| | (1) | (2) | (3) | (1) | (2) | (3) |
|-------------------------|---|--------------------|---|---|--------------------|---|
| | $\frac{R_{ur}^2 +}{(R_{ur}^2 - R_r^2)}$ | $2 \cdot R_{ur}^2$ | $R^2 \text{ for } \beta_{ols} = \beta_{iv}$ | $\frac{R_{ur}^2 +}{(R_{ur}^2 - R_r^2)}$ | $2 \cdot R_{ur}^2$ | $R^2 \text{ for } \beta_{ols} = \beta_{iv}$ |
| Probability of Election | 1.05 (.23) | 0.63 (.29) | (.46) | 1.69 (.96) | 1.49 (1.00) | (3.07) |
| Vote Share | 0.68 (.64) | 0.48 (.74) | - (.99) | 2.05 (1.00) | 2.05 (1.00) | (3.01) |
| Vote Distance to Cutoff | 7.74 | 6.05 | _ | 20.51 | 20.51 | _ |

Panel B: Individual Covariate and

Fixed-Effects Models

(City Councilor) (.21)(2.11)(.23)(1.00)(1.00)(24.86)Vote Distance to Cutoff 2.64 1.56 1.21 1.21

(Mayor) (.23)(.29)

column 3 in each panel.

(.64)(1.00)

(1.00)(1.51)

Note: In each panel, I compare the unrestricted coefficient for the model in the panel title $(\tilde{\beta})$ against the restricted

Panel A: Individual Covariate Models

coefficient for the bivariate model (β^0). The different outcomes are summarized across rows. Columns 1 and 2 in each panel display conditions for R_{max} calculations in the row just above table content. The first value in each cell

not cap it for calculations of the necessary R^2 to yield $\beta_{ols} = \beta_{iv}$; therefore, some nonsensical $R^2 > 1$ appear in

is the δ for each model. R^2 values are reported inside parentheses. While I cap R^2 at one for δ calculations, I do