Chapter 4

- 4.1 Normal approximation (Laplace's method)
- 4.2 Large-sample theory
- 4.3 Counter examples
 - includes examples of difficult posteriors for MCMC, too
- 4.4 Frequency evaluation*
- 4.5 Other statistical methods*

Normal approximation (Laplace approximation)

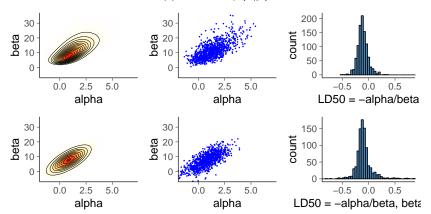
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 - Laplace used this (before Gauss) to approximate the posterior of binomial model to infer ratio of girls and boys born

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- if $\hat{\theta}$ is at mode, then $f'(\hat{\theta}) = 0$
- often when $n \to \infty$, $\frac{f^{(3)}(\hat{\theta})}{3!}(\theta \hat{\theta})^3 + \dots$ is small

Multivariate Taylor series

Multivariate series expansion

$$f(\theta) = f(\hat{\theta}) + \frac{df(\theta')}{d\theta'}_{\theta' = \hat{\theta}} (\theta - \hat{\theta}) + \frac{1}{2!} (\theta - \hat{\theta})^T \frac{d^2 f(\theta')}{d\theta'^2}_{\theta' = \hat{\theta}} (\theta - \hat{\theta}) + \dots$$

• Taylor series expansion of the log posterior around the posterior mode $\hat{\theta}$

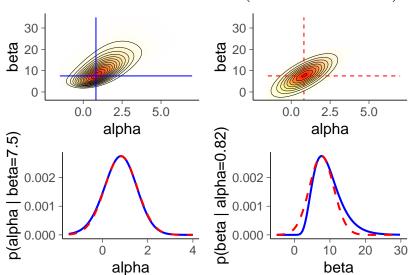
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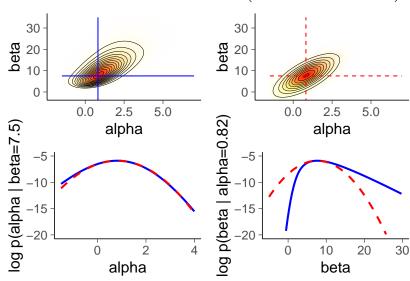
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Normal approximation

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where $I(\theta)$ is called *observed information*

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 $Hessian H(\theta) = -I(\theta)$

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- $I(\hat{\theta})$ is the second derivatives at the mode and thus describes the curvature at the mode
- if the mode is inside the parameter space, $I(\hat{\theta})$ is positive
- if θ is a vector, then $I(\theta)$ is a matrix

 BDA3 Ch 4 has an example where it is easy to compute first and second derivatives and there is easy analytic solution to find where the first derivatives are zero

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 - e.g. in R, demo4_1.R:

```
\begin{array}{lll} \mbox{bioassayfun} & <-\mbox{ function}(w,\mbox{ df})\ \{ & \mbox{ } z <-\mbox{ } w[1] + w[2]*df\$x \\ & -\mbox{sum}(\mbox{ df}\$y*(z) - \mbox{ df}\$n*\mbox{log1p}(\mbox{exp}(z))) \\ \} \\ \mbox{theta0} & <-\mbox{ } c(0\,,0) \\ \mbox{optimres} & <-\mbox{ optim}(w0,\mbox{ bioassayfun}\,,\mbox{ } gr=\mbox{NULL},\mbox{ df1}\,,\mbox{ hessian=T)} \\ \mbox{thetahat} & <-\mbox{ optimres}\$\mbox{par} \\ \mbox{Sigma} & <-\mbox{ solve}(\mbox{optimres}\$\mbox{hessian}) \\ \end{array}
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 - uses autodiff for gradients
 - uses finite differences of gradients to compute Hessian

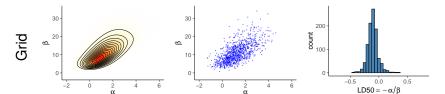
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 - second order autodiff in progress

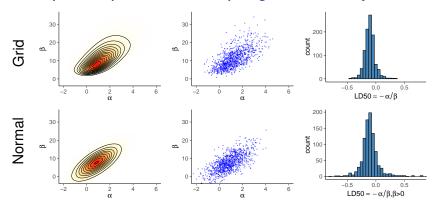
 Optimization and computation of Hessian requires usually much less density evaluations than MCMC

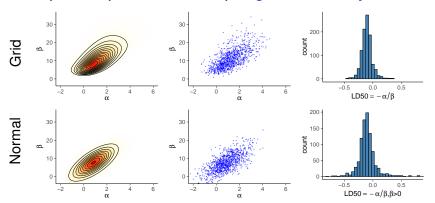
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 - Rasmussen & Williams: Gaussian Processes for Machine Learning
 - CS-E4895 Gaussian Processes (in spring)

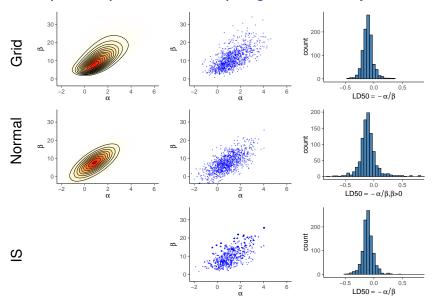
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- Accuracy can be improved by importance sampling (Ch 10)

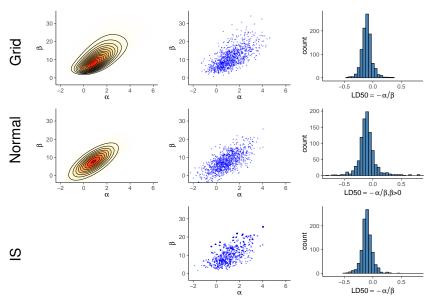






But the normal approximation is not that good here: Grid $sd(LD50) \approx 0.1$, Normal $sd(LD50) \approx .75!$





Grid sd(LD50) \approx 0.1, IS sd(LD50) \approx 0.1

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 - in Bioassay example k = 0.57, which is ok
- CmdStan(R) has Laplace algorithm
 - since version 2.33 (2023)
 - + Pareto-k diagnostic via posterior package
 - + importance resampling (IR) via posterior package

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```
real<lower=, upper=0> theta;
```

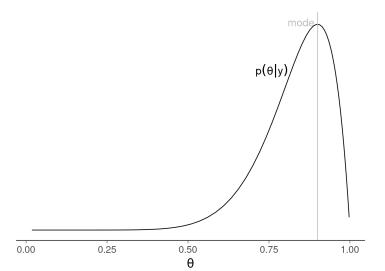
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 - density of the transformed parameter needs to include Jacobian of the transformation (BDA3 p. 21)

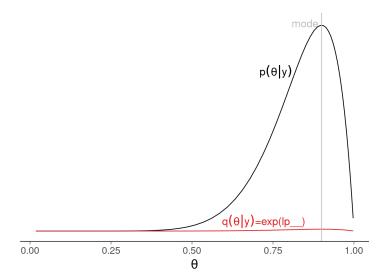
Binomial model $y \sim Bin(\theta, N)$, with data y = 9, N = 10

With Beta(1, 1) prior, the posterior is Beta(9 + 1, 1 + 1)



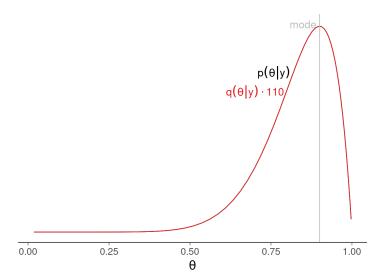
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Stan computes only the unnormalized posterior $q(\theta|y)$



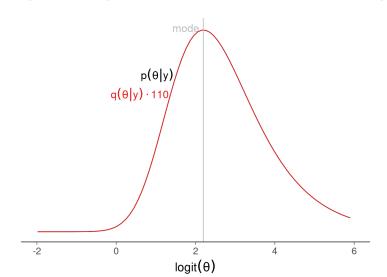
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For illustration purposes we normalize Stan result $q(\theta|y)$

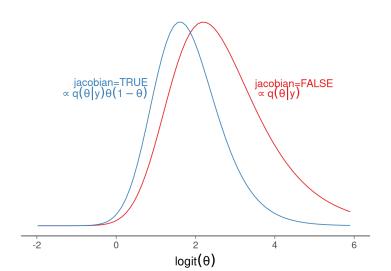


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Beta(9 + 1, 1 + 1), but x-axis shows the unconstrained $logit(\theta)$

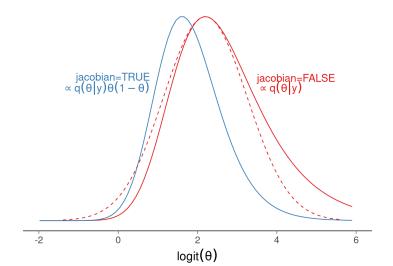


...but we need to take into account the absolute value of the determinant of the Jacobian of the transformation $\theta(1-\theta)$



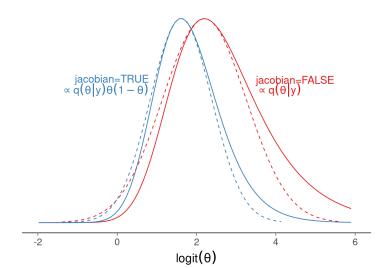
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Let's compare a wrong normal approximation...

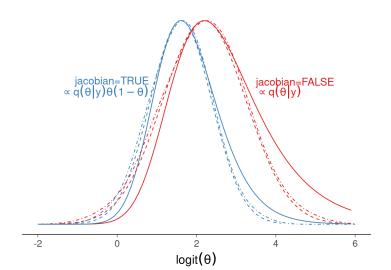


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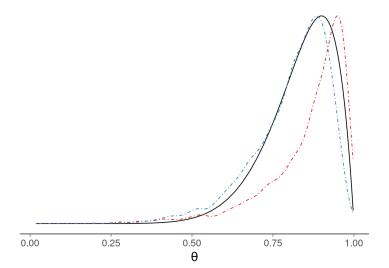
Let's compare a wrong normal approximation and correct one



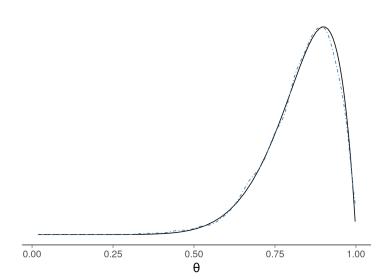
Let's compare a wrong normal approximation and correct one Sample from both approximations and show KDEs for draws



Let's compare a wrong normal approximation and correct one Inverse transform draws and show KDEs



Laplace approximation can be further improved with importance resampling



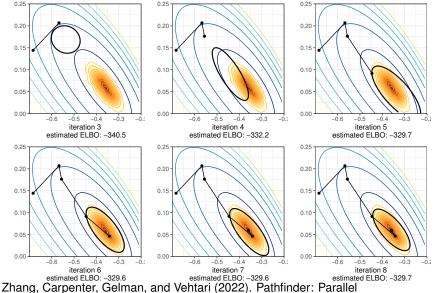
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- Instead of mode and Hessian at mode, e.g.
 - variational inference (Ch 13)
 - CS-E4820 Machine Learning: Advanced Probabilistic Methods
 - CS-E4895 Gaussian Processes
 - Stan has the ADVI algorithm (not very good implementation)
 - Stan has Pathfinder algorithm (CmdStanR github version)
 - instead of normal, methods with flexible flow transformations
 - expectation propagation (Ch 13)
 - speed of these is usually between optimization and MCMC
 - stochastic variational inference can be even slower than MCMC

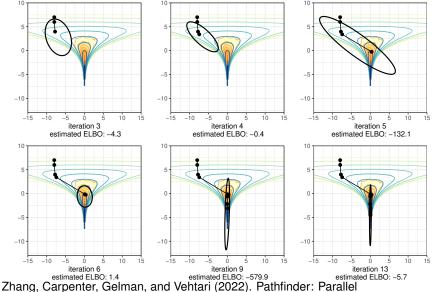
Pathfinder: Parallel quasi-Newton variational inference.



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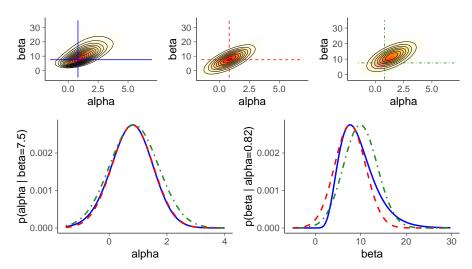
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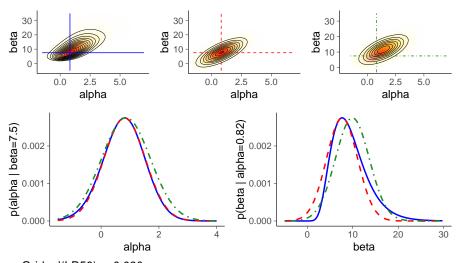
Distributional approximations

Exact, Normal at mode, Normal with variational inference



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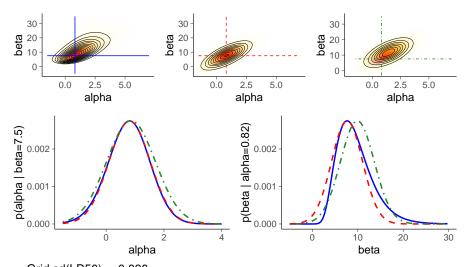
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 - with increasing number of posterior dimensions, the stochastic divergence estimate gets worse and flows have problems, too (Dhaka, Catalina, Andersen, Welandawe, Huggins, and Vehtari, 2021)

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 - see counter examples

- Assume "true" underlying data distribution f(y)
 - observations y₁,..., y_n are independent samples from the joint distribution f(y)
 - "true" data distribution f(y) is not always well defined
 - in the following we proceed as if there were true underlying data distribution
 - for the theory the exact form of f(y) is not important as long at it has certain regularity conditions

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- Problem also for other inference methods like MCMC

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- But a finite data from this data generating process may lack the joint height and weight observations, and thus the the finite data likelihood doesn't have information about ρ

Asymptotic identifiability vs finite data case

- If we randomly would measure both height and weight, asymptotically the correlation ρ would be identifiable
- But a finite data from this data generating process may lack the joint height and weight observations, and thus the the finite data likelihood doesn't have information about ρ
- If the likelihood is weakly informative for some parameters, priors and integration are more important

- If the number of parameter increases as the number of observation increases
 - in some models number of parameters depends on the number of observations
 - e.g. time series models $y_t \sim N(\theta_t, \sigma^2)$ and θ_t has prior in time
 - posterior of θ_t does not converge to a point, if additional observations do not bring enough information

- Aliasing (valetoisto in Finnish)
 - special case of under-identifiability where likelihood repeats in separate points
 - . e.g. mixture of normals

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 For MCMC makes the convergence diagnostics more difficult, as it is difficult to identify aliasing from other multimodality

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 - asymptotic results assume that probability sums to 1
 - e.g. Binomial model, with Beta(0,0) prior and observation y = n
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- Should have a positive prior probability/density where needed

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 - if θ_0 is on the edge of the parameter space, Taylor series expansion has to be truncated, and normal approximation does not necessarily hold
 - e.g. $y_i \sim N(\theta, 1)$ with a restriction $\theta \geq 0$ and assume that $\theta_0 = 0$
 - posterior of θ is left truncated normal distribution with $\mu = \bar{y}$
 - in the limit $n \to \infty$ posterior is half normal distribution
- Can be easy or difficult for MCMC

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 - Calibration
 - α %-posterior interval has the true value in α % cases
 - α %-predictive interval has the true future values in α % cases
 - approximate calibration with shorter intervals for likely true values more important than exact calibration with very bad intervals for all possible values.

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- Confidence interval is defined to have true value inside the interval in $\alpha\%$ cases of repeated data generation from the data generating mechanism
 - doesn't need be useful to have perfect calibration

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- A lot of machine learning is not pure frequentist or Bayesian