

Non-parametric tests

12.2, 12.4 of Ross Textbook

Non-parametric tests

- We make no assumptions of the form of the distribution function unlike previous cases where we assume Normal or Binomial.
- Generically denote distribution as $F(X)$. Note form of $F(X)$ is not known
- Possible hypothesis that can be tested in such cases
 - What is the median of $F(X)$?
 - Sign test
 - Is the distribution around the median similar
 - Sign rank test
 - Given samples of two distributions: Are they likely to be from the same or different distributions?
 - Two sample test

Sign test

Let X_1, \dots, X_n denote a sample from a continuous distribution F and suppose that we are interested in testing the hypothesis that the median of F , call it m , is equal to a specified value m_0 . That is, consider a test of

$$H_0 : m = m_0 \quad \text{versus} \quad H_1 : m \neq m_0$$

where m is such that $F(m) = .5$.

$$T = \frac{\# \text{ of } X_i\text{'s less than } m_0}{n} \quad \hat{m} = \text{median of } (X_1, \dots, X_n) \quad \hat{F}(x) = \frac{1}{n} \sum_{i=1}^n \mathbb{I}(X_i \leq x)$$

$$\underbrace{P_{H_0}(T)} \sim \text{Binomial}(n, \frac{1}{2}) \quad T = \hat{F}(m_0)$$

T-statistic

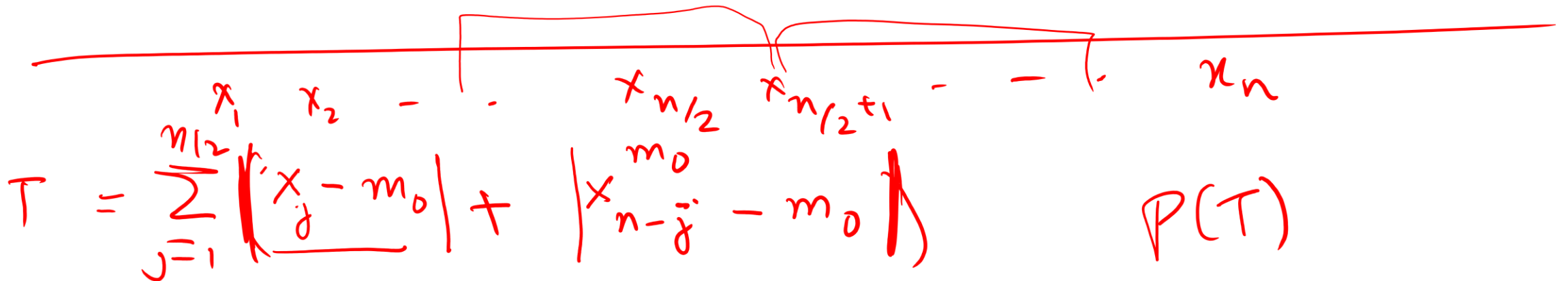
- T = sum of the sign of $\underline{m}_0 - X_i$

$$\underline{T} = \sum_i \underline{I}_i \quad \underline{I}_i = \begin{cases} 1 & \text{if } X_i < m_0 \\ 0 & \text{if } X_i \geq m_0 \end{cases}$$

$$P(I_i) \sim \text{Bernoulli}(\frac{1}{2})$$

- What is the distribution of T under the null hypothesis?

$$P_{H_0}(T) \sim \text{Binomial}(n, \frac{1}{2})$$


$$T = \sum_{j=1}^{n/2} |X_j - m_0| + \sum_{j=1}^{n/2} |X_{n-j+1} - m_0| \quad P(T)$$

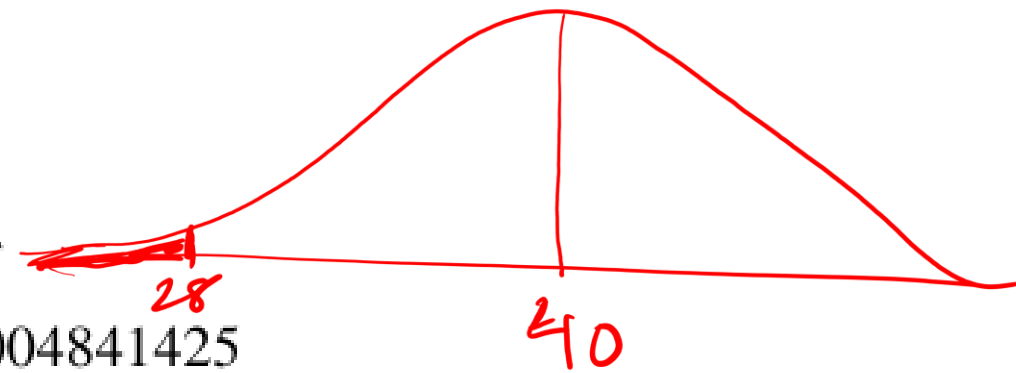
Example 12.2.b. A financial institution has decided to open an office in a certain community if it can be established that the median annual income of families in the community is greater than \$90,000. To obtain information, a random sample of 80 families was chosen, and the family incomes determined. If 28 of these families had annual incomes below and 52 had annual incomes above \$90,000, is this significant enough to establish, say, at the 5 percent level of significance, that the median annual income in the community is greater than \$90,000?

$$\begin{aligned} m_0 &= 90^k \\ n &= 80 \\ T &= 28 \end{aligned}$$

Solution. We need to see if the data are sufficient to enable us to reject the null hypothesis when testing

$$H_0 : m \leq 90 \quad \text{versus} \quad H_1 : m > 90$$

$$p\text{-value} = P(\text{Bin}(80, 1/2) \leq 28) = \underline{\text{pbinom}(28, 80, 1/2)} = \underline{0.004841425}$$



and so the null hypothesis that the median income is less than or equal to \$90,000 is rejected. ■

Signed rank test

Given n sample X_1, \dots, X_n from unknown distribution F , we are interested in the hypothesis that F is symmetric about a given median m_0 , that is,

- $H_0: P(X > m_0 + a) = P(X < m_0 - a)$, for all a

Let $Y_i = X_i - m_0$, $i = 1, \dots, n$ and rank (that is, order) the absolute values $|Y_1|, |Y_2|, \dots, |Y_n|$. Set, for $j = 1, \dots, n$,

$$I_j = \begin{cases} 1 & \text{if the } j\text{th smallest value comes from a data value that is smaller} \\ & \text{than } m_0 \\ 0 & \text{otherwise} \end{cases}$$

test statistic

$$T = \sum_{j=1}^n j I_j$$

Example 12.3.a. If $n = \underline{4}$, $m_0 = \underline{2}$, and the data values are $X_1 = 4.2$, $X_2 = 1.8$, $X_3 = 5.3$, $X_4 = 1.7$, then the rankings of $|X_i - 2|$ are $\underline{.2}$, $\underline{.3}$, $\underline{2.2}$, $\underline{3.3}$. Since the first of these values — namely, $\underline{.2}$ — comes from the data point X_2 , which is less than 2, it follows that $I_1 = 1$. Similarly, $I_2 = 1$, and I_3 and I_4 equal 0. Hence, the value of the test statistic is $T = 1 + 2 = 3$. ■

X_i	4.2	1.8	5.3	1.7
	- 2	- 2	- 2	- 2
	<hr style="width: 50%; margin: 0 auto;"/>			
Y_i	2.2	-0.2	3.3	-0.3
$ Y_i $	0.2	0.3	2.2	0.3
	1	1	0	0
	=			

$$P_{H_0}(T \leq 3)$$

Distribution of test statistic $P_{H_0}(T)$ under the null hypothesis?

Expected value and variance of T under H_0

smaller - probability that the j th absolute difference is from an $x_k < m_0$.

$$P\{\underline{I_j} = \underline{1}\} = \underline{\frac{1}{2}} = P\{I_j = 0\}, \quad j = 1, \dots, n \quad E[I_j] = \frac{1}{2}, \quad \underline{\text{Var}(I_j) = \frac{1}{4}}$$

Hence, we can conclude that under H_0 ,

$$\underline{E[T]} = E \left[\sum_{j=1}^n j I_j \right] \\ = \sum_{j=1}^n \frac{j}{2} = \underline{\frac{n(n+1)}{4}}$$

$$\text{Var}(T) = \text{Var} \left(\sum_{j=1}^n j I_j \right) \\ = \sum_{j=1}^n j^2 \text{Var}(I_j) \\ = \sum_{j=1}^n \frac{j^2}{4} = \underline{\frac{n(n+1)(2n+1)}{24}}$$

$P_{H_0}(T)$ = approximately normal for large n with mean and variance as above. But we can do better..

An exact computation of probability $P_{H_0}(T)$ recursively

$$\underline{P_k(i)} = \underline{P_{H_0}} \left\{ \sum_{j=1}^k jI_j \leq \underline{i} \right\} \quad P_{H_0}(T \leq i) = P_{H_0}($$

$$= P_{H_0} \left\{ \sum_{j=1}^k jI_j \leq i \mid \underline{I_k = 1} \right\} \underline{P_{H_0}\{I_k = 1\}}$$

$$+ \underline{P_{H_0}} \left\{ \sum_{j=1}^k jI_j \leq i \mid \underline{I_k = 0} \right\} \underline{P_{H_0}\{I_k = 0\}},$$

$$= P_{H_0} \left\{ \sum_{j=1}^{k-1} jI_j \leq i - \underline{k} \mid \underline{I_k = 1} \right\} \underline{P_{H_0}\{I_k = 1\}}$$

$$+ \underline{P_{H_0}} \left\{ \sum_{j=1}^{k-1} jI_j \leq i \mid \underline{I_k = 0} \right\} \underline{P_{H_0}\{I_k = 0\}}$$

$$= P_{H_0} \left\{ \sum_{j=1}^{k-1} jI_j \leq i - k \right\} \underline{P_{H_0}\{I_k = 1\}} + \underline{P_{H_0}} \left\{ \sum_{j=1}^{k-1} jI_j \leq \underline{i} \right\} \underline{P_{H_0}\{I_k = 0\}}$$

Continued..

$$\underline{P_{H_0}\{I_k = 1\}} = \underline{P_{H_0}\{I_k = 0\}} = \underline{\frac{1}{2}}$$

we see that

$$\underline{P_k(i) = \frac{1}{2}P_{k-1}(i - k) + \frac{1}{2}P_{k-1}(i)}$$

$$P_k(i) = P_{k-1}(i-k)P\{I_k=1\} + P_{k-1}(i)P\{I_k=0\}$$

Base Case:

$$\underline{P_1(i)} = \begin{cases} 0 & i < 0 \\ \frac{1}{2} & i = 0 \\ 1 & i \geq 1 \end{cases}$$

$$P(I_1 < 0) = 0$$

$$P(I_1 \leq 0) = \frac{1}{2}$$

$$P(I \leq 1) = 1$$

$$P_1(i) = P_{H_0}(T \leq i)$$

Example

Compute: $P_4(3)$

$$\begin{aligned} &= P_{H_0} \left(\sum_{j=1}^4 j I_j \leq 3 \right) = \frac{1}{2} P_3(-1) + \frac{1}{2} P_3(3) \\ &= 0 + \frac{1}{2} [P_2(0) + P_2(3)] \\ &\quad + \frac{1}{2} \cdot \frac{1}{2} [P_1(-2) + P_1(0) + P_1(1) + P_1(3)] \end{aligned}$$

HW

- How to extend paired-t-test to the non-parametric case?

Are two distributions equal?

- Let F and G be two continuous distributions of unknown form
- Given
 - n samples X_1, \dots, X_n from F
 - m samples Y_1, \dots, Y_m from G
- Null hypothesis: $H_0: F = G$
- Test is called: Rank-sum test, Mann-Whitney test, Wilcoxon test

Rank order the $n+m$ items.

R_i = rank of the data value X_i

Test statistic:

$$T = \sum_{i=1}^n R_i$$

Example 12.4.a. An experiment designed to compare two treatments against corrosion yielded the following data in pieces of wire subjected to the two treatments.

Treatment 1 65.2, 67.1, 69.4, 78.2, 74, 80.3

Treatment 2 59.4, 72.1, 68, 66.2, 58.5

(The data represent the maximum depth of pits in units of one thousandth of an inch.) The ordered values are 58.5, 59.4, 65.2*, 66.2, 67.1*, 68, 69.4*, 72.1, 74*, 78.2*, 80.3* with an asterisk noting that the data value was from sample 1. Hence, the value of the test statistic is $T = 3 + 5 + 7 + 9 + 10 + 11 = 45$. ■

Distribution of test-statistic under the null hypothesis $P_{H_0}(T)$

- Again we will compute recursively.
- Let $P(n, m, t) = P_{H_0}(T \leq t)$

self-study.

Either the last item in the rank is one of the N X_i s, or it is one of the M Y_j s. Under the null hypothesis, this probability:

$$P(N, M, K) = \frac{N}{N+M} P(N-1, M, K-N-M) \\ + \frac{M}{N+M} P(N, M-1, K)$$

Starting with the boundary condition

$$P(1, 0, K) = \begin{cases} 0 & K \leq 0 \\ 1 & K > 0 \end{cases}, \quad P(0, 1, K) = \begin{cases} 0 & K < 0 \\ 1 & K \geq 0 \end{cases}$$

Example 12.4.b. Suppose we wanted to determine $P(2, 1, 3)$. We use Equation (12.4.3) as follows:

$$P(2, 1, 3) = \frac{2}{3} P(1, 1, 0) + \frac{1}{3} P(2, 0, 3)$$

and

$$P(1, 1, 0) = \frac{1}{2} P(0, 1, -2) + \frac{1}{2} P(1, 0, 0) = 0$$

$$\begin{aligned} P(2, 0, 3) &= P(1, 0, 1) \\ &= P(0, 0, 0) = 1 \end{aligned}$$

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$$P(6,5,45) = \frac{6}{11}P(5,5,34) + \frac{5}{11}P(6,4,45) = \dots$$

Wilcoxon rank sum test

data: x and y

$W = 24$, p-value = 0.1255

Errors in Hypothesis testing

- Type-I error: Rejecting H_0 even when H_0 is true.
 - The probability with which it happens is called significant level α
- Type-II error: Accepting H_0 when it is false

Summary of hypothesis testing

- Follow this framework:
 - Formulate null and alternative hypothesis
 - Collect data
 - Decide on test statistic
 - Identify distribution of test statistic under null hypothesis
 - Apply p-value or critical region test to accept or reject null hypothesis
- We applied this framework on
 - Mean of Gaussian with unknown variance is μ_0
 - Are means of two normal distributions with shared unknown variance same?
 - Difference in means of two normal with unknown variance from paired observations

Summary..

- Parameter p of Bernoulli is p_0
- Non-parametric tests
 - Median is a given value
 - Distribution is symmetric around a median
 - Are two distributions equal

Topics not covered.

- Goodness of fit tests
- Test on sequences