Linear Models for Classification

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Classification

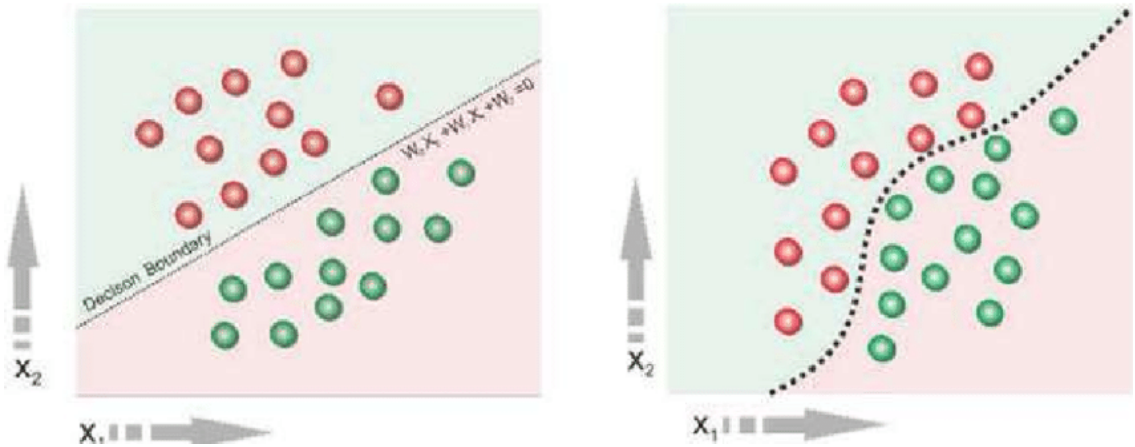
- Given an observation vector $\mathbf{x}$, assign it to one of the K classes
- The target can take one of the K discrete labels $\mathcal{C}_k$, $k = 1, 2, \cdots K$
  - Classify $\mathbf{x}$ =(45Kg, 4.8ft) as adult or kid
  - Classify a given image as face or nonface
  - Classify a speech waveform as one of the 40 phonemes
- Discriminant function to create separating hyperplane between classes

  Eg. Least squares, Fisher discriminant, perceptron, SVM
- Probabilistic approaches to estimate posterior probabilities $P[\mathcal{C}_k/\mathbf{x}]$
  - Discriminative models: Directly estimate $P[\mathcal{C}_k/\mathbf{x}] = f(\mathbf{x}, \mathbf{w})$

    Eg. Logistic regression, DNN classifiers
  - Generative models: Arrive at the posterior from joint density $p(\mathbf{x}, \mathcal{C}_k)$

    Eg. PDF Estimation: GMM (RV), HMM (RP)

Linear Separability



- Classes separable by a linear hyperplane
- Given the data, evaluate the equation of the decision boundary
  - Evaluate the weight vector $\mathbf{w}$, and bias parameter w_0
- Linear discriminant, Fisher discriminant, Perceptron and logistic regression

Linear Discriminant Functions (Binary Classification)

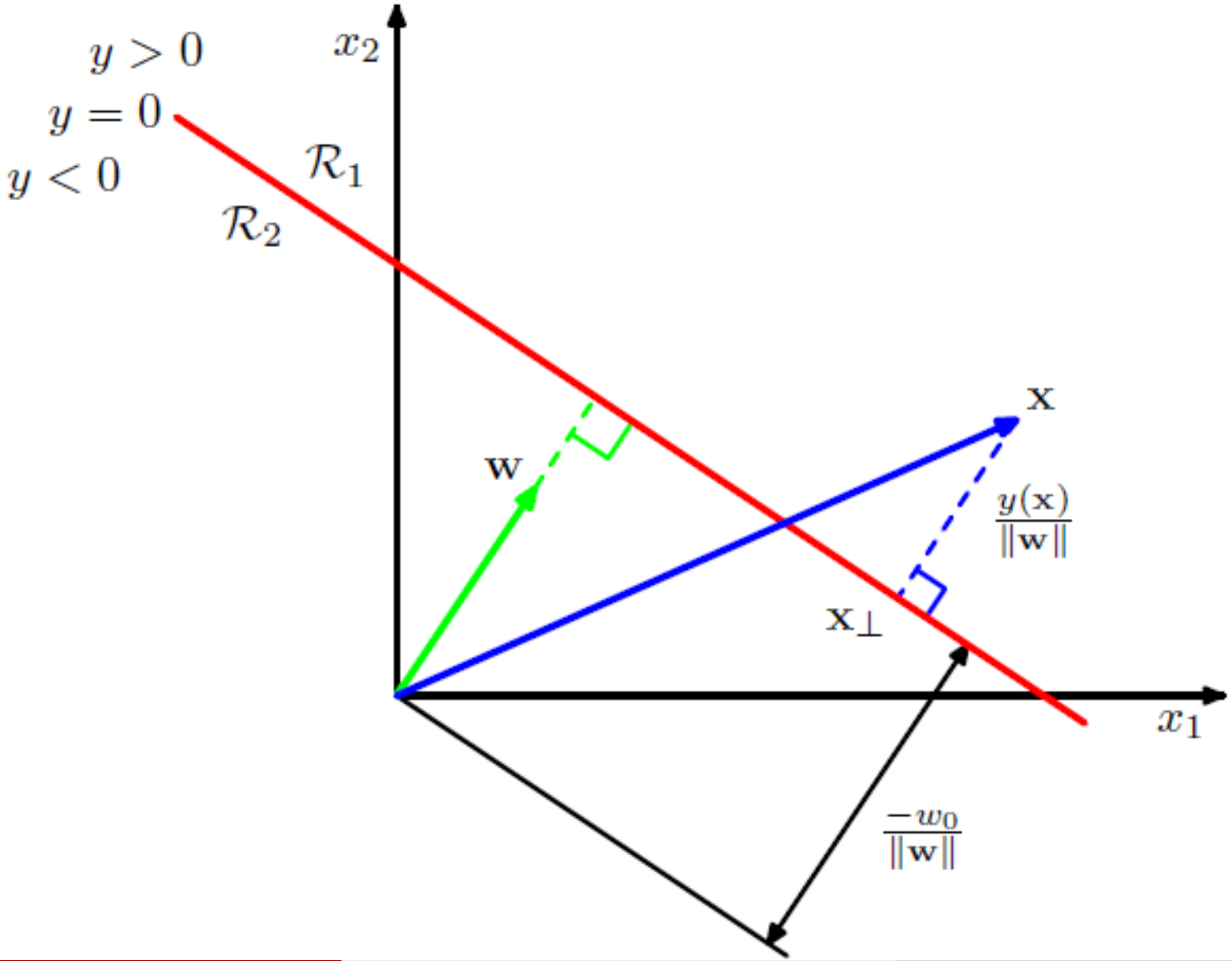
- For an input vector $\mathbf{x} \in \mathbb{R}^D$, Linear Discriminant Function is given by

$$y(\mathbf{x}) = \mathbf{w}^{\mathsf{T}}\mathbf{x} + w_0$$

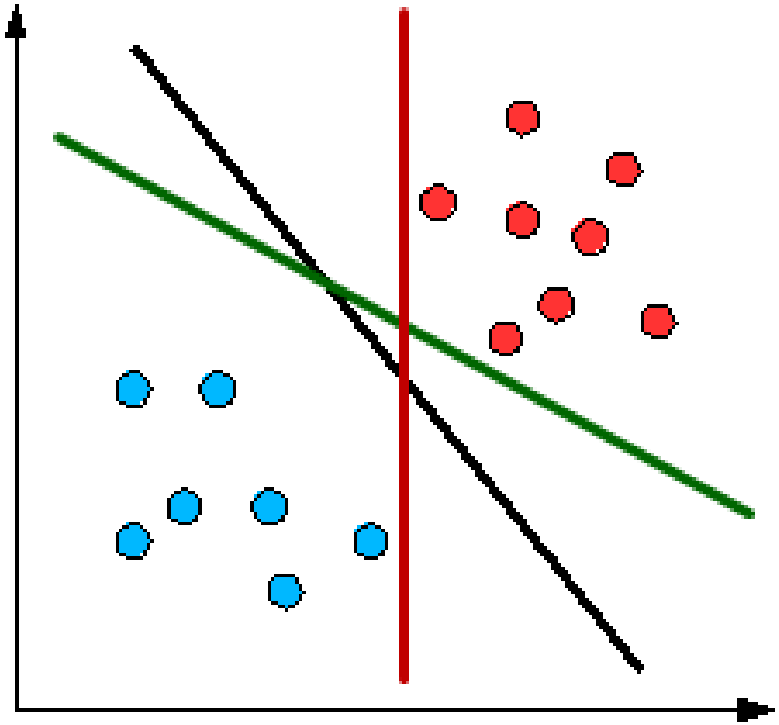
- Decision Rule: Assign $\mathbf{x}$ to $\mathcal{C}_1$ if $y(\mathbf{x}) \geq 0$, otherwise assign $\mathbf{x}$ to $\mathcal{C}_2$
- Decision boundary, $y(\mathbf{x}) = 0$, corresponds to a $D-1$ dim. hyperplane
- Weight vector $\mathbf{w}$ is orthogonal to every vector on the decision surface
- Weight vector $\mathbf{w}$ determines the orientation of the decision surface
- Bias parameter w_0 determines the location of the decision surface
- Normal distance from origin to the decision surface is given by

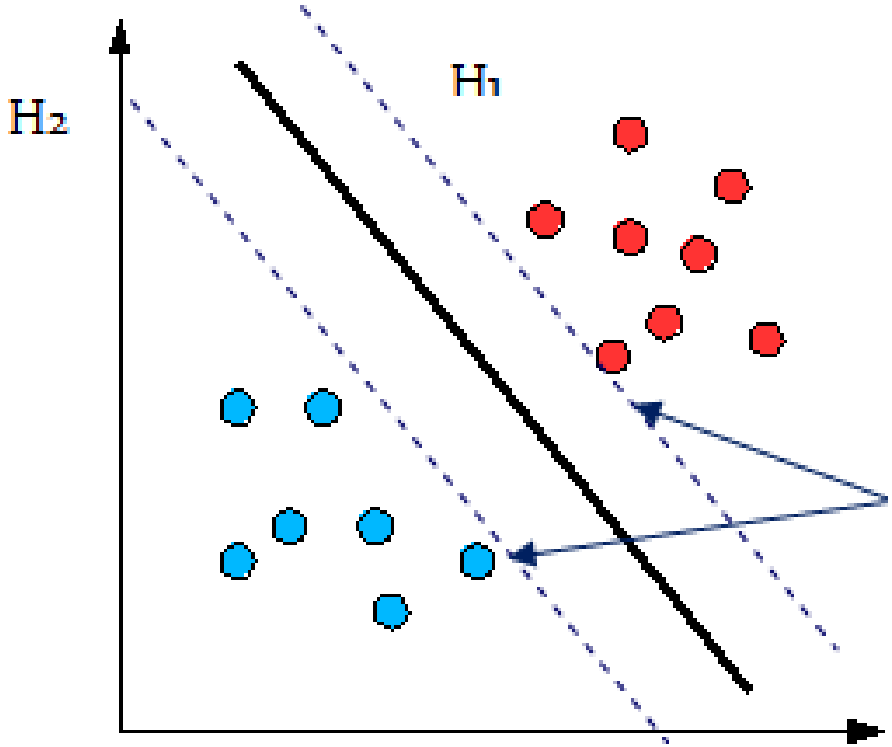
$$\frac{\mathbf{w}^{\mathsf{T}}\mathbf{x}}{\|\mathbf{w}\|} = -\frac{w_0}{\|\mathbf{w}\|}$$

Geometry of Decision Boundary in 2D



Effect of Orientation **w** and Position w_0

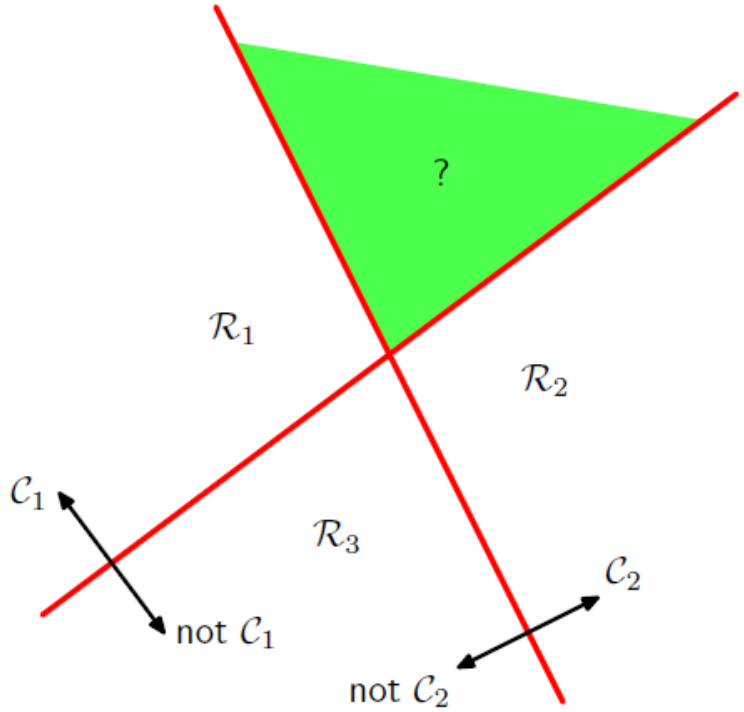




- Multiple linear discriminants can separate the classes
  - **w** controls their orientation
  - w_0 controls their position

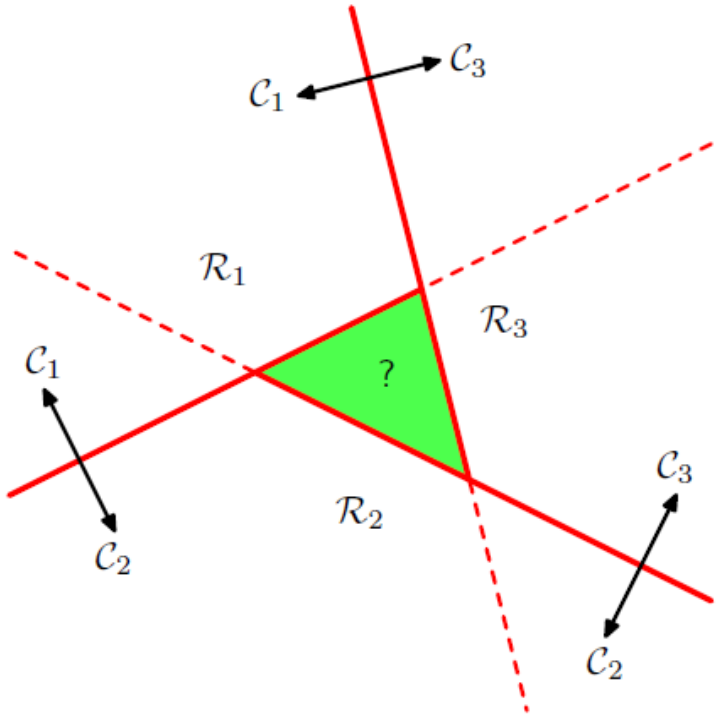
Extending to Multiple Classes

One Against Rest



K-1 Decision Surfaces

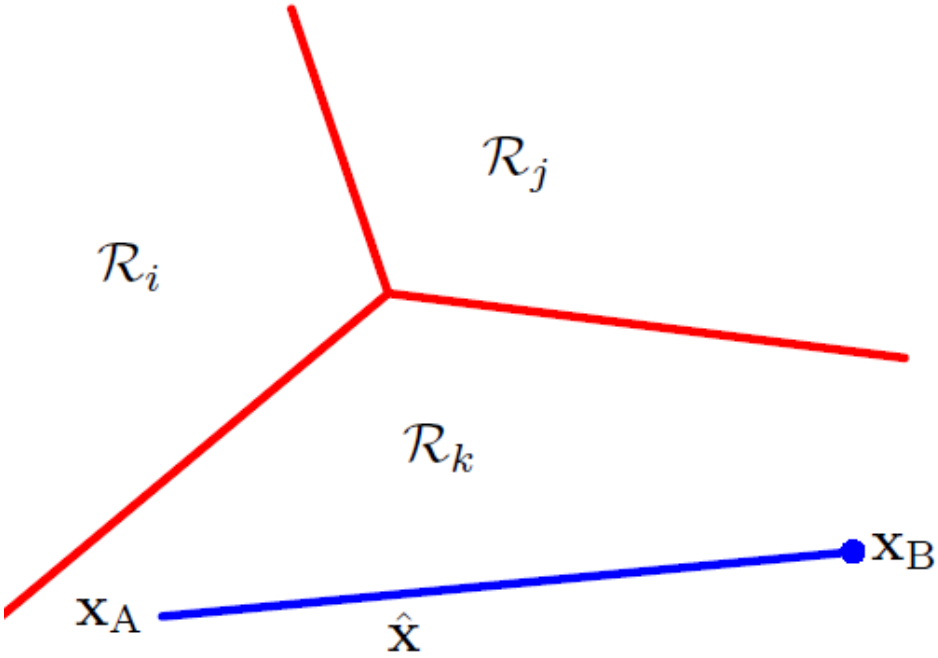
One Against One



$K(K-1)/2$ Decision Surfaces

- Both approaches lead to ambiguous regions!
- Decision regions should be singly connected and convex.

Multiclass Linear Discriminant



- K-Linear discriminant functions

  $y_k(\mathbf{x}) = \mathbf{w}_k^\mathsf{T}\mathbf{x} + w_{k0} \quad k = 1, 2, \cdots, K$

- Decision rule: Assign $\mathbf{x}$ to $\mathcal{C}_k$ if

  $y_k(\mathbf{x}) > y_j(\mathbf{x}), \quad \forall j \neq k$

- Decision boundary $\mathcal{C}_k$ & $\mathcal{C}_j$:

  $$y_k(\mathbf{x}) = y_j(\mathbf{x})$$
  $$(\mathbf{w}_k - \mathbf{w}_j)^\mathsf{T} x + (w_{k0} - w_{j0}) = 0$$

- Decision regions are convex
- For two classes, we can either employ a single discriminant or two discriminants

One-Hot Encoding of Targets

Pet type	Pet Code
cats	0
dogs	1
birds	2

categorical data

Pet type	is cat	is dog	is bird
cats	1	0	0
dogs	0	1	0
birds	0	0	1

one hot encoding

- One-hot encoding prevents ordinal relationship
  - Categorical encoding enforces a nonexisting order (cat $<$dog $<$ bird)
- Offers improved model interpretability.
  - One-hot encoding can explicitly learn separate weights for each class
  - Categorical encoding combines multiple classes into a single dimension
- Prevents model bias (classes with higher class labels)
- One-hot encoding may lead to high cardinality of the target labels.

Least Squares for Classification

- Let us apply Least Squares Regression approach to classification
- Training data: $\{(\mathbf{x}_1, \mathbf{t}_1), (\mathbf{x}_2, \mathbf{t}_2), \cdots (\mathbf{x}_N, \mathbf{t}_N)\}$
  - $\mathbf{x}_n \in \mathbb{R}^D$ denotes input observation vectors
  - $\mathbf{t}_n$, with 1-of-K binary coding, denotes the class label (K - Classes)
- Each class $\mathcal{C}_k$ is described by its own linear discriminant

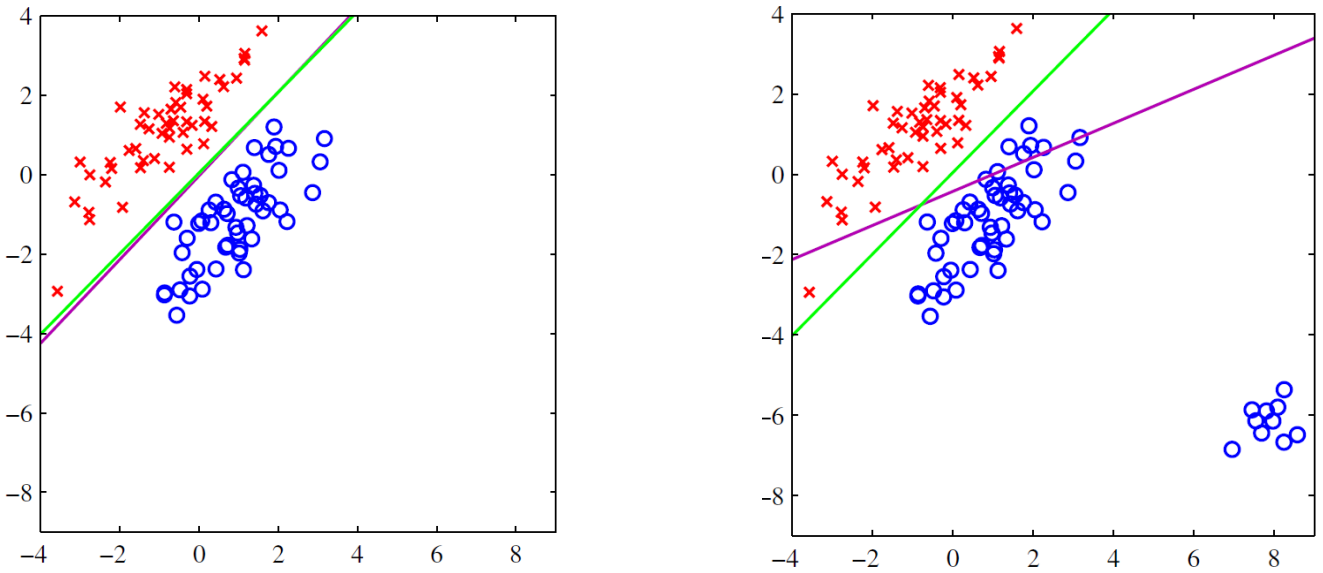
$$y_k(\mathbf{x}) = \mathbf{w}_k^{\mathsf{T}}\mathbf{x} + w_{k0} \quad k = 1, 2, \cdots, K$$

- Regressed target for $\mathbf{x}_n$: $\hat{\mathbf{t}}_n = \mathbf{y}(\mathbf{x}_n) = \mathbf{W}^{\mathsf{T}}\mathbf{x}_n$ (Augmented)
- Estimate $\mathbf{W}$ to minimize sum of squares error

$$J(\mathbf{W}) = \frac{1}{2}\,\mathrm{Tr}\left\{(\mathbf{X}\mathbf{W} - \mathbf{T})^{\mathsf{T}}(\mathbf{X}\mathbf{W} - \mathbf{T})\right\}$$

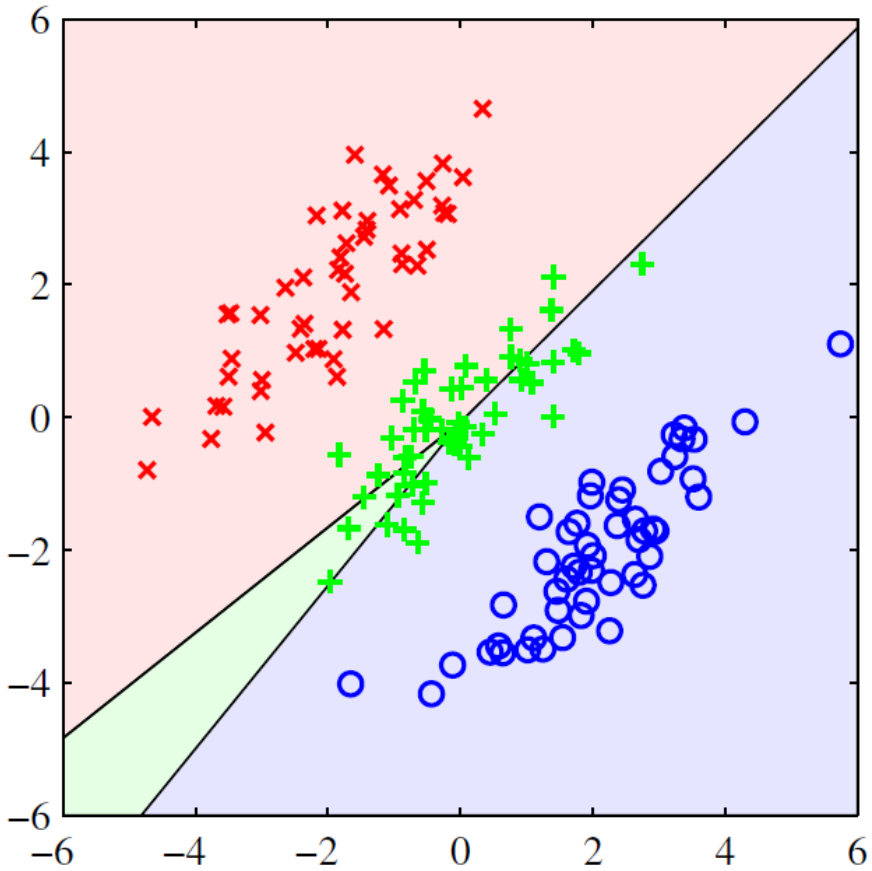
- Setting $\nabla_{\mathbf{W}} J(\mathbf{W}) = \mathbf{0}$, we get $\boxed{\mathbf{W}_* = \left(\mathbf{X}^{\mathsf{T}}\mathbf{X}\right)^{-1}\mathbf{X}^{\mathsf{T}}\mathbf{T}}$

Issues with LS

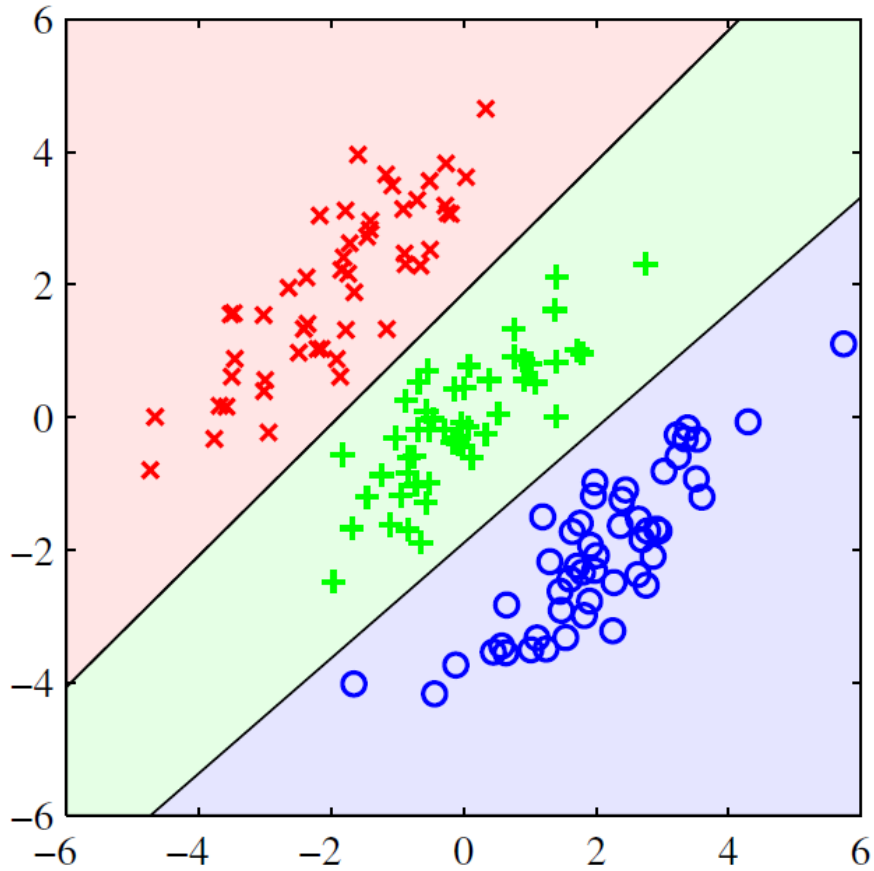


- Least squares solutions lack robustness to outliers
- Additional data-points resulted in significant change in boundary
- Sum of squares error penalizes predictions that are "too correct"
- Attempts to achieve "many-to-one" mapping through linearity!
- LS approach failed even for linearly separable classes

Issues with LS



Least Squares Classifier



Logistic Regression

Homework

- **Property of LS:** If every target in the training set satisfies some linear constraint

$$\mathbf{a}^{\mathsf{T}}\mathbf{t}_n + b = 0, \qquad \forall n$$

for some arbitrary constants $\mathbf{a}$ and b, then the model prediction for any value of $\mathbf{x}$ satisfies the same constraint.

$$\mathbf{a}^{\mathsf{T}}\mathbf{y}(\mathbf{x}) + b = 0$$

- If we use 1-of-K coding for targets, then the predictions sum to 1.

$$\sum_{k=1}^{K} y_k(\mathbf{x}) = 1$$

- However, $y_k(\mathbf{x})$ cannot be interpreted as posterior probability. They can be negative!

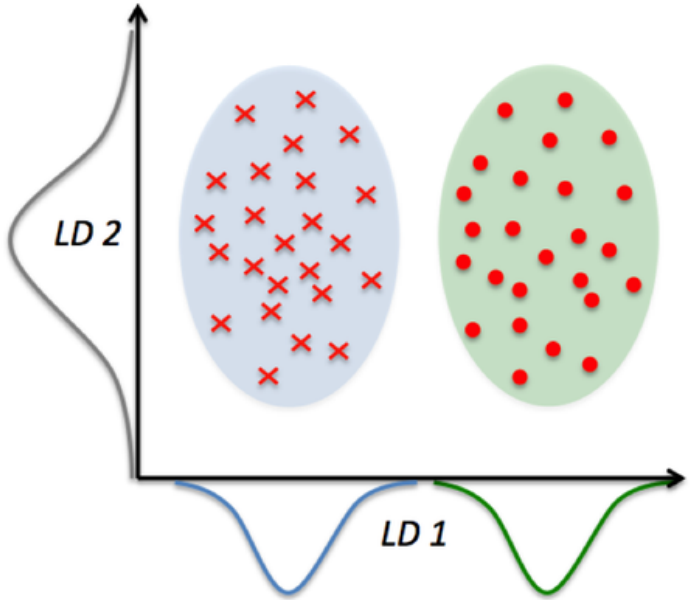
Linear Discriminant Analysis

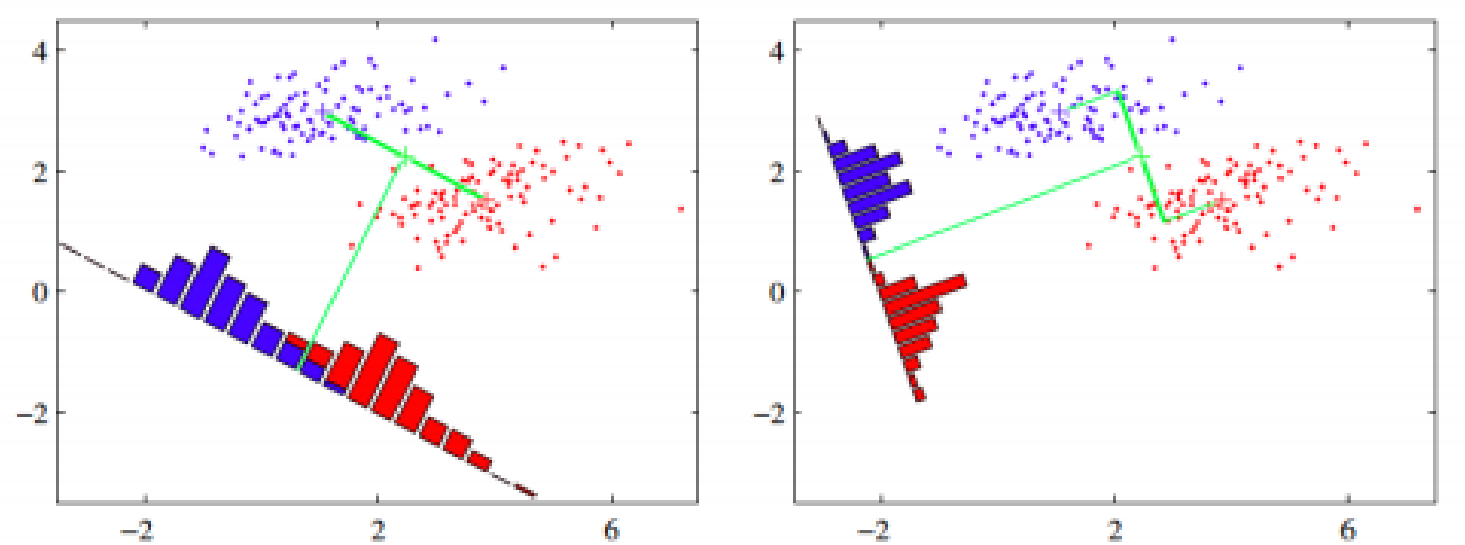
- Dimensionality reduction interpretation to LS classifier (2 classes)
  - Project the $\mathbf{x} \in \mathbb{R}^D$ on to real line (1-D): $y(\mathbf{x}) = \mathbf{w}^\mathsf{T}\mathbf{x}$
  - Assign $\mathbf{x}$ to $\mathcal{C}_1$ if $y(\mathbf{x}) > -w_0$ and to $\mathcal{C}_2$ otherwise
  - Projecting to 1-D, in general, leads to loss of information
  - Adjust $\mathbf{w}$ to select a projection that maximizes the class separation.
- Consider N_1 points from $\mathcal{C}_1$ and N_2 points from $\mathcal{C}_2$ in D-dim space
  - Mean vectors in original space: $\mathbf{m}_1 = \frac{1}{N_1}\sum_{n\in\mathcal{C}_1}\mathbf{x}_n \quad \mathbf{m}_2 = \frac{1}{N_2}\sum_{n\in\mathcal{C}_2}\mathbf{x}_n$
  - Let the projected means be $\mu_1 = \mathbf{w}^\mathsf{T}\mathbf{m}_1$ and $\mu_2 = \mathbf{w}^\mathsf{T}\mathbf{m}_2$
  - Choose $\mathbf{w}$ to maximize the separation between projected means

$$\mu_2 - \mu_1 = \mathbf{w}^\mathsf{T}(\mathbf{m}_2 - \mathbf{m}_1) \quad \textbf{st.} \quad \mathbf{w}^\mathsf{T}\mathbf{w} = 1$$

  - The direction for $\mathbf{w}$ can be shown to be $\mathbf{w} \propto (\mathbf{m}_2 - \mathbf{m}_1)$
- This approach is not optimal for nondiagonal covariances

LDA - Illustration





Fisher Discriminant Analysis

- For nondiagonal covariances, spread of data should also be considered
- Project the data in a direction that
  - maximizes seperation between the means of the projected classes
  - minimizes the variance within each projected class
- The objective function is given by

$$J(\mathbf{w}) = \frac{(\mu_2 - \mu_1)^2}{\sigma_1^2 + \sigma_2^2} = \frac{\mathbf{w}^{\mathsf{T}} \mathbf{S}_B \mathbf{w}}{\mathbf{w}^{\mathsf{T}} \mathbf{S}_W \mathbf{w}}$$

  - $\mathbf{S}_B$ is *between-class* covariance: $\mathbf{S}_B = (\mathbf{m}_2 - \mathbf{m}_1)(\mathbf{m}_2 - \mathbf{m}_1)^{\mathsf{T}}$
  - $\mathbf{S}_W$ is *within-class* covariance: $\mathbf{S}_W = \sum_{k=1}^{2} \sum_{n \in \mathcal{C}_k} (\mathbf{x}_n - \mathbf{m}_k)(\mathbf{x}_n - \mathbf{m}_k)^{\mathsf{T}}$
- Optimal direction of $\mathbf{w}$ can be obtained by maximizing $J(\mathbf{w})$

Solution to Fisher Criterion

- Equating $\nabla J(\mathbf{w}) = \mathbf{0}$, we obtain

$$(\mathbf{w}^T\mathbf{S}_B\mathbf{w})\mathbf{S}_W\mathbf{w} = (\mathbf{w}^T\mathbf{S}_W\mathbf{w})\mathbf{S}_B\mathbf{w}$$

  - $\mathbf{w}^T\mathbf{S}_B\mathbf{w}$ and $\mathbf{w}^T\mathbf{S}_W\mathbf{w}$ contribute to only scalar multiples
  - The direction of $\mathbf{w}$ is given by $\mathbf{S}_W\mathbf{w} \propto \mathbf{S}_B\mathbf{w}$
  - $\mathbf{S}_B\mathbf{w}$ is always in the direction of $\mathbf{m}_2 - \mathbf{m}_1$
  - The direction of the fisher discriminant: $\boxed{\mathbf{w} \propto \mathbf{S}_W^{-1}(\mathbf{m}_2 - \mathbf{m}_1)}$
- Multiplication by $\mathbf{S}_W^{-1}$ can be interpreted as data whitening.
- FDA also is sensitive to outliers even if they are "too correct"
- FD is strictly not a "discriminant" in true sense.
  - It just projects binary class data to 1-D
  - We need to arrive at a threshold w_0 to perform classification

Homework: Relation to Least Squares

- In LS approach, linear discriminant is determined to make model predictions as close as possible to target values
- In FDA, the discriminant is derived to achieve maximum class separation in the projected space
- If we take targets for $\mathcal{C}_1$ and $\mathcal{C}_2$ as $\frac{N}{N_1}$ and $-\frac{N}{N_2}$, respectively, where $N = N_1 + N_2$, show that LS approach yields the same solution as FD.

Extending FDA to Multiple Classes

- Assumption: Data dimensionality $D > K$ Number of classes
- For multi-class case, it is not enough to project data to 1-D
- Project the data to a lower dimensional space $D' < D$
  - Let $\mathbf{W} : \mathbb{R}^{D \times D'}$ be the linear map to achieve dimensionality reduction
  - The lower dimensional representation of $\mathbf{x}_n \in \mathbb{R}^D$ is $\mathbf{y}_n = \mathbf{W}^{\mathsf{T}}\mathbf{x}_n \in \mathbb{R}^{D'}$
  - Maximize between-class spread and minimize within-class spread in the projected space
  - Let $\mathbf{S}_W$ and $\mathbf{S}_B$ denote within-class and between-class covariance in original D-dim space
  - Let $\mathbf{\Sigma}_W$ and $\mathbf{\Sigma}_B$ be their counterparts in projected space
  - Estimate $\mathbf{W}$ to maximize $J(\mathbf{W}) = \mathrm{Tr}\left\{\mathbf{\Sigma}_W^{-1}\mathbf{\Sigma}_B\right\}$
- The columns of $\mathbf{W}$ are given by eigenvectors corresponding to the D' largest eigenvalues of $\mathbf{S}_W^{-1}\mathbf{S}_B$

Statistics in Original D-dim Space

- Class specific statistics in the original D-dim space

$$\mathbf{m}_k = \frac{1}{N_k} \sum_{n \in \mathcal{C}_k} \mathbf{x}_n \qquad \mathbf{S}_k = \sum_{n \in \mathcal{C}_k} (\mathbf{x}_n - \mathbf{m}_k)(\mathbf{x}_n - \mathbf{m}_k)^{\mathsf{T}}$$

- Define *within-class* covariance as $\mathbf{S}_W = \sum_{k=1}^{K} \mathbf{S}_k$
- Overall data statistics (without considering class labels)

$$\mathbf{m}_T = \frac{1}{N} \sum_{n=1}^{N} \mathbf{x}_n = \frac{1}{N} \sum_{k=1}^{K} N_k \mathbf{m}_k \qquad \mathbf{S}_T = \sum_{n=1}^{N} (\mathbf{x}_n - \mathbf{m}_T)(\mathbf{x}_n - \mathbf{m}_T)^{\mathsf{T}}$$

- Assumption: Define *between-class* covariance using $\mathbf{S}_T = \mathbf{S}_W + \mathbf{S}_B$

$$\mathbf{S}_B = \sum_{k=1}^{K} N_k (\mathbf{m}_k - \mathbf{m}_T)(\mathbf{m}_k - \mathbf{m}_T)^{\mathsf{T}}$$

Statistics in Projected D'-dim Space

- The projected data-points in $\mathbb{R}^{D'}$ are given by $\mathbf{y}_n = \mathbf{W}^{\mathsf{T}}\mathbf{x}_n$
- The class-specific statistics in the projected space are

$$\boldsymbol{\mu}_k = \frac{1}{N_k}\sum_{n\in\mathcal{C}_k}\mathbf{y}_n \qquad \boldsymbol{\Sigma}_k = \sum_{n\in\mathcal{C}_k}(\mathbf{y}_n - \boldsymbol{\mu}_k)(\mathbf{y}_n - \boldsymbol{\mu}_k)^{\mathsf{T}}$$

- Relation between the statistics of original and projected space

$$\boldsymbol{\mu}_k = \mathbf{W}^{\mathsf{T}}\mathbf{m}_k \quad \boldsymbol{\Sigma}_k = \mathbf{W}^{\mathsf{T}}\mathbf{S}_k\mathbf{W} \quad \boldsymbol{\Sigma}_W = \mathbf{W}^{\mathsf{T}}\mathbf{S}_W\mathbf{W} \quad \boldsymbol{\Sigma}_B = \mathbf{W}^{\mathsf{T}}\mathbf{S}_B\mathbf{W}$$

- Estimate $\mathbf{W}$ to maximize $J(\mathbf{W}) = \mathrm{Tr}\left\{\left(\mathbf{W}^{\mathsf{T}}\mathbf{S}_W\mathbf{W}\right)^{-1}\left(\mathbf{W}^{\mathsf{T}}\mathbf{S}_B\mathbf{W}\right)\right\}$
  - Solution leads to the famous eigenvalue problem: $\mathbf{S}_W\mathbf{W} \propto \mathbf{S}_B\mathbf{W}$
  - Eigenvectors corresponding to D' largest eigenvalues forms $\mathbf{W}$
- D' is bounded by rank of $\mathbf{S}_B$, and can be at most $K-1$

The Perceptron

- Issues with least squares and linear discriminants
  - Linear relation cannot achieve *many-to-one* mapping
  - No inbuilt mechanism to ignore *too-correct* outliers
- Perceptron employs a *nonlinearity* to estimate target: $y(\mathbf{x}) = f(\mathbf{w}^{\mathsf{T}}\mathbf{x})$
- $f(.)$ is a hard-limiting nonlinear function given by

$$f(a) = \begin{cases} +1, & a \geq 0 \\ -1, & a < 0 \end{cases}$$

- Desired targets are encoded as $+1$ for $\mathcal{C}_1$ and -1 for $\mathcal{C}_2$.
- Estimate $\mathbf{w}$ s.t. $\mathbf{w}^{\mathsf{T}}\mathbf{x}_n \geq 0$ for $x_n \in \mathcal{C}_1$, and $\mathbf{w}^{\mathsf{T}}\mathbf{x}_n < 0$ for $\mathbf{x}_n \in \mathcal{C}_2$
- Both the targets and estimates are discrete
- Number of misclassified examples is piece-wise constant

Perceptron Learning Criterion

- Perceptron criterion: defined over the set of misclassified examples $\mathcal{M}$

$$J_p(\mathbf{w}) = -\sum_{n \in \mathcal{M}} (\mathbf{w}^T \mathbf{x}_n) t_n$$

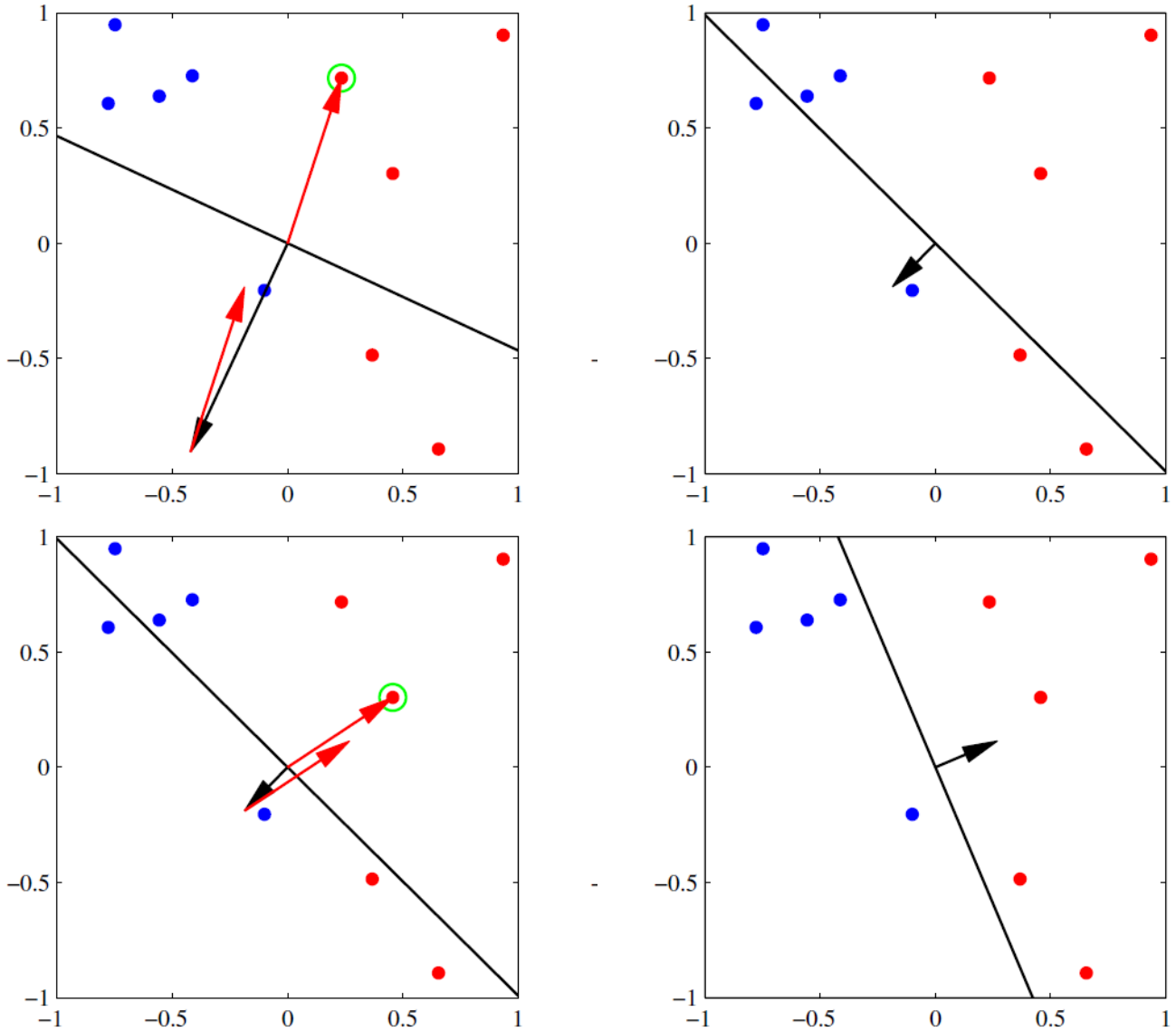
- SGD can be employed to update the weight vector

$$\mathbf{w}^{new} = \mathbf{w}^{old} - \eta \nabla J_p(\mathbf{w}) = \mathbf{w}^{old} + \eta \sum_{n \in \mathcal{M}} \mathbf{x}_n t_n$$

  - If x_n is correctly classified, do not update the weight vector
  - If $\mathbf{x}_n \in \mathcal{C}_1$ is misclassified: $\mathbf{w}^{new} = \mathbf{w}^{old} + \mathbf{x}_n$
  - If $\mathbf{x}_n \in \mathcal{C}_2$ is misclassified: $\mathbf{w}^{new} = \mathbf{w}^{old} - \mathbf{x}_n$

- Perceptron algorithm is guaranteed to converge in finite steps, for linearly separable classes. [Demo]
- Perceptron algorithm is not vulnerable to *too-correct* outliers

Perceptron Weight Updates



Perceptron Convergence Theorem (Upper Bound of $\|\mathbf{w}\|^2$)

- Starting from a random initial guess $\mathbf{w}_0$, perform K updates:

$$\mathbf{w}_K = \mathbf{w}_{K-1} + \mathbf{x}_K t_K = \mathbf{w}_0 + \sum_{k=1}^{K} \mathbf{x}_k t_k$$

  where $\mathbf{x}_k$ is the randomly selected misclassified point at k^{th} iteration
- Letting $\mathbf{w}_0 = \mathbf{0}$, the squared norm of $\mathbf{w}_K$ is bounded by

$$\|\mathbf{w}_K\|^2 = \left\| \sum_{k=1}^{K} \mathbf{x}_k t_k \right\|^2 \leq \sum_{k=1}^{K} \|\mathbf{x}_k\|^2$$

- Let α be the maximum sq. norm in the training data: $\alpha = \max_n \|\mathbf{x}_n\|^2$
- The norm of the weight vector at K^{th} iteration is upper bounded by

$$\|\mathbf{w}_K\|^2 \leq K\alpha$$

Perceptron Convergence Theorem (Lower Bound of $\|\mathbf{w}\|^2$)

- Let $\mathbf{w}_*$ be one of the solutions that separates classes exactly
  - Since $\mathbf{w}_*$ is a solution, $\mathbf{w}_* \mathbf{x}_n t_n \geq 0$ for all n in the dataset
- Consider projection of weight vector $\mathbf{w}_K$ onto $\mathbf{w}_*$

$$\mathbf{w}_*^{\mathsf{T}} \mathbf{w}_K = \sum_{k=1}^{K} \mathbf{w}_*^{\mathsf{T}} \mathbf{x}_k t_k$$

- Let β be the minimum projection onto $\mathbf{w}_*$: $\beta = \min_n \mathbf{w}_* \mathbf{x}_n t_n$
- Squared norm of the weight vector $\mathbf{w}_K$ is lower bounded by

$$\mathbf{w}_*^{\mathsf{T}} \mathbf{w}_K \geq K\beta \qquad \|\mathbf{w}_K\|^2 \geq \frac{1}{\|\mathbf{w}_*\|^2} K^2 \beta^2$$

- The range of the squared norm of the weight vector is given by

$$K^2 \beta' \leq \|\mathbf{w}_K\|^2 \leq K\alpha$$

- Lower bound is quadratic in K, upper bound is linear in K.

Limitations of Perceptron

- Perceptron algorithm converges only for linearly separable data
  - Number of iterations K could be very large
  - Error does not decrease monotonically $J_P(\mathbf{w}_{k+1}) \lesseqgtr J_P(\mathbf{w}_k)$
  - The final solution need not be optimal- one of the decision boundaries.
- Does not converge if the classes are not linearly separable - XOR
- Does not offer probabilistic interpretation or confidence intervals.
- Cannot be readily extended to multiclass problems

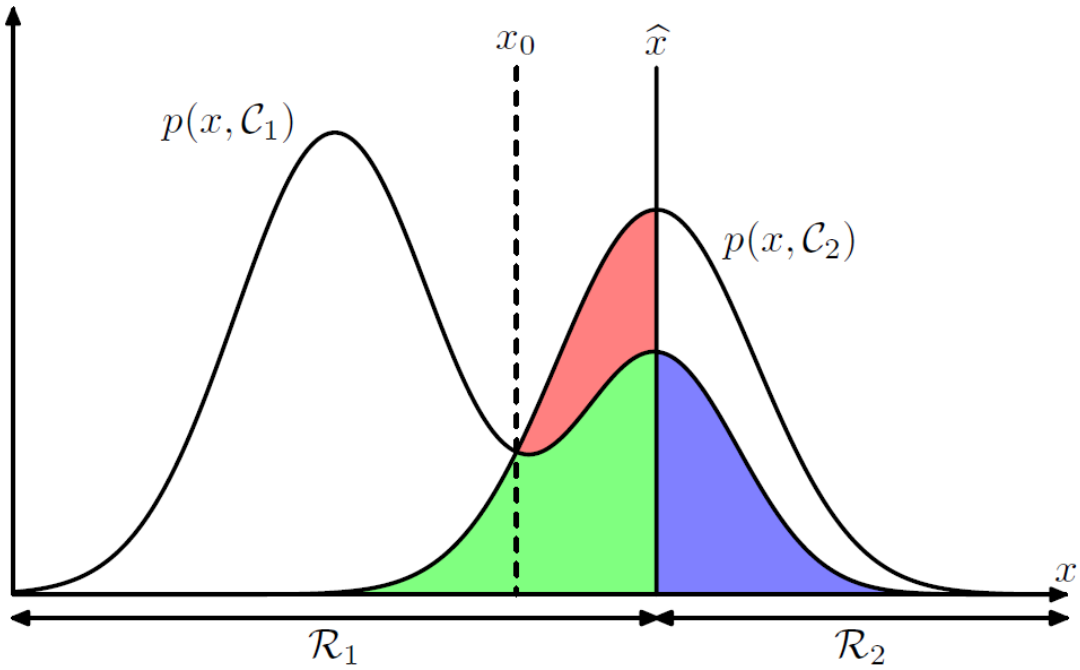
Decision Theory

- Probability theory offers a mathematical tool to deal with uncertainty.
- Supervised learning: Predict target $\mathbf{t}$ from observed input $\mathbf{x}$
- In such a case, $p(\mathbf{x}, \mathbf{t})$ provides a complete summary of uncertainty
- In classification task, $\mathbf{t}$ can take discrete labels $\mathcal{C}_k \quad k = 1, 2, \cdots, K$
  - *Inference* step: Given a test point $\mathbf{x}_t$, estimate $P[\mathcal{C}_k/\mathbf{x}_t]$
  - *Decision* step: Assign the test-point $\mathbf{x}_t$ to one of the classes
- Decision rule divides the input space into K decision regions $\mathcal{R}_k$
- Choose *decision boundaries* to minimize the probability of error
- The probability of error for binary classification is given by

$$P[\text{mistake}] = P[\mathbf{x} \in \mathcal{R}_1, \mathcal{C}_2] + P[\mathbf{x} \in \mathcal{R}_2, \mathcal{C}_1]$$

- Divide input space into $\mathcal{R}_1$ and $\mathcal{R}_2$ s.t. $P[\text{mistake}]$ is minimized

Choice of Decision Regions (Boundaries)



- Let $x = \hat{x}$ be decision boundary. $\mathcal{C}_1 : x \leq \hat{x}$ and $\mathcal{C}_2 : x > \hat{x}$
- Errors arise from area under the blue, red and green regions
  - Blue: Points from $\mathcal{C}_1$ but misclassified as $\mathcal{C}_2$
  - Green + Red: Points from $\mathcal{C}_2$, but misclassified as $\mathcal{C}_1$
  - Adjust the boundary $\hat{x}$ to minimize the area under blue+green+red
  - Area under blue+green remains constant irrespective of choice of $\hat{x}$

Minimizing Classification Error

$$
\begin{aligned}
P[\text{mistake}] &= \int_{\mathcal{R}_1} p(\mathbf{x}, \mathcal{C}_2) d\mathbf{x} + \int_{\mathcal{R}_2} p(\mathbf{x}, \mathcal{C}_1) d\mathbf{x} \\
&= 1 + \int_{\mathcal{R}_1} p(\mathbf{x}, \mathcal{C}_2) d\mathbf{x} - \int_{\mathcal{R}_1} p(\mathbf{x}, \mathcal{C}_1) d\mathbf{x} \\
&= 1 - \int_{\mathcal{R}_1} \Big(p(\mathbf{x}, \mathcal{C}_1) - p(\mathbf{x}, \mathcal{C}_2) \Big) d\mathbf{x}
\end{aligned}
$$

- Choose region $\mathcal{R}_1$ such that $p(\mathbf{x}, \mathcal{C}_1) > p(\mathbf{x}, \mathcal{C}_2)$
- Decision rule: Assign x $\mathcal{C}_1$ if $P[\mathcal{C}_1/\mathbf{x}] > P[\mathcal{C}_2/\mathbf{x}]$
- For general case of K classes, it is easier to maximize $p(\text{correct})$

$$
P[\text{correct}] = \sum_{k=1}^{K} \int_{\mathcal{R}_k} p(\mathbf{x}, \mathcal{C}_k) d\mathbf{x}
$$

- Decision Rule: Assign x to the class with highest posterior $P[\mathcal{C}_k/\mathbf{x}]$

Expected Loss

- In reality, the errors may have varying degrees of consequences
- Penalty for Healthy vs Cancer classification could be

$$\begin{array}{c} \\ H \\ C \end{array} \begin{array}{c} \begin{array}{cc} H & C \end{array} \\ \begin{pmatrix} 0 & 1 \\ 100 & 0 \end{pmatrix} \end{array}$$

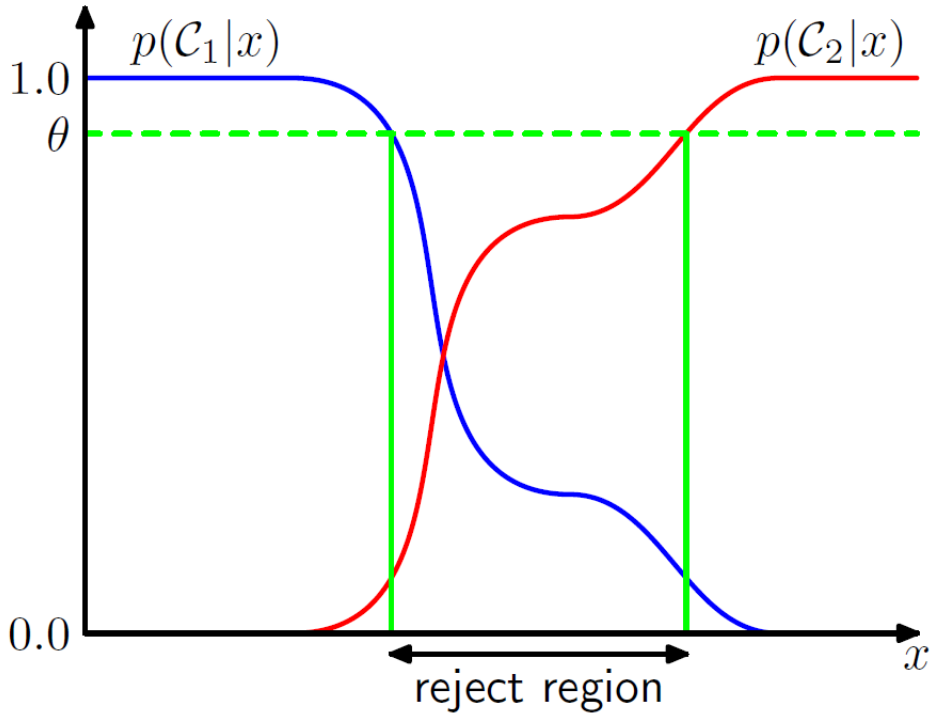
- Expected loss can be obtained by weighing error with penalty

$$\mathbb{E}[L] = \sum_k \sum_j \int_{\mathcal{R}_j} L_{kj} p(\mathbf{x}, \mathcal{C}_k) d\mathbf{x}$$

- L_{kj} denotes loss associated with wrongly assigning $x \in \mathcal{C}_k$ to $\mathcal{C}_j$
- Choose region $\mathcal{R}_j$ to minimize $\sum_k L_{kj} p(\mathbf{x}, \mathcal{C}_k)$
- Decision Rule: Assign x to $\mathcal{C}_j$ that minimizes $\sum_k L_{kj} p[\mathcal{C}_k/\mathbf{x}]$
- This is a trivial assignment once posterior probabilities are estimated

Reject Option

- Errors arises from regions where $\max_k p[\mathcal{C}_k/\mathbf{x}] << 1$
  - That means all the posteriors are in a similar range
  - In those regions, the classifier is relatively uncertain
- In such cases, it is better to avoid decision-making
  - Reject the test samples $\mathbf{x}$ for which $\max_k p[\mathcal{C}_k/\mathbf{x}] < \theta$



Homework: Expected Loss with Reject Option

- Consider a classification problem in which the loss incurred when an input vector from class $\mathcal{C}_k$ is classified as belonging to class $\mathcal{C}_j$ is given by the loss matrix L_{kj}, and for which the loss incurred in selecting the reject option is λ. Find the decision criterion that will give the minimum expected loss. Verify that this reduces to the reject criterion discussed earlier when the loss matrix is given by $L_{kj} = 1 - I_{kj}$. What is the relationship between λ and the rejection threshold θ?

Inference & Decision Stages

- Classification problem can be broken into inference and decision stages
- Generative models
  - Inference: Estimate posterior $p[\mathcal{C}_k/\mathbf{x}]$ from the joint density $p(\mathbf{x}, \mathcal{C}_k)$

$$p[\mathcal{C}_k/\mathbf{x}) = \frac{p[\mathcal{C}_k]p(\mathbf{x}/\mathcal{C}_k)}{p(\mathbf{x})}$$

  - Such a model can generate synthetic examples of a class from $p(\mathbf{x}, \mathcal{C}_k)$
- Discriminative models
  - Inference: Estimate posterior $p[\mathcal{C}_k/\mathbf{x}]$ as a parametric function $y(\mathbf{x}, \mathbf{w})$
  - Decision: $p[\mathcal{C}_k/\mathbf{x}]$ can be used for decision with loss and reject option
- Discriminant functions
  - Find a function $y(\mathbf{x}, \mathbf{w})$ that maps input $\mathbf{x}$ to a class label
  - Inference and decision stages cannot be separated

Pros & Cons

Feature	Generate	Discriminative	Discriminant
Computation	High	Moderate	Low
Data Req.	Very high	Moderate	Low
Outlier detection	Yes $p(\mathbf{x})$	No	No
Accuracy	Reasonable	Higher	Low
Minimizing Risk	Easy	Easy	Not SF
Reject option	Easy	Easy	Not SF
Modifying Priors	Easy	Not SF	No
Model Fusion	Easy	Easy	Not SF

$$
\begin{aligned}
P[\mathcal{C}_k/\mathbf{x}_A, \mathbf{x}_B] &\propto p[\mathcal{C}_k]p(\mathbf{x}_A, \mathbf{x}_B/\mathcal{C}_k) \\
&\propto p[\mathcal{C}_k]p(\mathbf{x}_A/\mathcal{C}_k)p(\mathbf{x}_B/\mathcal{C}_k) \qquad \text{Cond. Ind.} \\
&\propto \frac{p[\mathcal{C}_k/\mathbf{x}_A]p[\mathcal{C}_k/\mathbf{x}_B]}{p[\mathcal{C}_k]}
\end{aligned}
$$

Maximum Likelihood Density Estimation

- Consider data $X = \{x_1, x_2, \cdots x_N\}$ drawn from unknown distribution
- Data distribution $p_D(x_n)$ be approximated by a parametric form
  - Gaussian assumption: $p_D(x_n) \sim p_M(x_n/\mu, \sigma^2) = \mathcal{N}(x_n/\mu, \sigma^2)$
  - Laplacian assumption: $p_D(x_n) \sim p_M(x_n/\mu, b) = \mathcal{L}(x_n/\mu, b)$
- Assume that the observations x_n are drawn i.i.d from Gaussian
- The likelihood function can be written as

$$p_M(X/\mu, \sigma^2) = \prod_{n=1}^{N} p_M(x_n/\mu, \sigma^2)$$

$$= \prod_{n=1}^{N} \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x_n - \mu)^2}{2\sigma^2}\right)$$

- Estimate μ and σ to maximize the log-likelihood - $\log p_M(X/\mu, \sigma^2)$

ML Estimates of Gaussian Distribution

- μ and σ can be determined by equating the derivatives to zero

$$\hat{\mu} = \frac{1}{N}\sum_{n=1}^{N} x_n \qquad \hat{\sigma^2} = \frac{1}{N}\sum_{n=1}^{N}(x_n - \hat{\mu})^2$$

- Checking for *bias* in estimation: Taking expectation over estimates

$$\mathbb{E}[\hat{\mu}] = \mathbb{E}\left[\frac{1}{N}\sum_{n=1}^{N} x_n\right] = \frac{1}{N}\sum_{n=1}^{N}\mathbb{E}[x_n] = \mu \qquad \text{unbiased}$$

$$\mathbb{E}[\hat{\sigma^2}] = \frac{1}{N}\sum_{n=1}^{N}\mathbb{E}[(x_n - \hat{\mu})^2] = \frac{N-1}{N}\sigma^2 \qquad \text{biased}$$

- Variance is underestimated in ML approach. $\hat{\sigma^2} < \sigma^2$
- The variance estimation can be corrected as $\tilde{\sigma^2} = \frac{1}{N-1}\sum_{n=1}^{N}(x_n - \hat{\mu})^2$

Naive Bayes Classifier

- When the input $\mathbf{x}_n \in R^D$ comes from multivariate distribution $p_D(\mathbf{x})$
  - We can impose a multivariate density model on $\mathbf{x}_n$, e.g $p_M(\mathbf{x}_n/\boldsymbol{\mu}, \boldsymbol{\Sigma})$
  - Or assume that the dimensions of $\mathbf{x}_n$ are conditionally independent
- Naive Bayes classifier assumes statistical independence of dimensions

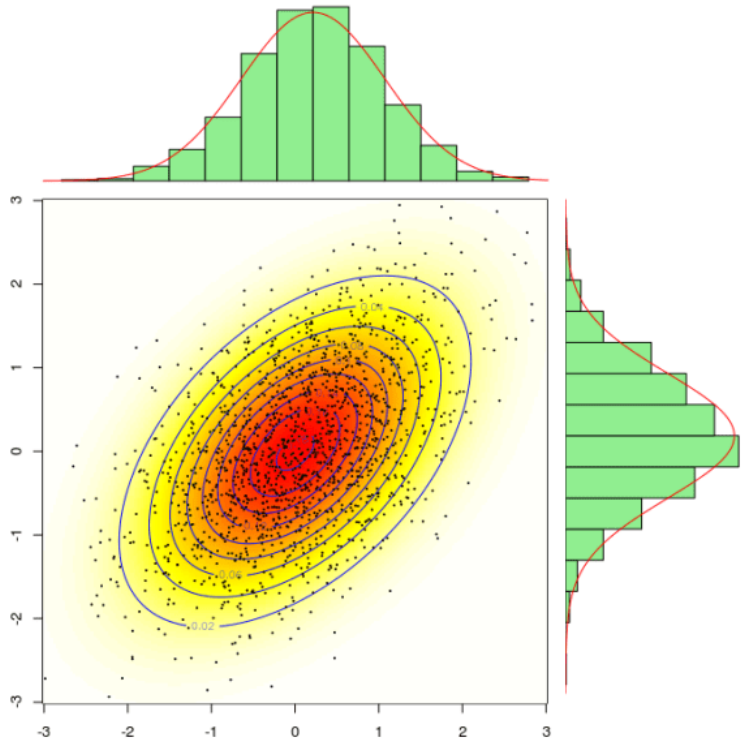
$$p(\mathbf{x}_n/C_k) = \prod_{d=1}^{D} p(x_{nd}/C_k)$$

- Under Gaussian assumption for marginal distributions

$$p_D(x_{nd}/C_k) \sim p_M(x_{nd}/C_k, \mu_d, \sigma_d^2) = \mathcal{N}(x_{nd}/\mu_d, \sigma_d^2) \quad d = 1, 2, \cdots, D$$

- Correlations among the dimensions are not taken into account

Marginals of a Multivariate Distribution



- Given joint, we can uniquely determine marginals.
- Given marginals, we cannot uniquely determine joint!

Homework

- Assuming that the observed multidimensional data $\mathbf{x}_n$ follow a multivariate Gaussian distribution with parameters $\boldsymbol{\mu}$ and $\boldsymbol{\Sigma}$, evaluate their ML estimates.
- Analyze the conditions under which these estimates approximate those obtained with statistical independence assumption.

ML approach to Generative Models

- Training Data: $\mathcal{D} = \{(\mathbf{x}_1, t_1), (\mathbf{x}_2, t_2), \cdots, (\mathbf{x}_n, t_n), \cdots, (\mathbf{x}_N, t_N)\}$
  - $\mathbf{x}_n \in \mathbb{R}^D$ are input variables, $t_n \in \{0, 1\}$ are targets
  - Let N_1 points are from $\mathcal{C}_1$ and N_2 points are from $\mathcal{C}_2$: $N = N_1 + N_2$
- Generative models require estimation of joint density $p(\mathbf{x}_n, \mathcal{C}_k)$

$$p(\mathbf{x}_n, \mathcal{C}_k) = P[\mathcal{C}_k] p(\mathbf{x}_n / \mathcal{C}_k)$$

- Let the priors for $\mathcal{C}_1$ and $\mathcal{C}_2$ be π and $1 - \pi$, respectively
- Let class conditional densities $p(\mathbf{x}_n / \mathcal{C}_k)$ follow Gaussian distributions

$$p(\mathbf{x}_n / \mathcal{C}_k) \sim \mathcal{N}(\mathbf{x}_n / \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

- The joint densities are given by

$$\begin{aligned} \text{For } \mathbf{x}_n \in \mathcal{C}_1 \quad p(\mathbf{x}_n, \mathcal{C}_1) &= \pi \mathcal{N}(\mathbf{x}_n / \boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1) \qquad t_n = 1 \\ \text{For } \mathbf{x}_n \in \mathcal{C}_2 \quad p(\mathbf{x}_n, \mathcal{C}_2) &= (1 - \pi) \mathcal{N}(\mathbf{x}_n / \boldsymbol{\mu}_2, \boldsymbol{\Sigma}_2) \quad t_n = 0 \end{aligned}$$

Generative Models - ML Estimates

- Estimate $\theta = (\pi, \boldsymbol{\mu}_1, \boldsymbol{\mu}_2, \boldsymbol{\Sigma}_1, \boldsymbol{\Sigma}_2)$ to maximize the likelihood function

$$p(\mathcal{D}/\theta) = \prod_{n=1}^{N} \left[\pi \mathcal{N}(\mathbf{x}_n/\boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1)\right]^{t_n} \left[(1-\pi)\mathcal{N}(\mathbf{x}_n/\boldsymbol{\mu}_2, \boldsymbol{\Sigma}_2)\right]^{1-t_n}$$

- It is convenient to maximize log-likelihood $\log p(\mathbf{t}/\theta)$
- The terms involving π in log-likelihood $\mathcal{L}(\theta)$ are

$$\sum_{n=1}^{N} t_n \log \pi + (1 - t_n) \log(1 - \pi)$$

  - Setting the derivative of $\mathcal{L}(\theta)$ w.r.t to π to zero,

$$\pi = \frac{N_1}{N} \implies p[\mathcal{C}_1] = \frac{N_1}{N} \text{ and } p[\mathcal{C}_2] = \frac{N_2}{N}$$

Generative Models - ML Estimates

- $\boldsymbol{\mu}_1, \boldsymbol{\mu}_2, \boldsymbol{\Sigma}_1, \boldsymbol{\Sigma}_2$ can be evaluated in a similar manner

$$\boldsymbol{\mu}_1 = \frac{1}{N_1} \sum_{n \in \mathcal{C}_1} \mathbf{x}_n \qquad \boldsymbol{\Sigma}_1 = \frac{1}{N_1} \sum_{n \in \mathcal{C}_1} (\mathbf{x}_n - \boldsymbol{\mu}_1)(\mathbf{x}_n - \boldsymbol{\mu}_1)^{\mathsf{T}}$$

$$\boldsymbol{\mu}_2 = \frac{1}{N_2} \sum_{n \in \mathcal{C}_2} \mathbf{x}_n \qquad \boldsymbol{\Sigma}_2 = \frac{1}{N_2} \sum_{n \in \mathcal{C}_2} (\mathbf{x}_n - \boldsymbol{\mu}_2)(\mathbf{x}_n - \boldsymbol{\mu}_2)^{\mathsf{T}}$$

- Equivalent to estimating individual CCDs
- CCD estimation of one class does not involve data from other class
- Offers a descriptive model but need not be discriminative
- Decision boundary is locus of all points satisfying

$$P[C_1/\mathbf{x}] = P[C_2/\mathbf{x}]$$

Home Work - Shared Covariance Matrix

- Let both the classes share the same covariance matrix

$$\boldsymbol{\Sigma}_1 = \boldsymbol{\Sigma}_2 = \boldsymbol{\Sigma}$$

- The shape of the distributions is the same for both classes
- The ML solution to the shared covariance matrix is given by

$$\boldsymbol{\Sigma} = \frac{N_1}{N}\boldsymbol{\Sigma}_1 + \frac{N_2}{N}\boldsymbol{\Sigma}_2$$

- Shared covariance is the weighted combination of class-specific covariances.

Decision Boundary(Shared Covariance)

- The decision boundary is locus of points satisfying $P[\mathcal{C}_1/\mathbf{x}] = P[\mathcal{C}_2/\mathbf{x}]$

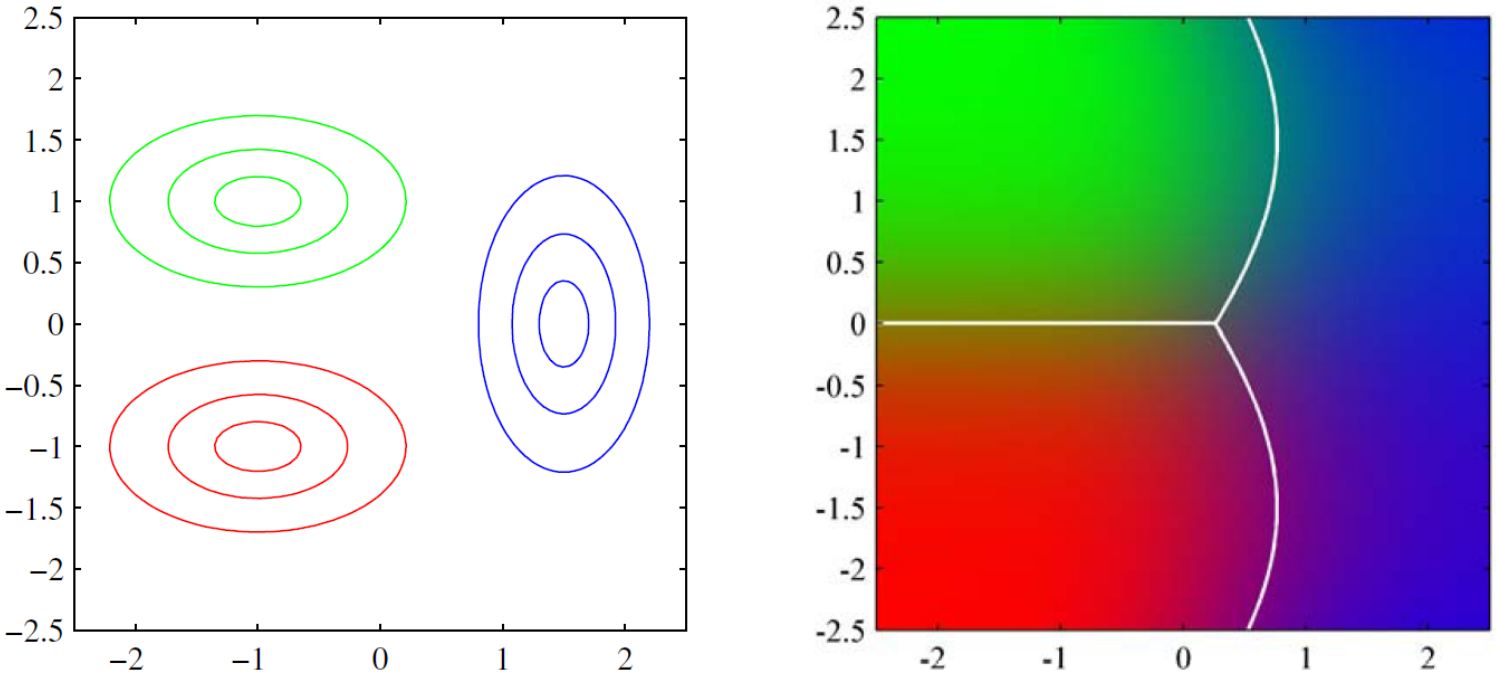
$$P[\mathcal{C}_1]\ \mathcal{N}(\mathbf{x}/\boldsymbol{\mu}_1, \boldsymbol{\Sigma}) = p[\mathcal{C}_2]\ \mathcal{N}(\mathbf{x}/\boldsymbol{\mu}_2, \boldsymbol{\Sigma})$$

- The quadratic terms cancel because of shared covariance
- The decision boundary is a linear in x: $\mathbf{w}^{\mathsf{T}}\mathbf{x} + w_0$

$$\mathbf{w} = \boldsymbol{\Sigma}^{-1}(\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2) \quad w_0 = -\frac{1}{2}\boldsymbol{\mu}_1^{\mathsf{T}}\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}_1 + \frac{1}{2}\boldsymbol{\mu}_2^{\mathsf{T}}\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}_2 + \log\frac{P[\mathcal{C}_1]}{P[\mathcal{C}_2]}$$

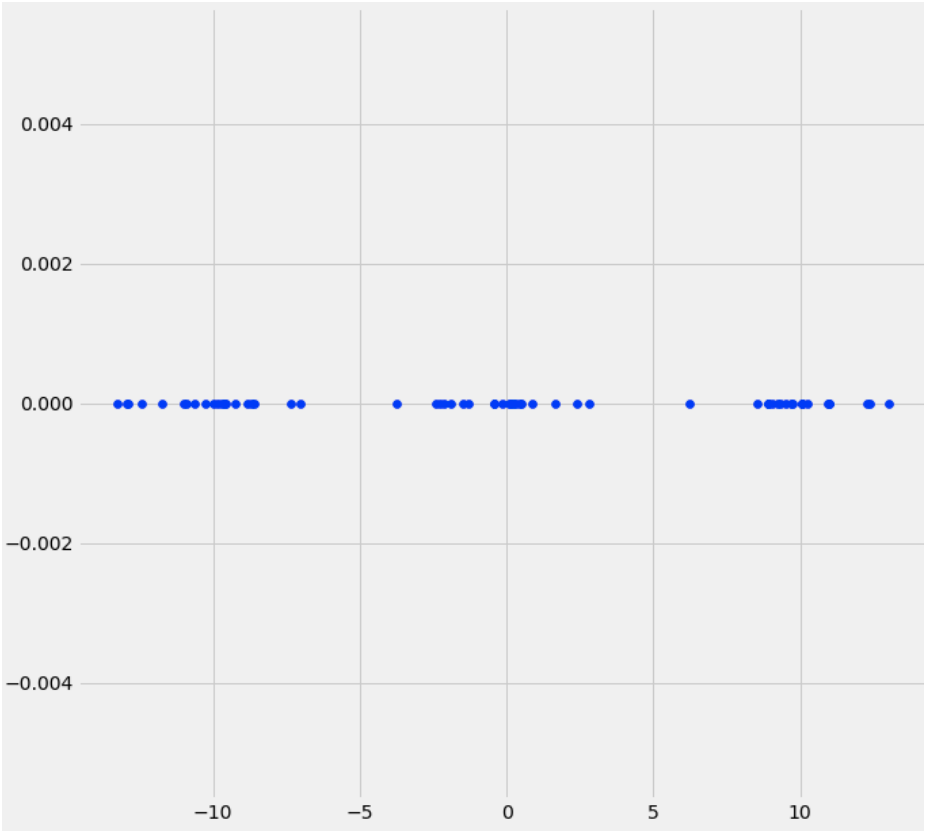
- Priors affect only the bias parameters, not the orientation
- The orientation is determined by ML estimates of CCD parameters
- The decision boundary would be quadratic if covariance is not shared

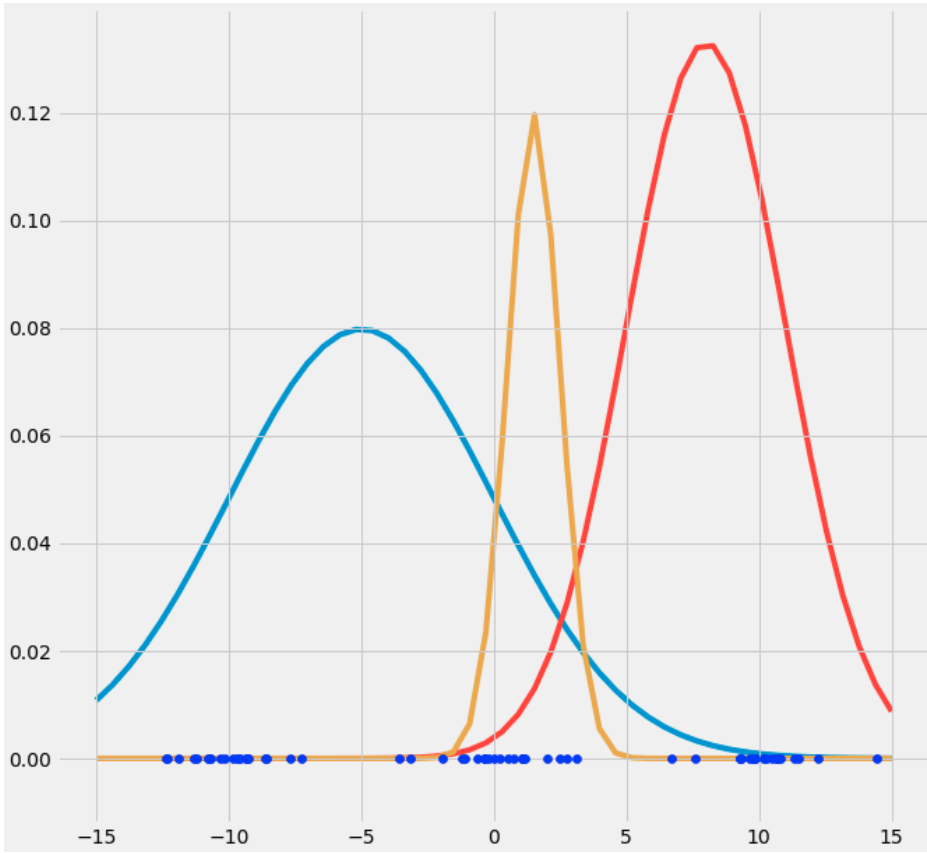
Illustration of Decision Boundaries



- Red and Green classes share the same covariance matrix
  - Decision boundary is linear
- Blue has a different covariance - decision boundaries are quadratic
- Nonlinear decision boundaries can be modeled with pdfs having higher order moments!

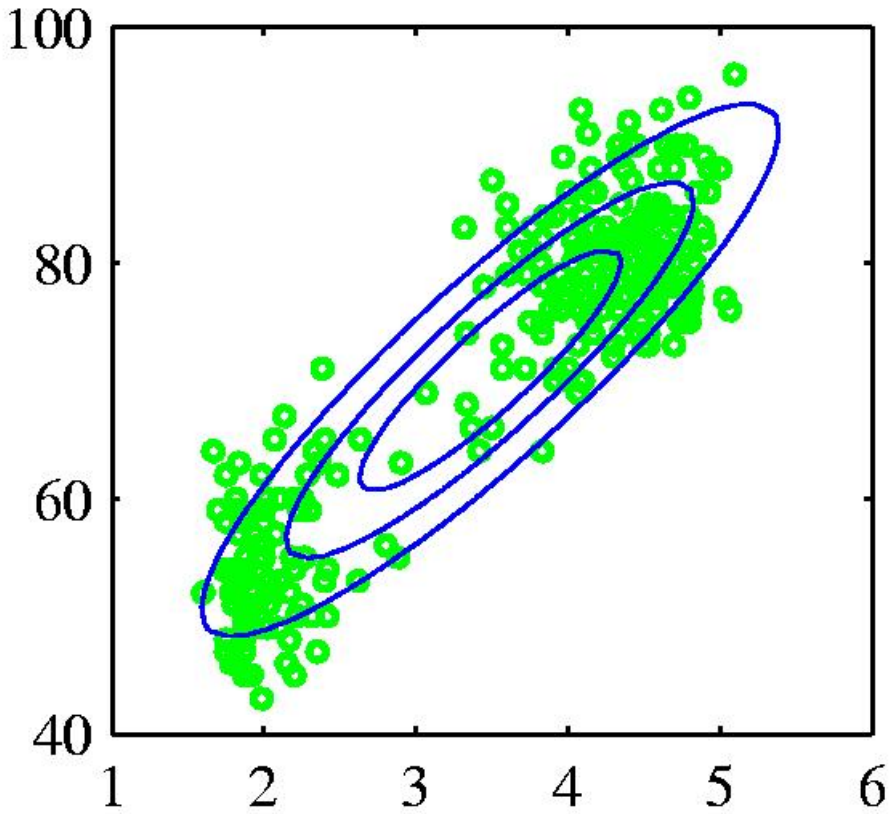
When one Gaussian is not enough!

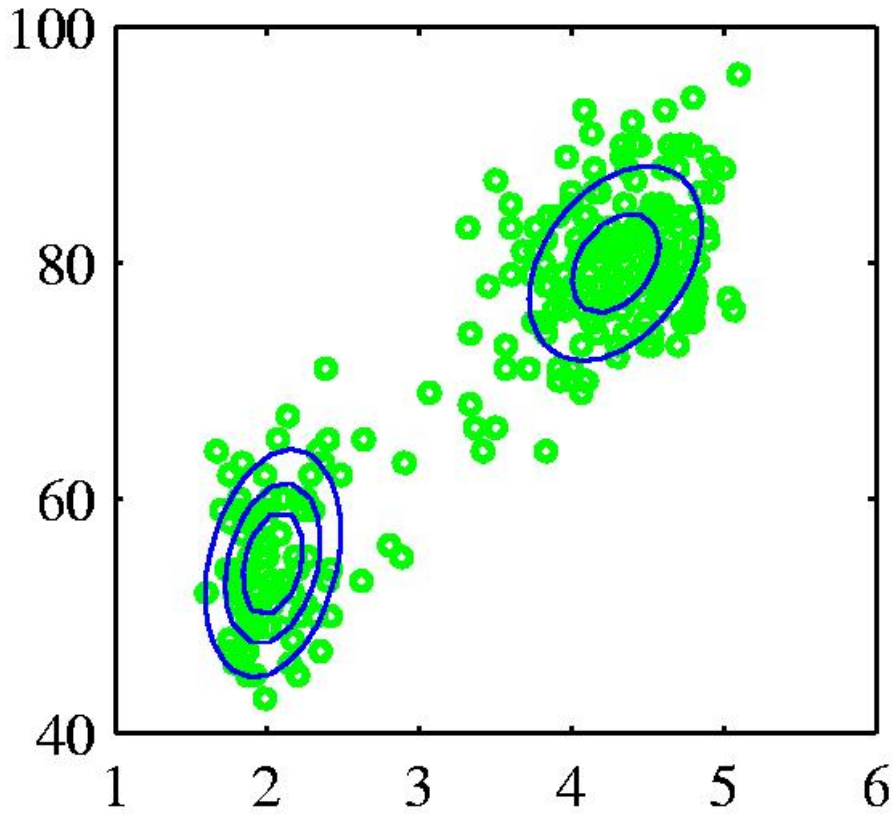




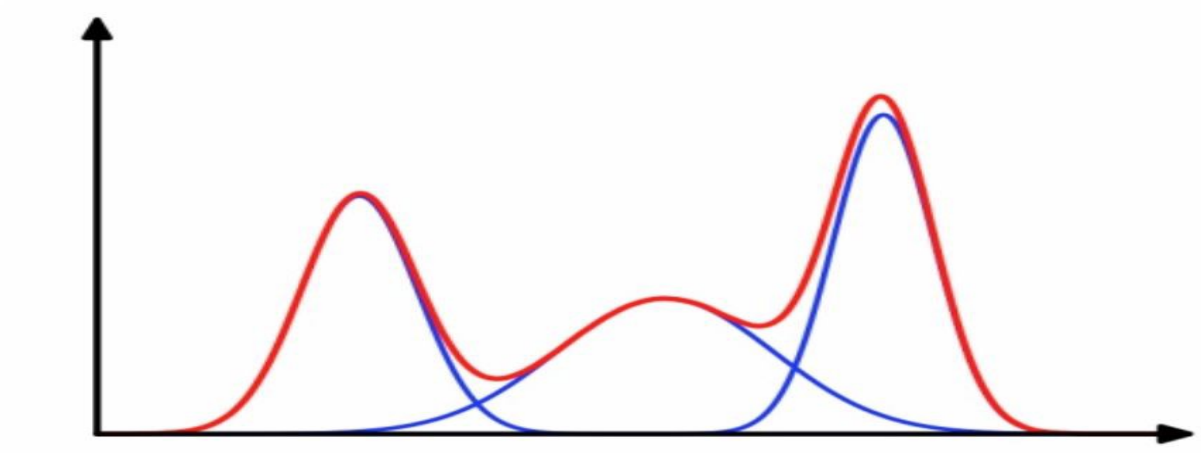
- Real-world datasets can have multimodal distribution
- Single Gaussian may not be a good approximation

2-D Illustration





Gaussian Mixture Model (GMM)



- PDF can be written as a linear weighted sum of Gaussian distributions
- Linear superposition of Gaussians

$$p(x) = \sum_{m=1}^{M} \pi_m \mathcal{N}(x/\mu_m, \sigma_m^2)$$

- PDF should be positive and integrate to one

$$\pi_m \geq 0 \qquad \& \qquad \sum_{m=1}^{M} \pi_m = 1$$

Sampling from GMM

- Parameters of the GMM $(\pi_m, \mu_m, \sigma_m^2), \quad m = 1, 2, \cdots, M$
- Data generation process
  - Roll an M-sided dice and observe the outcome (side k)
  - Generate a datapoint from the k^{th} Gaussian component (μ_k, σ_k^2)
  - Repeat the above two steps for generate multiple datapoints
- Given the observed data, we need to estimate the underlying pdf
  - Number of Gaussian components M is not known
  - M is typically determined empirically
  - Parameters of the GMM are estimated by maximizing the ML criterion

ML Formulation for GMM

- Consider data $X = \{x_1, x_2, \cdots x_N\}$ drawn from unknown distribution
- Let the unknown distribution $p(x_n)$ be approximated by a GMM

$$p(x_n) = \sum_{m=1}^{M} \pi_m \mathcal{N}(x_n/\mu_m, \sigma_m^2)$$

- Assuming i.i.d sampling, the likelihood function can expressed as

$$\mathcal{L}(\theta/X) = P(X/\theta) = \prod_{n=1}^{N} \sum_{m=1}^{M} \pi_m \mathcal{N}(x_n/\mu_m, \sigma_m^2)$$

- GMM parameters can be estimated by maximizing the log-likelihood

$$J(\theta) = \sum_{n=1}^{N} \log \left(\sum_{m=1}^{M} \pi_m \mathcal{N}(x_n/\mu_m, \sigma_m^2) \right) + \lambda \left(\sum_{m=1}^{M} \pi_m - 1 \right)$$

Estimating μ_k

$$J(\theta) = \sum_{n=1}^{N} \log \left(\sum_{m=1}^{M} \pi_m \mathcal{N}(x_n/\mu_m, \sigma_m^2) \right) + \lambda \left(\sum_{m=1}^{M} \pi_m - 1 \right)$$

- Taking partial derivative w.r.t μ_k

$$\begin{aligned}
\frac{\partial J(\theta)}{\partial \mu_k} &= \sum_{n=1}^{N} \frac{1}{\sum_{m=1}^{M} \pi_m \mathcal{N}(x_n/\mu_m, \sigma_m^2)} \pi_k \frac{\partial}{\partial \mu_k} \mathcal{N}(x_n/\mu_k, \sigma_k^2) \\
&= \sum_{n=1}^{N} \frac{\pi_k \mathcal{N}(x_n/\mu_k, \sigma_k^2)}{\sum_{m=1}^{M} \pi_m \mathcal{N}(x_n/\mu_m, \sigma_m^2)} \left(\frac{x_n - \mu_k}{\sigma_k^2} \right) \\
&= \sum_{n=1}^{N} \gamma_{nk} \left(\frac{x_n - \mu_k}{\sigma_k^2} \right)
\end{aligned}$$

Interpretation of γ_{nk}

$$\gamma_{nk} = \frac{\pi_k \mathcal{N}(x_n/\mu_k, \sigma_k^2)}{\sum_{m=1}^{M} \pi_m \mathcal{N}(x_n/\mu_m, \sigma_m^2)}$$

- γ_{nk}: Responsibility of k^{th} Gaussain in generating n^{th} datapoint
- All components should collectively share the responsibility

$$\sum_{k=1}^{M} \gamma_{nk} = 1$$

- Effective number of points generated by k^{th} component

$$\sum_{n=1}^{N} \gamma_{nk} = N_k \qquad \sum_{k=1}^{M} N_k = N$$

- After estimation, γ_{nk} can be interpreted as posterior probability

Estimates of GMM Parameters

- Estimated mean of the k^{th} component

$$\hat{\mu}_k = \frac{1}{N_k} \sum_{n=1}^{N} \gamma_{nk} x_n \qquad k = 1, 2, \cdots M$$

- Estimated variance of the k^{th} component

$$\hat{\sigma}_k^2 = \frac{1}{N_k} \sum_{n=1}^{N} \gamma_{nk} (x_n - \hat{\mu}_k)^2 \qquad k = 1, 2, \cdots M$$

- Estimated weight of the k^{th} component (prior)

$$\hat{\pi}_k = \frac{N_k}{N} \qquad k = 1, 2, \cdots M$$

- Are they in closed form? Can we directly compute them from data?

EM Algorithm for GMM

- Given data $X = \{x_1, x_2, \cdots x_N\}$
- **Initialization**: Choose the number of mixtures M
  - Randomly choose means (μ_k), variances (σ_k^2) and mixture weights (π_k)
  - Evaluate log-likelihood with initial conditions $p(X/\theta)$
- **Expectation step**: Evaluate the responsibilities using the current θ

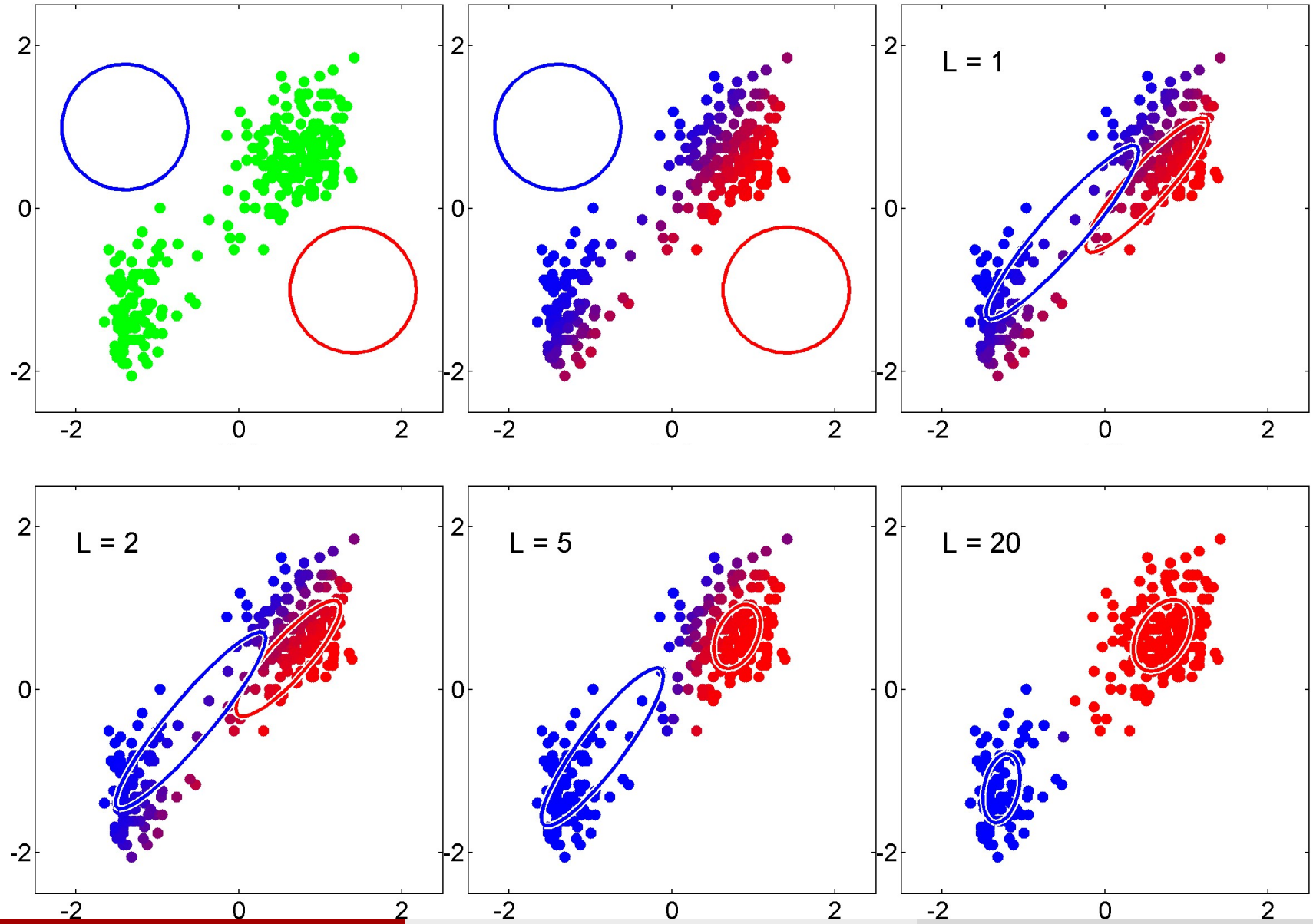
$$\gamma_{nk} = \frac{\pi_k \mathcal{N}(x_n/\mu_k, \sigma_k^2)}{\sum_{m=1}^{M} \pi_m \mathcal{N}(x_n/\mu_m, \sigma_m^2)} \qquad N_k = \sum_{n=1}^{N} \gamma_{nk}$$

- **Maximization step**: Update parameters with current γ_{nk}

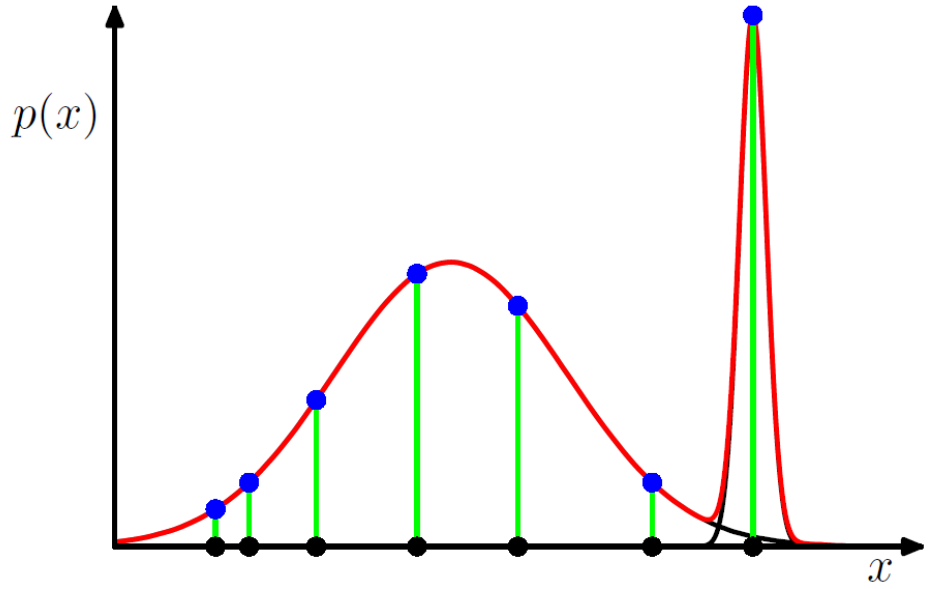
$$\hat{\mu}_k^{new} = \frac{1}{N_k} \qquad \hat{\sigma}_k^{new} = \frac{1}{N_k} \sum_{n=1}^{N} \gamma_{nk}(x_n - \hat{\mu}_k)^2 \qquad \hat{\pi}_k^{new} = \frac{N_k}{N}$$

- Compute log-likelihood with updated parameters
- Repeat E-step and M-step until convergence

Illustration of EM Iterations

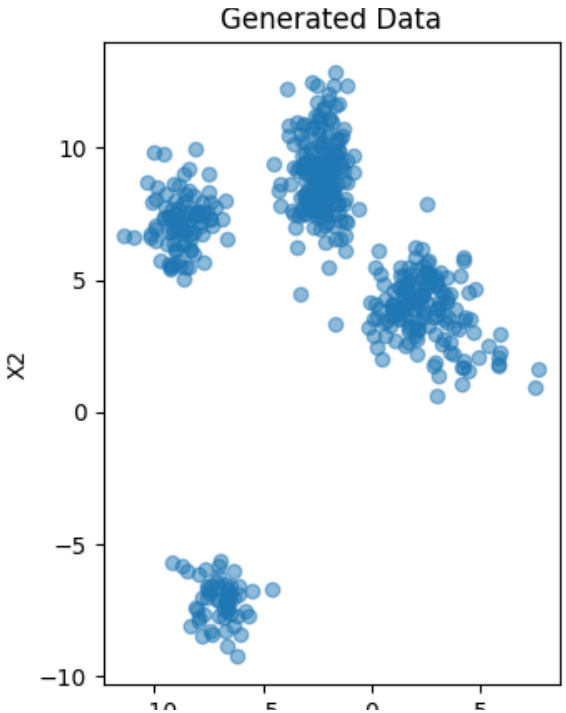


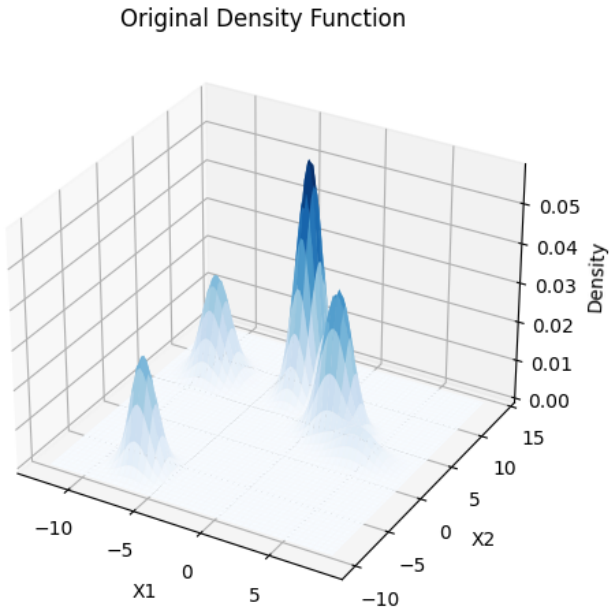
Singularities in Likelihood Function

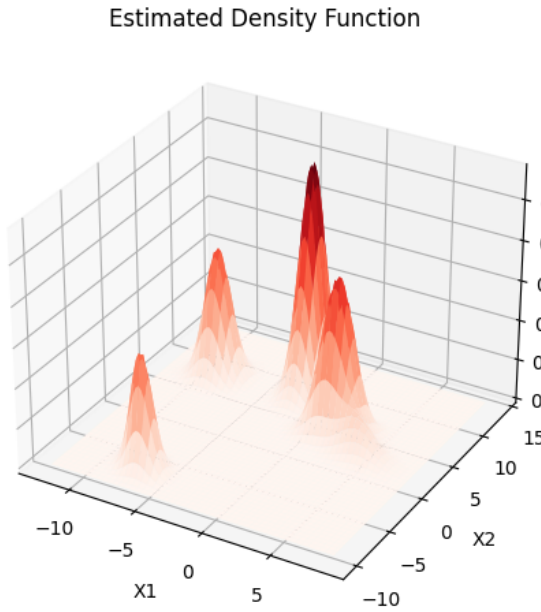


- One component has finite variance
- Some Gaussian components can collapse to specific data points
- Shrinking variance can lead to ever-increasing log-likelihood!
- ML approach can result in severe over-fitting (local maximum)
- Poor initialization, outliers, smaller dataset, more components

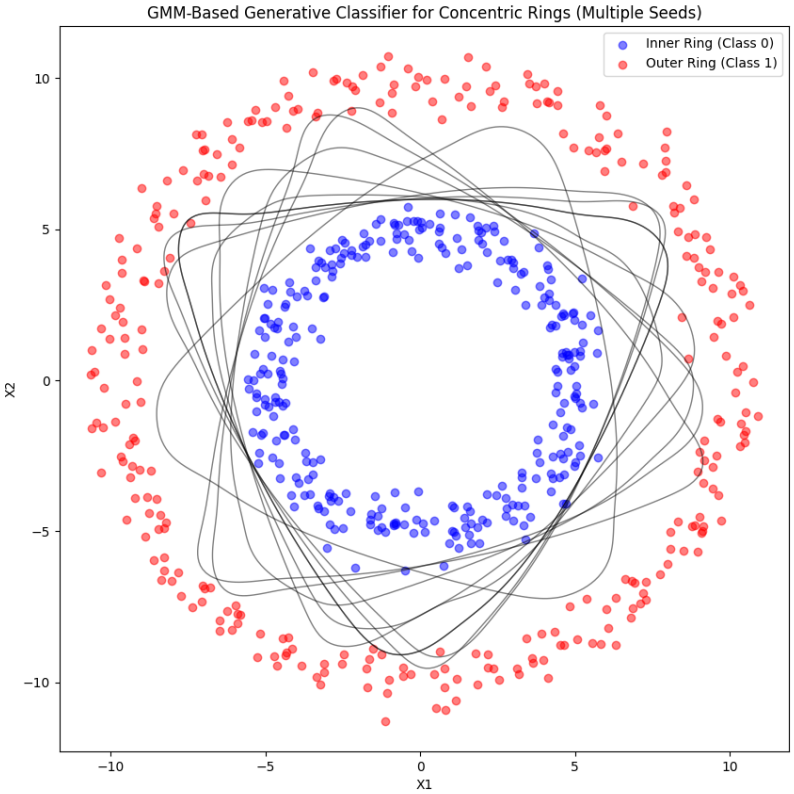
GMM Density Estimation - Illustration

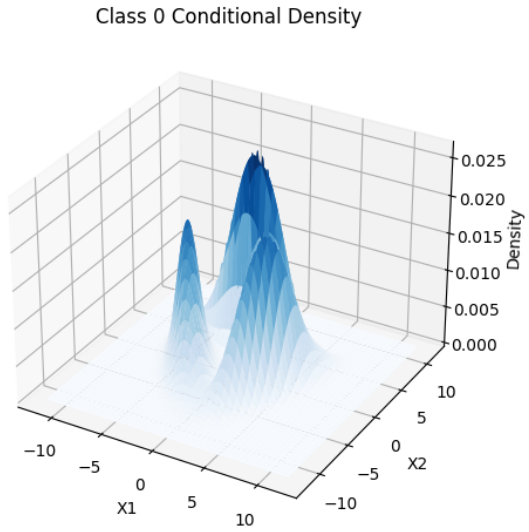


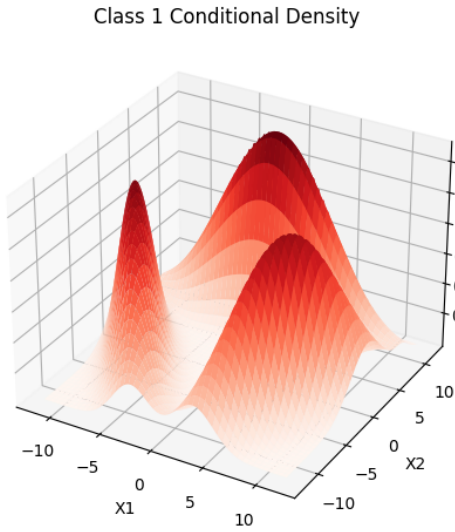




Generative Modeling using GMM







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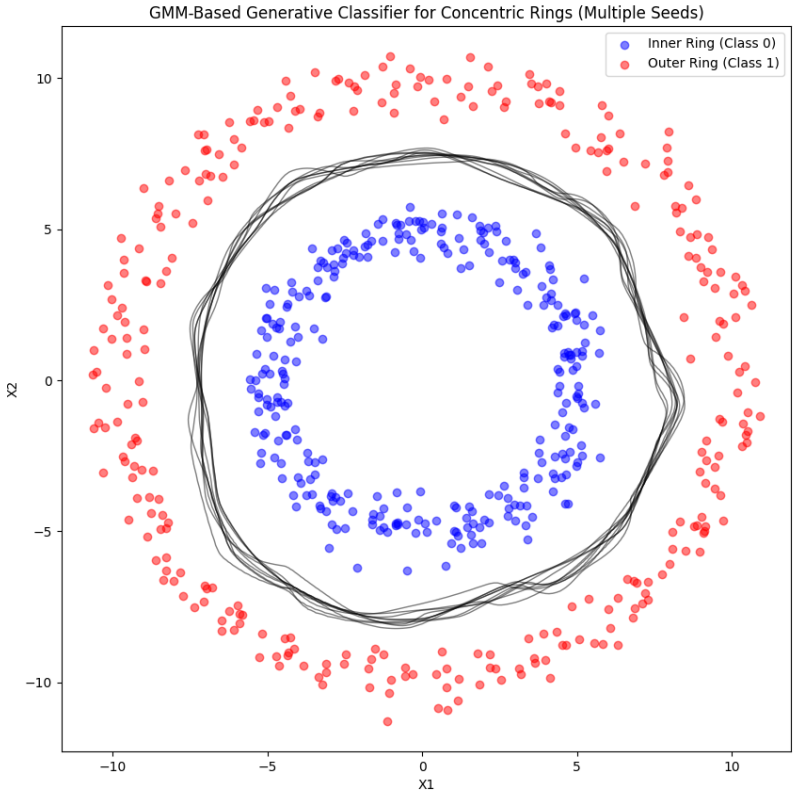
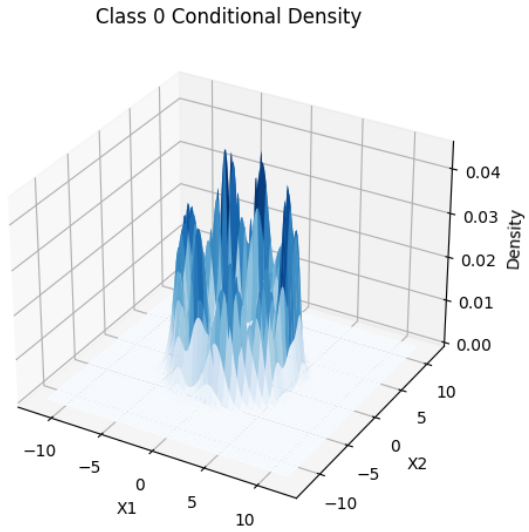
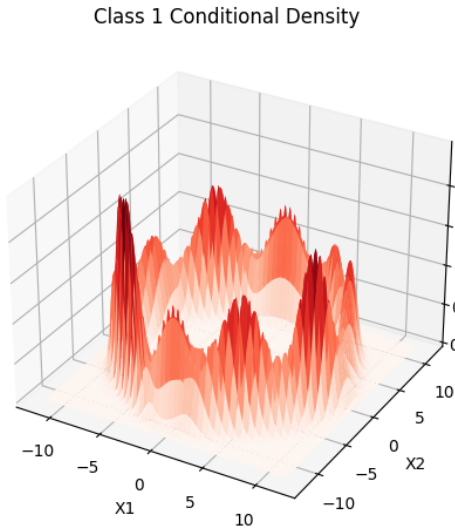






Summary of Generative Models

- Generative models are aimed at modeling the joint density function
  - Prior $P[\mathcal{C}_k]$, likelihood or CCD $p(\mathbf{x}/\mathcal{C}_k)$, and partition function $p(\mathbf{x})$
- Statistical approaches use ML approaches to estimate CCD
  - A parametric form is imposed on the CCD - Gaussian, Laplacian, etc
  - The parameters are estimated from the observed data.
- Naive Bayes assumes that dimensions are statistically independent

$$p(\mathbf{x}/\mathcal{C}_k) = \prod_{i=1}^{d} p(x_i/\mathcal{C}_k) \qquad p(x_i/\mathcal{C}_k) \sim \mathcal{N}(x_i/\mu_i, \sigma_i^2)$$

- Multivariate Gaussian to model correlated dimensions

$$p(\mathbf{x}/\mathcal{C}_k) \sim \mathcal{N}(\mathbf{x}/\boldsymbol{\mu}, \boldsymbol{\Sigma})$$

- Gaussian mixture to model multi-modal distributions

Recap: Formulation of Generative Models

- Training Data: $\mathcal{D} = \{(\mathbf{x}_1, t_1), (\mathbf{x}_2, t_2), \cdots, (\mathbf{x}_n, t_n), \cdots, (\mathbf{x}_N, t_N)\}$
  - $\mathbf{x}_n \in \mathbb{R}^D$ are input variables, $t_n \in \{0, 1\}$ are targets
  - Let N_1 points are from $\mathcal{C}_1$ and N_2 points are from $\mathcal{C}_2$: $N = N_1 + N_2$
- Generative models require estimation of joint density $p(\mathbf{x}_n, \mathcal{C}_k)$

$$p(\mathbf{x}_n, \mathcal{C}_k) = P[\mathcal{C}_k] p(\mathbf{x}_n / \mathcal{C}_k)$$

- Let the priors for $\mathcal{C}_1$ and $\mathcal{C}_2$ be π and $1 - \pi$, respectively
- Let class conditional densities $p(\mathbf{x}_n / \mathcal{C}_k)$ follow Gaussian distributions

$$p(\mathbf{x}_n / \mathcal{C}_k) \sim \mathcal{N}(\mathbf{x}_n / \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

- Estimate $\theta = (\pi, \boldsymbol{\mu}_1, \boldsymbol{\mu}_2, \boldsymbol{\Sigma}_1, \boldsymbol{\Sigma}_2)$ to maximize the likelihood function

$$p(\mathcal{D}/\theta) = \prod_{n=1}^{N} [\pi \mathcal{N}(\mathbf{x}_n / \boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1)]^{t_n} [(1 - \pi) \mathcal{N}(\mathbf{x}_n / \boldsymbol{\mu}_2, \boldsymbol{\Sigma}_2)]^{1 - t_n}$$

Probabilistic Discriminative Models - Binary Classifier

- Discriminative models impose a parametric function on posterior

$$P[\mathcal{C}_1/\mathbf{x}] = y(\mathbf{x}, \mathbf{w}) \qquad P[\mathcal{C}_2/x] = 1 - y(\mathbf{x}, \mathbf{w})$$

- The function $y(\mathbf{x}, \mathbf{w})$ has to be chosen to satisfy axioms of probability
- The posterior probability can be expressed as

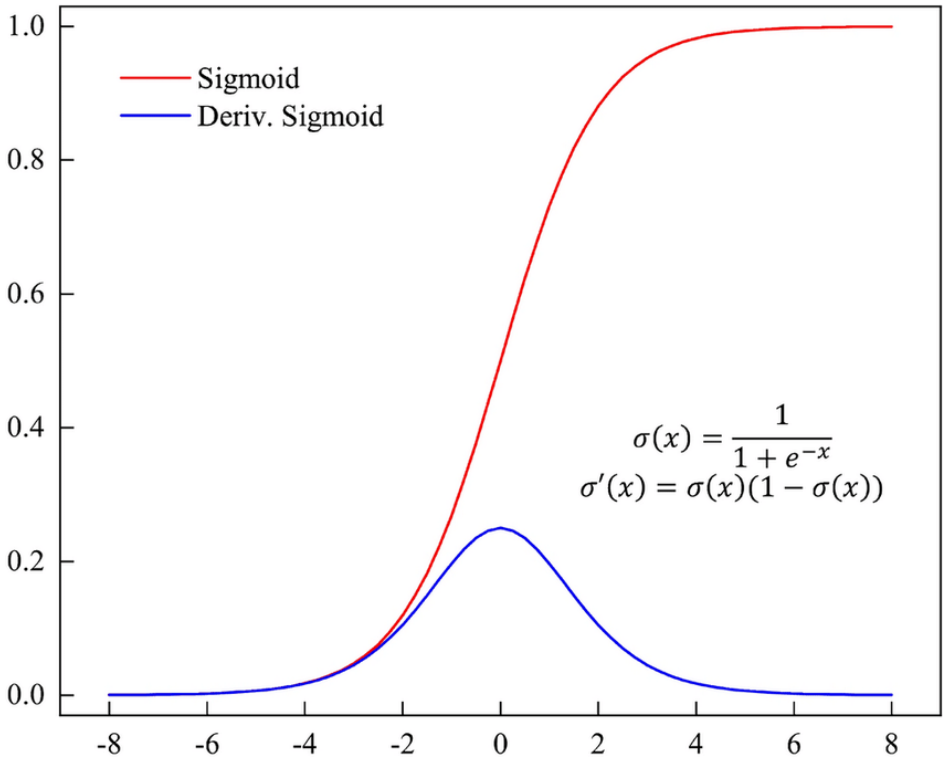
$$p[\mathcal{C}_1/\mathbf{x}] = \frac{P[\mathcal{C}_1]p(\mathbf{x}/\mathcal{C}_1)}{P[\mathcal{C}_1]p(\mathbf{x}/\mathcal{C}_1) + P[\mathcal{C}_2]p(\mathbf{x}/\mathcal{C}_2)}$$

- Let $a = \log \frac{P[\mathcal{C}_1]p(\mathbf{x}/\mathcal{C}_1)}{P[\mathcal{C}_2]p(\mathbf{x}/\mathcal{C}_2)}$ be parameterized as $a = \mathbf{w}^{\mathsf{T}}\mathbf{x}$
- Posterior probability of $\mathcal{C}_1$ can be expressed as Sigmoid over a

$$P[\mathcal{C}_1/\mathbf{x}] = \frac{1}{1 + e^{-a}} = \sigma(a)$$

- Posterior probability of $\mathcal{C}_2$ is given by $P[\mathcal{C}_2/\mathbf{x}] = 1 - \sigma(a)$

Sigmoid or Logistic Function



Logistic Regression

- Training Data: $\mathcal{D} = \{(\mathbf{x}_1, t_1), (\mathbf{x}_2, t_2), \cdots, (\mathbf{x}_n, t_n), \cdots, (\mathbf{x}_N, t_N)\}$
  - $\mathbf{x}_n \in \mathbb{R}^D$ are input variables, $t_n \in \{0, 1\}$ are targets: $P[\mathcal{C}_1/\mathbf{x}_n]$
  - Let the target for $\mathcal{C}_1$ be 1 and $\mathcal{C}_2$ be 0
- Let the posterior probability be estimated as

$$\hat{P}[\mathcal{C}_1/\mathbf{x}_n] = y(\mathbf{x}_n, \mathbf{w}) = \sigma(\mathbf{w}^\mathsf{T}\mathbf{x}_n) \qquad \hat{P}[\mathcal{C}_2/\mathbf{x}_n] = 1 - \sigma(\mathbf{w}^\mathsf{T}\mathbf{x}_n)$$

- Assuming that data points are i.i.d., the likelihood of data is given by

$$P(\mathbf{t}/\mathbf{X}, \mathbf{w}) = \prod_{n=1}^{N} y_n^{t_n}(1 - y_n)^{1-t_n}$$

- $\mathbf{w}$ can be estimated by minimizing the negative log of the likelihood, also referred to as *binary cross-entropy* loss

$$J(\mathbf{w}) = -\log P(\mathbf{t}/\mathbf{X}, \mathbf{w}) = -\sum_{1}^{N}\{t_n \log y_n + (1 - t_n)\log(1 - y_n)\}$$

Parameter Estimation

- Model parameters $\mathbf{w}$ can be updated using gradient descent

$$\mathbf{w}^{new} = \mathbf{w}^{old} - \eta \nabla J(\mathbf{w})$$

- The 1st and 2nd order derivatives of the loss function are given by

$$\nabla J(\mathbf{w}) = \sum_{n=1}^{N} \mathbf{x}_n (y_n - t_n) = \mathbf{X}^{\mathsf{T}} (\mathbf{y} - \mathbf{t})$$

$$\mathbf{H} = \nabla^2 J(\mathbf{w}) = \sum_{n=1}^{N} \mathbf{x}_n y_n (1 - y_n) \mathbf{x}_n^{\mathsf{T}} = \mathbf{X}^{\mathsf{T}} \mathbf{R} \mathbf{X}$$

  where $\mathbf{R}$ is a diagonal matrix with entries $R_{nn} = y_n(1 - y_n)$

- Hessian matrix varies depending on parameters $\mathbf{w}$ through $\mathbf{R}$
- Since $0 < y_n < 1$, Hessian matrix $\mathbf{H}$ is positive definite: $\mathbf{u}^{\mathsf{T}} \mathbf{H} \mathbf{u} > 0$
- Error function $J(\mathbf{w})$ is convex in $\mathbf{w}$ $\implies$ admits unique minimum

Iterative Reweighted Least Squares

$$\begin{aligned}
\mathbf{w}^{new} &= \mathbf{w}^{old} - \left(\mathbf{X}^{\mathsf{T}}\mathbf{R}\mathbf{X}\right)^{-1}\mathbf{X}^{\mathsf{T}}(\mathbf{y}-\mathbf{t}) \\
&= \left(\mathbf{X}^{\mathsf{T}}\mathbf{R}\mathbf{X}\right)^{-1}\left(\mathbf{X}^{\mathsf{T}}\mathbf{R}\mathbf{X}\mathbf{w}^{old} - \mathbf{X}^{\mathsf{T}}(\mathbf{y}-\mathbf{t})\right) \\
&= \left(\mathbf{X}^{\mathsf{T}}\mathbf{R}\mathbf{X}\right)^{-1}\mathbf{X}^{\mathsf{T}}\mathbf{R}\left(\mathbf{X}\mathbf{w}^{old} - \mathbf{R}^{-1}(\mathbf{y}-\mathbf{t})\right) \\
&= \left(\mathbf{X}^{\mathsf{T}}\mathbf{R}\mathbf{X}\right)^{-1}\mathbf{X}^{\mathsf{T}}\mathbf{R}\mathbf{z}
\end{aligned}$$

where $\mathbf{z} = \left(\mathbf{X}\mathbf{w}^{old} - \mathbf{R}^{-1}(\mathbf{y}-\mathbf{t})\right)$.

- The solution takes the form of normal equations of weighted LS.
- However, the weighing matrix $\mathbf{R}$ is not constant, but depends on $\mathbf{w}$
- Hence, the normal equations need to be applied iteratively.

Extending to Multiple Classes

- Discriminative models impose a parametric function on posterior

$$P[\mathcal{C}_k/\mathbf{x}] = y(\mathbf{x}, \mathbf{W})$$

- The function $y(.)$ should satisfy the axioms of probability
- The posterior probability of the k^{th} class is given by (Bayes)

$$P[\mathcal{C}_k/\mathbf{x}] = \frac{P[\mathcal{C}_k]p(\mathbf{x}/\mathcal{C}_k)}{\sum_{j=1}^{K} P[\mathcal{C}_j]p(\mathbf{x}/\mathcal{C}_j)}$$

- Let $a_k = \log P[\mathcal{C}_k]p(\mathbf{x}/\mathcal{C}_k)$ be parameterized as $a_k = \mathbf{w}_k^{\mathsf{T}}\mathbf{x}$
- Posterior probability can be expressed as a softmax over activations a_k

$$P[\mathcal{C}_k/\mathbf{x}] = \frac{\exp(a_k)}{\sum_{j=1}^{K} \exp(a_j)} \qquad \text{Softmax Fn.}$$

Multiclass Logistic Regression (Homework)

- Consider multiclass data examples denoted by $X = \{(\mathbf{x}_{1:N}, \mathbf{t}_{1:N})\}$
  - $\mathbf{x}_n \in \mathbb{R}^D$ represents input observations
  - $\mathbf{t}_n$ is a K-dim one-hot vector denoting the class posteriors
- For this case, the posterior probabilities can be estimated as

$$y(\mathbf{x}_n, \mathbf{w}_k) = \hat{P}[\mathcal{C}_k/\mathbf{x}_n] = \frac{\exp(a_{nk})}{\sum_{j=1}^{K} \exp(a_{nj})} \qquad a_{nk} = \mathbf{w}_k^{\mathsf{T}} \mathbf{x}_n$$

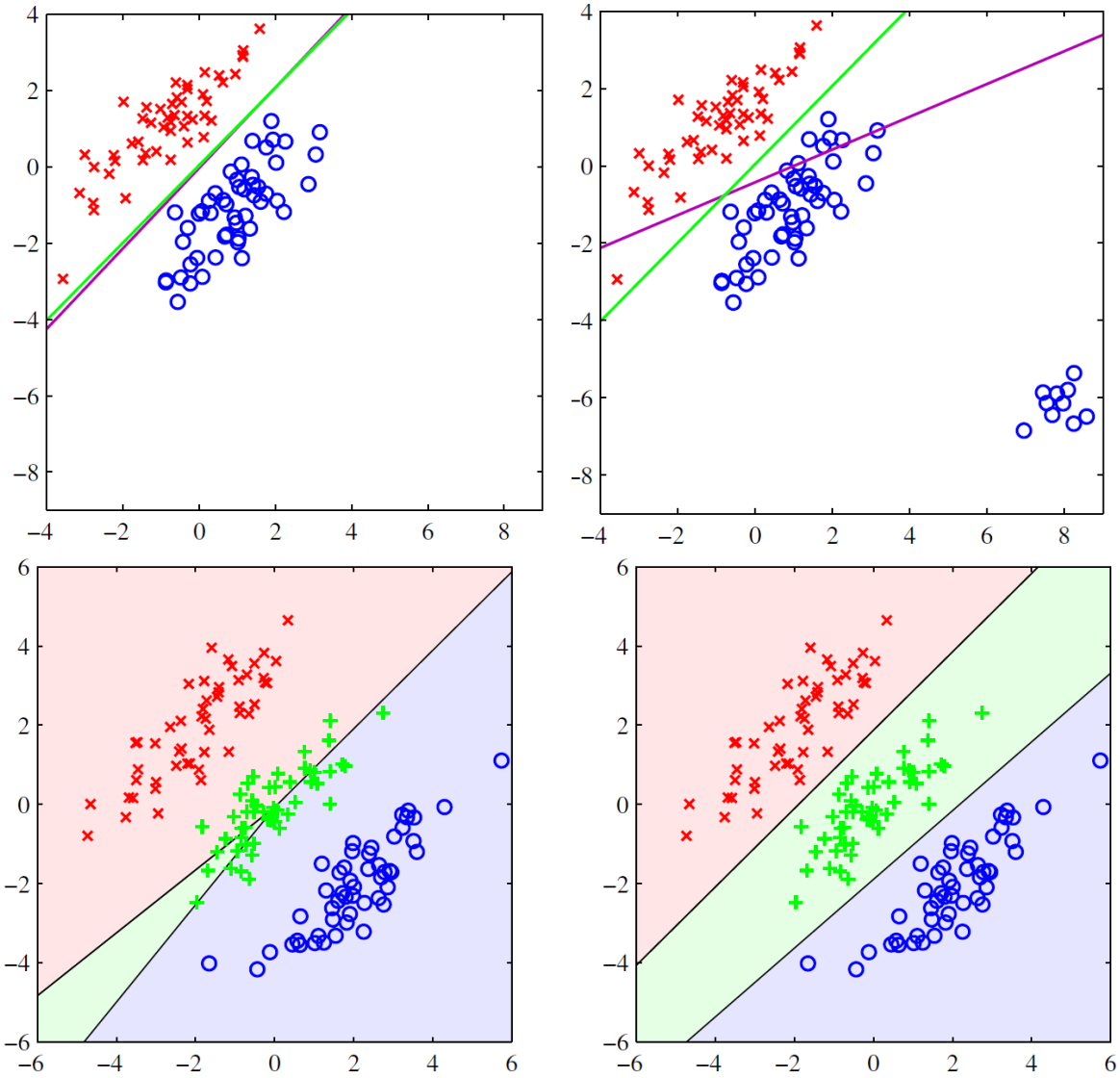
- The likelihood function can be written as

$$P[\mathbf{T}/\mathbf{X}, \mathbf{w}_1, \mathbf{w}_2, \cdots, \mathbf{w}_K] = \prod_{n=1}^{N} \prod_{k=1}^{K} y_{nk}^{t_{nk}}$$

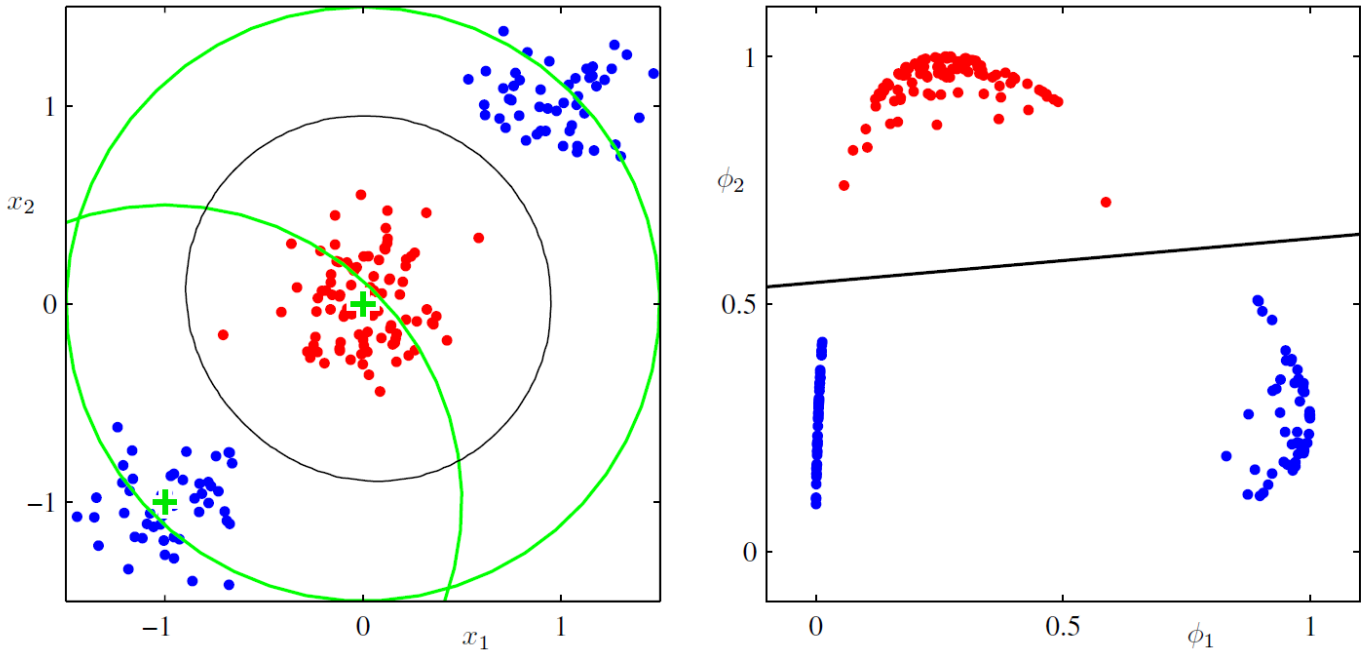
$$J(\mathbf{W}) = -P[\mathbf{T}/\mathbf{X}, \mathbf{W}] = -\sum_{n=1}^{N} \sum_{k=1}^{K} t_{nk} \log y_{nk} \qquad \text{(CE)}$$

- Derive update equations for multiclass logistic regression

Illustration of Logistic Regression



Nonlinear Decision Boundary (Transformed Space $\phi(.)$)



- Two Gaussian kernels are used to transform the data
- In general, we need to design the kernel ϕ from the data.
- DNN can be used to learn the optimal transformation from the data
- Last layer of a DNN classifier performs logistic regression

Summary of Linear Classifiers

- Assumption: Classes are separable by linear hyperplanes
- Linear discriminant functions model the separating hyperplanes
  - Least squares, Fisher discriminant, Perceptron, SVM (later)
  - May not work even if classes are separable because of outliers
- Generative models estimate posteriors from priors and CCDs
  - ML or MAP estimators are employed to model CCD
  - Don't consider other class examples while estimating CCD
- Discriminative models directly estimate posterior probabilities
  - Binary classes - logistic activation - binary cross entropy
  - Multiple classes - softmax activation - cross entropy
  - Rely on discriminative features - vulnerable to adversarial examples
- If decision boundary is not linear, apply these techniques on $\phi(\mathbf{x})$
  - Neural networks offer *a way* of learning $\phi(\mathbf{x})$ from data

Thank You!