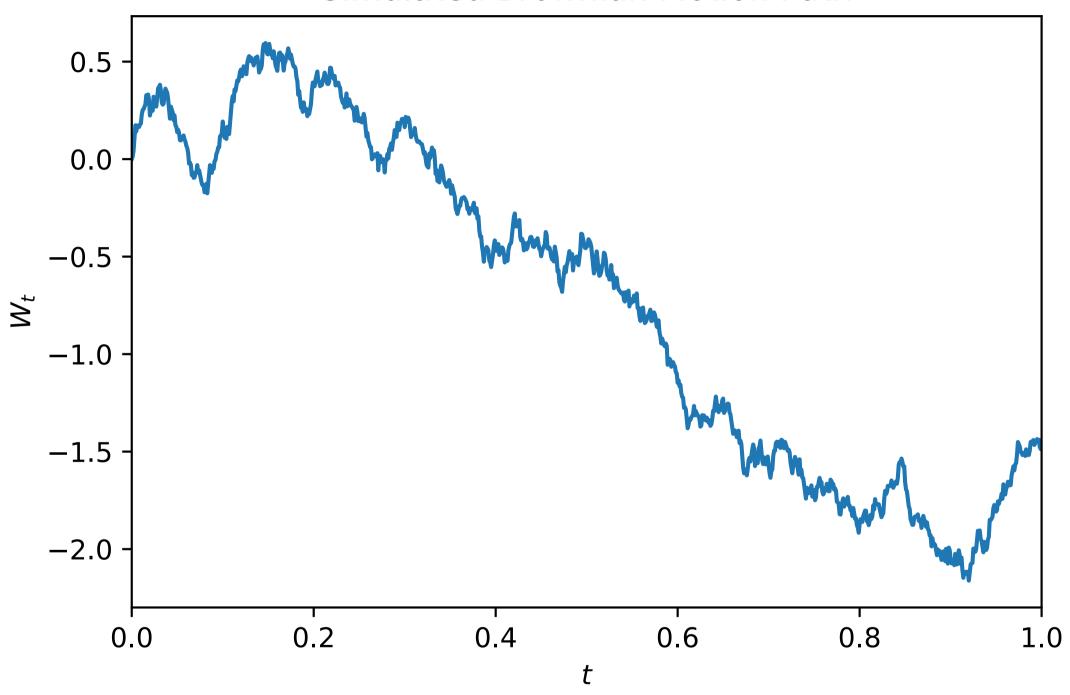
where the Z_i are independent standard normal random numbers and we use X_i and $X_{i\delta t}$ interchangeably.

- In order to implement this, we need a way of drawing random samples from the normal distribution.
- Computers are deterministic machines. They cannot generate true random numbers.
- Instead, they construct sequences of pseudo-random numbers from a specified distribution that *look* random, in the sense that they pass certain statistical tests.
- E.g., NumPy's np.random.randn(d0[, d1, ...]) constructs an array of standard normal pseudo random numbers.
- Random number generators use a *seed* value for initialization. Given the same seed, the same pseudo-random sequence will be returned.
- NumPy picks the seed automatically. To force it to use a specific seed, use np.random.seed(n). Putting this line at the beginning of your Monte-Carlo program ensures that you get exactly the same results every time the program is run.

Example 1: Simulating Brownian Motion

```
In [1]: import numpy as np
        import matplotlib.pyplot as plt
        import pandas as pd
        %matplotlib inline
In [2]: def bmsim(T, N, X0=0, mu=0, sigma=1):
            Simulate a Brownian motion path.
             11 11 11
            deltaT = float(T)/N
            tvec = np.linspace(0, T, N+1)
            z = np.random.randn(N+1) #N+1 is one more than we need, actually. This way we won't have to grow dX.
            dX = mu*deltaT + sigma*np.sqrt(deltaT)*z  #X[j+1]-X[j]=mu*deltaT + sigma*np.sqrt(deltaT)*z[j].
            dX[0] = 0.
            X = np.cumsum(dX)
            X += X0
             return tvec, X
In [3]: np.random.seed(0)
        tvec, W = bmsim(1, 1000)
        W = pd.Series(W, index=tvec)
        W.plot()
        plt.title('Simulated Brownian Motion Path')
        plt.xlabel("$t$"); plt.ylabel("$W_t$");
        plt.savefig("img/BMpath.svg"); plt.close()
```

Simulated Brownian Motion Path

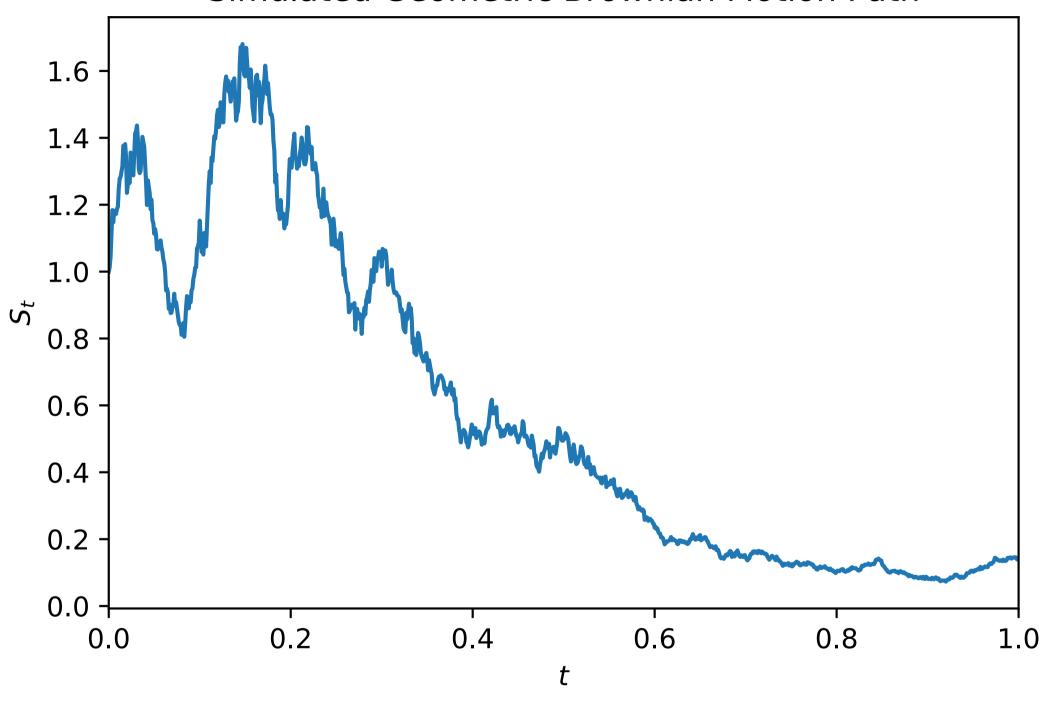


Example 2: Simulating GBM

• The Euler scheme for the GBM $dS_t = S_t \mu dt + S_t \sigma dW_t$ is $S_{i+1} = S_i + S_i \mu \delta t + S_i \sigma \sqrt{\delta t} Z_i.$

```
In [4]: def gbmsim(T, N, S0=1, mu=0, sigma=1):
             "Simulate a Geometric Brownian motion path."
            deltaT = float(T)/N
            tvec = np.linspace(0, T, N+1)
            z = np.random.randn(N+1) #Again one more than we need. This keeps it comparable to bmsim.
            S = np.zeros_like(z)
            S[0] = S0
            for j in xrange(0, N): #Note: we can no longer vectorize this, because S[:, j] is needed for S[:, j+1].
                S[j+1] = S[j] + mu*S[j]*deltaT + sigma*S[j]*np.sqrt(deltaT)*z[j+1]
             return tvec, S
In [5]: np.random.seed(0)
        tvec, S = gbmsim(1, 1000)
        S = pd.Series(S, index=tvec)
        S.plot()
        plt.title('Simulated Geometric Brownian Motion Path')
        plt.xlabel("$t$"); plt.ylabel("$S_t$")
        plt.savefig("img/GBMpath.svg"); plt.close()
```

Simulated Geometric Brownian Motion Path



- In the case of BM, the Euler scheme correctly reproduces the distribution of the W_{t_i} .
- This is not true in general: in Example 2 above, the Euler approximation

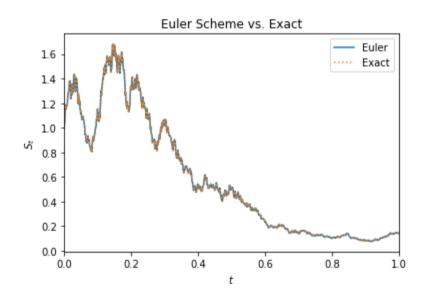
$$S_{i+1} = S_i + S_i \mu \delta t + S_i \sigma \sqrt{\delta t} Z_i$$

implies that the distribution of $S_{t+\delta t}-S_t$ is normal, not log-normal as it should be.

- Under mild conditions, the error introduced by discretization will disappear as $\delta t \to 0$.
- In the case of GBM, this error can be avoided altogether: let $X_t \equiv \log S_t$. By Ito's lemma,

$$dX_t = \nu dt + \sigma dW_t, \quad \nu = \mu - \frac{1}{2}\sigma^2,$$

so we can simulate X_t instead and then take the exponential.



The Black-Scholes Model

- Black and Scholes assumed the following model:
 - The stock $\{S_t\}_{t \in [0,T]}$ follows GBM:

$$dS_t = \mu S_t dt + \sigma S_t dW_t.$$

- The stock pays no dividends.
- Cash bond price $B_t = e^{rt} \iff dB_t = rB_t dt$; riskless lending and borrowing at the same rate r.
- European style derivative with price C_t and payoff $C_T = (S_T)$.
- Trading may occur continuously, with no transaction costs.
- No arbitrage opportunities.
- The problem is to find the price C_t , $t \in [0, T]$.

• It can be shown that the FTAP holds in continuous time as well: if the market is arbitrage free, then there exists a risk neutral measure $\mathbb Q$ under which all assets earn the risk free rate (on average), and the price of a claim is the discounted expected payoff under $\mathbb Q$. If the market is complete, then $\mathbb Q$ is unique. This gives us a pricing formula for general European claims:

$$C_t = e^{-r(T-t)} \mathbb{E}^{\mathbb{Q}} \left[C_T | \mathcal{F}_t \right].$$

• This implies that if we can simulate the stock price under the measure \mathbb{Q} , then we can price the claim by Monte Carlo simulation.

• In the BS model, it can be shown that under the risk-neutral measure \mathbb{Q} ,

$$dS_t = rS_t dt + \sigma S_t dW_t^{\mathbb{Q}},$$

where $W_t^{\mathbb{Q}}$ is a \mathbb{Q} -Brownian Motion.

• Note that by Ito's Formula, the discounted stock price $\tilde{S}_t \equiv e^{-rt}S_t =: f(t, S_t)$ satisfies

$$d\tilde{S}_{t} = \dot{f}(t, S_{t})dt + f'(t, S_{t})dS_{t} + \frac{1}{2}f''(t, S_{t})\sigma^{2}S_{t}^{2}dt$$

$$= -re^{-rt}S_{t}dt + e^{-rt}dS_{t} + 0$$

$$= -r\tilde{S}_{t}dt + e^{-rt}(rS_{t}dt + \sigma S_{t}dW_{t}^{\mathbb{Q}})$$

$$= \sigma\tilde{S}_{t}dW_{t}^{\mathbb{Q}}.$$

• I.e., \tilde{S}_t is an Ito process without drift, and thus a martingale. This is the reason $\mathbb Q$ is also called the equivalent martingale measure.

- We can extend the BS model by assuming that the underlying pays a continuous dividend at rate δ (realistic only for indices, not individual stocks). Then a position of 1 share generates an instantaneous dividend stream $\delta S_t dt$, in addition to the capital gains dS_t .
- Note that only the holder of the underlying receives the dividend.
- The pricing formula remains the same, but now the risk-neutral dynamics of S_t are $dS_t = (r \delta)S_t dt + \sigma S_t dW_t^{\mathbb{Q}}.$
- The expected growth rate of the underlying under \mathbb{Q} is $r-\delta$, so the expected return from holding it (capital gains plus dividend yield) is r.
- The price of a call is now

$$C_t = e^{-\delta(T-t)} S_t \Phi(d_1) - e^{-r(T-t)} K \Phi(d_2),$$

where

$$d_{1,2} = \frac{\log(S_t/K) + [(r-\delta) \pm \frac{1}{2}\sigma^2](T-t)}{\sigma\sqrt{T-t}}.$$

Monte Carlo Pricing

The goal in Monte Carlo simulation is to obtain an estimate of

$$\theta \equiv \mathbb{E}[X],$$

for some random variable X with finite expectation. The assumption is that we have a means of sampling from the distribution of X, but no closed-form expression for θ .

• Suppose we have a sample $\{X_i\}_{i\in\{1,\ldots,n\}}$ of independent draws for X, and let

$$\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i.$$

- The sample average \bar{X}_n is an unbiased estimator of θ : $\mathbb{E}[\bar{X}_n] = \theta$.
- The weak law of large numbers states that

$$\bar{X}_n \stackrel{p}{\to} \theta,$$

where the arrow denotes convergence in probability; i.e., as the sample size grows, the X_n becomes a better and better estimate of θ .

- Thus, our strategy is to use a computer to draw n (pseudo) random numbers X_i from the distribution of X, and then estimate θ as the sample mean of the X_i .
- *n* is called the number of *replications*.
- For finite n, the sample average will be an approximation to θ .
- It is usually desirable to have an estimate of the accuracy of this approximation. Such an estimate can be obtained from the *central limit theorem* (CLT), which states that

$$\sqrt{n}(\bar{X}_n - \theta) \stackrel{d}{\to} N(0, \sigma^2),$$

provided that σ^2 , the variance of X, is finite. The arrow denotes convergence in distribution; this implies that for large n, \bar{X}_n has approximately a normal distribution.

• Of course σ^2 is unknown, but we can estimate it as

$$\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n \left(\bar{X}_n - X_i \right)^2.$$

• A 95% confidence interval (CI) is an interval $[c_l, c_u]$ such that

$$\mathbb{P}[c_l \le \theta \le c_u] = 0.95.$$

• The CLT implies that, in the limit as $n \to \infty$,

$$\mathbb{P}[-1.96\sigma \le \sqrt{n}(\bar{X}_n - \theta) \le 1.96\sigma] = 0.95 \Leftrightarrow$$

$$\mathbb{P}[\bar{X}_n - 1.96\frac{\sigma}{\sqrt{n}} \le \theta \le \bar{X}_n + 1.96\frac{\sigma}{\sqrt{n}}] = 0.95.$$

- Hence $c_l=\bar{X}_n-1.96\frac{\sigma}{\sqrt{n}}$ and $c_u=\bar{X}_n+1.96\frac{\sigma}{\sqrt{n}}$ is an asymptotically valid CI.
- Note that c_l and c_u are random variables; we should interpret this as "before the experiment is performed, there is a 95% chance that a CI computed according to this formula will contain θ ". After performing the experiment, this statement is not valid anymore; the interval is now fixed, and contains θ with probability either 0 or 1.
- The unknown parameter σ can be consistently estimated by $\sqrt{\hat{\sigma}^2}$.

Application: Asian Options

- The payoff of Asian options depends on the average price of the underlying, \bar{S}_T . Types:
 - Average price Asian call with payoff $(\bar{S}_T K)^+$;
 - Average price Asian put with payoff $(K \bar{S}_T)^+$;
 - Average strike Asian call with payoff $(S_T \bar{S}_T)^+$;
 - Average strike Asian put with payoff $(\bar{S}_T S_T)^+$.
- It is important to specify how the average is computed: the continuous, arithmetic, and geometric averages are, respectively,

$$\frac{1}{T} \int_0^T S_t dt, \qquad \frac{1}{N} \sum_{i=1}^N S_{t_i} \quad \text{and} \quad \left(\prod_{i=1}^N S_{t_i}\right)^{1/N},$$

where the t_i are a set of N specified dates.

- Exact Black-Scholes type pricing formulas for Asian options only exist in special cases (e.g., the geometric average Asian call, see next week), so we rely on Monte Carlo simulation.
- Our pricing formula

$$C_t = e^{-r(T-t)} \mathbb{E}^{\mathbb{Q}} \left[C_T | \mathcal{F}_t \right]$$

is exactly in the required form.

As an example, consider pricing an arithmetic average price call with payoff

$$C_T = (\bar{S}_T - K)^+, \text{ where } \bar{S}_T = \frac{1}{N} \sum_{i=1}^N S_{t_i},$$

which cannot be priced analytically.

• The payoff is path-dependent, so we need to simulate the entire asset price path, not just S_T .

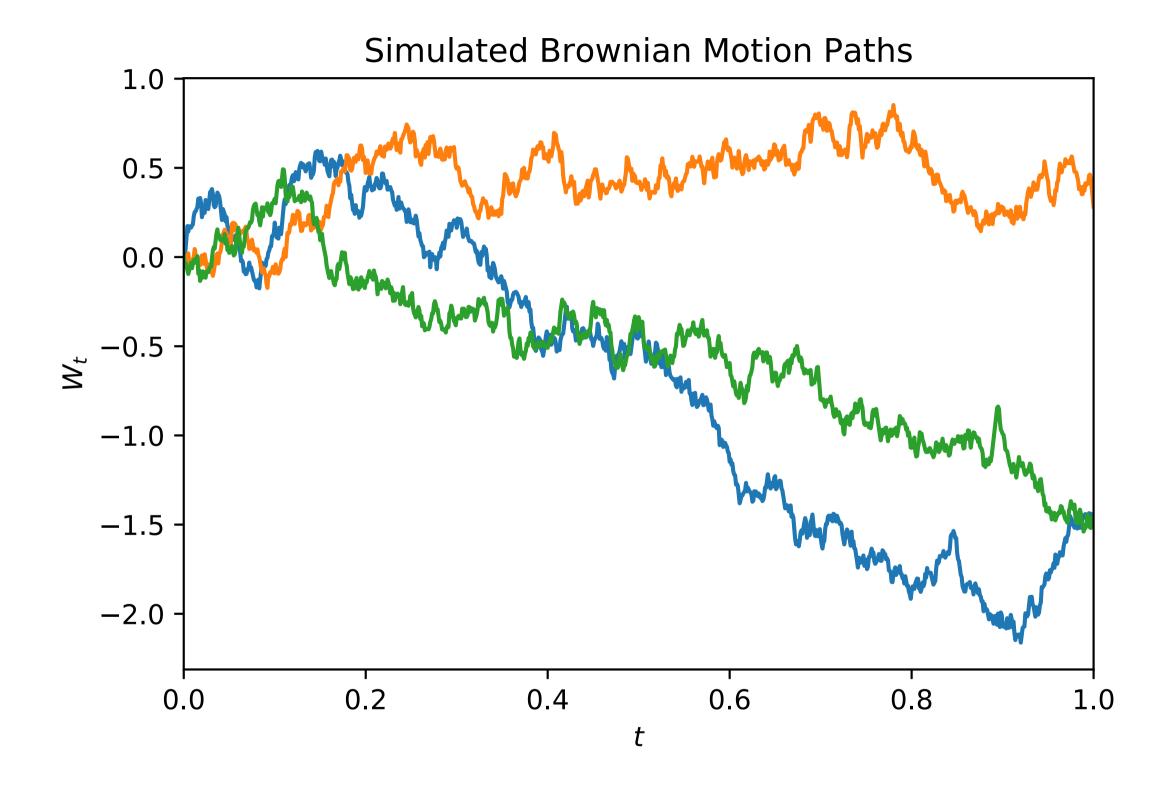
```
In [7]: from scipy.stats import norm
        def asianmc(S0, K, T, r, sigma, delta, N, numsim=10000):
            Monte Carlo price of an arithmetic average Asian call.
            X0 = np.log(S0)
            nu = r-delta-.5*sigma**2
            payoffs = np.zeros(numsim)
            for j in xrange(numsim):
                _, X = bmsim(T, N, X0, nu, sigma) #Convention: underscore holds value to be discarded.
                S = np.exp(X)
                payoffs[j] = \max(S[1:].mean()-K, 0.)
            g = np.exp(-r*T)*payoffs
            C = g.mean(); s = g.std()
            zq = norm.ppf(0.975)
            Cl = C-zq/np.sqrt(numsim)*s
            Cu = C+zq/np.sqrt(numsim)*s
            return C, Cl, Cu
In [8]: S0 = 11; K = 10; T = 3/12.; r = 0.02; sigma = .3; delta = 0.01; N = 10
        np.random.seed(0)
        CO, Cl, Cu = asianmc(SO, K, T, r, sigma, delta, N); CO, Cl, Cu
```

Out[8]: (1.0927262054551385, 1.0747653929130998, 1.1106870179971773)

Code Optimization

- Our code for pricing the Asian option is likely inefficient, because it contains a loop.
- The code can be 'vectorized' to speed it up.
- First step: simulate a bunch of Brownian paths in one shot.
- The resulting code is actually almost identical:

```
In [9]: def bmsim_vec(T, N, X0=0, mu=0, sigma=1, numsim=1): #Note new input: numsim, the number of paths.
             Simulate `numsim` Brownian motion paths.
             deltaT = float(T)/N
             tvec = np.linspace(0, T, N+1)
             z = np.random.randn(numsim, N+1) #(N+1) -> (numsim, N+1)
             dX = mu*deltaT + sigma*np.sqrt(deltaT)*z
             dX[:, 0] = 0. \#dX[0] -> dX[:, 0]
             X = np.cumsum(dX, axis=1) #cumsum(dX) -> cumsum(dX, axis=1)
             X += X0
             return tvec, X
In [10]: np.random.seed(0)
         tvec, W = bmsim_vec(1, 1000, numsim=3)
         W = pd.DataFrame(W.transpose(), index=tvec)
         W.plot().legend().remove()
         plt.title('Simulated Brownian Motion Paths')
         plt.xlabel("$t$"); plt.ylabel("$W_t$");
         plt.savefig("img/BMpaths.svg"); plt.close()
```



• Here is the vectorized code for the Asian option:

• Let's see if it works:

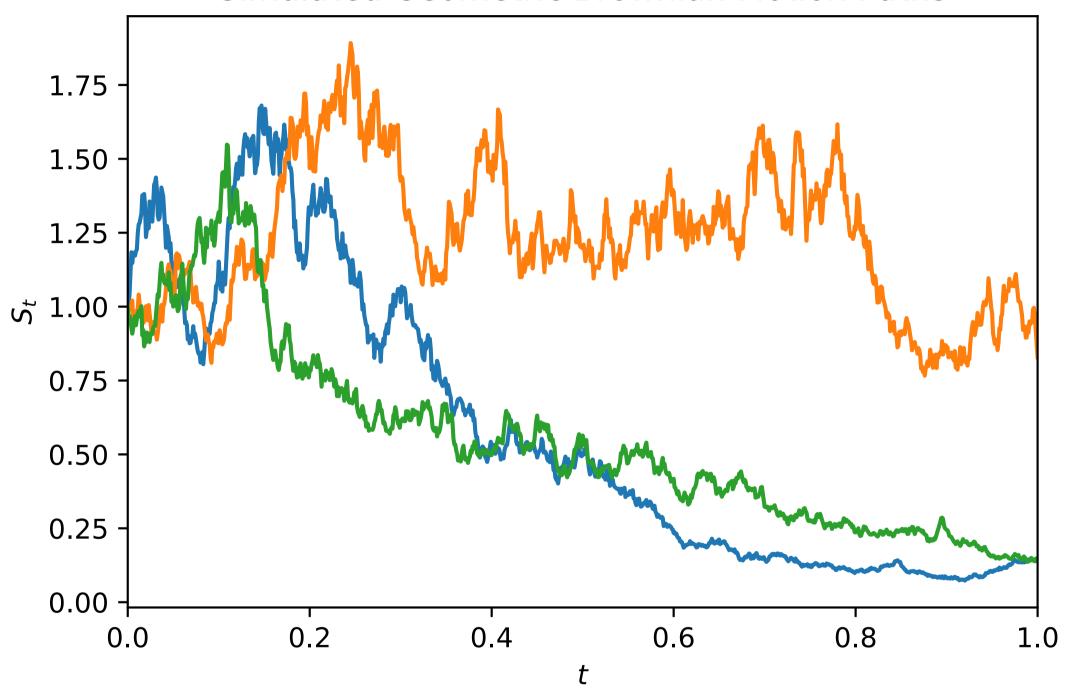
```
In [12]: np.random.seed(0)
    CO_vec, _, _ = asianmc_vec(S0, K, T, r, sigma, delta, N)
    np.allclose(C0_vec, C0)
Out[12]: True
```

• And time it:

- Our code for the Euler scheme can likewise be adjusted to compute many paths in one shot.
- We're still stuck with the loop over t though, which cannot be vectorized because S_{i+1} depends on S_i .
- We'll use Numba's JIT compiler to speed it up further.

```
In [15]: from numba import jit
         @jit(nopython=True)
         def gbmsim_vec(T, N, S0=1, mu=0, sigma=1, numsim=1, seed=0):
             "Simulate `numsim` Geometric Brownian motion paths."
             deltaT = float(T)/N
             tvec = np.linspace(0, T, N+1)
             np.random.seed(seed) #Note: with jit-compiled functions, the RNG must be seeded INSIDE the compiled code.
             z = np.random.randn(numsim, N+1)
             S = np.zeros_like(z)
             S[:, 0] = S0
             for j in xrange(0, N):
                 S[:, j+1]=S[:, j] + mu*S[:, j]*deltaT + sigma*S[:, j]*np.sqrt(deltaT)*z[:, j+1]
             return tvec, S
In [16]:
         tvec, S = gbmsim_vec(1, 1000, numsim=3, seed=0)
         S = pd.DataFrame(S.transpose(), index=tvec)
         S.plot().legend().remove()
         plt.title('Simulated Geometric Brownian Motion Paths')
         plt.xlabel("$t$"); plt.ylabel("$S_t$")
         plt.savefig("img/GBMpaths.svg"); plt.close()
```

Simulated Geometric Brownian Motion Paths



• The compiled code produces the same results:

• But it is quite a bit faster:

```
In [18]:  %%timeit #Cell magic (for timing the entire cell).

for k in xrange(10): #10 paths.
    gbmsim(1, 1000)

10 loops, best of 3: 22.9 ms per loop

In [19]:  %timeit gbmsim_vec(1, 1000, numsim=10)

The slowest run took 521.25 times longer than the fastest. This could mean that an intermediate result is being cached.
    1 loop, best of 3: 580 µs per loop
```