## Chapter 5

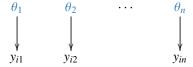
- 5.1 Lead-in to hierarchical models
- 5.2 Exchangeability (useful concept)
- 5.3 Bayesian analysis of hierarchical models (we use Stan/brms for computation)
- 5.4 Hierarchical normal model (we use Stan/brms for computation)
- 5.5 Example: parallel experiments in eight schools (useful discussion on benefits of hierarchical model)
- 5.6 Meta-analysis (can be skipped)
- 5.7 Weakly informative priors for hierarchical variance parameters

### Hierarchical model

- In simple model: posterior for the parameters
- In hierarchical model: posterior for the prior parameters

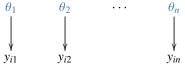
### Hierarchical model

- Example: CVD treatment effectiveness
  - in hospital j the survival probability is  $\theta_i$
  - observations  $y_{ij}$  tell whether patient i survived in hospital j

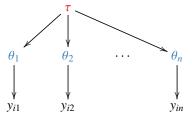


### Hierarchical model

- Example: CVD treatment effectiveness
  - in hospital j the survival probability is  $\theta_j$
  - observations  $y_{ij}$  tell whether patient i survived in hospital j



• sensible to assume that  $\theta_i$  are similar



- natural to think that  $\theta_i$  have common population distribution
- $\theta_j$  is not directly observed and the population distribution is unknown

### Hierarchical model: terms

Level 1: observations given parameters  $p(y_{ij} | \theta_j)$ 

$$p(\theta_j \mid \tau)$$
  $\theta_1$   $\theta_2$  ...  $\theta_n$  parameters  $p(y_{ij} \mid \theta_j)$   $y_{i1}$   $y_{i2}$   $y_{in}$  observations

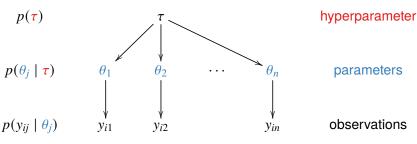
Joint posterior

$$p(\theta, \tau \mid y) \propto p(y \mid \theta, \tau)p(\theta, \tau)$$
  
  $\propto p(y \mid \theta)p(\theta \mid \tau)p(\tau)$ 

### Hierarchical model: terms

Level 1: observations given parameters  $p(y_{ij} \mid \theta_j)$ 

Level 2: parameters given hyperparameters  $p(\theta_j \mid \tau)$ 

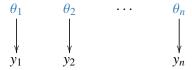


Joint posterior

$$p(\theta, \tau \mid y) \propto p(y \mid \theta, \tau)p(\theta, \tau)$$
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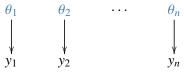
### Compare

"Separate model" (model with separate/independent effects)

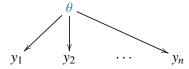


## Compare

"Separate model" (model with separate/independent effects)

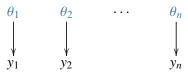


"Joint model" (model with a common effect / pooled model)

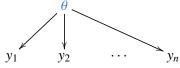


## Compare

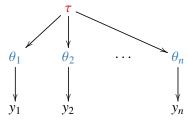
"Separate model" (model with separate/independent effects)



"Joint model" (model with a common effect / pooled model)



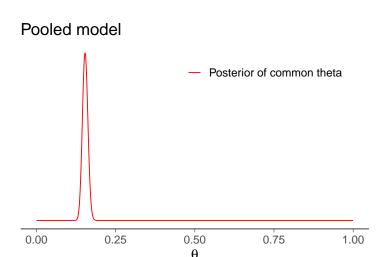
Hierarchical model

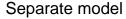


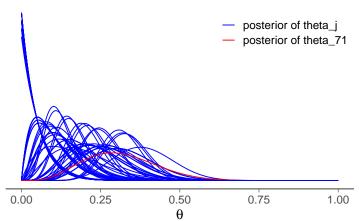
- Medicine testing
- Type F344 female rats in control group given placebo
  - count how many get endometrial stromal polyps
  - familiar binomial model example

- Medicine testing
- Type F344 female rats in control group given placebo
  - count how many get endometrial stromal polyps
  - familiar binomial model example
- Experiment has been repeated 71 times

0/20	0/20	0/20	0/20	0/20	0/20	0/20	0/19	0/19	0/19
0/19	0/18	0/18	0/17	1/20	1/20	1/20	1/20	1/19	1/19
1/18	1/18	2/25	2/24	2/23	2/20	2/20	2/20	2/20	2/20
2/20	1/10	5/49	2/19	5/46	3/27	2/17	7/49	7/47	3/20
3/20	2/13	9/48	10/50	4/20	4/20	4/20	4/20	4/20	4/20
4/20	10/48	4/19	4/19	4/19	5/22	11/46	12/49	5/20	5/20
6/23	5/19	6/22	6/20	6/20	6/20	16/52	15/46	15/47	9/24
4/14									





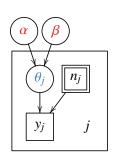


• Hierarchical binomial model for rats prior parameters  $\alpha$  and  $\beta$  are unknown

$$\theta_j \mid \alpha, \beta \sim \text{Beta}(\theta_j \mid \alpha, \beta)$$

$$y_j \mid n_j, \theta_j \sim \text{Bin}(y_j \mid n_j, \theta_j)$$

- Joint posterior  $p(\theta_1, \ldots, \theta_J, \boldsymbol{\alpha}, \boldsymbol{\beta} \mid y)$ 
  - multiple parameters

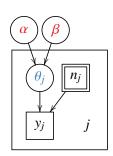


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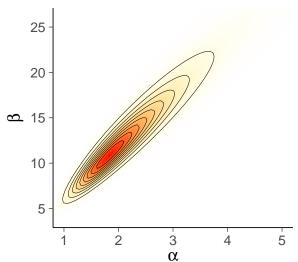
$$y_j \mid n_j, \theta_j \sim \text{Bin}(y_j \mid n_j, \theta_j)$$

- Joint posterior  $p(\theta_1, \dots, \theta_J, \boldsymbol{\alpha}, \boldsymbol{\beta} \mid y)$ 
  - multiple parameters
  - factorize  $\prod_{j=1}^{J} p(\theta_j \mid \alpha, \beta, y) p(\alpha, \beta \mid y)$

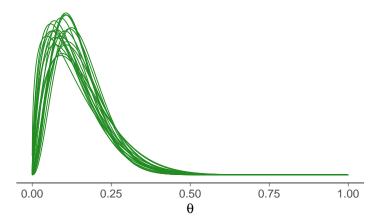


- Population prior Beta( $\theta_j \mid \alpha, \beta$ )
- Hyperprior  $p(\alpha, \beta)$ ?
  - $\alpha, \beta$  both affect the location and scale
  - BDA3 has  $p(\alpha, \beta) \propto (\alpha + \beta)^{-5/2}$ 
    - diffuse prior for location and scale (BDA3 p. 110)
- demo5\_1

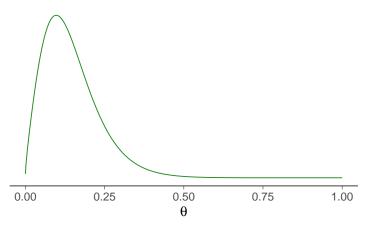
# The marginal of $\alpha$ and $\beta$



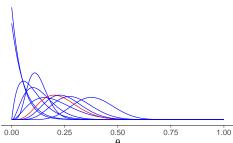
Beta( $\alpha, \beta$ ) given posterior draws of  $\alpha$  and  $\beta$ 



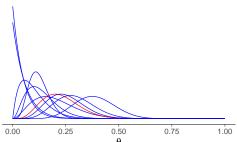
Population distribution (prior) for  $\boldsymbol{\theta}_j$ 



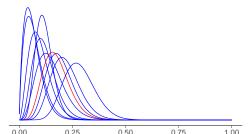






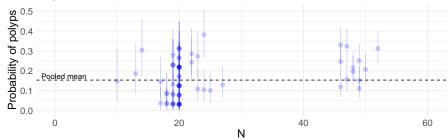


#### Hierarchical model

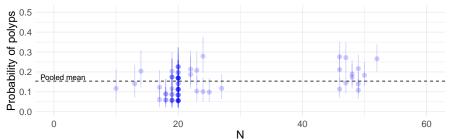


## Hierarchical model and group size: Rats

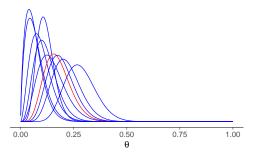




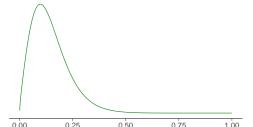
#### Hierarchical



Hierarchical model



Population distribution (prior) for  $\boldsymbol{\theta}_j$ 

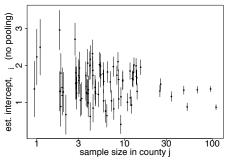


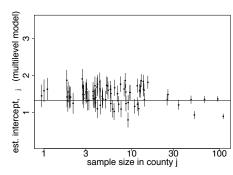
## Hierarchical model and group size: Radon

919 home radon levels in 85 counties in Minnesota:

Separate

Hierarchical





## Diet effect on chicken weights (at age 12 days)

- A typical treatment effect analysis
- Models
  - a separate model, in which each diet is modeled individually
  - a pooled model, in which all measurements are combined and there is no distinction between diets
  - a hierarchical model

```
model {
  // Priors
  for (diet in 1:N_diets) {
    mu_diet[diet] ~ normal(mu_0, sd);
  mu_0 \sim normal(0, 1);
  sd \sim normal(0, 1)
  sigma ~ normal(0, 1);
  // Observation model
  for (obs in 1:N_observations) {
    weight[obs] ~ normal(mu_diet[diet_idx[obs]], sigma);
```

```
model {
  // Priors
  for (diet in 1:N_diets) {
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    // Observation model
    weight ~ normal(mu_diet[diet_idx], sigma);
}
```

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model {
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  sigma ~ normal(0, 1);
  // Observation model
  weight ~ normal(mu_diet[diet_idx], sigma);
brms formula
brm(weight \sim 1 + (1 \mid Diet),
```

```
model {
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  mu_diet ~ normal(mu_0, sd);
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brms formula
brm(weight ~ 1 + (1 | Diet),
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  sigma \sim normal(0, 1);
  // Observation model
  weight ~ normal(mu_diet[diet_idx], sigma);
brms formula
brm(weight ~ 1 + (1 | Diet), data=Chick12,
```

```
model {
 // Priors
 mu_diet ~ normal(mu_0, sd);
 mu_0 \sim normal(0, 1);
 sd \sim normal(0, 1)
  sigma ~ normal(0, 1);
  // Observation model
 weight ~ normal(mu_diet[diet_idx], sigma);
brms formula
brm(weight ~ 1 + (1 | Diet), data=Chick12,
    prior=c(prior(normal(0,1), class="Intercept"), # p(mu_0)
            prior(normal(0,1), class="sd"),
                                            # p(tau)
            prior(normal(0,1), class="sigma"))
                                                  # p(sigma)
```

#### Stan vs brms

#### Stan without loops

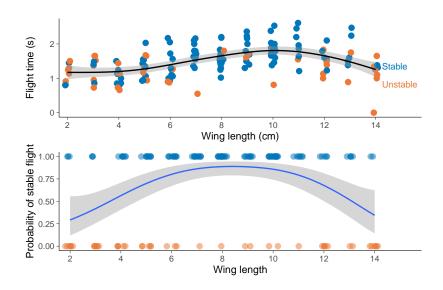
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```

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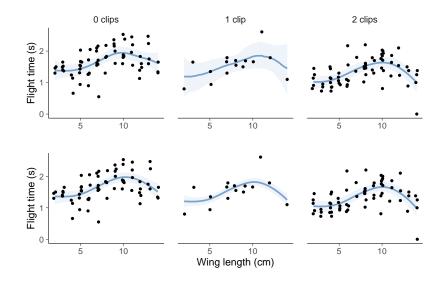
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```

# Paper helicopters



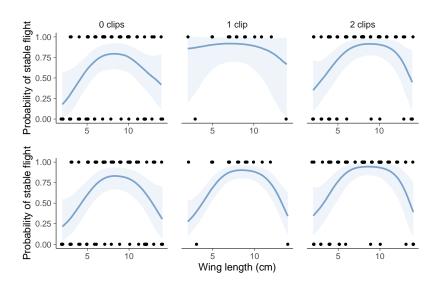
## Paper helicopters: flight time

#### Separate model vs. hierarchical model



# Paper helicopters: stability

#### Separate model vs. hierarchical model



## Paper helicopters: brms

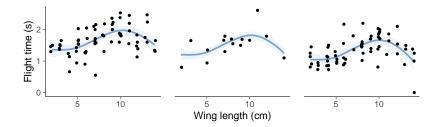
#### Flight time

```
flight_time ~ s(wing_length) + s(wing_length, by = nclips)
```

# Paper helicopters: brms

#### Flight time

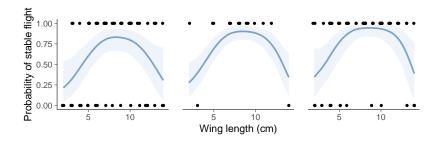
```
flight_time \sim s(wing_length) + s(wing_length, by = nclips)
```



# Paper helicopters: brms

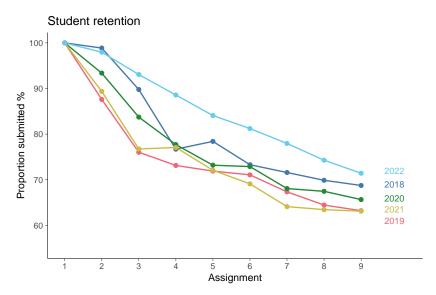
#### Stability

```
stable_flight ~ s(wing_length) + s(wing_length, by = nclips),
family = bernoulli()
```



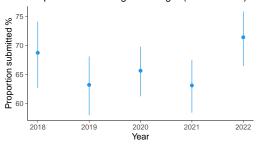
#### Student retention

Was year 2022 better than earlier year?



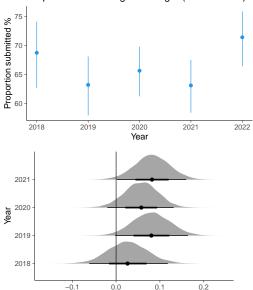
# Student retention separate model

Proportion submitting 9th assign. (90% interval)



# Student retention separate model

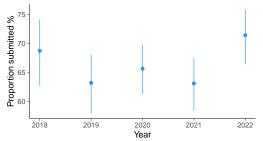
Proportion submitting 9th assign. (90% interval)



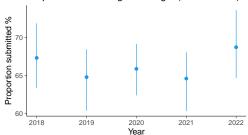
2022 retention was better →

## Student retention separate vs hierarchical model

Proportion submitting 9th assign. (90% interval)

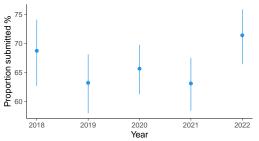




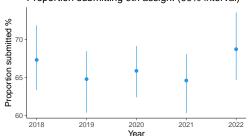


### Student retention separate vs hierarchical model

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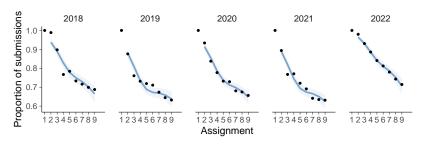






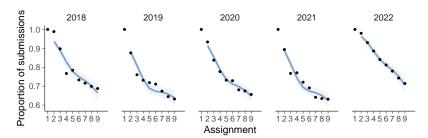
nstudents | trials(nstudents1) ~ 1 + (1 | year), family=binomial()

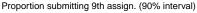
### Student retention latent spline model

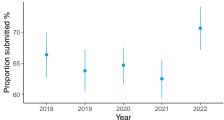


```
nstudents \mid trials(nstudents1) \sim s(assignment, k=4) + (assignment \mid year), \\ family=binomial()
```

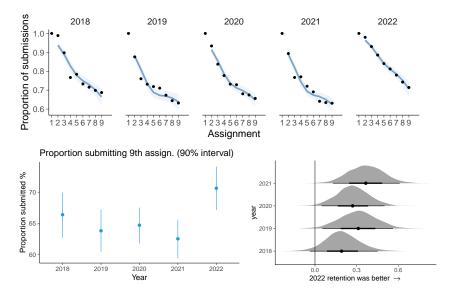
## Student retention latent spline model

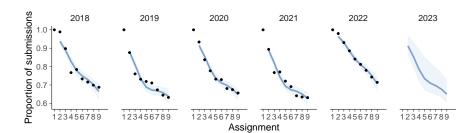


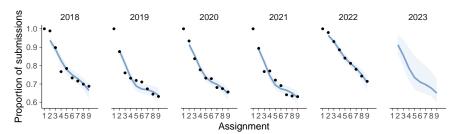


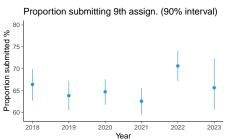


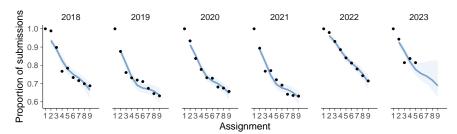
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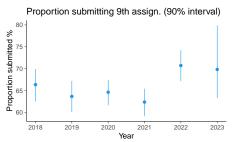


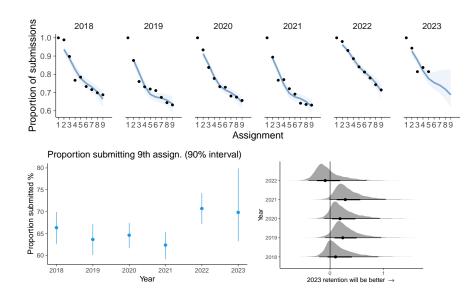












# Centered vs non-centered parameterization

HMC divergences are more likely when using hierarchical models

Hierarchical model code from the course demos

Hierarchical model code from the course demos

```
data {
  int<lower=0> N;
                                    // number of observations
                                    // number of groups
  int<lower=0> K;
  array[N] int<lower=1, upper=K> x; // discrete group indicators
                                    // real valued observations
  vector[N] y;
parameters {
  real mu0:
                                    // prior mean
  real<lower=0> sigma0;
                                    // prior std constrained to be pos.
  vector[K] mu;
                                    // group means
  real<lower=0> sigma;
                                    // common std constrained to be pos.
```

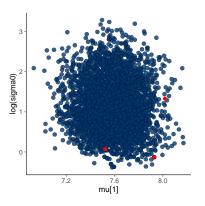
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  real<lower=0> sigma0;
                                    // prior std constrained to be pos.
  vector[K] mu;
                                    // group means
  real<lower=0> sigma;
                                    // common std constrained to be pos.
model {
 mu0 ~ normal(10, 10);
                                    // weakly informative prior
  sigma0 \sim normal(0, 10);
                                    // weakly informative prior
 mu ~ normal(mu0, sigma0);
                                 // population prior with unknown param.
  sigma ~ lognormal(0, .5);
                                 // weakly informative prior
  y ~ normal(mu[x], sigma);
                            // observation model
```

First data with many observations per group: 3 summer months with each having 71 observations.

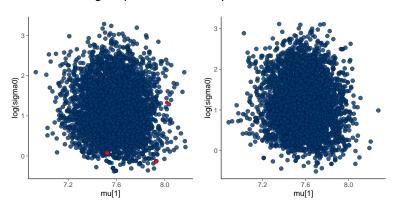
First data with many observations per group: 3 summer months with each having 71 observations.

A few divergences that are not clustered.



First data with many observations per group: 3 summer months with each having 71 observations.

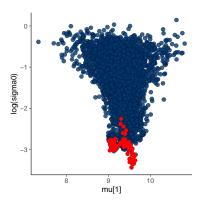
And decreasing step size a little helps.



Second data with a few observations per group: 71 years with each having 3 observations.

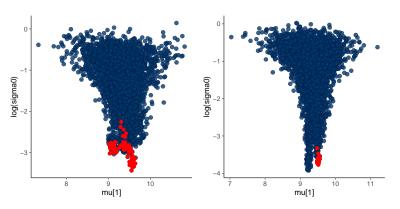
Second data with a few observations per group: 71 years with each having 3 observations.

Many divergences that are clustered.



Second data with a few observations per group: 71 years with each having 3 observations.

And decreasing step size doesn't remove the problem.



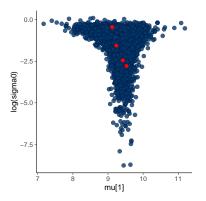
#### Transformation

```
parameters {
 real mu0;
                                   // prior mean
 real<lower=0> sigma0;
                                   // prior std constrained to be pos.
 vector[K] z;
                                   // latent variable
                                   // common std constrained to be pos.
 real<lower=0> sigma;
transformed parameters {
 vector[K] mu = mu0 + sigma0 * z; // group means
model {
 mu0 ~ normal(10, 10); // weakly informative prior
 sigma0 ~ normal(0, 10); // weakly informative prior
 z ~ normal(0, 1); // unit normal
 sigma ~ lognormal(0, .5); // weakly informative prior
 y ~ normal(mu[x], sigma); // observation model
```

Second data with a few observations per group: 71 years with each having 3 observations.

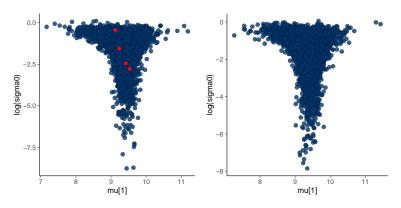
Second data with a few observations per group: 71 years with each having 3 observations.

A few divergences that are not clustered.



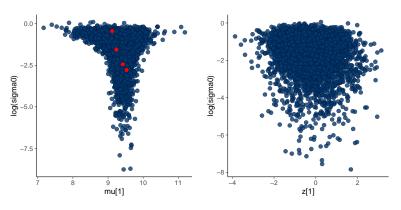
Second data with a few observations per group: 71 years with each having 3 observations.

And decreasing step size a little helps.



Second data with a few observations per group: 71 years with each having 3 observations.

Because we're actually sampling z and not mu



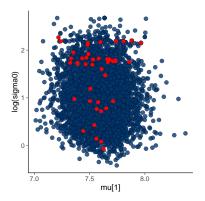
#### No free lunch

- non-centered parameterization is good when likelihood is weak
- non-centered parameterization is bad when likelihood is strong

First data with many observations per group: 3 summer months with each having 71 observations.

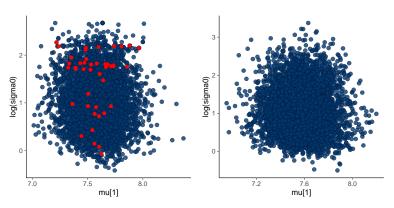
First data with many observations per group: 3 summer months with each having 71 observations.

Many divergences that are not clustered.



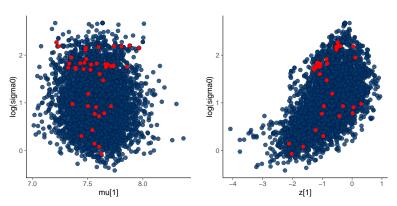
First data with many observations per group: 3 summer months with each having 71 observations.

But decreasing step size a lot helps.

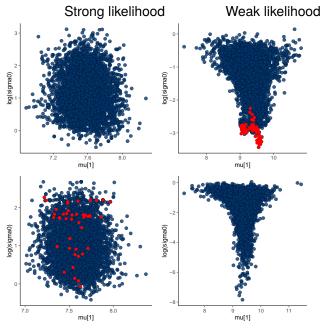


First data with many observations per group: 3 summer months with each having 71 observations.

Now the posterior for z is problematic.



# Centered vs. non-centered parameterization



#### brms and rstanarm

- brms and rstanarm use non-centered parameterization
  - as hierarchical models and Bayesian inference is most useful when likelihood is weak
- There can be need for both centered and non-centered parameterization in the same model
  - automation not easy, but research goes on

- Justifies why we can use
  - a joint model for data
  - a joint prior for a set of parameters
- Less strict than independence

- Exchangeability: Parameters  $\theta_1, \ldots, \theta_J$  (or observations  $y_1, \ldots, y_J$ ) are exchangeable if the joint distribution p is invariant to the permutation of indices  $(1, \ldots, J)$
- e.g.

$$p(\theta_1, \theta_2, \theta_3) = p(\theta_2, \theta_3, \theta_1)$$

• Exchangeability implies symmetry: If there is no information which can be used *a priori* to separate  $\theta_j$  form each other, we can assume exchangeability. ("Ignorance implies exchangeability")

- Exchangeability does not mean that the results of the experiments could not be different
  - e.g. if we know that the experiments have been in two different laboratories, and we know that the other laboratory has better conditions for the rats, but we do not know which experiments have been made in which laboratory
  - a priori experiments are exchangeable
  - model could have unknown parameter for the laboratory with a conditional prior for rats assumed to come form the same place (clustering model)

# Exchangeability and additional information

- Example: bioassay
  - $y_i$  number of dead animals are not exchangeable alone

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  - $x_i$  dose is additional information

## Exchangeability and additional information

- Example: bioassay
  - y<sub>i</sub> number of dead animals are not exchangeable alone
  - x<sub>i</sub> dose is additional information
  - $(x_i, y_i)$  exchangeable and logistic regression was used

$$p(\alpha, \beta \mid y, n, x) \propto \prod_{i=1}^{n} p(y_i \mid \alpha, \beta, n_i, x_i) p(\alpha, \beta)$$

- Example: hierarchical rats example
  - all rats not exchangeable

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  - all rats not exchangeable
  - in a single laboratory rats exchangeable

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- Example: hierarchical rats example
  - all rats not exchangeable
  - in a single laboratory rats exchangeable
  - laboratories exchangeable
  - → hierarchical model

# Partial or conditional exchangeability

- Conditional exchangeability
  - if y<sub>i</sub> is connected to an additional information x<sub>i</sub>, so that y<sub>i</sub> are not exchangeable, but (y<sub>i</sub>, x<sub>i</sub>) exchangeable use joint model or conditional model (y<sub>i</sub> | x<sub>i</sub>).

#### Partial or conditional exchangeability

- Conditional exchangeability
  - if  $y_i$  is connected to an additional information  $x_i$ , so that  $y_i$  are not exchangeable, but  $(y_i, x_i)$  exchangeable use joint model or conditional model  $(y_i | x_i)$ .
- Partial exchangeability
  - if the observations can be grouped (a priori), then use hierarchical model

• The simplest form of the exchangeability (but not the only one) for the parameters  $\theta$  conditional independence

$$p(x_1,\ldots,x_J\mid\theta)=\prod_{i=1}^J p(x_i\mid\theta)$$

#### Exchangeability - Counter example

- A six sided die with probabilities  $\theta_1, \ldots, \theta_6$ 
  - without additional knowledge  $\theta_1, \ldots, \theta_6$  exchangeable
  - due to the constraint  $\sum_{j=1}^6 \theta_j$ , parameters are not independent and thus joint distribution can not be presented as iid

 See more examples in the BDA3 notes - Exchangeability vs. independence