

**2.49** From 2.161, we have:

$$\begin{aligned}
St(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Lambda}, \nu) &= \int_0^\infty \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, (\eta\boldsymbol{\Lambda})^{-1}) \text{Gam}(\eta|\nu/2, \nu/2) d\eta \\
&= \int_0^\infty \frac{1}{(2\pi)^{D/2} |(\eta\boldsymbol{\Lambda})^{-1}|^{1/2}} \exp\left\{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T (\eta\boldsymbol{\Lambda})(\mathbf{x} - \boldsymbol{\mu})\right\} \frac{1}{\Gamma(\nu/2)} \left(\frac{\nu}{2}\right)^{\nu/2} \eta^{\nu/2-1} \exp\left(-\frac{\nu\eta}{2}\right) d\eta
\end{aligned}$$

And from 2.162, we have:

$$= \frac{\Gamma(\nu/2 + D/2)}{\Gamma(\nu/2)} \frac{|\boldsymbol{\Lambda}|^{1/2}}{(\pi\nu)^{D/2}} \left(1 + \frac{\Delta^2}{\nu}\right)^{-D/2-\nu/2}$$

Finding Expectation:

$$\begin{aligned}
\mathbb{E}[\mathbf{x}] &= \int_{\mathbf{x}} St(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Lambda}, \nu) \mathbf{x} d\mathbf{x} \\
&= \int_{\mathbf{x}} \int_0^\infty \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, (\eta\boldsymbol{\Lambda})^{-1}) \text{Gam}(\eta|\nu/2, \nu/2) \mathbf{x} d\eta d\mathbf{x} \\
&= \int_0^\infty \left( \int_{\mathbf{x}} \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, (\eta\boldsymbol{\Lambda})^{-1}) \mathbf{x} d\mathbf{x} \right) \text{Gam}(\eta|\nu/2, \nu/2) d\eta \\
&= \int_0^\infty (\boldsymbol{\mu}) \text{Gam}(\eta|\nu/2, \nu/2) d\eta \\
&= \boldsymbol{\mu} \int_0^\infty \text{Gam}(\eta|\nu/2, \nu/2) d\eta \\
&= \boldsymbol{\mu}
\end{aligned}$$

Unexplained: why does 2.164 require  $\nu > 1$ ? As per page 100, for a finite integral it is required that  $\nu/2 > 0 \implies \nu > 0$ .

Finding covariance:

$$cov[\mathbf{x}] = \mathbb{E}[\mathbf{x}\mathbf{x}^T] - \mathbb{E}[\mathbf{x}]\mathbb{E}[\mathbf{x}]^T$$

$$\begin{aligned}
\mathbb{E}[\mathbf{x}\mathbf{x}^T] &= \int_{\mathbf{x}} St(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Lambda}, \nu) \mathbf{x}\mathbf{x}^T d\mathbf{x} \\
&= \int_{\mathbf{x}} \int_0^\infty \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, (\eta\boldsymbol{\Lambda})^{-1}) Gam(\eta|\nu/2, \nu/2) \mathbf{x}\mathbf{x}^T d\eta d\mathbf{x} \\
&= \int_0^\infty \left( \int_{\mathbf{x}} \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, (\eta\boldsymbol{\Lambda})^{-1}) \mathbf{x}\mathbf{x}^T d\mathbf{x} \right) Gam(\eta|\nu/2, \nu/2) d\eta \\
&= \int_0^\infty (\eta\boldsymbol{\Lambda})^{-1} + \boldsymbol{\mu}\boldsymbol{\mu}^T Gam(\eta|\nu/2, \nu/2) d\eta \\
&= \int_0^\infty (\eta\boldsymbol{\Lambda})^{-1} Gam(\eta|\nu/2, \nu/2) d\eta + \int_0^\infty \boldsymbol{\mu}\boldsymbol{\mu}^T Gam(\eta|\nu/2, \nu/2) d\eta \\
&= \boldsymbol{\Lambda}^{-1} \int_0^\infty \eta^{-1} Gam(\eta|\nu/2, \nu/2) d\eta + \boldsymbol{\mu}\boldsymbol{\mu}^T \int_0^\infty Gam(\eta|\nu/2, \nu/2) d\eta \\
&= \boldsymbol{\Lambda}^{-1} \int_0^\infty \eta^{-1} Gam(\eta|\nu/2, \nu/2) d\eta + \boldsymbol{\mu}\boldsymbol{\mu}^T (1)
\end{aligned}$$

Focusing on the integral,

$$\begin{aligned}
&\int_0^\infty \eta^{-1} Gam(\eta|\nu/2, \nu/2) d\eta \\
&= \int_0^\infty \eta^{-1} \frac{1}{\Gamma(\nu/2)} (\nu/2)^{\nu/2} \eta^{(\nu/2-1)} exp(-(\nu/2)\eta) d\eta \\
&= \frac{1}{\Gamma(\nu/2)} (\nu/2)^{\nu/2} \int_0^\infty \eta^{-1} \eta^{(\nu/2-1)} exp(-(\nu/2)\eta) d\eta \\
&= \frac{1}{\Gamma(\nu/2)} (\nu/2)^{\nu/2} \int_0^\infty \eta^{(\nu/2-2)} exp(-(\nu/2)\eta) d\eta
\end{aligned}$$

Let  $z = (\nu/2)\eta$ .  $\frac{d\eta}{dz} = \frac{2}{\nu}$ . Limits of  $z$  stay the same.

$$\begin{aligned}
&= \frac{1}{\Gamma(\nu/2)} (\nu/2)^{\nu/2} \int_0^\infty \left( \frac{2z}{\nu} \right)^{(\nu/2-2)} exp(-z) dz \left( \frac{2}{\nu} \right) \\
&= \frac{1}{\Gamma(\nu/2)} \left( \frac{\nu}{2} \right)^{\nu/2} \left( \frac{2}{\nu} \right)^{(\nu/2-2)} \left( \frac{2}{\nu} \right) \int_0^\infty z^{(\nu/2-2)} exp(-z) dz
\end{aligned}$$

$$= \frac{\nu}{2\Gamma(\nu/2)} \int_0^\infty z^{(\nu/2-2)} \exp(-z) \, dz$$

Using 1.141, this becomes:

$$\begin{aligned} &= \frac{\nu}{2\Gamma(\nu/2)} \Gamma(\nu/2 - 1) \\ &= \frac{\nu}{2(\nu/2 - 1)\Gamma(\nu/2 - 1)} \Gamma(\nu/2 - 1) \\ &= \frac{\nu}{2(\nu/2 - 1)} \\ &= \frac{\nu}{\nu - 2} \end{aligned}$$

Substituting all the results into the expression for covariance, we get,

$$\begin{aligned} \text{cov}[\mathbf{x}] &= \mathbb{E}[\mathbf{x}\mathbf{x}^T] - \mathbb{E}[\mathbf{x}]\mathbb{E}[\mathbf{x}]^T \\ &= \mathbf{\Lambda}^{-1} \frac{\nu}{(\nu - 2)} + \boldsymbol{\mu}\boldsymbol{\mu}^T - \boldsymbol{\mu}\boldsymbol{\mu}^T \\ &= \mathbf{\Lambda}^{-1} \frac{\nu}{(\nu - 2)} \end{aligned}$$

Note:  $\nu > 2$  as covariance matrix has to be positive definite.

Finding mode by taking derivative of Student's t-distribution w.r.t  $\mathbf{x}$ :

$$\begin{aligned} \frac{\partial}{\partial \mathbf{x}} St(\mathbf{x}|\boldsymbol{\mu}, \mathbf{\Lambda}, \nu) &= \frac{\partial}{\partial \mathbf{x}} \int_0^\infty \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, (\eta\mathbf{\Lambda})^{-1}) \text{Gam}(\eta|\nu/2, \nu/2) \, d\eta \\ &= \int_0^\infty \frac{\partial}{\partial \mathbf{x}} \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, (\eta\mathbf{\Lambda})^{-1}) \text{Gam}(\eta|\nu/2, \nu/2) \, d\eta \\ &= \int_0^\infty \text{Gam}(\eta|\nu/2, \nu/2) \left( \frac{\partial \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, (\eta\mathbf{\Lambda})^{-1})}{\partial \mathbf{x}} \right) \, d\eta \end{aligned}$$

Setting it to  $\mathbf{0}$ , we get the same result as the mode of the Gaussian distribution, which is at  $\mathbf{x} = \boldsymbol{\mu}$ .