

Particle Filters

an overview

Matthias Mühlich

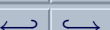
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Particle Filters a tutorial

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1 Introduction

An increasing number of researchers is using a family of techniques and algorithms called

- *condensation algorithms*
- *bootstrap filtering*
- *particle filters*
- *interacting particle approximations*
- *sequential Monte Carlo methods*
- *SIS, SIR, ASIR, RPF, ...*

Time scale: last 10 years [e.g. Isard & Blake 1996; Kitagawa 1996; Gordon, Salmond & Smith 1993]

The question of this talk is: What is behind all that?



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General Classification of Filter Strategies

Gaussian models:

- Kalman filter
- extended Kalman filter
- linear-update filter / linear regression filter / statistical linearization filter
 - unscented filter
 - central difference filter
 - divided difference filter
- assumed density filter / moment matching



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Mixture of Gaussian models:

- assumed density filter / pseudo-Bayes
- Gaussian-sum filter

Nonparametric models:

- **particle filter class**
- histogram filter



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Some Basic Remarks

- various applications: computer vision (i.e. tracking), control theory, econometrics (stock markets, monetary flow, interest rates), ...
- we deal with discrete time systems only
- no out-of-sequence measurements
- we are mainly interested in estimating the state at time k from measurements up to time $k' = k$ (opposite: smoothing ($k' > k$) and prediction ($k' < k$); furthermore k' need not be fixed...)
- no restrictions to linear processes or Gaussian noise!



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Overview of this Talk

- The Dynamic System Model
 - Bayesian Filter Approach
 - Optimal and Suboptimal Solutions
 - The Particle Filter
 - Experiments and Summary
-
- states of a system and state transition equation
 - measurement equation



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Overview of this Talk

- The Dynamic System Model
 - **Bayesian Filter Approach**
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- estimation of the state
 - probabilistic modelling
 - Bayesian filter



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Overview of this Talk

- The Dynamic System Model
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- **Optimal and Suboptimal Solutions**
- The Particle Filter
- Experiments and Summary

- filtered pdf can be written down easily, but it is not always tractable (\rightarrow ugly integrals ...)
- conditions under which optimal solutions exist: Kalman filter and grid-based filter
- what can be done in other cases: suboptimal approaches



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- standard particle filter
 - various improved versions



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Overview of this Talk

- The Dynamic System Model
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- Optimal and Suboptimal Solutions
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- Experiments and Summary

– some experimental data and conclusion



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2 Dynamic System

A dynamic system can be modelled with two equations:

State Transition or Evolution Equation

$$\mathbf{x}_k = f_k(\mathbf{x}_{k-1}, \mathbf{u}_{k-1}, \mathbf{v}_{k-1})$$

$f(\cdot, \cdot, \cdot)$: evolution function (possible non-linear)

$\mathbf{x}_k, \mathbf{x}_{k-1} \in \mathbb{R}^{n_x}$: current and previous state

$\mathbf{v}_{k-1} \in \mathbb{R}^{n_v}$: state noise (usually *not* Gaussian)

$\mathbf{u}_{k-1} \in \mathbb{R}^{n_u}$: known input

Note: state only depends on previous state, i.e. first order Markov process



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Measurement Equation

$$\mathbf{z}_k = h_k(\mathbf{x}_k, \mathbf{u}_k, \mathbf{n}_k)$$

$h(\cdot, \cdot, \cdot)$: measurement function (possible non-linear)

$\mathbf{z}_k \in \mathbb{R}^{n_z}$: measurement

$\mathbf{x}_k \in \mathbb{R}^{n_x}$: state

$\mathbf{n}_k \in \mathbb{R}^{n_n}$: measurement noise (usually *not* Gaussian)

$\mathbf{u}_k \in \mathbb{R}^{n_u}$: known input

(dimensionality of state, measurement, input, state noise, and measurement noise can all be different!)



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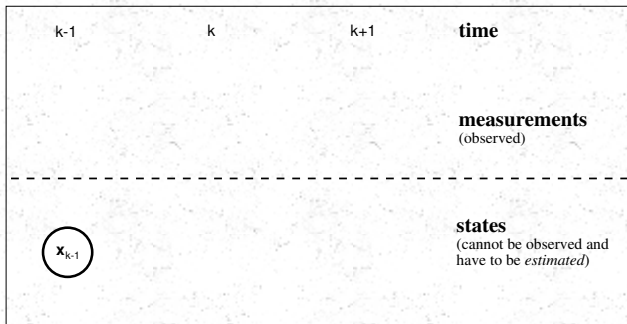


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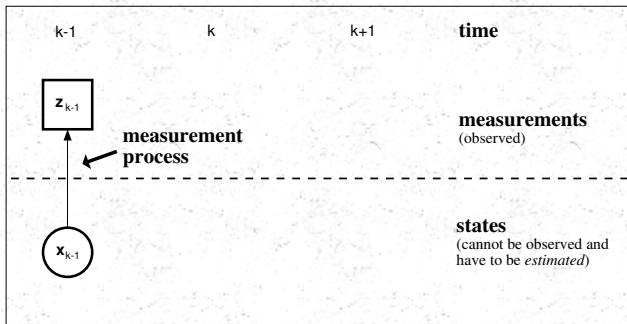


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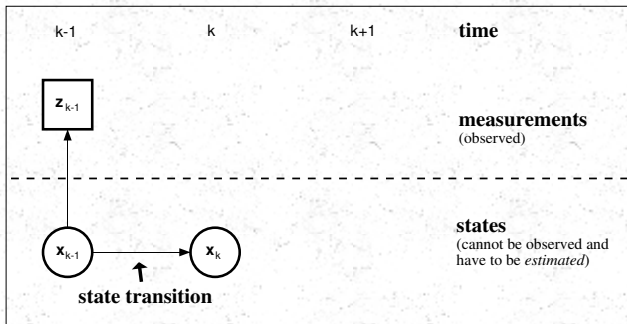


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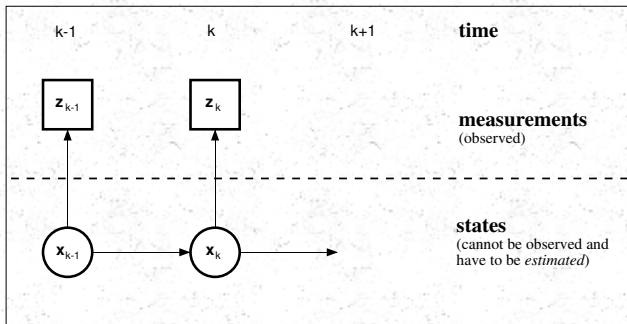


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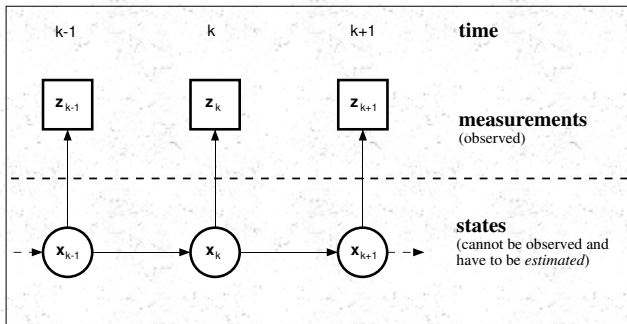


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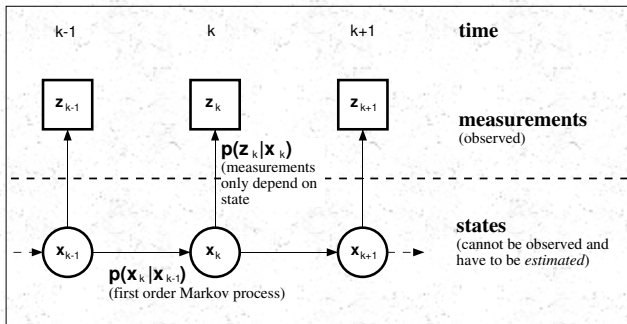


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Assumptions:

The observations are conditionally independent given the state: $p(\mathbf{z}_k | \mathbf{x}_k)$.

Hidden Markov Model (HMM):

$p(\mathbf{x}_0)$ given and $p(\mathbf{x}_k | \mathbf{x}_{k-1})$ defines state transition probability for $k \geq 1$.



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3 Bayesian Filters

Estimating the Posterior

Bayesian approach: We attempt to construct the posterior pdf of the state given all measurements.

⇒ can be termed a complete solution to the estimation problem because all available information is used; from the pdf, an optimal estimate can theoretically be found for any criterion.

in detail: We seek estimates of \mathbf{x}_k based on all available measurements up to time k (abbreviated as $\mathbf{z}_{1:k}$) by constructing the posterior $p(\mathbf{x}_k|\mathbf{z}_{1:k})$.

Assumption: initial state pdf (prior) $p(\mathbf{x}_0)$ is given



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The Use of Knowing the Posterior

Let $f_k : \mathbb{R}^{(k+1) \times n_x} \rightarrow \mathbb{R}$ be any arbitrary (integrable) function that can depend

- on all components of the state \mathbf{x}
- on the whole trajectory in state-space

Examples: This function can be an estimator for the current state or for future observations.

Then we can compute its expectation using

$$\mathbb{E}[f_k(\mathbf{x}_{0:k})] = \int f(\mathbf{x}_{0:k}) p(\mathbf{x}_{0:k} | \mathbf{z}_{1:k}) d\mathbf{x}_{0:k}$$

MMSE estimate of state: $\hat{\mathbf{x}} = \mathbb{E}[\mathbf{x}_k]$. Other estimates that can be computed: median, modes, confidence intervals, kurtosis, ...



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Recursive Filters

recursive filters (i.e. sequential update of previous estimate) \leftrightarrow batch processing (computation with all data in one step)

not only faster: allow on-line processing of data (lower storage costs, rapid adaption to changing signals characteristics)

essentially consist of two steps:

prediction step: $p(\mathbf{x}_{k-1}|\mathbf{z}_{1:k-1}) \rightarrow p(\mathbf{x}_k|\mathbf{z}_{1:k-1})$
(usually deforms / translates / spreads state pdf due to noise)

update step: $p(\mathbf{x}_k|\mathbf{z}_{1:k-1}), \mathbf{z}_k \rightarrow p(\mathbf{x}_k|\mathbf{z}_{1:k})$
(combines likelihood of current measurement with predicted state; usually concentrates state pdf)



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General Prediction-Update Framework

Assume that pdf $p(\mathbf{x}_{k-1}|\mathbf{z}_{1:k-1})$ is available at time $k-1$.

Prediction step: (using Chapman-Kolmogoroff equation)

$$p(\mathbf{x}_k|\mathbf{z}_{1:k-1}) = \int p(\mathbf{x}_k|\mathbf{x}_{k-1})p(\mathbf{x}_{k-1}|\mathbf{z}_{1:k-1})d\mathbf{x}_{k-1} \quad (1)$$

This is the prior of the state \mathbf{x}_k at time k *without knowledge of the measurement \mathbf{z}_k* , i.e. the probability given only previous measurements.

Update step: (compute posterior pdf from predicted prior pdf and new measurement)

$$p(\mathbf{x}_k|\mathbf{z}_{1:k}) = \frac{p(\mathbf{z}_k|\mathbf{x}_k)p(\mathbf{x}_k|\mathbf{z}_{1:k-1})}{p(\mathbf{z}_k|\mathbf{z}_{1:k-1})} \quad (2)$$



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Let us prove formula (2) (just in order to train calculations with joint and conditional probabilities...)

$$\begin{aligned} & p(\mathbf{x}_k | \mathbf{z}_{1:k}) \\ &= \frac{p(\mathbf{z}_{1:k} | \mathbf{x}_k) p(\mathbf{x}_k)}{p(\mathbf{z}_{1:k})} \end{aligned}$$

(Bayes rule)



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Let us prove formula (2) (just in order to train calculations with joint and conditional probabilities. . .)

$$\begin{aligned} & p(\mathbf{x}_k | \mathbf{z}_{1:k}) \\ &= \frac{p(\mathbf{z}_{1:k} | \mathbf{x}_k) p(\mathbf{x}_k)}{p(\mathbf{z}_{1:k})} \\ &= \frac{p(\mathbf{z}_k, \mathbf{z}_{1:k-1} | \mathbf{x}_k) p(\mathbf{x}_k)}{p(\mathbf{z}_k, \mathbf{z}_{1:k-1})} \end{aligned}$$

(separate $p(\mathbf{z}_{1:k})$ into $p(\mathbf{z}_k, \mathbf{z}_{1:k-1})$)



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Let us prove formula (2) (just in order to train calculations with joint and conditional probabilities. . .)

$$\begin{aligned} & p(\mathbf{x}_k | \mathbf{z}_{1:k}) \\ &= \frac{p(\mathbf{z}_{1:k} | \mathbf{x}_k) p(\mathbf{x}_k)}{p(\mathbf{z}_{1:k})} \\ &= \frac{p(\mathbf{z}_k, \mathbf{z}_{1:k-1} | \mathbf{x}_k) p(\mathbf{x}_k)}{p(\mathbf{z}_k, \mathbf{z}_{1:k-1})} \\ &= \frac{p(\mathbf{z}_k | \mathbf{z}_{1:k-1}, \mathbf{x}_k) p(\mathbf{z}_{1:k-1} | \mathbf{x}_k) p(\mathbf{x}_k)}{p(\mathbf{z}_k | \mathbf{z}_{1:k-1}) p(\mathbf{z}_{1:k-1})} \end{aligned}$$

(factorize joint probability: $p(a, b|c) = p(a|b, c) \cdot p(b|c)$ and $p(a, b) = p(a|b) \cdot p(b)$)



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$$\begin{aligned} & p(\mathbf{x}_k | \mathbf{z}_{1:k}) \\ &= \frac{p(\mathbf{z}_{1:k} | \mathbf{x}_k) p(\mathbf{x}_k)}{p(\mathbf{z}_{1:k})} \\ &= \frac{p(\mathbf{z}_k, \mathbf{z}_{1:k-1} | \mathbf{x}_k) p(\mathbf{x}_k)}{p(\mathbf{z}_k, \mathbf{z}_{1:k-1})} \\ &= \frac{p(\mathbf{z}_k | \mathbf{z}_{1:k-1}, \mathbf{x}_k) p(\mathbf{z}_{1:k-1} | \mathbf{x}_k) p(\mathbf{x}_k)}{p(\mathbf{z}_k | \mathbf{z}_{1:k-1}) p(\mathbf{z}_{1:(k-1)})} \\ &= \frac{p(\mathbf{z}_k | \mathbf{z}_{1:k-1}, \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{z}_{1:k-1}) p(\mathbf{z}_{1:k-1}) p(\mathbf{x}_k)}{p(\mathbf{z}_k | \mathbf{z}_{1:k-1}) p(\mathbf{z}_{1:k-1}) p(\mathbf{x}_k)} \end{aligned}$$

(Bayes rule)



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Let us prove formula (2) (just in order to train calculations with joint and conditional probabilities...)

$$\begin{aligned} & p(\mathbf{x}_k | \mathbf{z}_{1:k}) \\ &= \frac{p(\mathbf{z}_{1:k} | \mathbf{x}_k) p(\mathbf{x}_k)}{p(\mathbf{z}_{1:k})} \\ &= \frac{p(\mathbf{z}_k, \mathbf{z}_{1:k-1} | \mathbf{x}_k) p(\mathbf{x}_k)}{p(\mathbf{z}_k, \mathbf{z}_{1:k-1})} \\ &= \frac{p(\mathbf{z}_k | \mathbf{z}_{1:k-1}, \mathbf{x}_k) p(\mathbf{z}_{1:k-1} | \mathbf{x}_k) p(\mathbf{x}_k)}{p(\mathbf{z}_k | \mathbf{z}_{1:k-1}) p(\mathbf{z}_{1:(k-1)})} \\ &= \frac{p(\mathbf{z}_k | \mathbf{z}_{1:k-1}, \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{z}_{1:k-1}) p(\mathbf{z}_{1:k-1}) p(\mathbf{x}_k)}{p(\mathbf{z}_k | \mathbf{z}_{1:k-1}) p(\mathbf{z}_{1:k-1}) p(\mathbf{x}_k)} \\ &= \frac{p(\mathbf{z}_k | \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{z}_{1:k-1})}{p(\mathbf{z}_k | \mathbf{z}_{1:k-1})} \end{aligned}$$

(independence of observations; cancelling out terms)



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The Structure of the Update Equation

$$p(\mathbf{x}_k | \mathbf{z}_{1:k}) = \frac{p(\mathbf{z}_k | \mathbf{x}_k) \cdot p(\mathbf{x}_k | \mathbf{z}_{1:k-1})}{p(\mathbf{z}_k | \mathbf{z}_{1:k-1})}$$

$$\text{posterior} = \frac{\text{likelihood} \cdot \text{prior}}{\text{evidence}}$$

prior: given by prediction equation

likelihood: given by observation model

evidence: the normalizing constant in the denominator

$$p(\mathbf{z}_k | \mathbf{z}_{1:k-1}) = \int p(\mathbf{z}_k | \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{z}_{1:k-1}) d\mathbf{x}_k$$



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This theoretically allows an optimal Bayesian solution (in the sense of computing the posterior pdf).

Problem: only a conceptual solution; integrals are not tractable.

But: in some restricted cases, an optimal solution is possible. Two optimal solutions (under restrictive assumptions):

- (standard) Kalman filter
- grid-based filter



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4 Kalman Filter

Introduction

Assumptions:

- posterior at time $k-1$, i.e. $p(\mathbf{x}_{k-1}|\mathbf{z}_{k-1})$, is Gaussian
- dynamic system characterized by

$$\mathbf{x}_k = \mathbf{F}_k \mathbf{x}_{k-1} + \mathbf{G}_k \mathbf{v}_{k-1}$$

$$\mathbf{z}_k = \mathbf{H}_k \mathbf{x}_k + \mathbf{J}_k \mathbf{n}_k$$

- both noise vectors Gaussian (covariance matrices are \mathbf{Q}_{k-1} and \mathbf{R}_k)

Then new posterior $p(\mathbf{x}_k|\mathbf{z}_k)$ is Gaussian, too, and can be computed using simple linear equations.



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optimal solution, but *highly restrictive* assumptions must hold

Prediction Equation

At time $k-1$: $p(\mathbf{x}_{k-1}|\mathbf{z}_{1:k-1}) = \mathcal{N}(\mathbf{m}_{k-1|k-1}, \mathbf{P}_{k-1|k-1})$

Inserting into (1) yields

$$p(\mathbf{x}_k|\mathbf{z}_{1:k-1}) = \mathcal{N}(\mathbf{m}_{k|k-1}, \mathbf{P}_{k|k-1})$$

with

$$\mathbf{m}_{k|k-1} = \mathbf{F}_k \mathbf{m}_{k-1|k-1}$$

and

$$\mathbf{P}_{k|k-1} = \mathbf{G}_k \mathbf{Q}_{k-1} \mathbf{G}_k^T + \mathbf{F}_k \mathbf{P}_{k-1|k-1} \mathbf{F}_k^T$$



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Update Equation

Inserting into (2) yields

$$p(\mathbf{x}_k | \mathbf{z}_{1:k}) = \mathcal{N}(\mathbf{m}_{k|k}, \mathbf{P}_{k|k})$$

with

$$\mathbf{m}_{k|k} = \mathbf{m}_{k|k-1} + \mathbf{K}_k (\mathbf{z}_k - \underbrace{\mathbf{H}_k \mathbf{m}_{k|k-1}}_{\text{estimated } \hat{\mathbf{z}}_k})$$

and

$$\mathbf{P}_{k|k} = \mathbf{P}_{k|k-1} - \mathbf{K}_k \mathbf{H}_k \mathbf{P}_{k|k-1}$$

Kalman Gain:

$$\mathbf{K}_k = \mathbf{P}_{k|k-1} \mathbf{H}_k^T \underbrace{(\mathbf{H}_k \mathbf{P}_{k|k-1} \mathbf{H}_k^T + \mathbf{J}_k \mathbf{R}_k \mathbf{J}_k^T)^{-1}}_{\text{Cov}[\hat{\mathbf{z}}_k]}$$



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5 Grid-Based Filter

Introduction

Assumptions:

- state space is discrete
- number of different states (N_s) is limited
(Note: implicitly includes discreteness)

Suppose at time $k-1$ we have states \mathbf{x}^i with $i = 1, \dots, N_s$.
Conditional probability of these states:

$$Pr(\mathbf{x}_{k-1} = \mathbf{x}^i | \mathbf{z}_{1:k-1}) = w_{k-1|k-1}^i$$

Then the (old) posterior at time $k-1$ is given by:

$$p(\mathbf{x}_{k-1} | \mathbf{z}_{1:k-1}) = \sum_{i=1}^{N_s} w_{k-1|k-1}^i \delta(\mathbf{x}_{k-1} - \mathbf{x}^i)$$



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Results (Summary)

Both the (new) prior and the (new) posterior have the same structure: a sum of weighted Dirac peaks:

$$p(\mathbf{x}_k | \mathbf{z}_{1:k-1}) = \sum_{i=1}^{N_s} w_{k|k-1}^i \delta(\mathbf{x}_{k-1} - \mathbf{x}^i)$$

$$p(\mathbf{x}_k | \mathbf{z}_{1:k}) = \sum_{i=1}^{N_s} w_{k|k}^i \delta(\mathbf{x}_{k-1} - \mathbf{x}^i)$$

Note: extension to different sets of states for each time step

$$\{\mathbf{x}^i\} : i = 1, \dots, N_s \quad \longrightarrow \quad \{\mathbf{x}_k^i\} : i = 1, \dots, N_{s,k}$$

with time-varying index k is easily possible; the 'allowed' states need not be constant.



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Prediction Equation (in Detail)

Inserting into (1) yields

$$\begin{aligned} p(\mathbf{x}_k | \mathbf{z}_{1:k-1}) &= \sum_{i=1}^{N_s} \sum_{j=1}^{N_s} w_{k-1|k-1}^j p(\mathbf{x}^i | \mathbf{x}^j) \delta(\mathbf{x}_{k-1} - \mathbf{x}^i) \\ &= \sum_{i=1}^{N_s} w_{k|k-1}^i \delta(\mathbf{x}_{k-1} - \mathbf{x}^i) \end{aligned}$$

where $w_{k|k-1}^i = \sum_{j=1}^{N_s} w_{k-1|k-1}^j p(\mathbf{x}^i | \mathbf{x}^j)$

(new) prior weights = old posterior weights,
reweighted using state transition probabilities



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Update Equation (in Detail)

Inserting into (2) yields

$$p(\mathbf{x}_k | \mathbf{z}_{1:k}) = \sum_i^{N_s} w_{k|k}^i \delta(\mathbf{x}_{k-1} - \mathbf{x}^i)$$

where $w_{k|k}^i = \frac{w_{k|k-1}^i p(\mathbf{z}_k | \mathbf{x}^i)}{\sum_j^{N_s} w_{k|k-1}^j p(\mathbf{z}_k | \mathbf{x}^j)}$.

Note: denominator only needed for normalization

posterior weights = prior weights, reweighted using likelihoods



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6 Particle Filter

Suboptimal Approximations

If we want to preserve Kalman filter principle. . .

- Extended Kalman Filter (EKF)
- Unscented Kalman Filter (UKF)

...we get better results,

BUT: *we cannot get rid off Gaussian approximations*



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EKF / UKF:

All these approaches fail if we have

- bimodal / multimodal pdfs
- heavily skewed pdfs

We need a more general scheme to tackle these problems.



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Particle Filter – General Concept

Many different names (do you remember the **introduction?**) but the general concept is rather simple:

PARTICLE FILTER:

If we cannot solve the integrals required for a Bayesian recursive filter analytically . . . we represent the posterior probabilities by a set of randomly chosen weighted samples.

Note: “randomly chosen” \equiv “Monte Carlo”
(we are playing roulette / throwing the dice)

Increasing number of samples \Rightarrow (almost sure) convergence to true pdf



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Sequential Importance Sampling (SIS)

SIS is the basic framework for most particle filter algorithms. Let

$\{\mathbf{x}_{0:k}^i\}$: set of support points (samples, particles)

$i = 1, \dots, N_s$

(whole trajectory for each particle!)

w_k^i : associated weights, normalized to $\sum_i w_k^i = 1$

Then:

$$p(\mathbf{x}_k | \mathbf{z}_{1:k}) \approx \sum_{i=1}^{N_s} w_k^i \delta(\mathbf{x}_{0:k} - \mathbf{x}_{0:k}^i)$$

(discrete weighted approximation to the true posterior)



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SIS (continued)

Usually we cannot draw samples \mathbf{x}_k^i from $p(\cdot)$ directly. Assume we sample directly from a (different) *importance function* $q(\cdot)$. Our approximation is still correct (up to normalization) if

$$w_k^i \propto \frac{p(\mathbf{x}_{0:k}^i | \mathbf{z}_{1:k})}{q(\mathbf{x}_{0:k}^i | \mathbf{z}_{1:k})}$$

The trick: we can choose $q(\cdot)$ freely!

If the importance function is chosen to factorize such that

$$q(\mathbf{x}_{0:k} | \mathbf{z}_{1:k}) = q(\mathbf{x}_k | \mathbf{x}_{0:k-1}, \mathbf{z}_{1:k}) q(\mathbf{x}_{0:k-1} | \mathbf{z}_{1:k-1})$$

then one can augment old particles $\mathbf{x}_{0:k-1}^i$ by $\mathbf{x}_k \sim q(\mathbf{x}_k | \mathbf{x}_{0:k-1}, \mathbf{z}_{1:k})$ to get new particles $\mathbf{x}_{0:k}^i$.



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SIS (continued)

Weight update (after some lengthy computations...):

$$w_k^i = w_{k-1}^i \frac{p(\mathbf{z}_k | \mathbf{x}_k^i) p(\mathbf{x}_k^i | \mathbf{x}_{k-1}^i)}{q(\mathbf{x}_k^i | \mathbf{x}_{0:k-1}^i, \mathbf{z}_{1:k})} \quad (3)$$

Furthermore, if $q(\mathbf{x}_k | \mathbf{x}_{0:k-1}, \mathbf{z}_{1:k}) = q(\mathbf{x}_k | \mathbf{x}_{k-1}, \mathbf{z}_{1:k})$
(only dependent on *last* state and observations):

$$p(\mathbf{x} | \mathbf{z}_{1:k}) \approx \sum_{i=1}^{N_s} w_k^i \delta(\mathbf{x}_k - \mathbf{x}_k^i)$$

(and we need not preserve trajectories $\mathbf{x}_{0:k-1}^i$ and history of observations $\mathbf{z}_{1:k-1}$)



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SIS Algorithm – Pseudo Code

```
[ { $\mathbf{x}_k^i, w_k^i$ } $_{i=1}^{N_s}$  ] = SIS( { $\mathbf{x}_{k-1}^i, w_{k-1}^i$ } $_{i=1}^{N_s}$ ,  $\mathbf{z}_k$  )  
  FOR  $i = 1 : N_s$   
    draw  $\mathbf{x}_k^i \sim q(\mathbf{x}_k | \mathbf{x}_{k-1}^i, \mathbf{z}_k)$   
    update weights according to (3)  
  END FOR  
  normalize weights to  $\sum_{i=1}^{N_s} w_k^i = 1$ 
```



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PROBLEM: Degeneracy Problem

Problem with SIS approach: after a few iterations, most particles have negligible weight (the weight is concentrated on a few particles only)

Counter measures:

- brute force: many, many samples N_s
- good choice of importance density
- resampling

Note: amount of degeneracy can be estimated based on variance of weights [Liu 1996].



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Optimal Impotance Density:

It can be shown that the optimal importance density is given by

$$q(\mathbf{x}_k | \mathbf{x}_{k-1}, \mathbf{z}_k)_{opt} = p(\mathbf{x}_k | \mathbf{x}_{k-1}, \mathbf{z}_k)$$

Then

$$w_k^i = w_{k-1}^i \int p(\mathbf{z}_k | \mathbf{x}_k') p(\mathbf{x}_k' | \mathbf{x}_{k-1}^i) d\mathbf{x}_k'$$

Two major drawbacks: usually neither sampling from q_{opt} nor solving the integral in w_k^i is possible... (but in some special cases, it works)

Other alternative which is often convenient:

$q(\cdot) = p(\mathbf{x}_k | \mathbf{x}_{k-1})$ (prior). Easy to implement, but does not take measurements into account.



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Resampling Approaches

Basic idea of resampling:

Whenever degeneracy rises above threshold: replace old set of samples (+ weights) with new set of samples (+ weights), such that sample density better reflects posterior pdf.

This eliminates particles with low weight and chooses more particles in more probable regions.

Complexity: possible in $O(N_s)$ operations



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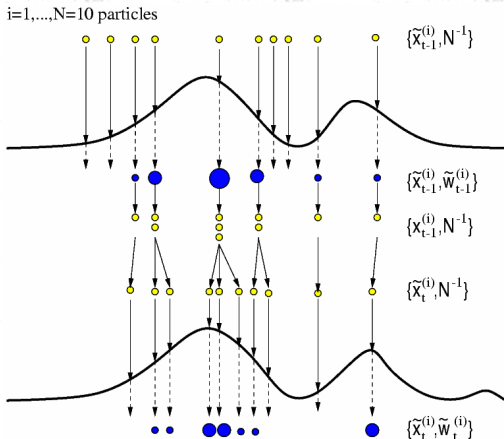
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The resampling principle:



(graphics taken from Van der Merwe et al.)



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General Particle Filter – Pseudo Code

```
[ { $\mathbf{x}_k^i, w_k^i$  }i=1Ns ] = PF ( { $\mathbf{x}_{k-1}^i, w_{k-1}^i$  }i=1Ns ,  $\mathbf{z}_k$  )  
  FOR  $i = 1 : N_s$   
    draw  $\mathbf{x}_k^i \sim q(\mathbf{x}_k | \mathbf{x}_{k-1}^i, \mathbf{z}_k)$   
    update weights according to (3)  
  END FOR  
  normalize weights to  $\sum_{i=1}^{N_s} w_k^i = 1$   
  IF degeneracy too high  
    resample { $\mathbf{x}_k^i, w_k^i$  }i=1Ns  
  END IF
```



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PROBLEM: Loss of Diversity

No degeneracy problem but new problem arises:

Particles with high weight are selected more and more often, others die out slowly
⇒ *loss of diversity or sample impoverishment*

For small process noise, all particles can collapse into a single point within a few iterations.

Other problem: resampling limits the ability to parallelize algorithm.



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Other Particle Filter Variants

Methods to counteract loss of diversity and degeneracy problem:

- resample-move algorithm
- regularization
- Rao-Blackwellisation
- multiple Monte-Carlo

Other particle filter variants found in the literature:

- sampling importance resampling (SIR)
- auxiliary sampling importance resampling (ASIR)
- regularized particle filter (RPF)
- ...



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7 Experiments

see videos...



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8 Summary

First of all: what I did *not* talk about. . .

- speed of convergence
- number of samples needed
- complexity issues / tricks for speed-up of algorithms
- advanced particle filter variants in detail

⇒ refer to the literature if you want to know more



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Advantages of particle filters (PFs):

- can deal with non-linearities
- can deal with non-Gaussian noise
- can be implemented in $O(N_s)$
- mostly parallelizable
- easy to implement
- in contrast to HMM filters (state-space discretized to N fixed states) : PFs focus adaptively on probable regions of state-space



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Thesis:

If you want to solve a filtering problem, then particle filters are the best filters you can use, much better than e.g. Kalman filters.

Right or wrong?



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WRONG!

Particle filters include a random element; they only converge to the true posterior pdf (almost surely) if $N_s \rightarrow \infty$.

Therefore: *If the assumptions for Kalman filters or grid-based filters are valid, no PF can outperform them!*

Additionally: depending on the dynamic model, Gaussian sum filters, unscented Kalman filters or extended Kalman filters may produce satisfactory results at lower computational cost.

(But you should at least try a PF; it is usually better than other suboptimal methods!)



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PF approaches proved their usefulness in a variety of applications.

But:

- choice of importance function $q(\cdot)$ is crucial in PF design
- large sample number N_s increases computational effort
- potential problems: *degeneracy* and *loss of diversity*

If these points are taken into account, then particle filters are an extremely powerful tool for filtering / estimation.

(“black box usage” vs “know what you’re doing!”)



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Thank you!

This presentation was made with \LaTeX .
(try to write București in Powerpoint. . .)



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