

# UNI HW1 : Abhay Pawar (asp2197)

(1)

$$(a) p(x_i = j | \pi, \eta) = \binom{x_i + \eta - 1}{x_i} \pi^{x_i} (1 - \pi)^{\eta - x_i}$$

Since,  $x_i$ 's are i.i.d  $\rightarrow$

$$p(x_1, x_2, \dots, x_N | \pi, \eta) = p(x_1 | \pi, \eta) p(x_2 | \pi, \eta) \dots p(x_N | \pi, \eta)$$

Joint likelihood

$$= \prod_{i=1}^N \binom{x_i + \eta - 1}{x_i} \pi^{\sum_{i=1}^N x_i} (1 - \pi)^{\eta N}$$

(b) Taking  $\ln$  of the joint pdf  $\rightarrow$

$$\ln(L) = \sum_{i=1}^N \ln \left[ \binom{x_i + \eta - 1}{x_i} \right] + \ln(\pi) \sum_{i=1}^N x_i + \eta N \ln(1 - \pi)$$

To find MLE w.r.t.  $\pi \rightarrow$

$$\frac{\partial \ln(L)}{\partial \pi} = 0 + \frac{\sum_{i=1}^N x_i}{\pi} + \frac{\eta N (-1)}{(1 - \pi)} = 0$$

$$\therefore \sum x_i - \pi \sum x_i = \eta N \pi$$

$$\hat{\pi}_{MLE} = \frac{\sum x_i}{(\eta N + \sum x_i)}$$

(c)  $\pi_{MAP} = \arg \max_{\pi} \ln p(\pi | x_i's, \eta)$

$$= \arg \max_{\pi} [\ln(x_i's | \pi, \eta) + \ln p(\pi)]$$

Ignoring third term because it is not dependent on  $\pi$ .

$$p(\pi) = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \pi^{a-1} (1-\pi)^{b-1}$$

$$\therefore \pi_{MAP} = \arg \max_{\pi} \left[ \ln(\pi) \sum_{i=1}^N x_i + \eta N \ln(1 - \pi) + (a-1) \ln(\pi) + (b-1) \ln(1 - \pi) \right]$$

Excluded all terms which were not a function of  $\pi$



Taking derivative -

$$\frac{\sum_{i=1}^N x_i}{\pi} - \frac{nN}{(1-\pi)} + \frac{a-1}{\pi} - \frac{(b-1)}{(1-\pi)} = 0$$

$$\frac{\sum x_i + a - 1}{\pi} - \frac{(nN + b - 1)}{(1-\pi)} = 0$$

$$\therefore \pi_{MAP} = \frac{\sum x_i + a - 1}{\sum x_i + a + b + nN - 2}$$

$$(d) p(\pi | x_i's, n) = \frac{p(x_i's | \pi, n) p(\pi)}{p(x_i | n)} = I$$

Writing only the numerator =  ~~$\frac{1}{n!} \left( \sum x_i + n - 1 \right)!$~~   $p(x_i's | \pi, n) p(\pi)$

$$= \frac{n!}{\prod_{i=1}^n x_i!} \pi^{\sum x_i} (1-\pi)^{nN} \times \frac{\Gamma(a+b) \pi^{a-1} (1-\pi)^{b-1}}{\Gamma(a)\Gamma(b)}$$

$$= K \times \pi^{\sum x_i + a - 1} (1-\pi)^{nN + b - 1}$$

(K is a term not involving  $\pi$ )

Also, the denominator of I does not depend on  $\pi$

$$\therefore p(\pi | x_i's, n) = K_1 \pi^{\sum x_i + a - 1} (1-\pi)^{nN + b - 1}$$

( $K_1$  is not a func of  $\pi$ ).

This is again a beta func. ~~with~~ with

$$a' = \sum x_i + a, \quad b' = nN + b$$

$$\therefore K_1 = \frac{\Gamma(a' + b')}{\Gamma(a') \Gamma(b')}$$

$$\therefore p(\pi | x_i's, n) = \frac{\Gamma(\sum x_i + nN + a + b)}{\Gamma(\sum x_i + a) \Gamma(nN + b)} \pi^{\sum x_i + a - 1} (1-\pi)^{nN + b - 1}$$

(e) For beta dist,  $\mu = a/(a+b)$  &  $\text{Var} = ab/[(a+b)^2(a+b+1)]$

$$\therefore E(\pi) = \frac{\sum x_i + a}{(\sum x_i + a + nN + b)}$$



$$\text{Var}(\pi) = \frac{(\sum x_i + a)(\sum N + b)}{(\sum x_i + a + \sum N + b)^2 (\sum x_i + a + b + \sum N + 1)}$$

Comparison -

$$E(\pi) = \frac{\sum x_i + a}{\sum x_i + a + \sum N + b} = \frac{\sum x_i (1 + a/\sum x_i)}{(\sum x_i + \sum N) (1 + \frac{a+b}{\sum x_i + \sum N})}$$

$$= \hat{\pi}_{ML} \times \left[ \frac{1 + a/\sum x_i}{1 + (a+b)/(\sum x_i + \sum N)} \right]$$

$\therefore E(\pi) > \hat{\pi}_{ML}$  when the multiplying factor is  $> 1$ .

Similarly,

$$E(\pi) = \pi_{MAP} \times \left[ \frac{1 + (a-1)/\sum x_i}{1 + (a+b-2)/(\sum x_i + \sum N)} \right]$$

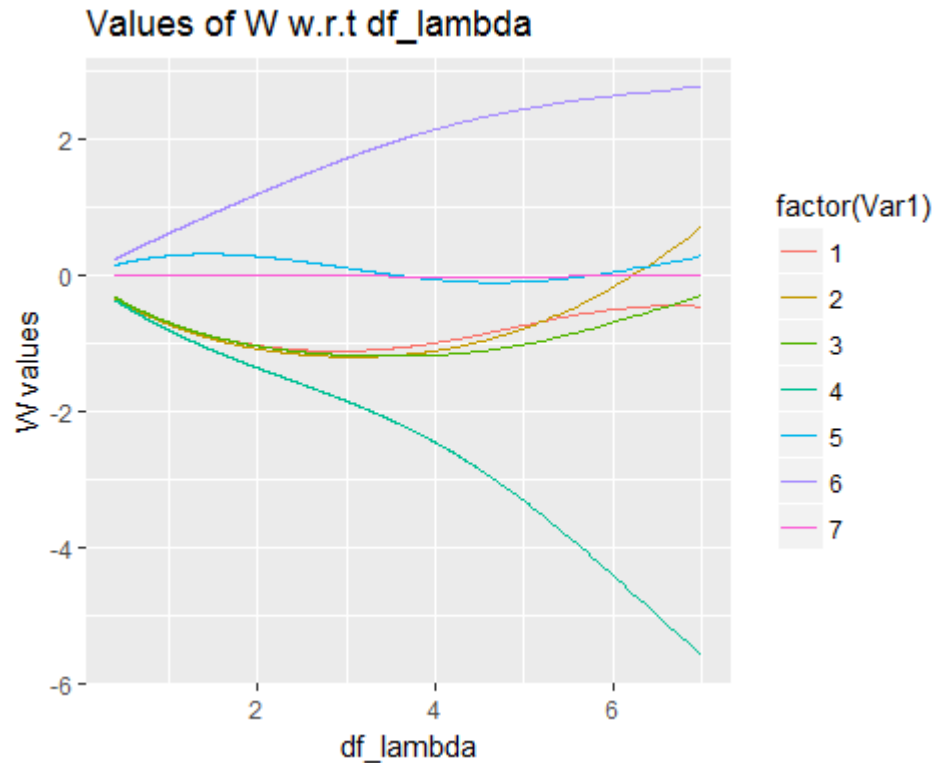
For  $N \rightarrow \infty$ ,

$$E(\pi) = \hat{\pi}_{ML} = \pi_{MAP}$$

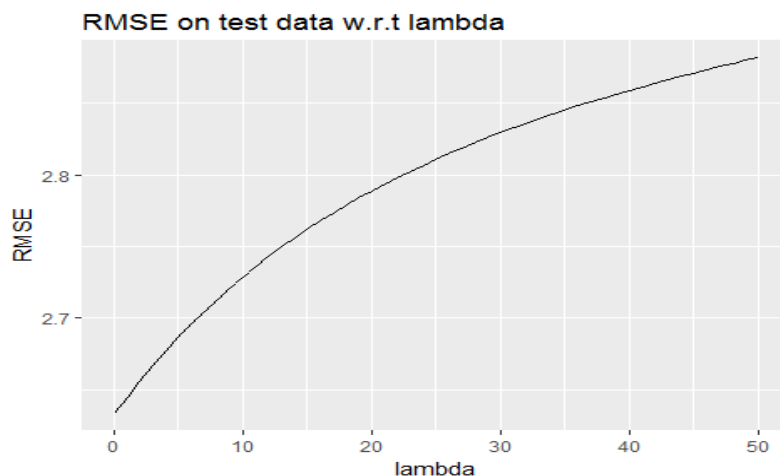
## Problem 2(coding):

### Part 1:

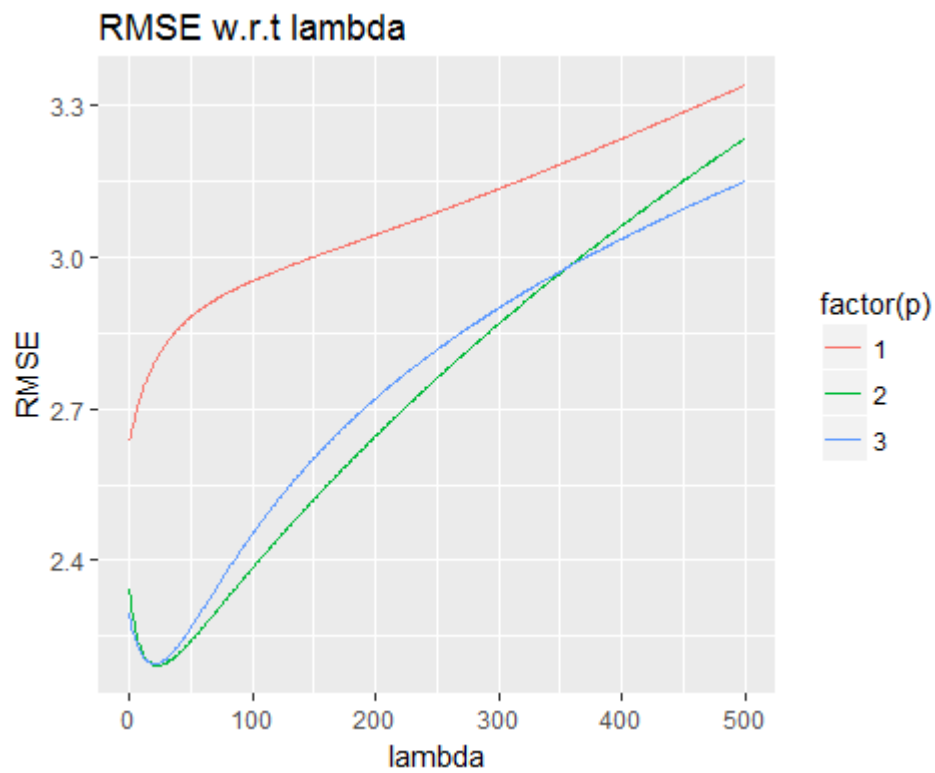
(a) & (b) Values for 4<sup>th</sup> and 6<sup>th</sup> dimension are the highest in magnitude. Since, all the dimensions are normalized,  $w$  of a dimension will be high when it is highly correlated with the dependent variable (miles/gallon) and can predict the value of dependent variable well. Dimension 4 (car weight) has a negative  $w$ , implying negative correlation and Dimension 6 has a positive  $w$ , implying positive correlation.



(c)  $\lambda=0$  gives the least value of RMSE and hence, least squares should be used. Ridge regression has worse RMSE than Least Squares solution.



(d) I will **choose  $p=2$**  because it gives minimum RMSE for some value of  $\lambda$  as compared to  $p=1$  and  $p=3$ . The **ideal value of  $\lambda$**  is no more zero and it is **21**. It implies that ridge regression with  $\lambda=21$  works better than just the OLS solution for  $p=2$ .



```

library(readr)
library(reshape2)
X_test <- read_csv("G:/Acads/ML for data science/hw1-data/X_test.csv",
                  col_names = FALSE)
X_train <- read_csv("G:/Acads/ML for data science/hw1-data/X_train.csv",
                  col_names = FALSE)
y_test <- read_csv("G:/Acads/ML for data science/hw1-data/y_test.csv",
                  col_names = FALSE)
y_train <- read_csv("G:/Acads/ML for data science/hw1-data/y_train.csv",
                  col_names = FALSE)

X_test_mat=as.matrix(X_test)
X_train_mat=as.matrix(X_train)
y_test_mat=as.matrix(y_test)
y_train_mat=as.matrix(y_train)

lambda=c(0:5000)
df_lambda=array(dim=length(lambda))
RMSE=array(dim=length(lambda))
W_mat=matrix(nrow=ncol(X_train),ncol=length(lambda))

#Calculating df_lambda, W_mat, RMSE on Y_test and y_pred on test data
for (i in 1:length(lambda)){
  W_mat[,i]=solve(lambda[i]*diag(ncol(X_train_mat))+((t(X_train_mat))%*%X_train_mat))%*%t(X_train_mat)%*%y_train_mat
  df_lambda[i]=sum(diag(X_train_mat%*%solve(lambda[i]*diag(ncol(X_train_mat))+t(X_train_mat)%*%X_train_mat)%*%t(X_train_mat)))
  y_pred=X_test_mat%*%W_mat[,i]
  RMSE[i]=sqrt(sum((y_pred-y_test_mat)^2)/length(y_pred))
}

W_mat_melt <- melt(W_mat, id=c())
W_mat_melt$df_lambda=0
for (i in 1:length(W_mat_melt$df_lambda)){
  W_mat_melt$df_lambda[i]=df_lambda[W_mat_melt$Var2[i]]
}
df_lambda[W_mat_melt$Var2[10]]
W_mat_melt$Var2[10]

ggplot(data=W_mat_melt)+geom_line(mapping=aes(y=value,x=df_lambda,colour=factor(Var1)))+
  labs(title="Values of W w.r.t df_lambda",x="df_lambda",y="W values")

rmse_frame=data.frame(RMSE,lambda)
ggplot(data=rmse_frame[1:51,])+geom_line(aes(y=RMSE,x=lambda))+labs(title="RMSE on test data w.r.t lambda")

#p=2
X_train_mat_cr=matrix(nrow=nrow(X_train_mat),ncol=13)
X_train_mat_cr[,1:7]=X_train_mat
X_test_mat_cr=matrix(nrow=nrow(X_test_mat),ncol=13)
X_test_mat_cr[,1:7]=X_test_mat
W_mat_cr=matrix(nrow=ncol(X_train_mat_cr),ncol=length(lambda))
for (i in 1:6){

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X_train_mat_cr[,i+7]=X_train_mat[,i]^2
X_test_mat_cr[,i+7]=X_test_mat[,i]^2
}
for (i in 1:length(lambda)){
  W_mat_cr[,i]=solve(lambda[i]*diag(ncol(X_train_mat_cr))+
((t(X_train_mat_cr))%*%X_train_mat_cr))%*%t(X_train_mat_cr)%*%y_train_mat
  #df_lambda[i]=sum(diag(X_train_mat_cr%*%solve(lambda[i]*diag(ncol(X_train_mat_cr))+t(X_train_m
at_cr)%*%X_train_mat_cr)%*%t(X_train_mat_cr)))
  y_pred=X_test_mat_cr%*%W_mat_cr[,i]
  RMSE[i]=sqrt(sum((y_pred-y_test_mat)^2)/length(y_pred))
}

rmse_frame_cr=data.frame(RMSE,lambda)
ggplot(data=rmse_frame_cr[1:501,])+geom_line(aes(y=RMSE,x=lambda))+labs(title="RMSE on test data
(p=2) w.r.t lambda")

#p=3
X_train_mat_cr3=matrix(nrow=nrow(X_train_mat),ncol=19)
X_train_mat_cr3[,1:13]=X_train_mat_cr
X_test_mat_cr3=matrix(nrow=nrow(X_test_mat),ncol=19)
X_test_mat_cr3[,1:13]=X_test_mat_cr
W_mat_cr3=matrix(nrow=ncol(X_test_mat_cr3),ncol=length(lambda))
for (i in 1:6){
  X_train_mat_cr3[,i+13]=X_train_mat[,i]^3
  X_test_mat_cr3[,i+13]=X_test_mat[,i]^3
}
for (i in 1:length(lambda)){
  W_mat_cr3[,i]=solve(lambda[i]*diag(ncol(X_train_mat_cr3))+((t(X_train_mat_cr3))%*%X_train_mat_
cr3))%*%t(X_train_mat_cr3)%*%y_train_mat
  #df_lambda[i]=sum(diag(X_train_mat_cr3%*%solve(lambda[i]*diag(ncol(X_train_mat_cr3))+t(X_train_
mat_cr3)%*%X_train_mat_cr3)%*%t(X_train_mat_cr3)))
  y_pred=X_test_mat_cr3%*%W_mat_cr3[,i]
  RMSE[i]=sqrt(sum((y_pred-y_test_mat)^2)/length(y_pred))
}

rmse_frame_cr3=data.frame(RMSE,lambda)
ggplot(data=rmse_frame_cr3[1:501,])+geom_line(aes(y=RMSE,x=lambda))+labs(title="RMSE on test dat
a(p=3) w.r.t lambda")
#Ideal value of lambda=21

rmse_frame$p=1
rmse_frame_cr$p=2
rmse_frame_cr3$p=3
rmse_all=rbind(rmse_frame[1:501,],rmse_frame_cr[1:501,],rmse_frame_cr3[1:501,])
ggplot(data=rmse_all)+geom_line(aes(y=RMSE,x=lambda,colour=factor(p)))+labs(title="RMSE w.r.t la
mbda")

```