

Last Time

- uncertainty quantification!
- epistemic / aleatoric
- modeling paradigms

lecture 10

EAI S'25

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This Time:

- practical methods for UQ

Announcement: midterm report due March 14<sup>th</sup> (Friday)

CREDIT: Notes inspired by Prof. Eric Nalisnick's lecture @ m<sup>2</sup>L

# Summary & UQ Methods

	<u>Frequentism</u>	<u>Bayesianism</u>
✓	data-dirch, easy comp. MLE $\hat{\theta} = \underset{\theta}{\operatorname{argmax}} p(y x; \theta)$	prior dist "jump starts" learning, posterior models uncertainty over $\theta$ params
✗	misked by sampling noise, dataset size, etc.	computation usually too costly for exact solution
frontiers:	$\Rightarrow$ "beef them up"	$\rightarrow$ approximate this!

## Practical Methods for UQ

### Frequentism

- 1) bootstrap aggregation ("bagging") "ensemble"
- 2) conformal prediction  $\leftarrow$  next week: Prof. Anushri Dixit will lecture

### Bayesianism

- 1) sample-then-optimize ensembling "ensemble"
- 2) Gaussian linear regression  $\Rightarrow$  Gaussian Processes (GPs)
- 3) Laplace Approximation

## BOOTSTRAP AGGREGATION (BAGGING)

Recall how frequentism assumes that the randomness comes from the data sampling process. But, we have fixed dataset!

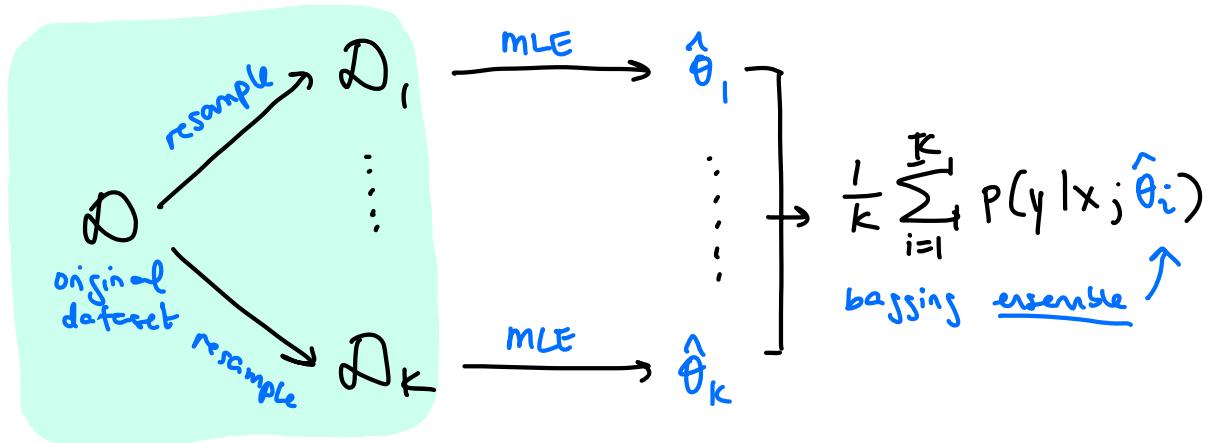
Bootstrapping synthesizes additional datasets by resampling from the <sup>fixed</sup> training set.

Dirac delta

$$\{\mathcal{D}_k\}_{k=1}^K \sim \frac{1}{N} \sum_{i=1}^N \delta[(x_i, y_i)]$$

"sample with replacement" ↗

Intuition: w/ equal probability I will sample a data pt. from the existing dataset, draw a new data pt & put it into a new dataset. Do this K times.



Further Reading: "Intro to the Bootstrap" by Efron & Tibshirani

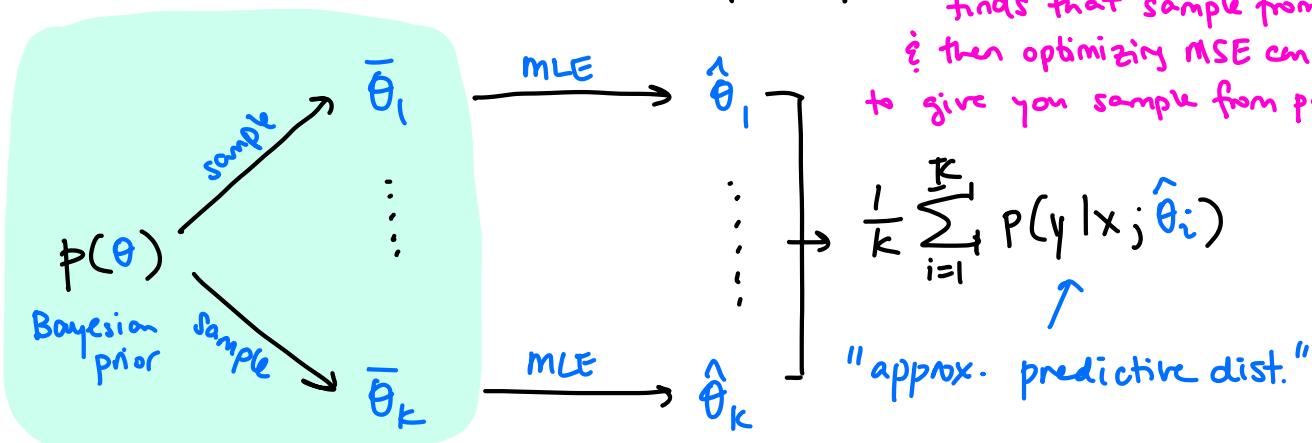
### Sample - then - Optimize Ensemble

Bayesians model randomness in the prior too. We can perform a bagging-like procedure but using samples from the prior to initialize training!

$$\{\bar{\theta}_k\}_{k=1}^K \sim p(\theta)$$

Matthews et. al 2017

finds that sample from prior & then optimizing MSE can be shown to give you sample from posterior.

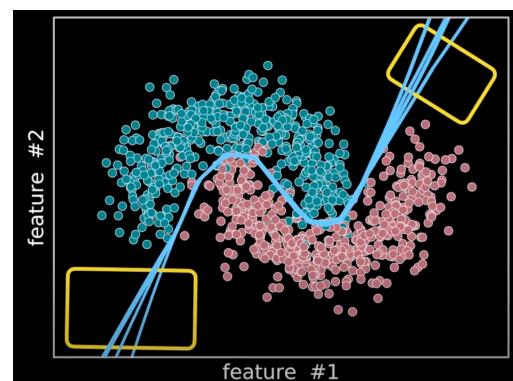


NOTE: Izmailov et. al. ICML 2021 finds surprisingly comparable results

btwn. this approach & high-fidelity Bayesian inference.

ASIDE: this is what generated these  $K$  decision boundaries from  $K$  NN's.  $\longrightarrow$

Re-initialize NN parameters, prior comes from the initialization scheme implemented in scikit-learn



← from Prof. Eric Nalisnick's lecture @ m2L.

### Gaussian (linear) Regression

Is there any way to model the Bayesian uncertainty explicitly exactly? i.e.  $p(\tilde{y} | \tilde{x}, \mathcal{D}) = \int_{\theta} p(\tilde{y} | \tilde{x}; \theta) p(\theta | \mathcal{D}) d\theta$

The core assumption / condition under which you can get this exactly is that all aspects of our model + world are Gaussian. The reason why this is helpful is b/c of the properties of Gaussians:

ONCE A GAUSSIAN, ALWAYS A GAUSSIAN

Specifically, we will start w/ regression problems but we will use a running example of linear models.

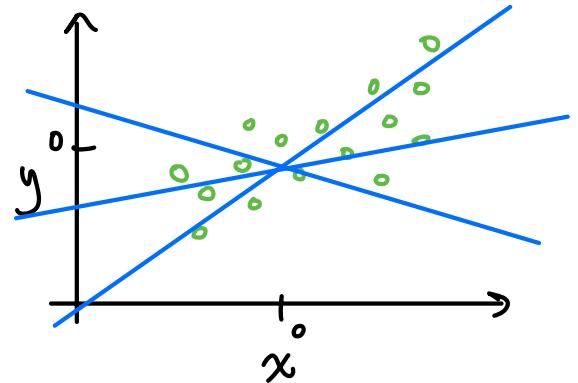
data:  $y \sim N(\theta^T x, \Sigma)$

$x \in \mathbb{R}^n$

mean

covariance

model:  $\hat{y} = \hat{\theta}^T x$



I could just fit a line here, but I also want some uncertainty estimate over alternative lines I could have chosen ( $p(\theta | \mathcal{D})$ )

PRIOR:  $p(\theta) = N(\mu_0, \Sigma_0)$

LIKELIHOOD (i.e. MODEL):  $p(y | x; \theta) = N(\theta^T x, \Sigma)$

POSTERIOR: 
$$p(\theta | \mathcal{D}) = \frac{p(\theta) \prod_{i=1}^N p(y_i | x_i; \theta)}{\int_{\bar{\theta}} p(\bar{\theta}) \prod_{i=1}^N p(y_i | x_i; \bar{\theta}) d\bar{\theta}}$$

Closed-form expression of the posterior which looks like Gaussian!

$$p(\theta | \mathcal{D}) = N(\mu_N, \Sigma_N)$$

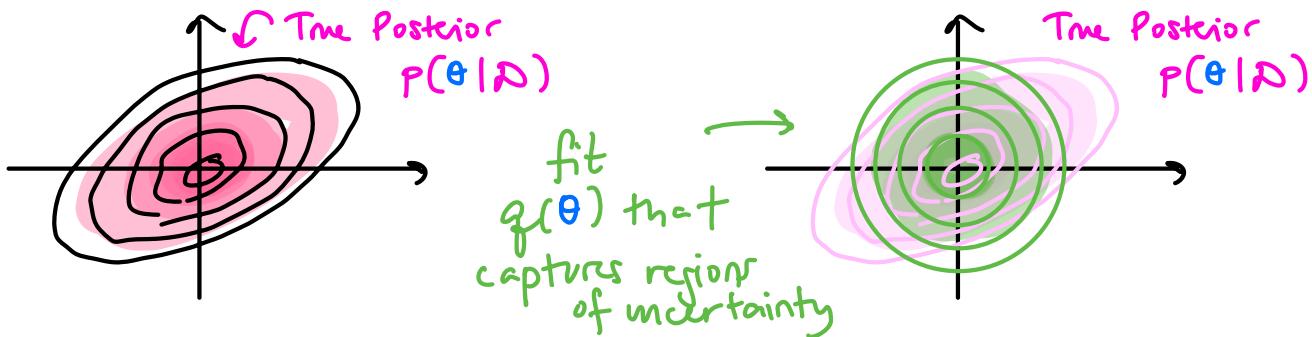
↑ a function of  $\Sigma_0, \Sigma, \theta$   
↑ function of  $\Sigma_N, \Sigma_0, \mu_0, \Sigma, \theta, y_i$

! ULTIMATELY, these operations are just matrix-vector mult & addition

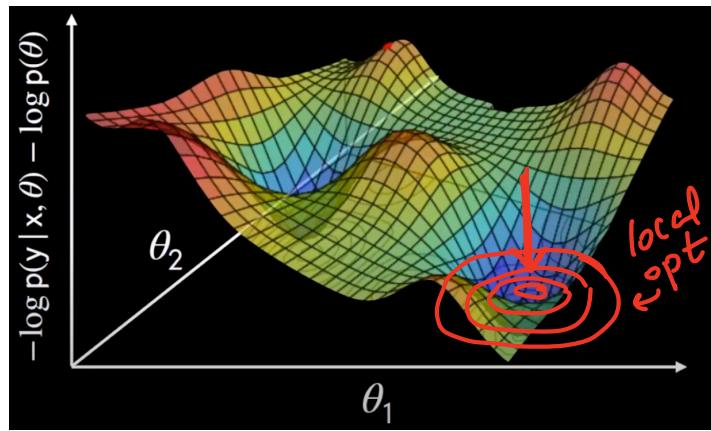
### LAPLACE APPROXIMATION (LA)

What if my models aren't Gaussian?

We want to compute a tractable approx to the true Bayesian posterior by taking the intractable posterior dist and fit a simple dist to it!



Laplace approx. is old technique w/ recent resurgence (~2021)



Normally we'd stop here ↓ & return this local opt, but Laplace says let's find a local distribution around this pt. estimate  $\hat{\theta}$

LA says that the posterior can be approx. by Gaussian dist where the mean is that  $\hat{\theta}$  and its variance is the inverse Hessian matrix.

$$p(\theta | \mathcal{D}) \approx \mathcal{N}(\hat{\theta}, H^{-1}(\hat{\theta}))$$

Q Where does this come from?

$$\begin{aligned} p(\theta | \mathcal{D}) &:= \frac{1}{\int p(\mathcal{D}|\theta) p(\theta) d\theta} p(\mathcal{D}|\theta) p(\theta) \\ &= \frac{1}{Z} h(\theta) \end{aligned} \quad // \text{Bayes Rule}$$

we want to approx. w/ Gaussian. First, note that:

$$Z = \int \exp[\log h(\theta)] d\theta$$

let  $\hat{\theta}$  be local min.

$$\text{FACT: } \int \exp[-\frac{1}{2} x^T A x] dx = \frac{\sqrt{(2\pi)^n}}{\sqrt{\det A}} \quad \begin{matrix} \leftarrow \text{dim. of } x \\ \text{integral of Gaus. func is closed form!} \end{matrix}$$

STEP 1: 2<sup>nd</sup> order Taylor Series Expansion about  $\hat{\theta}$

$$\begin{aligned} \log h(\theta) &\approx \log h(\hat{\theta}) + \boxed{\nabla_{\theta} \log h(\hat{\theta})^T (\theta - \hat{\theta})} = 0 @ \text{optimum.} \\ &\quad + \frac{1}{2} (\theta - \hat{\theta})^T \nabla_{\theta}^2 \log h(\hat{\theta}) (\theta - \hat{\theta}) \end{aligned}$$

$$\begin{aligned}\log h(\theta) &\approx \log h(\hat{\theta}) - \left( -\frac{1}{2}(\theta - \hat{\theta})^T \nabla_{\theta}^2 \log h(\hat{\theta})(\theta - \hat{\theta}) \right) \\ &= \log h(\hat{\theta}) - \left( \frac{1}{2}(\theta - \hat{\theta})^T \boxed{\Delta}(\theta - \hat{\theta}) \right) \\ &\quad \uparrow := -\nabla_{\theta}^2 \log h(\hat{\theta})\end{aligned}$$

STEP 2: Plug into integral!

$$\begin{aligned}&\int \exp[\log h(\theta)] d\theta \\ &\approx \int \exp[\log h(\hat{\theta}) - \left( \frac{1}{2}(\theta - \hat{\theta})^T \boxed{\Delta}(\theta - \hat{\theta}) \right)] d\theta \\ &= h(\hat{\theta}) \int \exp[-\left( \frac{1}{2}(\theta - \hat{\theta})^T \boxed{\Delta}(\theta - \hat{\theta}) \right)] d\theta \\ &\quad \uparrow := "x" \quad \uparrow := "A" \\ &= h(\hat{\theta}) \frac{\sqrt{(2\pi)^n}}{\sqrt{\det \Delta}} \text{ hessian matrix! eval @ } \hat{\theta}\end{aligned}$$

STEP 3: Plug approx back into  $P(\theta | \mathcal{D})$  posterior!

$$\begin{aligned}p(\theta | \mathcal{D}) &:= \frac{1}{\int p(\mathcal{D}|\theta)p(\theta)d\theta} \boxed{p(\mathcal{D}|\theta)p(\theta)} \\ &= \frac{1}{\sqrt{\det \Delta}} \cdot \exp\left[-\frac{1}{2}(\theta - \hat{\theta})^T \boxed{\Delta}(\theta - \hat{\theta})\right]\end{aligned}$$

This is just Gaussian density  $\theta \sim \mathcal{N}(\hat{\theta}, \Sigma)$   
 by def<sup>n</sup> of multivar Gaussian:  $\Sigma := \Delta^{-1}$

⚠ If  $\theta \in \mathbb{R}^m$  then the hessian  $\Delta \in \mathbb{R}^m \times \mathbb{R}^m$ . But, if  $\theta$  is weights of NN, this could be HUGE (e.g. GPT!)  
 ⇒ in practice, use only last layer  $\theta$ 's ? impose low-rank  $\Delta$  structure!